

# Longitudinal trends in the prevalence of detectable HIV viremia: Population-based evidence from rural KwaZulu-Natal, South Africa

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**Summary:** Using data from a large surveillance program in South Africa, we estimated the longitudinal change in the prevalence of detectable HIV viremia. Results show the need to report the prevalence of detectable viremia among all adults, irrespective of HIV status.

## Abstract

**Background:** The prevalence of detectable viremia has previously been used to infer the potential for ongoing HIV transmission. To date, no study has evaluated the longitudinal change in the prevalence of detectable viremia within the HIV-positive community (PDV<sub>+</sub>) and the entire population (PDV<sub>P</sub>) using data from a sub-Saharan African setting.

**Methods:** In 2011, 2013, and 2014, we obtained 6,752 HIV-positive and 15,415 HIV-negative test results from a population-based surveillance system in the KwaZulu-Natal province of South Africa. We quantified the PDV<sub>+</sub> as the proportion of the 6,752 HIV-positive results with a viral load >1,550 copies/mL and the PDV<sub>P</sub> as the proportion of the 6,752 HIV-positive *and* 15,415 HIV-negative results with a viral load >1,550 copies/mL.

**Results:** Between 2011 and 2014, the PDV<sub>+</sub> decreased by 16.5 percentage points (pp) for women (from 71.8% to 55.3%) and 10.6 pp for men (from 77.8% to 67.2%). However, a steady rise in the overall HIV prevalence, from 26.7% to 32.4%, offset the declines in the PDV<sub>+</sub> for both sexes. For woman, the PDV<sub>P</sub> decreased by only 2.1 pp, from 21.3% to 19.2%; but for men, the PDV<sub>P</sub> actually increased by 1.6 pp, from 14.6% to 16.2%, over the survey period.

**Discussion:** The PDV<sub>+</sub>, which is currently being tracked under the UNAID 90-90-90 targets, may not be accurate indicator of the potential for ongoing HIV transmission. There is a critical need for countries to monitor and report the prevalence of detectable viremia among all adults (PDV<sub>P</sub>), irrespective of HIV status.

**Keywords:** HIV, viral load, detectable viremia, prevalence, South Africa

## Introduction

By 2015, almost half of the 36.7 million people living with HIV were on combination antiretroviral therapy (ART) [1]. ART is expected to prevent the onward transmission of HIV by reducing the number of infected persons with detectable viremia [2, 3]. For this reason, the HIV-positive prevalence of detectable viremia (PDV<sub>+</sub>), which is the proportion of all infected persons with a recent viral load above a copies/mL threshold, has been promoted as a sensitive biological index of ART programme effectiveness. The PDV<sub>+</sub> has previously been used to monitor a community's uptake of ART [4, 5], and is central to the UNAIDS 90-90-90 targets to have 90% of all ART-initiated patients achieve undetectable viremia by the year 2020 [6]. In addition, the PDV<sub>+</sub> has been used to quantify the potential for ongoing HIV transmission within a well-defined community or geographic area [4, 5, 7-9]. An assumption underlying the use of this measure is that higher levels of ART coverage will lower the PDV<sub>+</sub> and thus reduce the incidence of HIV infection within the general population.

However, one key limitation of the PDV<sub>+</sub> is that it does not account for the relative sizes of the HIV-infected and HIV-uninfected populations [10]. This information is important because the risk of acquiring HIV will depend not only on the number of infected persons with detectable viremia (i.e., PDV<sub>+</sub>) but also on the number of infected persons in the general population (i.e., HIV prevalence), and the rate of sexual contact between them [10]. Thus, an improved biological index, which we call the population prevalence of detectable viremia (PDV<sub>P</sub>) [11], can be obtained by multiplying the PDV<sub>+</sub> with the HIV prevalence (see Figure S1 of the Supplement). Aggregated viral load indices that account for the HIV prevalence have gained traction in the literature [12-15], and we recently showed that the PDV<sub>P</sub> is significantly better than the PDV<sub>+</sub> at predicting the prospective risk of HIV infection [11].

As far as we know, time-trends in both the PDV<sub>+</sub> and the PDV<sub>P</sub> have not been evaluated and compared using data from a sub-Saharan African population. In 2011, 2013, and 2014, we obtained 6,752 HIV-positive and 15,415 HIV-negative test results from a population-based surveillance system in the KwaZulu-Natal province of South Africa. We quantified the PDV<sub>+</sub> as the proportion of the HIV-positive test results with a viral load >1,550 copies/mL and then quantified the PDV<sub>P</sub> as the proportion of the HIV-positive *and* HIV-negative test results with a viral load >1,550 copies/mL. Using this

population-based data, we have a unique opportunity to empirically estimate and compare the changes in both the  $PDV_+$  and  $PDV_P$  measures over time.

## Methods

### *Setting*

The Africa Health Research Institute (AHRI) maintains a population-based surveillance system in the Umkhanyakude district of the northern KwaZulu-Natal province. Most of the surveillance area is poor and rural, with several informal peri-urban settlements and a single urban township [16]. The area is 438 km<sup>2</sup> in size with a population of approximately 90,000 people and 11,000 households.

### *HIV surveillance survey*

AHRI has collected longitudinal data on households and individuals within the surveillance area since 2000. Every six months, trained field-workers visit a key-informant within the household to collect information on both resident and non-resident members. Biannual participation rates for household data collection are typically >95%. Nested within the AHRI cohort is the population-based HIV cohort. Field-workers have visited households every twelve months since 2004 and identified eligible participants older than 15 years for HIV testing. After obtaining consent, the field workers extract blood according to the UNAIDS and WHO *Guidelines for Using HIV Testing Technologies in Surveillance*. Of the eligible participants contacted, 78% agreed to be tested for HIV at least once in the three survey years. Participants from the AHRI and HIV cohorts were linked across the survey years and the data were stored in a SQL database server. The AHRI and HIV cohorts are described in greater detail elsewhere [16].

### *HIV incidence and ART usage*

The AHRI surveillance area is situated at the epicentre of the global AIDS epidemic. Between 2004 and 2011, the crude HIV incidence was 2.6 new infections per 100 person-years (95% confidence interval [CI]: 2.50–2.77) [17]. Incidence peaked at 6.6 new infections per 100 person-years in woman aged 24 years and at 4.1 new infections per 100 person-years in males aged 29 years [17]. Since 2005, the HIV prevalence

among men and woman aged 15–54 years has increased steadily from 21.7% in 2005 to 28.7% in 2010 [18]. The increase in HIV prevalence has been attributed to ART-associated reductions in mortality [19].

ART can be accessed for free at any of the 17 primary health-care clinics within or adjacent to the surveillance area [20]. When ART was first made available in 2004, the CD4+ T-cell count eligibility criteria was <200 cells/ $\mu$ L. In 2010, treatment eligibility was extended to pregnant woman with CD4+ T-cell counts <350 cells/ $\mu$ L and patients with active tuberculosis. All patients with CD4+ T-cell counts <350 cells/ $\mu$ L became eligible for ART in 2011. Approximately 32.2% (95% CI: 30.2–34.2) of the HIV-participants in our study area were on ART in 2011, which increased to 40.7% (95% CI: 38.6–42.7) in 2013.

### *Viral load measurements*

All of the 5,368 participants, aged 15 to 64 years, who tested HIV-positive in 2011 (n=2,401), 2013 (n=2,510), and 2014 (n=2,611) provided dried blood spot (DBS) samples. The total number of DBS samples was 7,522 since 32.4% of the 5,368 participants tested HIV-positive in more than one survey year. From all 7,522 DBS samples, we extracted nucleic acid with NucliSENS® EasyMag® (Bordeaux, France) and used the Generic HIV Viral Load kit (Biocentric) to quantify the viral load levels. As described in greater detail elsewhere [21], the quantification method has a lower detection limit of 1,550 copies/mL. Due to insufficient specimens, we had to exclude 770 (10.24%) viral load samples. For the final analysis, we therefore used a total of 6,752 viral load measurements from 4,991 unique participants who tested HIV-positive in 2011 (n=2,366), 2013 (n=2,135), and 2014 (n=2,251).

### *Prevalence of detectable viremia measures*

We calculated the  $PDV_+$  for the each survey year  $t$  as follows (we drop the subscript  $t$  as it is implicit throughout). Let  $v_i$  denote  $i$ th viral load measurement for  $i = 1, \dots, n^+$ , where  $n^+$  is the number of HIV-positive test results, and let  $y_i = 1$  if  $v_i > 1,550$  copies/mL otherwise  $y_i = 0$ . Then, the  $PDV_+ = \sum_{i=1}^{n^+} y_i / n^+$ , which is the number of viral load measurements  $>1,550$  copies/mL divided by the number of HIV-positive test results. This  $PDV_+$  measure is a true population estimator because the viral load measurements come from a representative sample of HIV-positive participants. For this

reason, our analysis avoids the sampling biases typically associated with facility-based studies in which patients self-select into care [10].

We calculated the  $PDV_P$  for each survey year as follows: let  $n^-$  denote the number of HIV-negative test results and let  $N$  denote the total number of HIV-positive and HIV-negative test results, with  $N = n^- + n^+$ . For all HIV-negative test results we denote  $y_i = 0$ . Then, the  $PDV_P = \sum_{i=1}^N y_i / N$ , which is the number of viral loads  $>1,550$  copies/mL divided by the total number of HIV-positive and HIV-negative test results.

We note that the number of HIV-negative test results for each survey year was determined with  $n_{as}^- = [n_{as}^+ - (n_{as}^+ \times H_{as})] / H_{as}$ , where  $H$  is the HIV prevalence and the subscripts  $a$  and  $s$  denote the age group and sex respectively. Overall, 15,415 HIV-negative test results were sampled from 11,522 unique participants. We used this proportional allocation approach [22] to determine  $n^-$  because 770 HIV-positive samples were excluded from the analysis due to insufficient specimens (as described in the previous section). Otherwise, we would underestimate the  $PDV_P$  if we did not sample the correct  $n^-$  using this approach.

### *Statistical analysis*

We performed summary statistics for the unadjusted and age-sex adjusted  $PDV_+$ ,  $PDV_P$ , and HIV prevalence measures by year. To statistically assess the change in the  $PDV_+$  and  $PDV_P$  measures over time, we used a generalized estimating equations (GEE) model with a logit link function. We chose a GEE model because 32.4% of the participants tested HIV-positive in more than one survey year. We fitted four regression models using data from the HIV-positive participants only (i.e.,  $PDV_+$ ) and from the HIV-positive and HIV-negative participants (i.e.,  $PDV_P$ ). For Model 1, we included a variable indicating the year of the HIV-positive (i.e., viral load measurement) or HIV-negative test result. For Model 2, we added a sex variable to the year variable of Model 1, and for Model 3 we added an age variable ( $>25$  years) to the Model 2 variables. For Model 4, we added a sex-year interaction term to the Model 3 variables to determine if the  $PDV_+$  and  $PDV_P$  measures changed significantly for men and woman over time.

## Results

For all participants with a viral load measurement, the median age was 35 (IQR: 27–45) years and 79% were female. For the HIV-positive and HIV-negative participants, the median age was 31 (IQR: 21–47) years and 69% were female, as shown in Table 1.

Results show that the adjusted  $PDV_+$  decreased by 13.86 percentage points, from 73.76% in 2011 to 64.38% in 2013, and then to 59.90% in 2014 (see Table 1 and Figure 1). During this time, the adjusted HIV prevalence increased from 26.73% in 2011 to 30.64% in 2013 and then to 32.36% in 2014. Thus, when we accounted for the HIV prevalence, the adjusted  $PDV_P$  decreased by only 0.92 percentage points, from 18.83% in 2011 to 18.80% in 2013 and then to 17.91% in 2014.

We observed marked differences in the adjusted  $PDV_+$  and  $PDV_P$  measures by sex over time, as shown in Figure 2. Between 2011 and 2014, the  $PDV_+$  for woman decreased by 16.5 percentage points (pp), from 71.8% to 55.3%, compared with a 10.6 pp decrease in the  $PDV_+$  for men, from 77.80% to 67.18% (Table S1 of the Supplement). However, woman had a higher HIV prevalence, 30.56% in 2011 and 35.61% in 2014, and therefore a higher  $PDV_P$ , which decreased by 2.1 pp, from 21.35% to 19.23% over the survey period. For men, the HIV prevalence rose sharply from 19.63% in 2011 to 27.05% in 2014, which offset the decline in their  $PDV_+$ . Thus, the  $PDV_P$  for men actually increased by 1.6 pp over the survey period, from 14.58% to 16.18%.

The GEE model results show that the odds of detectable viremia within the HIV-positive population ( $PDV_+$ ) was significantly lower in 2013 (0.647; 95% CI: [0.575, 0.727]; p-value <0.001) and 2014 (0.490; 95% CI: [0.436, 0.551]; p-value <0.001) when compared with 2011 (Table 2). In addition, the odds of detectable viremia was significantly lower in woman than men, but there was no difference between men and woman over time, as shown by the two interaction terms in Table 2 (p-values >0.266).

The odds of detectable viremia within the entire population ( $PDV_P$ ) was slightly lower in 2014 (0.911; 95% CI: [0.850, 0.977]; p-value=0.009), but not in 2013 (0.968; 95% CI: [0.908, 1.031]; p-value=0.31), when compared with 2011 (Table 3). Although the odd of detectable viremia was higher for woman, these odds declined significantly over time when compared with men. We found a similar result when we stratified our analysis by sex (see Table S2 of the Supplement).

## Discussion

Our study has quantified the temporal change in the HIV-positive prevalence of detectable viremia (PDV<sub>+</sub>) and the population prevalence of detectable viremia (PDV<sub>P</sub>) using data from a sub-Saharan African population. The results show that the PDV<sub>+</sub> decreased by almost 14 percentage points, from 73.8% to 59.9%, over the 2011–2014 survey period. In this regard, the 17 health-care clinics within or adjacent to our surveillance area have been effective in getting HIV-positive persons onto ART and then reducing their viral load levels over time. This is positive news for the global HIV treatment-as-prevention initiative as well as for our study community, which is considered to be at the epicentre of the global AIDS epidemic.

We compare our 40.1% prevalence of undetectable viremia in the HIV-positive community (i.e.,  $100 - \text{PDV}_+$ ) in 2014 with population-based studies undertaken in Malawi [23], Zambia [24], and Zimbabwe [25] in 2015/2016. In Malawi, the prevalence of undetectable viremia (<1,000 copies/mL) in the HIV-positive community was 67.6% (95% CI: 65.0–70.2%) among 15–64 year olds, 59.8% (95% CI: 57.4–62.2%) among 15–59 year olds in Zambia, and 60.4% (95% CI: 58.3–62.5%) among 15–64 year olds in Zimbabwe. These estimates are markedly higher than our PDV<sub>+</sub> result, despite a lower detection level. It is likely that these differences would be slightly smaller in 2015/2016, if our PDV<sub>+</sub> continued to decrease as it did over the survey period. Nevertheless, we acknowledge that our 40.1% estimate is well below the UNAIDS target of 73% (i.e.,  $90 \times 90 \times 90$ ) to be achieved by 2020.

In addition to quantifying a community's exposure to ART, the PDV<sub>+</sub> has also been used to infer the potential for ongoing HIV transmission at the population level [2-5, 7]. However, measures such as the PDV<sub>+</sub> have been criticized by Miller et al. [10] and others [11-15] because they do not account for the relative sizes of the infected and uninfected populations (i.e., HIV prevalence). Following this work, we multiplied the PDV<sub>+</sub> by the HIV prevalence to construct a measure called the population prevalence of detectable viremia (PDV<sub>P</sub>) [11]. This measure enabled us to account for the high HIV prevalence in the AHRI study area, which increased from 26.7% to 32.4% over the 2011–2014 period. Our results show that the steady rise in the HIV prevalence offset the gains made by the declining PDV<sub>+</sub>. Thus, the PDV<sub>P</sub> only decreased by less than one percentage point, from 18.8% in 2011 to 17.9% in 2014.

We also observed significant differences in the  $PDV_+$  and  $PDV_P$  measures by sex over time. For example, the  $PDV_+$  for woman decreased by 16.5 percentage points (pp) between 2011 and 2014, from 71.8% to 55.3%, when compared with a decrease of 10.6 pp for men, from 77.8% to 67.2%. Previous research has shown that women have more frequent contact with the health-care system, due in large part to their antenatal treatment-and-care needs, where they can initiate ART early and have their viral loads monitored [26, 27]. However, because women had a higher HIV prevalence they also had a higher overall  $PDV_P$ , which decreased by 2.1 pp, from 21.3% to 19.2%, over the survey period. Importantly, we found that men had a greater increase in their HIV prevalence over time, which offset the decline in their  $PDV_+$ . Thus, the  $PDV_P$  for men actually increased by 1.6 pp, from 14.6% in 2011 to 16.2% in 2014.

We have previously exploited the substantial space-time heterogeneity in ART scale-up over eight years to demonstrate independent reductions in the individual risk of HIV acquisition with increasing ART exposure [17, 28, 29]. In more recent work, we used viral load survey data from 2011 to show that the prospective risk of HIV acquisition (5-years of follow-up) was independently associated with the  $PDV_P$  (adjusted Hazard Ratio [aHR]=1.07, p-value<0.001) but not the  $PDV_+$  (aHR=1.005, p-value=0.4) [11]. Barring substantial changes in sexual behaviour, one might expect that the minimal change in the  $PDV_P$  would translate into a minimal change in the crude HIV incidence rate. In this regard, we report elsewhere that the crude HIV incidence rate has been relatively stable in the AHRI study population between 2008 and 2016 [30, 31]. Thus, at an ecological level, the HIV incidence rate corresponds with the  $PDV_P$ , rather than declining in relation to the marked decrease in the  $PDV_+$ . These findings, and the results from our earlier work [11], provide further empirical support for the  $PDV_P$ 's utility as a measure of the potential for HIV transmission.

The  $PDV_P$  will not capture all the fundamental phenomena that underlie HIV transmission dynamics within a population. To better quantify the potential for HIV transmission, it would be ideal to use population-based surveillance systems to collect information on the number and patterns of condomless sex acts. But reliable self-report data is often difficult to obtain and not all countries will have population-based surveillance systems, which are costly to establish and maintain. Public health-care facilities can be a more affordable and convenient source of data. However, two recent

studies have shown that facility-based  $PDV_+$  measures are poor indicators of the incidence of HIV infection [11, 12].

One potential limitation of the study is that 22% of the participants refused to take an HIV test during the survey period. In a previous study, Larmarange et al.[32] found that HIV-infected participants were significantly less likely than HIV-uninfected participants to consent to an HIV test during a single survey round. This refusal rate could potentially bias both the HIV prevalence and  $PDV_P$  measures downward. However, two recent studies have confirmed that survey nonparticipation in this community did not lead to large biases in the cross-sectional estimation of the HIV prevalence [33, 34]. Further, it is unlikely that the 22% refusal rate would bias the  $PDV_+$  measure, since viral load measurements were obtained from all of the HIV-positive test results.

The  $PDV_+$  has been promoted as a proxy for ART program effectiveness. In recent years, it has gained traction in light of the UNAIDS target to have 90% of all ART-initiated persons achieve and maintain undetectable viremia by the year 2020 [6]. But while the  $PDV_+$  may reflect an infected community's exposure to ART, it may not tell us enough about the potential for HIV transmission within the general population. Recent work has therefore begun to promote the  $PDV_P$  as a more sensitive biological measure for this purpose, primarily because it accounts for the underlying prevalence of HIV [10-15]. We therefore highlight the need for countries to monitor and report the prevalence of detectable viremia among all adults ( $PDV_P$ ), irrespective of HIV status.

## **Funding**

This work was supported by two National Institute of Health (NIH) grants (R01HD084233 and R01AI124389) as well as a UK Academy of Medical Sciences Newton Advanced Fellowship (NA150161). AV and FT were also supported by a South African Medical Research Council (SA MRC) Flagship grant (MRC-RFA-UFSP-01-2013/UKZN HIVEPI). Funding for the Africa Health Research Institute's Demographic Surveillance Information System and Population-based HIV Survey was received from the Wellcome Trust. JH was supported by two NIH grants (R01AI108490 and R01AI127232). TB was supported by the Alexander von Humboldt Foundation through the endowed Alexander von Humboldt Professorship funded by the German Federal Ministry of Education and Research, as well as by the Wellcome Trust, the European Commission, the Clinton Health Access Initiative, and the National Institutes of Health's Fogarty International Center (D43-TW009775).

## **Conflicts of Interest**

The authors declare that they have no conflicts of interest.

## References

- [1] UNAIDS. Global AIDS Update. 2016.
- [2] National HIV/AIDS Strategy for the United States. National HIV/AIDS Strategy for the United States 2011. Available from: <http://www.whitehouse.gov/sites/default/files/uploads/NHAS.pdf>.
- [3] US Centers for Disease Control. Guidance on Community Viral Load: A Family of Measures, Definitions, and Method for Calculation. 2011:1-46.
- [4] Das M, Chu PL, Santos G-M, Scheer S, Vittinghoff E, McFarland W, et al. Decreases in community viral load are accompanied by reductions in new HIV infections in San Francisco. *PloS One*. 2010;5(6):e11068.
- [5] Montaner JSG, Lima VD, Barrios R, Yip B, Wood E, Kerr T, et al. Association of highly active antiretroviral therapy coverage, population viral load, and yearly new HIV diagnoses in British Columbia, Canada: a population-based study. *Lancet*. 2010;376:532-9.
- [6] UNAIDS. 90–90–90 - An ambitious treatment target to help end the AIDS epidemic. 2014.
- [7] Castel ADA, Befus M, Willis S, Griffin A, West T, Hader S, et al. Use of the community viral load as a population-based biomarker of HIV burden. *AIDS*. 2012;26:345-53.
- [8] Wood E, Kerr T, Marshall BD, Li K, Zhang R, Hogg RS, et al. Longitudinal community plasma HIV-1 RNA concentrations and incidence of HIV-1 among injecting drug users: prospective cohort study. *BMJ*. 2009;338:b1649.
- [9] Henard S, Jeanmaire E, Nguyen Y, Yazdanpanah Y, Cheret A, Hoen B, et al. Is total community viral load a robust predictive marker of the efficacy of the TasP strategy? *JAIDS Journal of Acquired Immune Deficiency Syndromes*. 2012;61(3):400-2.
- [10] Miller WC, Powers KA, Smith MK, Cohen MS. Community viral load as a measure for assessment of HIV treatment as prevention. *Lancet Infectious Diseases*. 2013;13(5):459-64.
- [11] Tanser F, Vandormael A, Cuadros D, Phillips A, de Oliveira T, Tomita A, et al. Effect of population viral load on prospective HIV incidence in a hyper-endemic rural South African community: a population-based cohort study. *Science Translational Medicine (In Press)*. 2017.
- [12] Solomon SS, Mehta SH, McFall AM, Srikrishnan AK, Saravanan S, Laeyendecker O, et al. Community viral load, antiretroviral therapy coverage, and HIV incidence in India: a cross-sectional, comparative study. *Lancet HIV*. 2016;3(4):e183-e90.
- [13] Jain V, Byonanebye DM, Liegler T, Kwarisiima D, Chamie G, Kabami J, et al. Changes in population HIV RNA levels in Mbarara, Uganda during scale-up of HIV antiretroviral therapy access. *JAIDS*. 2014;65(3):327.
- [14] Jain V, Liegler T, Kabami J, Chamie G, Clark TD, Black D, et al. Assessment of population-based HIV RNA levels in a rural east African setting using a fingerprick-based blood collection method. *Clinical Infectious Diseases*. 2012:598-605.
- [15] Kelley CF, Rosenberg ES, O'Hara BM, Frew PM, Sanchez T, Peterson JL, et al. Measuring population transmission risk for HIV: an alternative metric of exposure risk in men who have sex with men (MSM) in the US. *PloS one*. 2012;7(12):e53284.
- [16] Tanser F, Hosegood V, Bärnighausen T, Herbst K, Nyirenda M, Muhwava W, et al. Cohort Profile: Africa centre demographic information system (ACDIS) and population-based HIV survey. *International Journal of Epidemiology*. 2008;37(5):956-62.

- [17] Tanser F, Barnighausen T, Grapsa E, Zaidi J, Newell ML. High coverage of ART associated with decline in risk of HIV acquisition in rural KwaZulu-Natal, South Africa. *Science*. 2013;339(6122):966-71.
- [18] Vandormael A, de Oliveira T, Tanser F, Bärnighausen T, Herbeck J. A high percentage of undiagnosed HIV cases in a rural and hyper-endemic South African community. Under Review. 2017.
- [19] Bor J, Herbst AJ, Newell ML, Barnighausen T. Increases in adult life expectancy in rural South Africa: valuing the scale-up of HIV treatment. *Science*. 2013;339(6122):961-5.
- [20] Houlihan CF, Bland RM, Mutevedzi PC, Lessells RJ, Ndirangu J, Thulare H, et al. Cohort profile: Hlabisa HIV treatment and care programme. *International Journal of Epidemiology*. 2011;40(2):318-26.
- [21] Viljoen J, Gampini S, Danaviah S, Valea D, Pillay S, Kania D, et al. Dried blood spot HIV-1 RNA quantification using open real-time systems in South Africa and Burkina Faso. *JAIDS*. 2010;55(3):290-8.
- [22] Lohr S. Sampling: design and analysis: Nelson Education; 2009.
- [23] MPHIA. Malawi population-based HIV impact assessment, MPHIA: 2015–2016. 2016.
- [24] ZAMPHIA. Zambia population-based HIV impact assessment, ZAMPHIA: 2015–2016. 2016.
- [25] ZIMPHIA. Zimbabwe population-based HIV impact assessment, ZIMPHIA: 2015–2016. 2016.
- [26] Bor J, Rosen S, Chimbindi N, Haber N, Herbst K, Mutevedzi T, et al. Mass HIV treatment and sex disparities in life expectancy: demographic surveillance in rural South Africa. *PLoS Medicine*. 2015;12(11):e1001905.
- [27] Rosen S, Fox MP. Retention in HIV care between testing and treatment in sub-Saharan Africa: a systematic review. *PLoS Medicine*. 2011;8(7):e1001056.
- [28] Vandormael A, Newell M-L, Bärnighausen T, Tanser F. Use of antiretroviral therapy in households and risk of HIV acquisition in rural KwaZulu-Natal, South Africa, 2004–12: a prospective cohort study. *The Lancet Global Health*. 2014;2(4):e209-e15.
- [29] Oldenburg CE, Bärnighausen T, Tanser F, Iwuji CC, De Gruttola V, Seage GR, et al. Antiretroviral therapy to prevent HIV acquisition in serodiscordant couples in a hyperendemic community in rural South Africa. *Reviews of Infectious Diseases*. 2016;63(4):548-54.
- [30] Vandormael A, Dobra A, Bärnighausen T, de Oliveira T, Tanser F. Incidence rate estimation, periodic testing and the limitations of the mid-point imputation approach. *International Journal of Epidemiology*. 2017.
- [31] Tanser F, Bärnighausen T, Dobra A, Sartorius B. Identifying ‘corridors of HIV transmission’ in a severely affected rural South African population: A case for a shift toward targeted prevention strategies *International Journal of Epidemiology* 2017;In Press.
- [32] Larmarange J, Mossong J, Bärnighausen T, Newell ML. Participation dynamics in population-based longitudinal HIV surveillance in rural South Africa. *PloS one*. 2015;10(4):e0123345.
- [33] Zaidi J, Grapsa E, Tanser F, Newell M-L, Bärnighausen T. Dramatic increases in HIV prevalence after scale-up of antiretroviral treatment: a longitudinal population-based HIV surveillance study in rural KwaZulu-Natal. *AIDS (London, England)*. 2013;27(14):2301.

[34] McGovern ME, Bärnighausen T, Salomon JA, Canning D. Using interviewer random effects to remove selection bias from HIV prevalence estimates. *BMC medical research methodology*. 2015;15(1):8.

Figure Legends:

**Figure 1:** Time trends in the HIV-positive prevalence of detectable viremia (PDV<sub>+</sub>), the population prevalence of detectable viremia (PDV<sub>P</sub>), and the HIV prevalence over the 2011–2014 survey period.

**Figure 2:** Time trends in the HIV-positive prevalence of detectable viremia (PDV<sub>+</sub>), the population prevalence of detectable viremia (PDV<sub>P</sub>), and the HIV prevalence over the 2011–2014 survey period for males (Panel A) and females (Panel B).

## Tables

**Table 1:** Summary statistics for the HIV-positive population only and the entire population (HIV-positive and HIV-negative participants) for the 2011, 2013, and 2014 survey years.

	Year					
	2011		2013		2014	
<i>HIV-positive population</i>						
Dried blood spot samples, N	2,401		2,510		2,611	
Successful viral load measurements, N (%)	2,366	98.54	2,135	85.06	2,251	86.21
Viral loads > 1,550 copies/mL	1,663		1,304		1,237	
HIV-positive prevalence of detectable viremia (PDV <sub>+</sub> )						
Unadjusted, Mean (95% CI)	70.29	(66.95–73.75)	61.08	(57.81–64.48)	54.95	(51.93–58.1)
Age-sex adjusted, Mean (95% CI)	73.76	(68.77–79.26)	64.38	(59.63–69.64)	59.90	(54.98–65.37)
Female, N (%)	1,877	79.33	1,690	79.16	1,794	79.70
Age, Median (IQR)	35	(27–45)	35	(27–44)	35	(28–45)
<i>HIV-positive and HIV-negative population</i>						
Observations, N	8,626		6,881		6,660	
Population prevalence of detectable viremia (PDV <sub>P</sub> )						
Unadjusted, Mean (95% CI)	19.28	(18.36–20.23)	18.95	(17.94–20.01)	18.57	(17.55–19.64)
Age-sex adjusted, Mean (95% CI)	18.83	(17.94–19.76)	18.80	(17.79–19.85)	17.91	(16.92–18.95)
HIV prevalence						
Unadjusted, Mean (95% CI)	27.43	(26.33–28.56)	31.03	(29.73–32.37)	33.80	(32.42–35.22)
Age-sex adjusted, Mean (95% CI)	26.73	(25.66–27.83)	30.64	(29.35–31.97)	32.36	(31.03–33.73)
Female, N (%)	5,832	67.61	4,775	69.39	4,730	71.02
Age, Median (IQR)	31	(21–47)	30	(20–47)	31	(21–47)

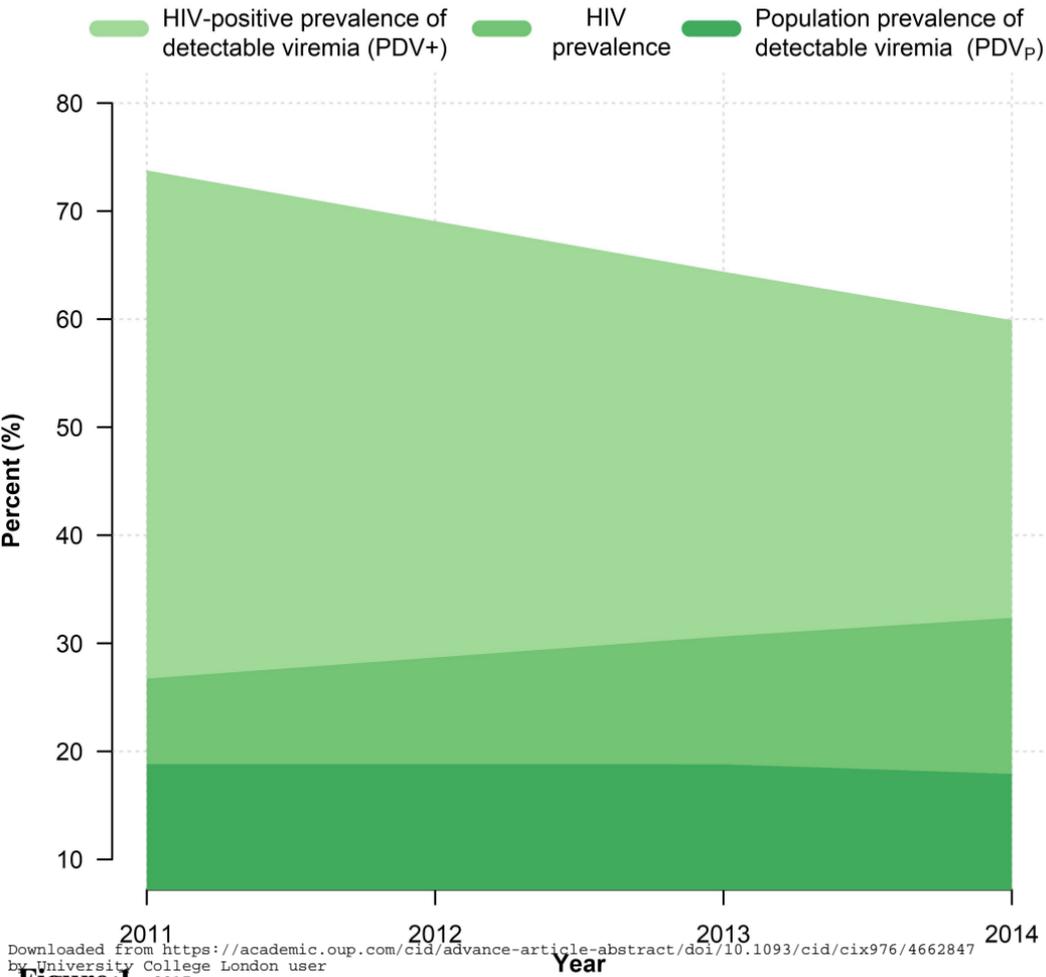
Table 1 shows unadjusted and age-sex adjusted results for the PDV<sub>+</sub>, PDV<sub>P</sub>, and HIV prevalence measures. The unadjusted PDV<sub>P</sub> is obtained by multiplying the PDV<sub>+</sub> by the HIV prevalence. For example, in 2011, there were 1,663 HIV-positive participants with a viral load >1,550 copies/mL. Therefore, the unadjusted PDV<sub>+</sub> = 1,663/2,366 = 70.29%, the HIV prevalence = 2,366/8,626 = 27.43%, and the PDV<sub>P</sub> = 1,663/8,626 = 19.28%. Multiplying the PDV<sub>+</sub> by the HIV prevalence (*H*) returns the PDV<sub>P</sub>: PDV<sub>+</sub> × *H* = 70.29% × 27.43% = 19.28%. We also report the age- and sex-adjusted PDV<sub>+</sub>, PDV<sub>P</sub>, and HIV prevalence measures.

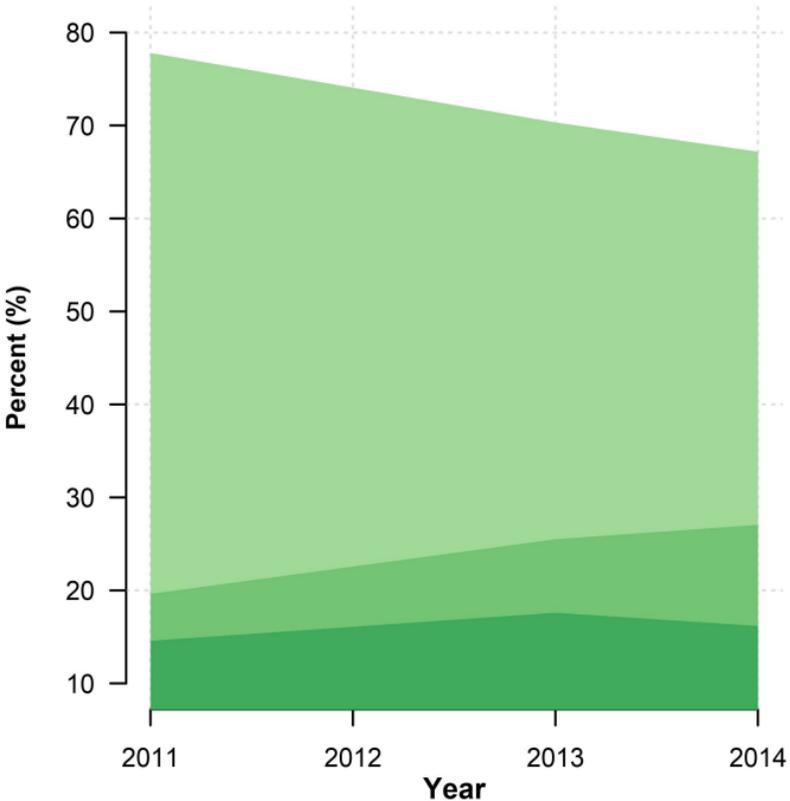
**Table 2:** Regression results showing the relative odds (odds ratio) of a detectable viral load for the HIV-positive population (PDV<sub>+</sub>), adjusting for year, age and sex.

	Model 1			Model 2			Model 3			Model 4		
	OR	95% CI	p-value									
Year (Ref: 2011)												
2013	0.647	(0.575,0.727)	<0.001	0.646	(0.575,0.726)	<0.001	0.649	(0.577,0.729)	<0.001	0.749	(0.565,0.993)	0.044
2014	0.490	(0.436,0.551)	<0.001	0.49	(0.436,0.551)	<0.001	0.495	(0.44,0.556)	<0.001	0.498	(0.379,0.654)	<0.001
Female				0.700	(0.611,0.801)	<0.001	0.680	(0.594,0.779)	<0.001	0.721	(0.573,0.908)	0.005
Age (>25 years)							0.605	(0.521,0.702)	<0.001	0.605	(0.521,0.702)	<0.001
2013 * Female										0.839	(0.616,1.144)	0.266
2014 * Female										0.993	(0.735,1.343)	0.966
Constant	2.508	(2.294,2.741)	<0.001	3.33	(2.89,3.837)	<0.001	5.167	(4.26,6.268)	<0.001	4.922	(3.848,6.295)	<0.001
HIV tests, N	6,752			6,752			6,752			6,752		
Participants, N	4,991			4,991			4,991			4,991		

**Table 3:** Regression results showing the relative odds (odds ratio) of a detectable viral load for the HIV-positive and HIV-negative population (PDV<sub>p</sub>) by year, adjusting for sex and age.

	Model 1			Model 2			Model 3			Model 4		
	OR	95% CI	p-value									
Year (Ref: 2011)												
2013	0.968	(0.908,1.031)	0.310	0.964	(0.905,1.026)	0.248	0.962	(0.903,1.024)	0.222	1.163	(1.018,1.328)	0.026
2014	0.911	(0.85,0.977)	0.009	0.904	(0.844,0.968)	0.004	0.878	(0.82,0.94)	<0.001	1.049	(0.903,1.219)	0.532
Female				1.708	(1.568,1.861)	<0.001	1.461	(1.337,1.595)	<0.001	1.674	(1.487,1.884)	<0.001
Age (>25 years)							3.029	(2.761,3.322)	<0.001	3.036	(2.767,3.33)	<0.001
2013 * Female										0.785	(0.675,0.912)	0.002
2014 * Female										0.798	(0.674,0.944)	0.009
Constant	0.246	(0.234,0.258)	<0.001	0.169	(0.156,0.183)	<0.001	0.089	(0.081,0.097)	<0.001	0.080	(0.071,0.089)	<0.001
HIV tests, N	22,167			22,167			22,167			22,167		
Participants, N	16,319			16,319			16,319			16,319		



**A: Males****B: Females**