

# Active galactic nuclei vs. host galaxy properties in the COSMOS field<sup>★</sup>

G. Lanzuisi<sup>1,2</sup>, I. Delvecchio<sup>3</sup>, S. Berta<sup>4,★★</sup>, M. Brusa<sup>1,2</sup>, A. Comastri<sup>2</sup>, R. Gilli<sup>2</sup>, C. Gruppioni<sup>2</sup>, S. Marchesi<sup>2,5</sup>,  
M. Perna<sup>1,2</sup>, F. Pozzi<sup>1,2</sup>, M. Salvato<sup>6</sup>, M. Symeonidis<sup>7</sup>, C. Vignali<sup>1,2</sup>, F. Vito<sup>8</sup>, M. Volonteri<sup>9</sup>, and G. Zamorani<sup>2</sup>.

<sup>1</sup> Dipartimento di Fisica e Astronomia, Università di Bologna, Viale Berti Pichat 6/2, 40127 Bologna, Italy  
e-mail: giorgio.lanzuisi2@unibo.it

<sup>2</sup> INAF–Osservatorio Astronomico di Bologna, via Ranzani 1, 40127 Bologna, Italy

<sup>3</sup> Department of Physics, University of Zagreb, Bijenička cesta 32, 10002 Zagreb, Croatia

<sup>4</sup> Department of Physics, Faculty of Science, University of Zagreb, Bijenička cesta 32, 10000 Zagreb, Croatia

<sup>5</sup> Department of Physics & Astronomy, Clemson University, Clemson, SC 29634, USA

<sup>6</sup> Max-Planck-Institut für extraterrestrische Physik, Giessenbachstrasse, 85748 Garching, Germany

<sup>7</sup> Mullard Space Science Laboratory, University College London, Holmbury St. Mary, Dorking, Surrey RH5 6NT, UK

<sup>8</sup> Department of Astronomy and Astrophysics, 525 Davey Laboratory, The Pennsylvania State University, U. Park, PA 16802, USA

<sup>9</sup> Institut d’Astrophysique de Paris, UPMC et CNRS, UMR 7095, 98bis Bd Arago, 75014 Paris, France

Received 25 October 2016 / Accepted 22 February 2017

## ABSTRACT

**Context.** The coeval active galactic nuclei (AGN) and galaxy evolution, and the observed local relations between super massive black holes (SMBHs) and galaxy properties suggest some sort of connection or feedback between SMBH growth (i.e., AGN activity) and galaxy build-up (i.e., star formation history).

**Aims.** We looked for correlations between average properties of X-ray detected AGN and their far-IR (FIR) detected, star forming host galaxies in order to find quantitative evidence for this connection, which has been highly debated in recent years.

**Methods.** We exploited the rich multiwavelength data set (from X-ray to FIR) available in the COSMOS field for a large sample (692 sources) of AGN and their hosts in the redshift range  $0.1 < z < 4$ . We use X-ray data to select AGN and determine their properties, such as X-ray intrinsic luminosity and nuclear obscuration, and broadband (from UV to FIR) SED fitting results to derive host galaxy properties, such as stellar mass ( $M_*$ ) and star formation rate (SFR).

**Results.** We find that the AGN 2–10 keV luminosity ( $L_X$ ) and the host 8–1000  $\mu\text{m}$  star formation luminosity ( $L_{\text{IR}}^{\text{SF}}$ ) are significantly correlated, even after removing the dependency of both quantities with redshift. However, the average host  $L_{\text{IR}}^{\text{SF}}$  has a flat distribution in bins of AGN  $L_X$ , while the average AGN  $L_X$  increases in bins of host  $L_{\text{IR}}^{\text{SF}}$  with logarithmic slope of  $\sim 0.7$  in the redshift range  $0.4 < z < 1.2$ . We also discuss the comparison between the full distribution of these two quantities and the predictions from hydrodynamical simulations. No other significant correlations between AGN  $L_X$  and host properties is found. On the other hand, we find that the average column density ( $N_{\text{H}}$ ) shows a clear positive correlation with the host  $M_*$  at all redshifts, but not with the SFR (or  $L_{\text{IR}}^{\text{SF}}$ ). This translates into a negative correlation with specific SFR at all redshifts. The same is true if the obscured fraction is computed.

**Conclusions.** Our results are in agreement with the idea, introduced in recent galaxy evolutionary models, that SMBH accretion and SFRs are correlated, but occur with different variability time scales. Finally, the presence of a positive correlation between  $N_{\text{H}}$  and host  $M_*$  suggests that the column density that we observe in the X-rays is not entirely due to the circumnuclear obscuring torus, but may also include a significant contribution from the host galaxy.

**Key words.** galaxies: active – galaxies: nuclei – galaxies: evolution – infrared: galaxies – X-rays: galaxies

## 1. Introduction

Super massive black hole (SMBH) growth and galaxy build up follow a similar evolution through cosmic history, with a peak at  $z \sim 2-3$  and a sharp decline toward the present age (see Madau & Dickinson 2014, for a review). Furthermore, at  $z = 0$ , SMBH and their hosts sit on tight relations that link the SMBH mass and the bulge properties of the host, such as luminosity, stellar mass, and velocity dispersion (Kormendy & Richstone 1995; Magorrian et al. 1998; Kormendy & Ho 2013). Therefore SMBH

growth and star formation history are likely related in some way during the cosmic time.

The key parameter that regulates both processes seems to be cold/molecular gas supply (Lagos et al. 2011; Vito et al. 2014; Delvecchio et al. 2015; Saintonge et al. 2016). Processes related to star formation (e.g., supernova and stellar winds) are known to produce galaxy-wide outflows that can regulate the in-fall of gas and therefore star formation itself (e.g., Genzel et al. 2011). However, more powerful mechanisms, globally known as “AGN feedback”, have been invoked in numerical and semi-analytic models of galaxy evolution (e.g., Granato et al. 2004; Di Matteo et al. 2005; Menci et al. 2008; Sijacki et al. 2015; Dubois et al. 2016; Pontzen et al. 2017) in order to reproduce the observed galaxy population, and particularly the high mass end of the galaxy mass function.

\* Full Table 1 is only available at the CDS via anonymous ftp to [cdsarc.u-strasbg.fr](http://cdsarc.u-strasbg.fr) (130.79.128.5) or via <http://cdsarc.u-strasbg.fr/viz-bin/qcat?J/A+A/602/A123>

\*\* Visiting scientist.

Observationally, the role of AGN activity in influencing the evolution of the global galaxy population is not clear yet. This issue has been investigated, and correlations searched for between average AGN and host properties, such as BH accretion rates (BHAR) or AGN luminosity (typically in the 2–10 keV band; hereafter  $L_X$ ) on the one hand, and star formation rates (SFRs) or IR luminosity (in the 8–1000  $\mu\text{m}$ ; hereafter  $L_{\text{IR}}$ ) on the other hand, taking advantage of the wealth of multiwavelength information collected in deep extragalactic surveys. However, different, somewhat contradictory results have been reported in recent years.

Several studies have found a flat distribution computing average  $L_{\text{IR}}$  in bins of  $L_X$  of X-ray selected sources (or SFR and BHAR, respectively, e.g., Shao et al. 2010; Rovilos et al. 2012; Rosario et al. 2012) at low redshift and luminosities. A significant positive correlation instead appears for luminous AGN ( $L_X > 10^{43-44}$  erg s $^{-1}$ ) and high redshifts ( $z > 1-2$ ), suggesting two different triggering mechanisms at high and low luminosities, via merger and secular evolution, respectively.

Other groups have found a linear correlation at all  $z$  and  $L_X$  when computing the average  $L_X$  in bins of  $L_{\text{IR}}$  (in log-log space) of IR selected sources (Chen et al. 2013; Delvecchio et al. 2015; Dai et al. 2015), with the ratio  $\text{Log}(\text{SFR}/\text{BHAR}) \sim 3$ , roughly consistent with the local  $M_{\text{bulge}}/\text{SMBH}$  value (Magorrian et al. 1998; Marconi & Hunt 2003). Finally, other authors have found no correlation at all, regardless of the  $L_X$  and  $z$  range (e.g., Mullaney et al. 2012; Stanley et al. 2015).

Looking at AGN obscuration, Rovilos et al. (2012) explored for the first time the possible relation between AGN column density ( $N_{\text{H}}$ ), as measured from the X-ray spectra, and host properties, and found no correlation on a sample of 65 sources in the XMM-CDFS survey (Comastri et al. 2011). Rosario et al. (2012) found similar result from hardness ratios (HR) on a larger sample in COSMOS, while Rodighiero et al. (2015) found a positive correlation between  $N_{\text{H}}$  and  $M_*$ , again based on HR, on a sample of  $z \sim 2$  AGN hosts in the same field.

From a technical point of view, these differences may partly arise from different biases and analysis methods. For example, given that only a small fraction of the X-ray detected sources are far-IR (FIR) detected, and vice versa, most of these studies rely on X-ray or FIR stacking in order to recover the average properties of large samples of AGN/host systems, or are limited to small subsamples. Mullaney et al. (2015) pointed out that modeling the SFR distribution of X-ray selected hosts as a log-normal distribution, and including upper limits, gives different results than computing the linear mean of the distribution (i.e., via stacking), which is instead driven upwards by the bright outliers.

Another issue was raised by Symeonidis et al. (2016), who showed that the intrinsic AGN SED in the FIR is cooler than usually assumed. Therefore, in some cases there is no “safe” photometric band which can be used to calculate the SFR without subtracting the AGN contribution. On the other hand, several of the works cited above take the FIR photometry directly (typically at 60  $\mu\text{m}$ ) in order to estimate SFR, thus potentially introducing a spurious correlation at high AGN luminosities.

Recently, from theoretical studies, a physical mechanism has been proposed to explain part of these contradictory results. Volonteri et al. (2015a,b) explain that these different observations are due to the way we analyze the data: the bivariate distribution of AGN and SF luminosities gives two very different results depending on the binning axis used. Hickox et al. (2014) reached similar conclusions starting from the simple assumptions that long-term AGN activity and SFR are perfectly

correlated, that the observed SFR is the average over  $\sim 100$  Myr, while the AGN activity, traced by X-ray emission, varies on much shorter time scales. In these models the different time scales involved in AGN and SF variability dilute the linear dependency between the two quantities if the rapidly variable AGN luminosity is used to build the subsamples to be studied “on average”. This result was also confirmed observationally by Dai et al. (2015) using shallow data from XMM-LSS.

Furthermore, Volonteri et al. (2015a) suggest that spatial scales are important; the BH accretion rate should be correlated with the nuclear ( $< 100$  pc) SFR, while it is less correlated with the total ( $< 5$  kpc) SFR, except for the most intense merger episodes, which are able to affect the whole host galaxy. Of course, the SFR that can be inferred from the FIR luminosity is the global, galaxy-scale SFR (with the exception of the local Universe; see, e.g., Diamond-Stanic & Rieke 2012), and this introduces another source of uncertainty in the observational comparison between BHAR and SFR.

Here we explore the possible correlations between AGN and host properties for a large sample of X-ray and FIR detected sources thanks to the extensive *Chandra*, *XMM-Newton*, and *Herschel* coverage on the COSMOS field (Scoville et al. 2007; Hasinger et al. 2007; Elvis et al. 2009; Lutz et al. 2011; Oliver et al. 2012). This approach avoids the uncertainties related to the stacking, and allows for a proper SED deconvolution, source by source. This of course limits the significance of our findings to the brightest, most accreting, and most star forming systems. These systems are, however, the most interesting ones, the ones for which there is less agreement in the literature on the presence of a correlation between AGN and SF, and also the ones for which theoretical models predict that the correlation should be stronger.

The paper is organized as follows: Sect. 2 describes the sample and source properties; Sect. 3 presents the analysis of  $L_X$  and  $L_{\text{IR}}$  distributions; in Sect. 4 we compare our results with a set of hydrodynamical simulations; in Sect. 5 we discuss correlations between nuclear obscuration and host properties and in Sect. 6 we discuss our results. Throughout the paper we assume a standard  $\Lambda$ CDM cosmology with  $H_0 = 70$  km s $^{-1}$  Mpc $^{-1}$ ,  $\Omega_{\Lambda} = 0.73$ , and  $\Omega_{\text{M}} = 0.27$  (Bennett et al. 2013).

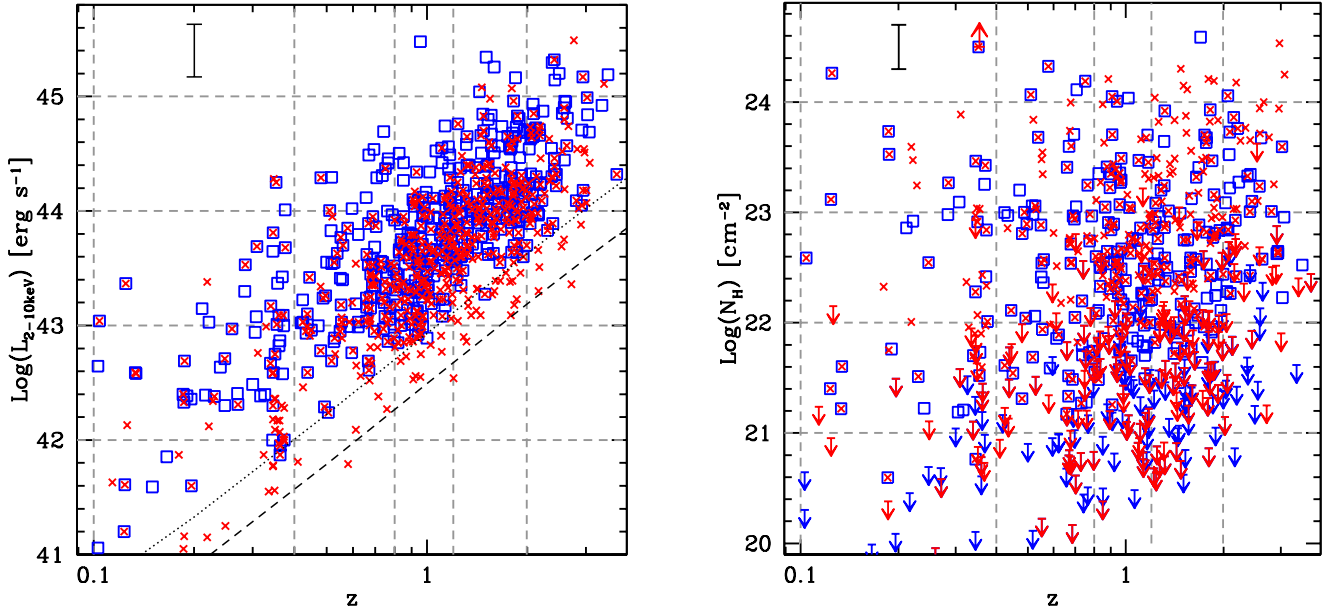
## 2. Sample

We performed X-ray spectral fitting for all the *Chandra* and *XMM-Newton* detected sources (from the catalogs of Brusa et al. 2010; and Civano et al. 2015, respectively) with more than 30 counts, in Marchesi et al. (2016) and Lanzuisi et al. (2013, 2015), respectively. This sample consist of 2333 individual sources (1949 *Chandra* and 1187 *XMM-Newton* sources, with 803 source in common<sup>1</sup>).

For all the *Herschel* detected sources in the COSMOS field ( $\sim 17\,000$  with at least a detection at  $> 3\sigma$  in one of the FIR bands, from 100 to 500  $\mu\text{m}$ ; Lutz et al. 2011), an SED deconvolution with three components – stellar emission, AGN torus emission, and SF-heated dust emission – performed using photometric points from the UV to sub-mm, is available from Delvecchio et al. (2014, 2015; hereafter D15), following the recipe described in Berta et al. (2013).

We then selected all the *XMM-Newton* and *Chandra* detected sources that have at least one FIR detection (and therefore SED deconvolution). The final sample comprises 692 X-ray and

<sup>1</sup> The *Chandra* data from Marchesi et al. (2016) are used for the sources in common, given the deeper coverage.



**Fig. 1.** *Left:* distribution of rest frame, absorption-corrected 2–10 keV luminosity vs. redshift for the *XMM–Newton* (blue squares) and *Chandra* (red crosses) detected sources that are also *Herschel* detected (X-FIR sample). The dotted (dashed) line marks the sensitivity limit of the *XMM–Newton* (*Chandra*) surveys. The redshift bins adopted in the text are marked by vertical gray dashed lines. *Right:* distribution of intrinsic column density vs. redshift for the sample. Symbols as in left panel. Arrows show upper limits for unobscured sources. The average  $1\sigma$  error bars on  $L_X$  and  $N_H$  are shown at the top left of both panels.

FIR detected sources (hereafter the X-FIR sample), 459 spectroscopic and 233 photometric, all of them with an available redshift (Civano et al. 2012; Brusa et al. 2010; Salvato et al. 2009; Marchesi et al. 2015). To date this is the largest sample of AGN/host systems for which X-ray spectral parameters, such as column density and absorption-corrected 2–10 keV luminosity, are known in combination with host properties such as  $M_*$  and SFR.

## 2.1. AGN properties

Figure 1, (left) shows the distribution of  $L_X$  vs. redshift for the X-FIR sample. The average  $1\sigma$  error bar on  $L_X$  is shown in the upper left corner. The absorption-corrected  $L_X$  is affected by uncertainties related to both the number of net counts (observed flux uncertainties) and the spectral shape of each source (uncertainties on  $N_H$  and spectral slope). Therefore, the errors have been derived, for each source using the equivalent in Sherpa (Fruscione et al. 2006) of the cflux model component in Xspec (Arnaud 1996), applied to the best-fit unabsorbed power law. The flux and errors are then computed in the observed band corresponding to 2–10 keV rest frame, and converted into luminosity.

The redshift bins that will be used in the following analysis are shown with vertical dashed lines. The intervals have been chosen with the aim of having a fairly large number of sources in each bin ( $\sim 80$ – $160$ ) with a reasonably narrow redshift interval. The  $L_X$  bins that will be used in the following (1 bin per dex) are shown as horizontal dashed lines.

Figure 1, (right) shows the column density distribution for the X-FIR sample. Arrows show sources for which the obscuration is constrained only by an upper limit. The average  $1\sigma$  error bar on  $N_H$  is shown in the upper left corner. The distribution of  $N_H$  from X-ray spectral analysis has a clear upper boundary

around Compton Thick (hereafter CT) column densities<sup>2</sup> due to the strong flux decrement associated with CT obscuration in the 2–10 keV band. Also, the minimum measurable  $N_H$  increases with redshift as the low energy cutoff due to obscuration move outside the observing band.

The global fraction of X-ray obscured sources (those with  $N_H > 10^{22}$  cm<sup>-2</sup>) in the X-FIR sample is  $\sim 50\%$ , higher than the typical obscured fraction (30–40%) of the X-ray samples in the *Chandra*- and *XMM–Newton*-COSMOS (Lanzuisi et al. 2015; Marchesi et al. 2016). Indeed, the FIR luminosity (and therefore *Herschel* detection rate) of type-2 AGN seems to be higher than for type-1 QSO (Chen et al. 2015).

## 2.2. Host properties

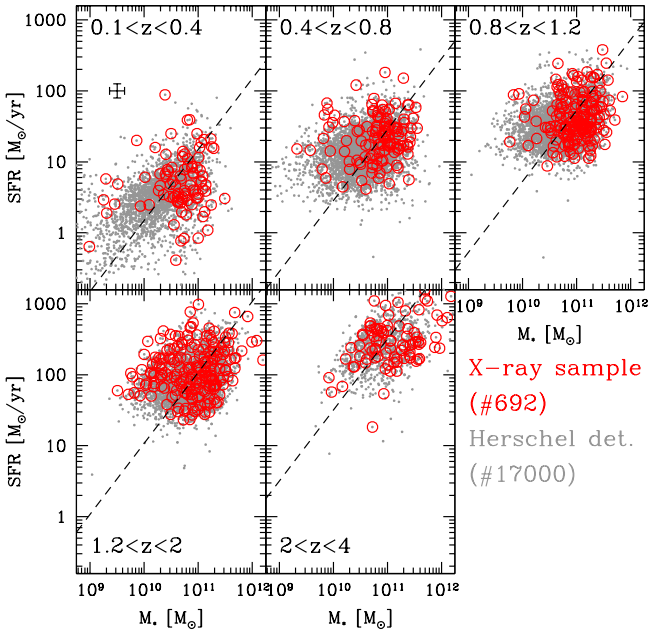
The host properties (SFR vs.  $M_*$ ) of the 692 sources in the X-FIR sample are shown in Fig. 2 (red circles) divided into five redshift bins as described above. The values are taken from D15: the SFR has been derived by converting the IR luminosity (rest 8–1000  $\mu\text{m}$ ) of the best-fitting galaxy SED (i.e., subtracting the AGN emission when present) with the SF law of Kennicutt (1998), scaled to a Chabrier (2003) initial mass function (IMF). The  $M_*$  is derived from the SED decomposition itself, and based on Bruzual & Charlot (2003) models, with a consistent IMF. Table 1 (full version available at the CDS) summarizes the multi-wavelength properties of the sources in the X-FIR sample.

The host properties of the sample of *Herschel* detected sources (from D15,  $\sim 17$  000 sources) are shown for comparison with gray dots. The average of the statistical  $1\sigma$  error bars resulting from the SED fit<sup>3</sup> are shown in the top left corner.

<sup>2</sup> Lanzuisi et al. (2015a,b) present the CT sources detected by *XMM–Newton*, while Lanzuisi et al. (in prep.) will present the ones detected by *Chandra*.

<sup>3</sup> Systematic errors like uncertainties related to the adopted IMF or SF law, are not included in the error budget.





**Fig. 2.** SFR vs.  $M_*$  distribution for the entire sample of *Herschel* detected sources ( $\sim 17\,000$  sources, gray points) and for the 692 sources with X-ray spectral analysis (X-FIR sample, red circles), divided into the five redshift bins defined in Sect. 2.1. The dashed lines in each panel mark the redshift-dependent MS of Withaker et al. (2012). The average  $1\sigma$  error bars are shown in the top left as a black cross.

The errors on  $M_*$  follow a log-normal distribution, with average  $\langle \text{err}(M_*) \rangle = 0.14$  dex and standard deviation  $\sigma = 0.09$ . The mean error on SFR is  $\langle \text{err}(\text{SFR}) \rangle = 0.10$  dex, and standard deviation of  $\sigma = 0.07$  as for  $L_{\text{IR}}^{\text{SF}}$  (see Sect. 3.1), since the SFR is derived from  $L_{\text{IR}}^{\text{SF}}$  adopting a Kennicutt (1998) law. The redshift-dependent MS of star forming galaxies, as described in Whitaker et al. (2012), is also shown in each panel. The FIR selected sources broadly follow the MS relation. However, the *Herschel*-based selection is sensitive to the most star forming systems, introducing a cut in SFR that moves towards higher values with increasing redshift (e.g., Rodighiero et al. 2011, D15).

X-ray detected AGN are preferentially found at the highest  $M_*$ , i.e., the fraction of X-ray detected sources increases as a function of  $M_*$ , in the first three redshift bins at least. This is a well-known effect (Kauffmann et al. 2003; Bundy et al. 2008; Brusa et al. 2009; Silverman et al. 2009; Mainieri et al. 2011; Santini et al. 2012; Delvecchio et al. 2014). Aird et al. (2012) suggested that it is the result of an observational bias, such that more massive galaxies (i.e., more massive BHs) can be detected at a given X-ray flux limit with a variety of accretion rates, while lower mass systems can be detected only if they have a high accretion rate. This, combined with a steep Eddington ratio distribution (i.e., sources with low Eddington ratio are much more common than sources with high Eddington ratio) can explain the observed  $M_*$  distribution (see also Bongiorno et al. 2012).

In our case there is a threshold at around  $\log M_* \sim 10.5 M_\odot$  in the first 3 redshift bins. A simple calculation shows that this value can be roughly derived from the X-ray flux limit of the *Chandra* and *XMM-Newton* surveys using standard values for bolometric corrections ( $k_{\text{Bol}} = 10\text{--}30$ ), Eddington ratios ( $\lambda_{\text{Edd}} \sim 0.05$ ), and BH-host mass ratios ( $M_*/M_{\text{BH}} = 1000\text{--}3000$ ). A more detailed study of the Eddington ratio distribution that can be derived from the  $M_*$  and  $L_X$  distributions will be presented in Suh et al. (2017).

Several studies in the local Universe suggest that the fraction of galaxies hosting an AGN also increases with IR luminosity (e.g., Lutz et al. 1998; Imanishi et al. 2010; Alonso-Herrero et al. 2012; Pozzi et al. 2012). We also tested that the observed threshold in mass is not driven by our requirement of *Herschel* detection; when using the  $M_*$ -SFR distribution of Bongiorno et al. (2012), computed for the full XMM-COSMOS catalog, a drop in the number of X-ray detected AGN below  $\log M_* = 10.2\text{--}10.4 M_\odot$  is visible up to  $z = 2.5$ .

The consequence of this selection effect is that the X-FIR sample has a  $M_*$  distribution shifted toward higher  $M_*$  with respect to the global *Herschel* sample (Fig. 3, top right); instead, the distribution of SFR for the X-FIR sample is roughly consistent with that of the global *Herschel* sample (Fig. 3, top left). This has important implications when measuring sSFR and MS offsets, for example (Fig. 3, bottom left and right): due to this selection effect the X-FIR sample has lower sSFR with respect to the MS of star forming galaxies (or to the *Herschel* sample) if the two samples are not properly mass-matched (Silverman et al. 2009; Xue et al. 2010).

### 3. $L_X$ vs. $L_{\text{IR}}$ distributions

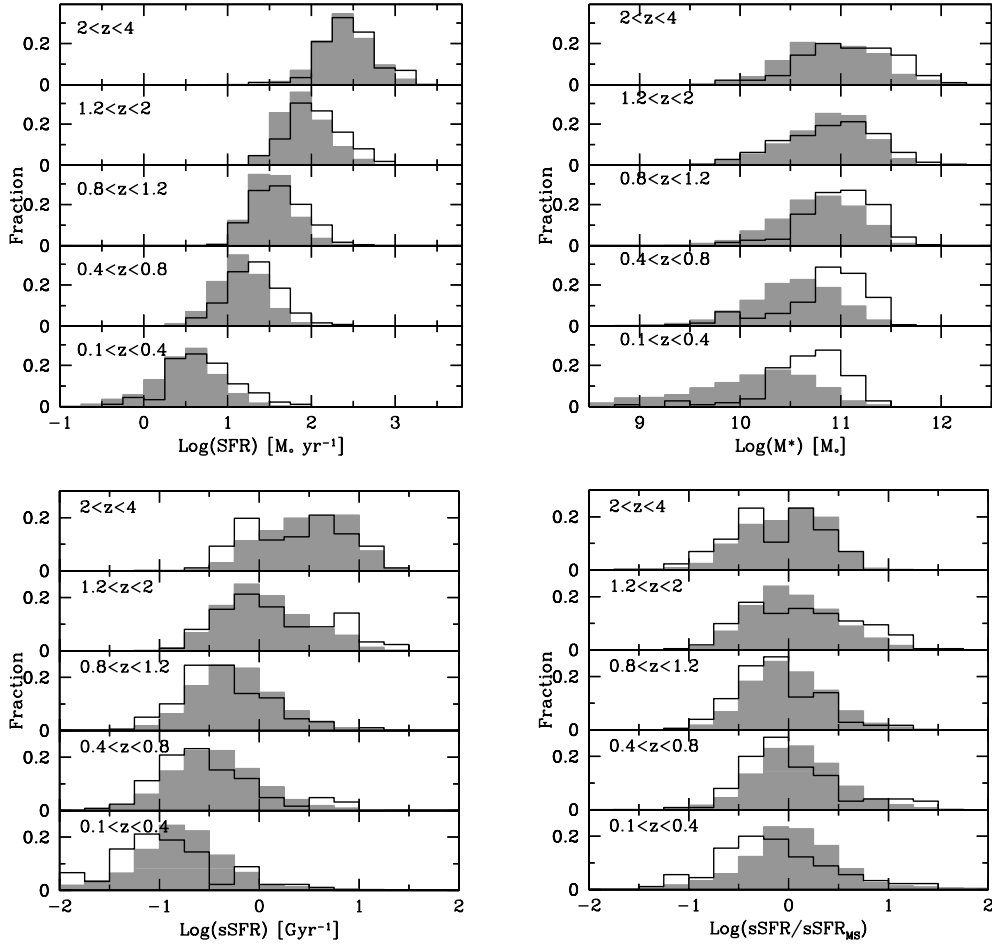
#### 3.1. Partial correlation analysis

The two quantities that have been used most often to look for BHAR-SFR correlations are the AGN luminosity, often represented by  $L_X$ , and the SF luminosity in the form of  $L_{\text{IR}}$  (or  $L_{60\ \mu\text{m}}$ ; Santini et al. 2012; Rosario et al. 2012; Chen et al. 2013). It is generally assumed that the total FIR luminosity is not significantly affected by any contamination from the AGN emission. However, recent studies have shown that AGN may contribute significantly to the IR emission and in some cases even in the FIR band (Symeonidis et al. 2016). Therefore, the SFR derived directly from FIR photometry can be overestimated, especially in high luminosity AGN hosts. Thanks to the SED decomposition available, we use in the following the  $L_{\text{IR}}$  computed only for the SF component (hereafter  $L_{\text{IR}}^{\text{SF}}$ ), after subtracting the AGN contribution, modeled with the SED templates of Fritz et al. (2006; see also Feltre et al. 2012). This avoids introducing a spurious correlation between AGN and SF luminosity, especially at the highest luminosities.

Clearly two luminosities are always correlated in any sample that is flux limited in both directions, due to the combination of the luminosity-distance effect and tendency of the sources to cluster at the flux limit (Malmquist bias, e.g., Feigelson & Berg 1983). Figure 4 shows the distribution of  $L_X$  vs.  $L_{\text{IR}}$  for the X-FIR sample.

The  $1\sigma$  errors on  $L_{\text{IR}}^{\text{SF}}$  follow a log-normal distribution with average value  $\langle \text{err}(L_{\text{IR}}^{\text{SF}}) \rangle = 0.10$  dex, and standard deviation of  $\sigma = 0.07$ . As mentioned in Sect. 2.1, the errors on the absorption-corrected luminosity follow a much broader distribution depending on the number of counts available and on the spectral shape. They range from  $\lesssim 0.1\text{--}0.2$  dex for bright, unobscured sources, to  $\sim 0.5\text{--}1.0$  dex for faint and highly obscured sources. The average value of the  $1\sigma$  error is  $\langle \text{err}(L_X) \rangle \sim 0.23$  dex, with standard deviation  $\sigma = 0.18$ . We show the average errors with a black cross in the left panel of Fig. 4, while the specific value for each source is used in the following analysis.

In order to look for intrinsic correlations between these two quantities, one possibility is to compute the partial Spearman rank correlation between two variables in the presence of a third, and to assess the statistical significance of such correlation (e.g., Macklin 1982). To derive the correlation coefficient between  $L_X$



**Fig. 3.** Fractional distribution of the host properties, SFR,  $M_*$ , sSFR, and MS offset, from top left to bottom right, for the X-FIR sample (black open histogram) and for the whole *Herschel* sample (gray filled histogram), in redshift bins.

**Table 1.** Multiwavelength properties of the 692 sources in the X-FIR sample.

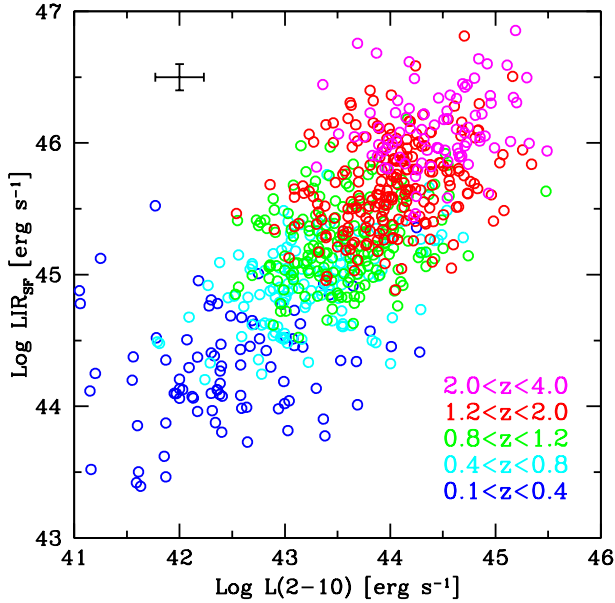
ID	RA	Dec	$z$	$\text{Log}(L_{\text{IR}}^{\text{SF}})$	$\text{Log}(M_*)$	SFR	$\text{Log}(N_{\text{H}})$	$\text{Log}(L_{\text{X}})$	$\text{Log}(L_{\text{Bol}})$	XID	CID
(1)	deg	deg	(4)	erg/s	$M_{\odot}$	$M_{\odot}/\text{yr}$	$\text{cm}^{-2}$	erg/s	erg/s	(11)	(12)
1846545	150.500	2.862	0.102s	$44.77 \pm 0.07$	$10.20 \pm 0.13$	15.8	$<20.77$	$41.05 \pm 0.27$	41.96	60095	lid2100
1883498	150.065	2.929	0.102s	$43.69 \pm 0.16$	$10.55 \pm 0.09$	1.3	$<20.42$	$42.64 \pm 0.06$	43.8	5617	lid385
89570	150.372	1.609	0.104s	$44.03 \pm 0.08$	$10.80 \pm 0.09$	2.8	$22.58^{+0.03}_{-0.03}$	$43.04 \pm 0.04$	44.29	2021	cid1678
1612003	150.550	2.628	0.113s	$43.37 \pm 0.28$	$10.58 \pm 0.09$	0.4	$<21.36$	$41.63 \pm 0.36$	42.6	–	lid3189
1197519	150.335	2.304	0.123s	$44.24 \pm 0.06$	$10.58 \pm 0.09$	4.7	$21.40^{+0.62}_{-0.29}$	$41.2 \pm 0.29$	42.11	1533	cid967
785579	150.162	1.991	0.124s	$44.06 \pm 0.06$	$10.96 \pm 0.15$	3.1	$<21.07$	$40.41 \pm 0.21$	41.26	–	cid2976
1472936	149.938	2.577	0.124s	$43.5 \pm 0.02$	$10.99 \pm 0.11$	0.8	$23.11^{+0.15}_{-0.16}$	$41.61 \pm 0.31$	42.58	54517	cid1272
725087	150.424	2.066	0.125s	$43.88 \pm 0.06$	$11.08 \pm 0.03$	1.6	$24.26^{+0.05}_{-0.04}$	$43.36 \pm 0.35$	44.71	2608	cid482
...											

**Notes.** Catalog entries are as follows: (1) Source ID from Capak et al. (2007); (2) and (3) right ascension and declination of the optical/IR counterpart; (4) redshift ( $s$  for spectroscopic or  $p$  photometric); (5)  $\text{Log}(L_{\text{IR}}^{\text{SF}})$  with  $1\sigma$  errors; (6)  $\text{Log}(M_*)$  with  $1\sigma$  errors; (7) SFR derived from  $L_{\text{IR}}^{\text{SF}}$ ; (8)  $\text{Log}(N_{\text{H}})$  with  $1\sigma$  errors or upper limits; (9)  $\text{Log}(L_{\text{X}})$  with  $1\sigma$  errors; (10)  $\text{Log}(L_{\text{Bol}})$  computed from  $L_{\text{X}}$  using Marconi et al. (2004); (11) and (12) XMM-COSMOS and *Chandra*-COSMOS IDs (from Brusa et al. 2010; and Marchesi et al. 2016, respectively). The full table is available at the CDS.

and  $L_{\text{IR}}^{\text{SF}}$ , conditioned by the distance,  $\rho(L_{\text{X}}, L_{\text{IR}}^{\text{SF}}, z)$ , we evaluate the Spearman coefficient  $\rho$  related to each pair of parameters and then combine them according to the expression

$$\rho(a, b, c) = \frac{\rho_{ab} - \rho_{ca}\rho_{bc}}{\sqrt{(1 - \rho_{ca}^2)(1 - \rho_{bc}^2)}} \quad (1)$$

(Conover 1980), which returns the partial correlation between  $a$  and  $b$ , corrected for the dependency on  $c$ . The resulting  $\rho$  is 0.15, and the associated confidence level, in terms of standard deviations, that the first two variables are correlated, independently of the influence of the third, is  $\sim 3.7\sigma$ , following Eq. (6) of Macklin (1982). Therefore, the two quantities appear to be significantly



**Fig. 4.**  $L_X$  vs.  $L_{\text{IR}}^{\text{SF}}$  for the X-FIR sample. Different colors represent different redshift bins: blue for  $0.1 < z < 0.4$ , cyan for  $0.4 < z < 0.8$ , green for  $0.8 < z < 1.2$ , red for  $1.2 < z < 2.0$ , and magenta for  $2 < z < 4.0$ . The average  $1\sigma$  errors on  $L_X$  and  $L_{\text{IR}}^{\text{SF}}$  are shown in the upper left corner.

correlated, after the effect of redshift on both of them is taken into account.

### 3.2. Redshift bins

The second approach, often used in the literature, is to define redshift bins that are as narrow as possible to minimize the distance effect, and to look for correlations between the two quantities. Thanks to the large sample collected in this work, we can divide the sample into five redshift bins. For every redshift bin, a large distribution in both luminosities can be observed, with the typical luminosity increasing with redshift (Fig. 4).

Most of the observational works mentioned in Sect. 1 looked for the distribution of average  $L_{\text{IR}}^{\text{SF}}$  in  $L_X$  bins, or average  $L_X$  in  $L_{\text{IR}}^{\text{SF}}$  bins (but see Gruppioni et al. 2016). Both hydrodynamical simulations (e.g., Volonteri et al. 2015a,b) and semi-analytical models (e.g., Neistein & Netzer 2014) show that, in the  $L_{\text{IR}}^{\text{SF}}-L_X$  plane, there may be the superimposition of a weak correlation for the bulk of the population, and a strong correlation only for the most extreme merger phases, corresponding to the highest  $L_X$  and  $L_{\text{IR}}^{\text{SF}}$ . If the underlying distribution shows such a complex shape, the results of the two approaches (average  $L_{\text{IR}}^{\text{SF}}$  in  $L_X$  bins or average  $L_X$  in  $L_{\text{IR}}^{\text{SF}}$  bins) may be very different.

In Fig. 5 (left) we show the result of plotting average  $L_{\text{IR}}^{\text{SF}}$  in bin of  $L_X$  (both in log scale) in five redshift bins. As can be seen, there is no correlation at all  $L_{\text{IR}}^{\text{SF}}$  and, as expected, there are no sources below the relation computed for a pure AGN template in Mullaney et al. (2011), similar to the one used in Netzer et al. (2009). Following this approach, we are therefore able to reproduce the results of Shao et al. (2010), Rosario et al. (2012), and others that claim no correlation between AGN activity and SF over several orders of magnitude in luminosity.

On the other hand, computing average  $L_X$  in  $L_{\text{IR}}^{\text{SF}}$  bins (in log scale) from the same bivariate distribution gives different results. Figure 5 (right) shows that at all redshifts the average  $L_X$  correlates with the  $L_{\text{IR}}^{\text{SF}}$  and the binned points are close to

the SFR/BHAR  $\sim 500$  ratio found in Chen et al. (2013; hereafter C13).

In both panels, we computed the error on the average  $L_X$  and  $L_{\text{IR}}^{\text{SF}}$  through a bootstrap re-sampling procedure, as done in several previous works. For each bin with  $N$  sources, we randomly extract  $N$  sources, allowing repetitions, and computed the mean value. The process is iterated  $10^4$  times, and the standard deviation of the mean is taken as the error on the average SFR.

The two approaches described above are the equivalent of computing the forward and inverse linear regression of one variable over the other. Table 2 reports the slopes  $\alpha$  and intercept  $\beta$ , and their associated errors, for each redshift bin in the log-log space of the least-squares (LS) fit<sup>4</sup> of  $L_{\text{IR}}^{\text{SF}}$  as a function of  $L_X$  (hereafter  $L_{\text{IR}}^{\text{SF}} | L_X$ ), and  $L_X$  as a function of  $L_{\text{IR}}^{\text{SF}}$  (hereafter  $L_X | L_{\text{IR}}^{\text{SF}}$ ), respectively<sup>5</sup>. Indeed, the slopes in the left panel are all consistent with 0 within  $\sim 2\sigma$  confidence level. On the other hand, LS fits of  $(L_X | L_{\text{IR}}^{\text{SF}})$  give steeper correlations at all  $z$  bins and slopes not consistent with 0 at  $\sim 3\sigma$  confidence level.

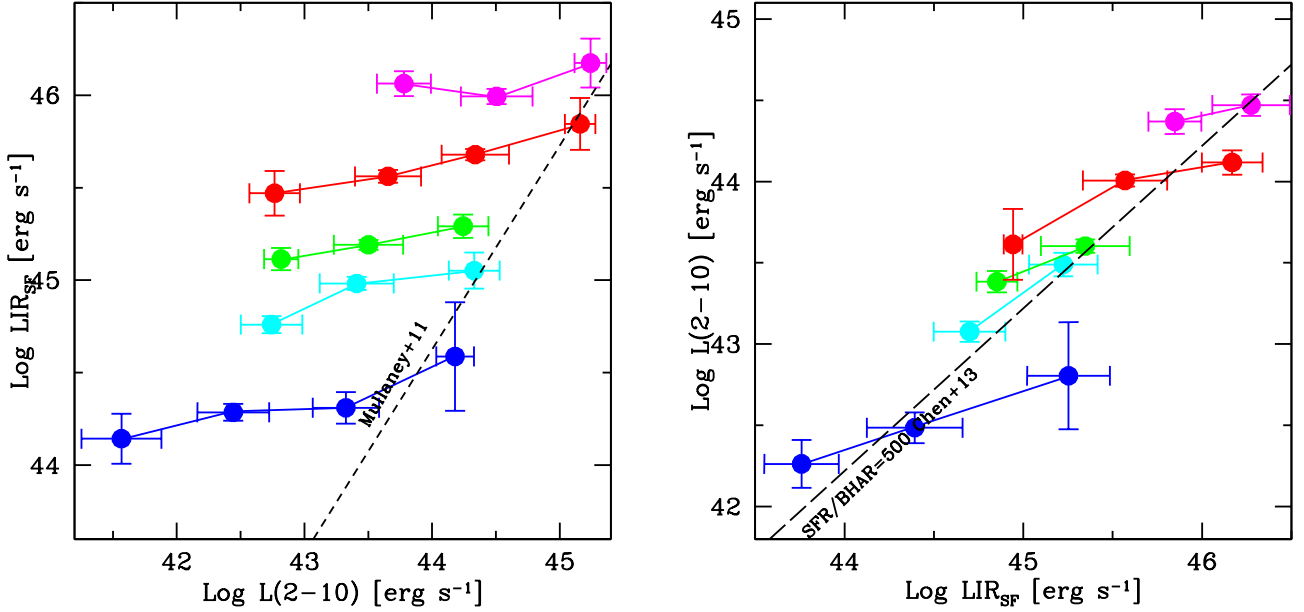
The SFR/BHAR  $\sim 500$  ratio plotted in Fig. 5 is the value found in C13 for a sample of 121 FIR selected AGN-hosts at  $0.25 < z < 0.8$ . For a comparison with their results we should look at our first two  $z$ -bins: While the  $z$ -bin  $0.1 \leq z < 0.4$  has a very flat ( $L_X | L_{\text{IR}}^{\text{SF}}$ ) slope, possibly due to the small volume sampled, the  $0.4 \leq z < 0.8$  interval shows a correlation with slope consistent with 1 at  $\sim 1\sigma$ , therefore in broad agreement with the C13 findings. Interestingly, we can extend up to  $0.8 \leq z < 1.2$  the redshift range for which a correlation roughly consistent with SFR/BHAR  $\sim 500$  can be found. Above this redshift interval, the slopes become flatter. Therefore, we found a strong (almost linear) correlation between  $\log L_X$  and  $\log L_{\text{IR}}^{\text{SF}}$  for  $(L_X | L_{\text{IR}}^{\text{SF}})$  at redshifts lower than the peak of the SF and AGN activity, i.e., between 4 and 8 Gyr ago, while at higher redshift the correlation is still present but weaker.

The exact value of the ratio SFR/BHAR in terms of  $L_X$  and  $L_{\text{IR}}^{\text{SF}}$  depends strongly on the assumptions made to scale between these quantities, i.e., the accretion efficiency and bolometric correction in the first case, and the SF law and initial mass function (IMF) in the second. Chen et al. (2013) derived the SFR from  $L_{\text{IR}}^{\text{SF}}$  using the Kennicutt (1998) relation modified for a Chabrier IMF (Chabrier 2003), and the BHAR from  $L_X$  using a constant  $k_{\text{bol}} = 22.4$  and accretion efficiency of 0.1. They use as reference the value of SFR/BHAR  $\sim 500$  derived from the  $M_{\text{Bulge}}/M_{\text{BH}}$  ratio observed in Marconi et al. (2004). The authors suggest it is a coincidence that the detected sources sit on the SFR/BHAR  $\sim 500$  ratio, due to the ratio between the X-ray and FIR flux limits in the Boötes field.

In the X-FIR sample in COSMOS we have a factor of  $\sim 10$  deeper X-ray data (taking into account the flux limit corresponding to our spectral analysis requirements), while the *Herschel* data are only a factor of 2–3 deeper ( $\sim 8$  mJy at  $110 \mu\text{m}$  and 8 mJy at  $250 \mu\text{m}$ ) than in Boötes. Nonetheless, our X-FIR detected sample sit close to the C13 relation. We note that in both cases the X-ray and FIR detected sources are a small minority of both the original X-ray and FIR samples (a small percent, up to  $\sim 20\%$ ), and the flux limit has an important role in the observed properties of the detected sources alone, as also discussed in C13.

<sup>4</sup> The LS fit is performed with the BCES code (Akritas & Bershady 1996), adopting  $10^4$  bootstrap re-samplings. Similar results are obtained using the LINMIX code (Kelly et al. 2007).

<sup>5</sup> In the first case slopes and intercepts refer to a relation in the form  $\text{Log } L_{\text{IR}}^{\text{SF}} = 45 + \alpha \times (\text{Log } L_X - 44) + \beta$ , while in the second in the form  $\text{Log } L_X = 44 + \alpha \times (\text{Log } L_{\text{IR}}^{\text{SF}} - 45) + \beta$ .



**Fig. 5.** *Left:* average  $\text{Log}(L_{\text{IR}}^{\text{SF}})$  in bins of  $\text{Log}(L_X)$  in five redshift bins. The short dashed line is the correlation derived in Mullaney et al. (2011) for a pure AGN SED. *Right:* average  $\text{Log}(L_X)$  in bins of  $\text{Log}(L_{\text{IR}}^{\text{SF}})$ . The long dashed line represents a constant SFR/BHAR of 500, from C13. In both panels the vertical error bars are computed through a bootstrap resampling procedure, while the horizontal error bars show the  $1\sigma$  dispersion of that bin.

**Table 2.** Slopes  $\alpha$  and intercept  $\beta$  of the linear LS fit of  $(L_{\text{IR}}^{\text{SF}} | L_X)$ ,  $(L_X | L_{\text{IR}}^{\text{SF}})$ , and of the bisector estimator in each redshift bin.

z-bin	LS $(L_{\text{IR}}^{\text{SF}}   L_X)$		LS $(L_X   L_{\text{IR}}^{\text{SF}})$		Bisector $(L_X, L_{\text{IR}}^{\text{SF}})$	
	$\alpha$	$\beta$	$\alpha$	$\beta$	$\alpha$	$\beta$
$0.1 \leq z < 0.4$	$0.07 \pm 0.06$	$-0.61 \pm 0.09$	$0.44 \pm 0.13$	$-1.24 \pm 0.21$	$1.28 \pm 0.45$	$-0.65 \pm 0.20$
$0.4 \leq z < 0.8$	$0.20 \pm 0.10$	$0.06 \pm 0.04$	$0.80 \pm 0.17$	$-0.68 \pm 0.05$	$1.21 \pm 0.34$	$-0.67 \pm 0.11$
$0.8 \leq z < 1.2$	$0.12 \pm 0.09$	$0.25 \pm 0.03$	$0.61 \pm 0.15$	$-0.56 \pm 0.04$	$1.30 \pm 0.25$	$-0.75 \pm 0.09$
$1.2 \leq z < 2.0$	$0.16 \pm 0.09$	$0.62 \pm 0.04$	$0.48 \pm 0.12$	$-0.34 \pm 0.08$	$1.29 \pm 0.15$	$-0.86 \pm 0.15$
$2.0 \leq z < 4.0$	$0.01 \pm 0.08$	$1.02 \pm 0.04$	$0.29 \pm 0.18$	$0.10 \pm 0.20$	$1.16 \pm 0.57$	$-0.85 \pm 0.17$

**Notes.** The first set of slopes and intercepts refers to a relation in the form  $\text{Log } L_{\text{IR}}^{\text{SF}} = 45 + \alpha \times (\text{Log } L_X - 44) + \beta$ , while the second and third in the form  $\text{Log } L_X = 44 + \alpha \times (\text{Log } L_{\text{IR}}^{\text{SF}} - 45) + \beta$ .

As discussed in Hickox et al. (2014) and Volonteri et al. (2015), a possible physical explanation for this behavior is that we are averaging a slowly changing quantity, such as the host SFR, of galaxies grouped on the basis of the rapidly changing AGN  $L_X$  (see Fig. 5, left panel). In the right panel, instead, the average  $L_X$  of a large sample of sources grouped on the basis of the slowly changing SFR allows us to recover the underlying, long-term correlation between AGN activity and SFR. In the same way, from a statistical point of view, it may be reasonable to interpret the  $L_X$  as the dependent variable in this context as it has larger uncertainties with respect to  $L_{\text{IR}}^{\text{SF}}$  (Hogg et al. 2010), both in terms of measurement errors (see Sect. 3.1) and noise (i.e., variability).

If we instead assume that in this case there are no “dependent” and “independent” variables (see, e.g., Tremaine et al. 2002; Novak et al. 2006), the two variables may need to be treated symmetrically. We used the BCES code again to derive slope and intercept, and their standard deviation, using a symmetric estimator such as the bisector regression<sup>6</sup> (Isobe et al. 1990). The results are shown in Fig. 6, while the slopes and

intercepts are listed in Table 2. At all redshift bins, the slopes of the linear regression, although always larger than 1, are consistent with 1 within a  $1\sigma$  confidence level.

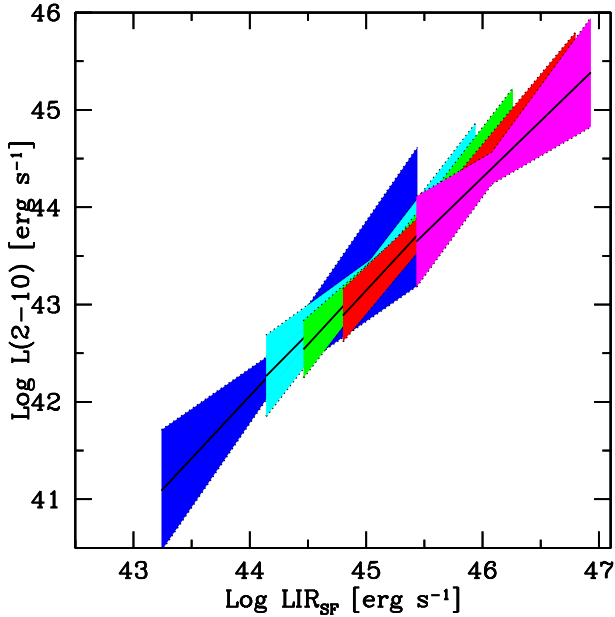
### 3.3. Effect of contamination

Since the *Herschel* PACS and SPIRE point spread functions (Pilbratt et al. 2010) are much larger than those in the optical and NIR bands, going from  $\sim 5''$  to  $\sim 36''$  FWHM (Poglitsch et al. 2010; Griffin et al. 2010), there is the possibility that the FIR flux of our sources is contaminated by unresolved neighbors (see, e.g., Scudder et al. 2016).

We verified the effect of contamination by excluding all the X-FIR sources with a second HST catalog entry from the ACS *F814W* (*I*-band) catalog (28.6 AB limiting magnitude, Scoville et al. 2007; Koekemoer et al. 2007). We choose a circular area of diameter  $8''$  around the optical position. While this distance is not great enough to ensure negligible contamination, it has been chosen in order to retain a sufficient number of sources to allow an analysis in all five redshift bins. The 146 “isolated” sources obtained in this way show the same behavior described above, with a flat distribution of average  $L_{\text{IR}}^{\text{SF}}$  computed in bin of  $L_X$ ,

<sup>6</sup> We recall that the BCES estimators, both the LS and the symmetric, are not immune from biases that arise from data truncation, which is the case for flux-limited samples (see Akritas & Bershady 1996).





**Fig. 6.** Linear regression for  $L_X$  and  $L_{\text{IR}}^{\text{SF}}$  computed for each redshift bin with the bisector estimator in BCES. The color code is the same as in Fig. 4.

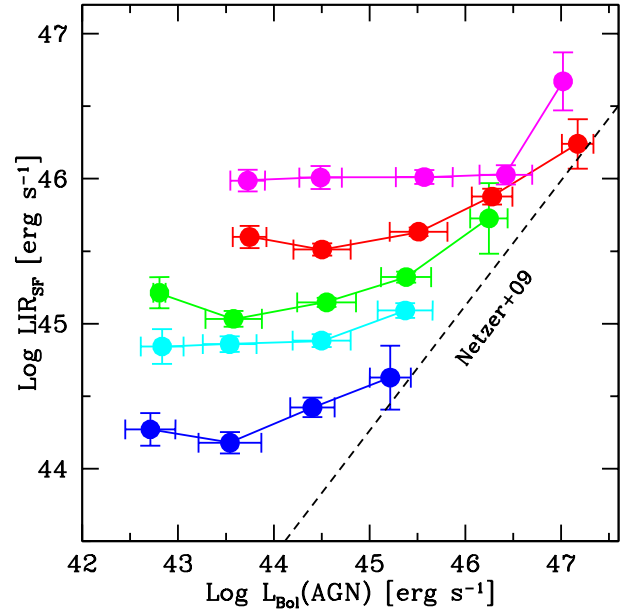
and an almost linear correlation of average  $L_X$  computed in bin of  $L_{\text{IR}}^{\text{SF}}$ .

We also verified that sources with a single PACS or SPIRE detection (more subject to contamination) do not affect our results. Indeed, excluding the 154 (out of 692) sources with only one detection (at  $3\sigma$ ) either in PACS or SPIRE photometry does not change the results presented in Sect. 3.2 and in the following paragraphs.

### 3.4. $L_{\text{Bol}}$ , $M_*$ , sSFR, and MS offset

Several authors have used the AGN bolometric luminosity ( $L_{\text{Bol}}$ ), instead of the  $L_X$ , to look for correlation with the  $L_{\text{IR}}$  or SFR. The  $L_{\text{Bol}}$  is generally derived from the  $L_X$  through a luminosity dependent bolometric correction (e.g., Marconi et al. 2004; Lusso et al. 2012). The net effect of this procedure is to stretch the horizontal axis of Fig. 5 (left) (the high  $L_X$  sources have a higher X-ray bolometric correction than the low  $L_X$  ones), while keeping the  $L_{\text{IR}}^{\text{SF}}$  fixed. In Fig. 7 we show the result of this approach (here we used the Marconi et al. 2004, luminosity dependent bolometric correction, but the Lusso et al. 2012, relation would have the same effect): in each redshift bin the sources populating the highest  $L_X$  bin are now spread in two  $L_{\text{Bol}}$  bins, and the last  $L_{\text{Bol}}$  bin at each redshift is now populated by a smaller number of more extreme sources. The relation found locally for AGN-dominated systems in Netzer et al. (2009) is also shown. Once again, we are able to reproduce results obtained in other works (Shao et al. 2010; Rosario et al. 2012). However, we are now confident that this result is not in disagreement with what is shown in Fig. 5 (right) and the apparent contradiction is only dependent on the way the data are analyzed and grouped, as shown in e.g., Volonteri et al. (2015) and discussed in Stanley et al. (2015) and Dai et al. (2015).

Finally, we found a flat distribution when computing average  $L_X$  in bins of  $M_*$ , sSFR, and MS offset and average  $M_*$ , sSFR, and MS offset in bins of  $L_X$  in all five redshift bins. Indeed, no significant partial correlation is found between any pair of these quantities following the approach described in Sect. 3.1 to take



**Fig. 7.** Average  $\text{Log}(L_{\text{IR}}^{\text{SF}})$  in bin of  $L_{\text{Bol}}$ , for the X-FIR sample. The dashed line is the relation found in Netzer et al. (2009) for AGN-dominated systems. Error bars are computed as in Fig. 5.

into account the redshift effect, which also affects  $M_*$ , SFR, and sSFR ( $\sigma \ll 1$  in all cases). We stress, however, that the range of  $M_*$  covered by our sample is limited to the very high mass end,  $10 < \text{Log}(M_*) < 12$  (to be compared with the underlying galaxy  $M_*$  distribution,  $7 < \text{Log}(M_*) < 12$ , in the same redshift interval shown in e.g., Laigle et al. 2016). Deeper X-ray surveys are needed to investigate the dependency of  $L_X$  with this crucial quantity.

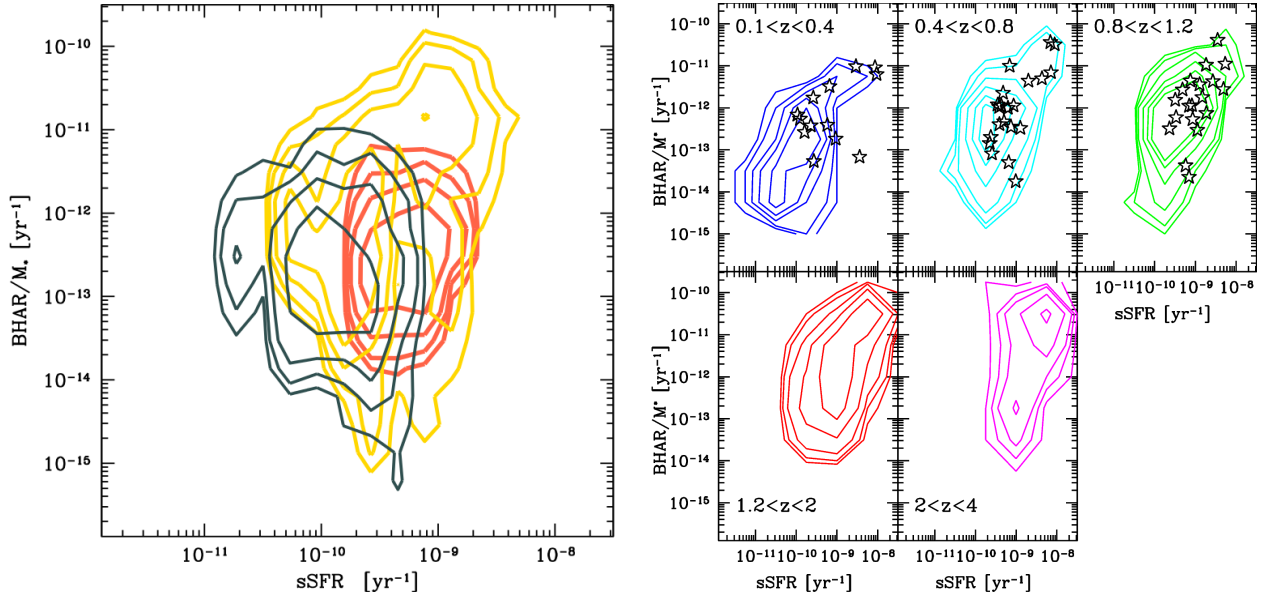
## 4. Theory and observations

### 4.1. Comparison with simulations

Here we compare our results with predictions from the simulations of galaxy mergers presented in Volonteri et al. (2015a). They are based on very high spatial and temporal resolution simulations, covering a wide range of initial mass ratios (1:1 to 1:10), several orbital configurations, and gas fraction (defined as  $M_{\text{gas}}/M_*$ ) in the range  $f_{\text{gas}} = 0.3-0.6$ . The very high resolution imposes a limit on the mass of the simulated galaxies, which typically have  $M_* \sim (2-8) \times 10^9 M_{\odot}$ , i.e., much smaller than the typical mass of our observed galaxies (see Fig. 3). The process is divided into three phases: the stochastic phase, which lasts until the second pericenter, where the galaxies behave as they do in isolation; the merger phase characterized by strong dynamical torques and angular momentum loss; and the remnant phase, which starts when the angular momentum returns to be constant in time. While the stochastic and remnant phases have the same duration (by construction), the merger phase is much shorter (typically 1/10 of the total).

To compare our data with this set of simulated galaxies, we converted the AGN bolometric luminosity into a BH accretion rate (BHAR), by assuming an efficiency of  $\eta = 0.1$  (e.g., Fabian & Iwasawa 1999) and dividing it by the host stellar mass to obtain a specific BHAR (sBHAR) relative to the host mass rather than to the BH mass. We chose to do so because, from an observational point of view, the determination of the  $M_*$  (from SED fitting) is much less uncertain (see Sect. 2.2 for the error





**Fig. 8.** *Left:* BHAR/ $M_*$  rate vs. sSFR contours obtained from the simulations presented in Volonteri et al. (2015a). In red the stochastic phase, in yellow the merger phase, and in black the remnant phase. *Right:* BHAR/ $M_*$  vs. sSFR contours observed in COSMOS in five redshift bins. Sources that are in a major merger state in the first three redshift bins are marked with black stars.

budget in our sample) than that of  $M_{\text{BH}}$ , and is available for both type-1 and type-2 AGN. This value is then compared with the sSFR for each source. The contours of global (within 5 kpc) sSFR vs. sBHAR obtained from the simulations for the three different phases (stochastic, merger, and remnant), are color-coded in Fig. 8 (left) in red, yellow, and black, respectively.

The results from the X-FIR sample are shown in Fig. 8 (right) for the five redshift bins. As can be seen the observed contours in the low redshift bins span a similar range of physical properties, with respect to simulations, with the bulk of the population concentrated between  $5 \times 10^{-11}$  and  $5 \times 10^{-9} \text{ yr}^{-1}$  in sSFR, and between  $10^{-14}$  and  $10^{-11} \text{ yr}^{-1}$  in sBHAR, and with a tail at higher sSFR and sBHAR, possibly produced by sources in the merger phase as in the simulations (yellow contours). Interestingly, the importance of this tail grows with increasing redshift, even if the selection effect in both directions must be taken into account.

We also exploited the deep HST ACS coverage in the COSMOS field to identify sources in the merger phase. We only selected sources that appear to be in a clear major merger phase, and overplotted them in Fig. 8 (right) as black stars in the first three redshift bins (above  $z \sim 1$  it becomes difficult to assess the AGN host morphology). This selection is not meant to be complete: not all the sources are covered by ACS, and it is not possible to recognize the host morphology for all of them, due to bright point-like AGN contribution, for example. However, it is interesting that AGN hosts clearly in merger state tend to cover the highest sSFR and sBHAR range, as predicted by simulations.

#### 4.2. Caveat

One caveat to be considered here is the fact that the simulations are performed at high- $z$ , starting at  $z = 3$  and ending after 1–3 Gyr depending on the merger dynamics (see Capelo et al. 2015, for details). By construction, the simulations have a relatively low gas fraction: 30% of the disk stellar mass. This is probably a low value for SF galaxies at these redshifts. Only one set of simulations has been performed with a higher gas fraction (60%) and, as expected, these simulated galaxies move toward

higher sSFR and sBHAR, as the contours of the observed high redshift sample do.

Another caveat is the fact that the simulations are performed for low mass galaxies. The typical  $M_*$  for these galaxies is in the range  $\text{Log}(M_*) = 9-9.5 (M_\odot)$ , i.e., in the low mass tail of the mass distribution even for the lowest redshift bin of the observed sample. Since the efficiency of SFR and BHAR is most probably mass-dependent, the comparison between different mass ranges may not be straightforward. Volonteri et al. (2015a) argue, however, that SFR and BHAR are self-similar, on the basis of the mass sequence of star forming galaxies and of the possible power-law dependence of the specific BHAR (Aird et al. 2012; Bongiorno et al. 2013, but see Kauffmann & Heckman 2009; Lusso et al. 2012; and Schulze et al. 2015).

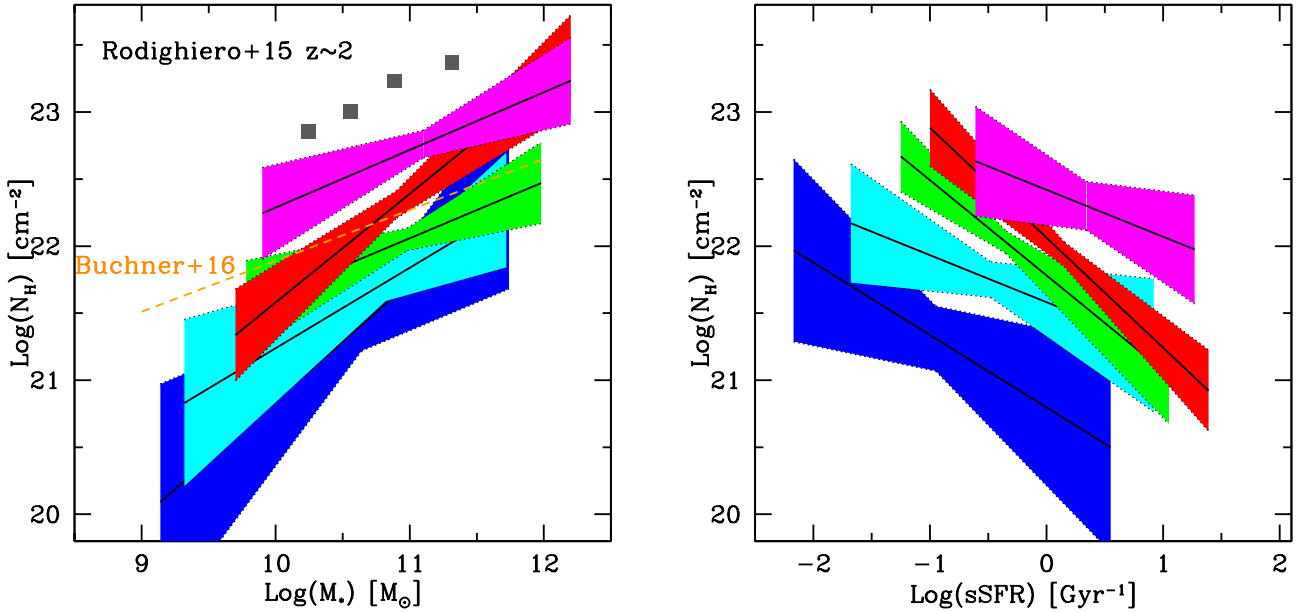
Finally, the simulations are not cosmological, in the sense that the gas mass is not replenished by cosmic inflows and gas accretion, as is the case for real galaxies. This leads to a possible underestimate of SFR and BHAR towards the end of the simulation when galaxies have converted a large fraction of their gas into stellar and BH mass (see also Vito et al. 2014).

## 5. Obscuration

### 5.1. $N_{\text{H}}$ and host properties

Here we discuss the possible correlations between the column density through the AGN line of sight, as measured by the X-ray  $N_{\text{H}}$ , and the host galaxy properties, such as  $M_*$ , SFR, sSFR, and MS offset. The partial correlation analysis described in Sect. 3.1 gives a significant positive correlation (at  $>4\sigma$  confidence level) between  $N_{\text{H}}$  and  $M_*$  in the entire sample once the distance effect is removed (both  $N_{\text{H}}$  and  $M_*$  tend to increase with redshift in two different ways, due to two different selection effects). We also find a significant negative correlation (at  $>5\sigma$  confidence level) between  $N_{\text{H}}$  and sSFR, while we do not find any significant correlation of  $N_{\text{H}}$  with SFR and MS offset.

As in the case of  $L_{\text{X}}$  vs.  $L_{\text{IR}}$ , the binning direction (or the variable chosen as independent) is relevant for the final distribution of  $N_{\text{H}}$  as a function of host properties and vice versa: computing



**Fig. 9.** Linear regression of  $N_{\text{H}}$  vs.  $M_*$  (left) and sSFR (right) in five redshift bins. The regression is performed using the `linmix` code, which also takes into account the  $N_{\text{H}}$  upper limits. The color-coding is the same as in Fig. 4. The gray squares in the left panel show results from Rodighiero et al. (2015) at  $z \sim 2$ , obtained from the HR of X-ray stacked images of FIR detected galaxies in the COSMOS field. The orange dashed line is the relation found in Buchner et al. (2017) for a sample of GRB hosts in a wide range of redshifts (see text).

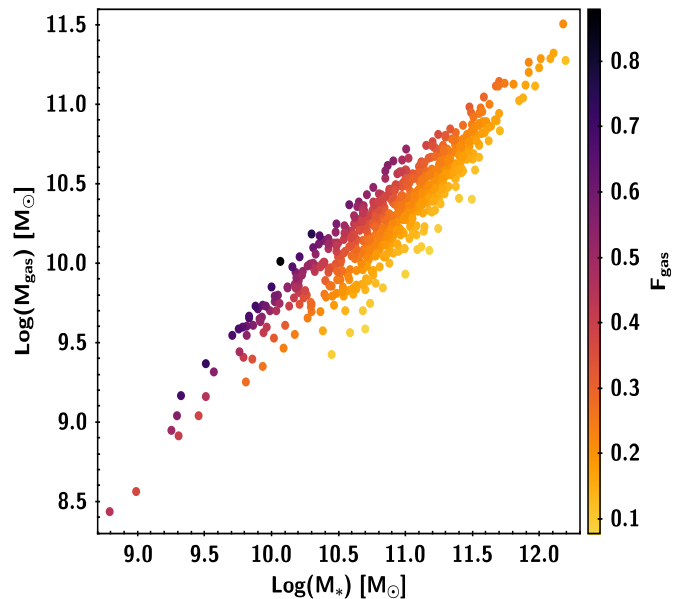
average SFR,  $M_*$ , sSFR, and MS offset in bins of  $N_{\text{H}}$  we found a remarkably flat distribution of all these quantities, in agreement with results from Shao et al. (2010), Rovilos et al. (2012), and Rosario et al. (2012), where the authors do not find any evolution of the average host properties in bins of  $N_{\text{H}}$ .

On the other hand, computing average  $N_{\text{H}}$  values in bins of  $M_*$  gives a positive trend in each redshift bin, while computing the average  $N_{\text{H}}$  in sSFR bins gives a negative trend, in agreement with partial correlation analysis. However, the situation in this case is complicated by the presence of upper limits in  $N_{\text{H}}$ , that makes the problem inherently asymmetric. We therefore performed the linear regression of  $(Y|X)$  with a Bayesian approach using the `linmix` code (Kelly et al. 2007), which is able to properly take into account the upper limits on  $N_{\text{H}}$ .

The result is shown in Fig. 9: the linear regression gives a clear positive correlation of  $N_{\text{H}}$  with the host stellar mass, increasing by 1–2 dex from low to high masses at all redshifts (slopes in the range  $\alpha = 0.42$ – $0.88$ ). An opposite result is found for the sSFR: the average  $N_{\text{H}}$  decreases typically by one order of magnitude or more, going from low to high sSFR (slopes in the range  $\alpha = -0.35$ – $-0.82$ ). Given that there is no trend of  $N_{\text{H}}$  with SFR, and that the sSFR is defined as  $\text{SFR}/M_*$ , the two relations are clearly connected.

A similar result between  $N_{\text{H}}$  and  $M_*$  was found in Rodighiero et al. (2015) for a sample of  $z \sim 2$  AGN hosts. In their analysis, however, the average  $N_{\text{H}}$  is globally  $\sim 1$  dex higher (gray squares in Fig. 9, left) because they derive  $N_{\text{H}}$  from the hardness ratio of the X-ray stacking, which also includes highly obscured, undetected AGN.

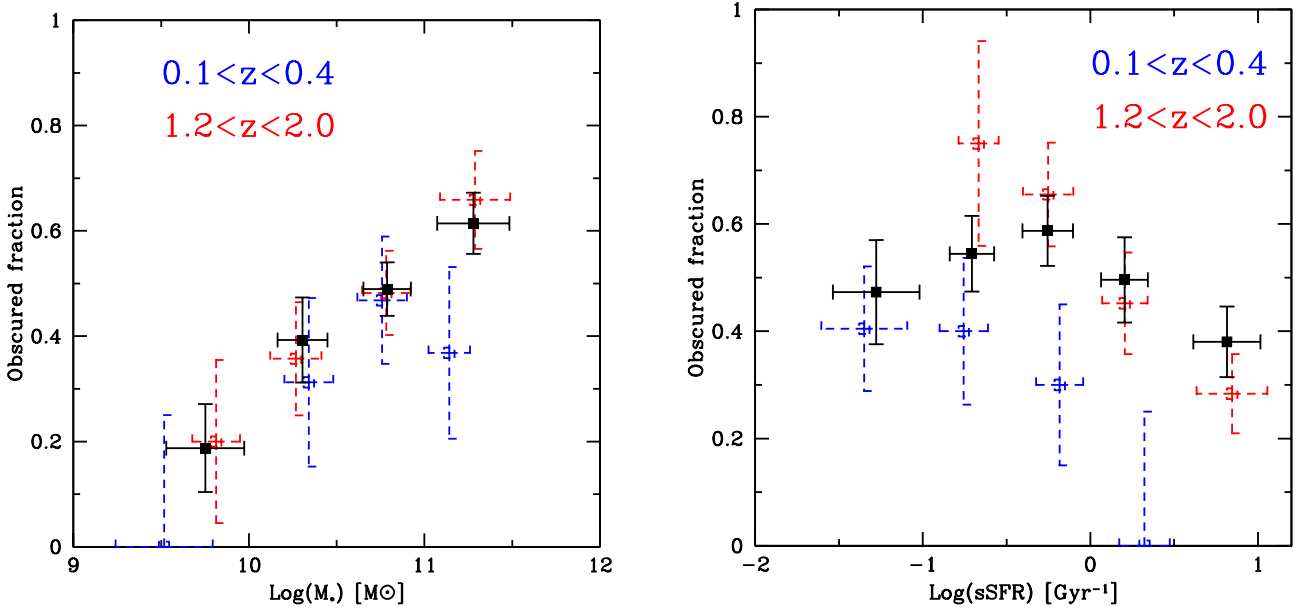
Interestingly, a recent study on the distribution of the obscuration observed in X-ray spectra of GRB, as a function of the host galaxy mass, found a similar trend in the redshift range  $1 \lesssim z \lesssim 5$  (Buchner et al. 2017, orange line in Fig. 9, left). Since for these sources the  $N_{\text{H}}$  from the GRB spectra probes only the host obscuration, the authors conclude that a large fraction of the obscuration observed in AGN, at least in the Compton thin



**Fig. 10.**  $\log M_*$  vs.  $\log M_{\text{gas}}$  as derived from Eq. (1) of Scoville et al. (2016). The sources are color-coded on the basis of their gas fraction.

regime, is not due to the nuclear torus, but to the galaxy-scale gas in the host.

These dependencies imply that at increasing galaxy mass there are more chances to have an additional component to the amount of gas and dust along the line of sight through the AGN. It is well established that the gas fraction is a strong decreasing function of the galaxy mass (e.g., Santini et al. 2014; Peng et al. 2015). However, it is possible to show that the total amount of gas is driven mainly by the total galaxy mass and not by the gas fraction. To this end, we computed gas mass for all our galaxies, following the empirical relation found in Scoville et al. (2016) (their Eq. (1)), that links  $M_*$ , sSFR offset from the MS, and molecular gas mass. This is shown in Fig. 10, where the sources



**Fig. 11.** Fraction of obscured sources as a function of  $M_*$  (left) and sSFR (right), for the entire sample (black points). The blue (red) dashed points show the results for the first (fourth) redshift bin, respectively.

are color-coded on the basis of their gas fraction. Even if at increasing  $M_*$  the gas fraction is smaller, the total amount of gas still increases with  $M_*$ .

The well-known mass-metallicity relation (e.g., Tremonti et al. 2004; Mannucci et al. 2010) goes in the direction of having more metals (responsible for X-ray absorption) with increasing  $M_*$ . In particular, going from  $\text{Log}(M_*) = 9.5$  to 11.5, there is an increase of a factor  $\sim 2$  in the metallicity, up to  $z \sim 2$  (Erb et al. 2006); however, this is not enough to explain the increase in average  $N_{\text{H}}$  observed here. Measuring the  $N_{\text{H}}$  with fixed metallicity (as is done here) for sources with such a range in metallicity translates into a factor of  $\sim 2$  difference in measured  $N_{\text{H}}$  for a given input obscuration.

## 5.2. Obscured fraction

To compare our results with the literature, we also looked at the fraction of obscured sources as a function of host properties. In Fig. 11 we show the fraction of obscured sources, defined as  $N_{\text{Obs}}/N_{\text{Tot}}$ , where  $N_{\text{Obs}}$  is the number of sources with a detection of  $N_{\text{H}}$  and  $N_{\text{H}} > 1 \times 10^{22} \text{ cm}^{-2}$ . As expected from what was shown in the previous section, the fraction of obscured sources increases with increasing  $M_*$  and decreases with sSFR (for  $\text{sSFR} > 1 \text{ Gyr}^{-1}$ ). The decrease in sSFR is partly hidden because we consider the full redshift interval ( $z = 0.1-4$ ), while Fig. 9 (right) shows that the range covered by the different subsamples shifts toward higher sSFR with redshift. For this reason we also show in Fig. 11 the results for the first and fourth bins as an example (blue and red dashed points, respectively).

Merloni et al. (2014) found a flat relation between the fraction of obscured sources and  $M_*$  in a sample of X-ray detected AGN from the XMM-COSMOS catalog. However, they limited their analysis to a narrow range in  $L_X$  (in order to cover a wide range redshift), while the obscured fraction is known to evolve strongly with  $L_X$  (e.g., Ueda et al. 2015).

Another group, instead, have found an increasing fraction of obscured sources as a function of sSFR and MS offset, in a sample of  $70 \mu\text{m}$  selected galaxies at  $0.3 < z < 1$ , interpreted as an indication of increasing gas fraction or density in the host,

which in turn would sustain the increased sSFR (e.g., Juneau et al. 2013; hereafter J13).

We note that the definition of obscured AGN adopted here and in J13 are different, and in the latter, are mostly based on the lack of X-ray detection; there are 64 sources (out of 99 AGN) classified as obscured AGN on the basis of the mass-excitation diagram selection (MEX; Juneau et al. 2011), and the X-ray non-detection. If these objects are indeed highly obscured, Compton-thick AGN, this population is mostly missed in our X-ray based sample.

Another possibility is that a fraction of the MEX-selected AGN are not actively/strongly accreting SMBHs. Indeed, a sizable fraction ( $\sim 30\%$ ) of the AGN selected in J13 through the MEX diagram has a host  $M_*$  below  $\text{Log}(M_*) = 10.5$ . As shown in Sect. 2.2, however, X-ray detected AGN are rare at low  $M_*$ . Therefore, all the sources that are X-ray undetected for reasons different from obscuration (variability, intrinsic weakness, contaminant non-AGN, etc.) would appear as obscured, low  $M_*$  host AGN (hence high sSFR), possibly affecting the observed trends.

## 6. Discussion

We collected a large sample of X-ray and FIR detected AGN and host systems in the COSMOS field, spanning  $\sim 4$  orders of magnitude in  $L_X$ ,  $N_{\text{H}}$ ,  $L_{\text{IR}}^{\text{SF}}$ ,  $M_*$ , and covering the redshift range  $0.1 < z < 4$ . We applied X-ray spectral analysis down to very low counts ( $> 30$  net counts) and adopted the SED decomposition results derived in D15, to recover both AGN and SF properties of each source. With this data set in hand, we demonstrated that it is possible to reproduce both the flat distribution of average  $L_{\text{IR}}^{\text{SF}}$  in bins of  $L_X$  and the steeper correlation of average  $L_X$  in bins of  $L_{\text{IR}}^{\text{SF}}$  reported in the literature in recent years (e.g., Shao et al. 2010; Rosario et al. 2012; Mullaney et al. 2012, C13; Stanley et al. 2015).

The apparently contradictory results found in the literature, and reproduced in Sect. 3.2, are due to the different results that are obtained when binning along one axis or the other, the equivalent of a forward or inverse linear regression (i.e.,  $L_{\text{IR}}^{\text{SF}} | L_X$  vs.  $L_X | L_{\text{IR}}^{\text{SF}}$ ), as proposed in Hickox et al. (2014) and

Volonteri et al. (2015), and found in Dai et al. (2015) and Stanley et al. (2015) (but see also McAlpine et al. 2017, for a possible alternative explanation).

Both from a physical and a statistical point of view, it seems more appropriate to consider the results from  $L_X|L_{\text{IR}}^{\text{SF}}$ , given the larger measurement uncertainties on  $L_X$ , and the shorter time scale variability of  $L_X$  with respect to  $L_{\text{IR}}^{\text{SF}}$ , which adds a further term of intrinsic scatter. Doing so, we found a linear correlation between  $L_X$  and  $L_{\text{IR}}^{\text{SF}}$  with a slope consistent with 1, at least in the redshift range 0.4–1.2, i.e., below the peak of the SF and BH accretion history. Beyond that and up to  $z = 4$ , the slope becomes significantly flatter,  $\alpha = 0.3\text{--}0.5$ .

The other possibility is to adopt a symmetrical approach, even if there is no general agreement on this (see Hogg et al. 2010, on the bisector method). In this case the result is a correlation with slope consistent with  $\sim 1$  at all redshifts. This would point toward an average one-to-one correlation between SF and BH accretion, in the last 12 Gyr of cosmic history.

Even more interesting is the full distribution of BH and host properties, such as  $L_X$  and  $L_{\text{IR}}^{\text{SF}}$  or sBHAR and sSFR, that can be only qualitatively compared, for the moment, with predictions from galaxy merger simulations, resulting in interesting similarities between observations and models.

We stress again that these results apply to the small subsample of AGN/host systems detected in both X-ray and FIR, which represents only  $\sim 20\%$  of the full X-ray sample and  $\sim 10\%$  of the AGN FIR sample. Indeed, one of the main reasons why it is so difficult for present observations to probe the AGN-SF connection is that X-ray and/or FIR detected systems span a limited range in AGN and SF activity, sampling only the high  $L_X$ /SFR tail of the possible correlation (e.g., Sijacki et al. 2015).

It is interesting, however, that we are able to reproduce the results obtained via stacking of samples where the vast majority of the sources are not detected (e.g., 20% of FIR detected AGN selected in X-ray in Shao et al. 2010). As suggested in Mullaney et al. (2015), the stacking analysis is the equivalent of a linear mean, and may be dominated by the brightest sources.

A crucial next step in the comparison between theory and observations will be to select the observed systems in different evolutionary stage, to reach a similar level of detail as in the current simulations. This will be feasible for large samples only at low redshift, while detailed and complete morphological studies in COSMOS (and other deep fields) data are already very difficult at  $z \gtrsim 1$ . From the theoretical point of view, more demanding galaxy merger simulations will be required in order to cover a mass range comparable to the one of observed systems with the same high resolution, and to possibly move toward a high redshift environment.

Finally, a positive correlation between  $N_{\text{H}}$  and  $M_*$ , and a similar negative correlation with sSFR, have been found at all redshift bins. A similar result was found by Rodighiero et al. (2015) in a large sample of high redshift galaxies, computing HR of stacked X-ray images. A recent study on GRB hosts has found a similar behavior (Buchner et al. 2017), implying that an important fraction (up to 40%) of the Compton thin obscuration found in AGN can be ascribed to galaxy scale gas (Buchner & Bauer 2017).

Several studies have found no correlation between column density and host properties (Rovilos et al. 2012; Rosario et al. 2012), while others (e.g., J13) have found a positive correlation of the fraction of obscured sources with sSFR. Further investigation in this direction will help to shed light on the role of the host in contributing to the obscuration through the AGN line of sight.

*Acknowledgements.* The authors thank the anonymous referee for valuable comments. G.L., M.B., and M.P. acknowledge financial support from the CIG grant “eEASY” No. 321913. G.L. acknowledges financial support from ASI-INAF 2014-045-R.O. I.D. acknowledges the European Union’s Seventh Framework programme under grant agreement 337595 (ERC Starting Grant, “CoSMas”). We acknowledge the contributions of the entire COSMOS collaboration consisting of more than 100 scientists. More information on the COSMOS survey is available at <http://cosmos.astro.caltech.edu/>. Based on observations obtained with *XMM-Newton*, an ESA science mission with instruments and contributions directly funded by ESA Member States and NASA, and data obtained from the *Chandra* Data Archive.

## References

- Aird, J., Coil, A. L., Moustakas, J., et al. 2012, *ApJ*, 746, 90  
 Akritas, M. G., & Bershad, M. A. 1996, *ApJ*, 470, 706  
 Alonso-Herrero, A., Pereira-Santaella, M., Rieke, G. H., & Rigopoulou, D. 2012, *ApJ*, 744, 2  
 Arnaud, K. A. 1996, *Astronomical Data Analysis Software and Systems V*, *ASP Conf. Ser.*, 101, 17  
 Bennett, C. L., Larson, D., Weiland, J. L., et al. 2013, *ApJS*, 208, 20  
 Berta, S., Lutz, D., Santini, P., et al. 2013, *A&A*, 551, A100  
 Bongiorno, A., Merloni, A., Brusa, M., et al. 2012, *MNRAS*, 427, 3103  
 Brusa, M., Fiore, F., Santini, P., et al. 2009, *A&A*, 507, 1277  
 Brusa, M., Civano, F., Comastri, A., et al. 2010, *ApJ*, 716, 348  
 Bruzual, G., & Charlot, S. 2003, *MNRAS*, 344, 1000  
 Buchner, J., & Bauer, F. E. 2017, *MNRAS*, 465, 4348  
 Buchner, J., Schulze, S., & Bauer, F. E. 2017, *MNRAS*, 464, 4545  
 Bundy, K., Georgakakis, A., Nandra, K., et al. 2008, *ApJ*, 681, 931  
 Cai, Z.-Y., Lapi, A., Xia, J.-Q., et al. 2013, *ApJ*, 768, 21  
 Capak, P., Aussel, H., Ajiki, M., et al. 2007, *ApJS*, 172, 99  
 Capelo, P. R., Volonteri, M., Doti, M., et al. 2015, *MNRAS*, 447, 2123  
 Chabrier, G. 2003, *ApJ*, 586, L133  
 Chen, C.-T. J., Hickox, R. C., Alberts, S., et al. 2013, *ApJ*, 773, 3 (C13)  
 Chen, C.-T. J., Hickox, R. C., Alberts, S., et al. 2015, *ApJ*, 802, 50  
 Civano, F., Elvis, M., Brusa, M., et al. 2012, *ApJS*, 201, 30  
 Civano, F., Marchesi, S., Elvis, M., et al. 2017, *ApJ*, submitted  
 Comastri, A., Ranalli, P., Iwasawa, K., et al. 2011, *A&A*, 526, L9  
 Cresci, G., Hicks, E. K. S., Genzel, R., et al. 2009, *ApJ*, 697, 115  
 Dai, Y. S., Wilkes, B. J., Bergeron, J., et al. 2015, ArXiv e-prints [[arXiv:1511.06761](https://arxiv.org/abs/1511.06761)]  
 Delvecchio, I., Gruppioni, C., Pozzi, F., et al. 2014, *MNRAS*, 439, 2736  
 Delvecchio, I., Lutz, D., Berta, S., et al. 2015, *MNRAS*, 449, 373  
 Diamond-Stanic, A. M., & Rieke, G. H. 2012, *ApJ*, 746, 168  
 Di Matteo, T., Springel, V., & Hernquist, L. 2005, *Nature*, 433, 604  
 Dubois, Y., Peirani, S., Pichon, C., et al. 2016, *MNRAS*, 463, 3948  
 Elvis, M., Civano, F., Vignali, C., et al. 2009, *ApJS*, 184, 158  
 Erb, D. K., Shapley, A. E., Pettini, M., et al. 2006, *ApJ*, 644, 813  
 Fabian, A. C., & Iwasawa, K. 1999, *MNRAS*, 303, L34  
 Feigelson, E. D., & Berg, C. J. 1983, *ApJ*, 269, 400  
 Feltre, A., Hatziminaoglou, E., Fritz, J., & Franceschini, A. 2012, *MNRAS*, 426, 120  
 Fritz, J., Franceschini, A., & Hatziminaoglou, E. 2006, *MNRAS*, 366, 767  
 Fruscione, A., McDowell, J. C., Allen, G. E., et al. 2006, *Proc. SPIE*, 6270, 1  
 Genzel, R., Newman, S., Jones, T., et al. 2011, *ApJ*, 733, 101  
 Granato, G. L., De Zotti, G., Silva, L., Bressan, A., & Danese, L. 2004, *ApJ*, 600, 580  
 Griffin, M. J., Abergel, A., Abreu, A., et al. 2010, *A&A*, 518, L3  
 Gruppioni, C., Berta, S., Spinoglio, L., et al. 2016, *MNRAS*, 458, 4297  
 Hasinger, G., Cappelluti, N., Brunner, H., et al. 2007, *ApJS*, 172, 29  
 Hickox, R. C., Mullaney, J. R., Alexander, D. M., et al. 2014, *ApJ*, 782, 9  
 Hogg, D. W., Bovy, J., & Lang, D. 2010, ArXiv e-prints [[arXiv:1008.4686](https://arxiv.org/abs/1008.4686)], unpublished  
 Imanishi, M., Maiolino, R., & Nakagawa, T. 2010, *ApJ*, 709, 801  
 Isobe, T., Feigelson, E. D., Akritas, M. G., & Babu, G. J. 1990, *ApJ*, 364, 104  
 Juneau, S., Dickinson, M., Alexander, D. M., & Salim, S. 2011, *ApJ*, 736, 104  
 Juneau, S., Dickinson, M., Bournaud, F., et al. 2013, *ApJ*, 764, 176 (J13)  
 Kauffmann, G., & Heckman, T. M. 2009, *MNRAS*, 397, 135  
 Kauffmann, G., Heckman, T. M., Tremonti, C., et al. 2003, *MNRAS*, 346, 1055  
 Kelly, B. C. 2007, *ApJ*, 665, 1489  
 Kennicutt, R. C., Jr. 1998, *ApJ*, 498, 541  
 Koekemoer, A. M., Aussel, H., Calzetti, D., et al. 2007, *ApJS*, 172, 196  
 Kormendy, J., & Ho, L. C. 2013, *ARA&A*, 51, 511  
 Kormendy, J., & Richstone, D. 1995, *ARA&A*, 33, 581  
 Lagos, C. D. P., Baugh, C. M., Lacey, C. G., et al. 2011, *MNRAS*, 418, 1649  
 Lanzuisi, G., Civano, F., Elvis, M., et al. 2013, *MNRAS*, 431, 978



- Lanzuisi, G., Ranalli, P., Georgantopoulos, I., et al. 2015a, *A&A*, 573, A137 (L15)
- Lanzuisi, G., Perna, M., Delvecchio, I., et al. 2015b, *A&A*, 578, A120
- Lusso, E., Comastri, A., Simmons, B. D., et al. 2012, *MNRAS*, 425, 623 (L12)
- Lutz, D., Spoon, H. W. W., Rigopoulou, D., Moorwood, A. F. M., & Genzel, R. 1998, *ApJ*, 505, L103
- Lutz, D., Poglitsch, A., Altieri, B., et al. 2011, *A&A*, 532, A90
- Macklin, J. T. 1982, *MNRAS*, 199, 1119
- Madau, P., & Dickinson, M. 2014, *ARA&A*, 52, 415
- Magorrian, J., Tremaine, S., Richstone, D., et al. 1998, *AJ*, 115, 2285
- Mainieri, V., Bongiorno, A., Merloni, A., et al. 2011, *A&A*, 535, A80
- Maiolino, R., Shemmer, O., Imanishi, M., et al. 2007, *A&A*, 468, 979
- Mannucci, F., Cresci, G., Maiolino, R., Marconi, A., & Gnerucci, A. 2010, *MNRAS*, 408, 2115
- Marchesi, S., Civano, F., Elvis, M., et al. 2016a, *ApJ*, 817, 34
- Marchesi, S., Lanzuisi, G., Civano, F., et al. 2016b, *ApJ*, 830, 100
- Marconi, A., & Hunt, L. K. 2003, *ApJ*, 589, L21
- Marconi, A., Risaliti, G., Gilli, R., et al. 2004, *MNRAS*, 351, 169
- McAlpine, S., Bower, R. G., Harrison, C. M., et al. 2017, *MNRAS*, 468, 3395
- Menci, N., Fiore, F., Puccetti, S., & Cavaliere, A. 2008, *ApJ*, 686, 219
- Merloni, A., Bongiorno, A., Brusa, M., et al. 2014, *MNRAS*, 437, 3550
- Mullaney, J. R., Alexander, D. M., Goulding, A. D., & Hickox, R. C. 2011, *MNRAS*, 414, 1082
- Mullaney, J. R., Pannella, M., Daddi, E., et al. 2012, *MNRAS*, 419, 95
- Mullaney, J. R., Alexander, D. M., Aird, J., et al. 2015, *MNRAS*, 453, L83
- Neistein, E., & Netzer, H. 2014, *MNRAS*, 437, 3373
- Novak, G. S., Faber, S. M., & Dekel, A. 2006, *ApJ*, 637, 96
- Oliver, S. J., Bock, J., Altieri, B., et al. 2012, *MNRAS*, 424, 1614
- Peng, Y., Maiolino, R., & Cochrane, R. 2015, *Nature*, 521, 192
- Pilbratt, G. L., Riedinger, J. R., Passvogel, T., et al. 2010, *A&A*, 518, L1
- Poglitsch, A., Waelkens, C., Geis, N., et al. 2010, *A&A*, 518, L2
- Pontzen, A., Tremmel, M., Roth, N., et al. 2017, *MNRAS*, 465, 547
- Pozzi, F., Vignali, C., Gruppioni, C., et al. 2012, *MNRAS*, 423, 1909
- Rodighiero, G., Daddi, E., Baronchelli, I., et al. 2011, *ApJ*, 739, L40
- Rodighiero, G., Brusa, M., Daddi, E., et al. 2015, *ApJ*, 800, L10
- Rosario, D. J., Santini, P., Lutz, D., et al. 2012, *A&A*, 545, A45
- Rovilos, E., Comastri, A., Gilli, R., et al. 2012, *A&A*, 546, A58
- Saintonge, A., Catinella, B., Cortese, L., et al. 2016, *MNRAS*, 462, 1749
- Salvato, M., Hasinger, G., Ilbert, O., et al. 2009, *ApJ*, 690, 1250
- Santini, P., Rosario, D. J., Shao, L., et al. 2012, *A&A*, 540, A109
- Santini, P., Maiolino, R., Magnelli, B., et al. 2014, *A&A*, 562, A30
- Schulze, A., Bongiorno, A., Gavignaud, I., et al. 2015, *MNRAS*, 447, 2085
- Scoville, N., Aussel, H., Brusa, M., et al. 2007, *ApJS*, 172, 1
- Scoville, N., Sheth, K., Aussel, H., et al. 2016, *ApJ*, 820, 83
- Scudder, J. M., Oliver, S., Hurley, P. D., et al. 2016, *MNRAS*, 460, 1119
- Shao, L., Lutz, D., Nordon, R., et al. 2010, *A&A*, 518, L26
- Sijacki, D., Vogelsberger, M., Genel, S., et al. 2015, *MNRAS*, 452, 575
- Silverman, J. D., Lamareille, F., Maier, C., et al. 2009, *ApJ*, 696, 396
- Stanley, F., Harrison, C. M., Alexander, D. M., et al. 2015, *MNRAS*, 453, 591
- Suh, H., Civano, F., Hasinger, G., et al. 2017, *ApJ*, submitted
- Symeonidis, M., Giblin, B. M., Page, M. J., et al. 2016, *MNRAS*, 459, 257
- Tremaine, S., Gebhardt, K., Bender, R., et al. 2002, *ApJ*, 574, 740
- Tremonti, C. A., Heckman, T. M., Kauffmann, G., et al. 2004, *ApJ*, 613, 898
- Ueda, Y., Akiyama, M., Hasinger, G., Miyaji, T., & Watson, M. G. 2014, *ApJ*, 786, 104
- Vito, F., Maiolino, R., Santini, P., et al. 2014, *MNRAS*, 441, 1059
- Volonteri, M., Capelo, P. R., Netzer, H., et al. 2015a, *MNRAS*, 449, 1470
- Volonteri, M., Capelo, P. R., Netzer, H., et al. 2015b, *MNRAS*, 452, L6
- Whitaker, K. E., van Dokkum, P. G., Brammer, G., & Franx, M. 2012, *ApJ*, 754, L29
- Xue, Y. Q., Brandt, W. N., Luo, B., et al. 2010, *ApJ*, 720, 368