Brief Assessment of Schizotypal Traits: A Multinational Study

Eduardo Fonseca-Pedrero*

Department of Educational Sciences, University of La Rioja, Logroño, Spain Center for Biomedical Research in the Mental Health Network (CIBERSAM), Oviedo, Spain

Javier Ortuño-Sierra

Department of Educational Sciences, University of La Rioja, Logroño, Spain

Beatriz Lucas Molina

Department of Psychology, University of Valencia, Valencia, Spain

Martin Debbané,

Faculty of Psychology and Educational Sciences, University of Geneva, Geneva, Switzerland Department of Clinical, Educational and Health Psychology, University College London, London, UK

Raymond C.K. Chan

Neuropsychology and Applied Cognitive Neuroscience Laboratory, CAS key Laboratory of Mental Health, Beijing, China

David C. Cicero

Department of Psychology, University of Hawaii at Manoa, Honolulu, HI, USA

Lisa C. Zhang

Department of Psychology, University of British Columbia, Vancouver, BC, Canada

Colleen Brenner

Department of Psychology, University of British Columbia, Vancouver, BC, Canada

Emma Barkus

School of Psychology, University of Wollongong, Wollongong, Australia

Richard J. Linscott

Department of Psychology, University of Otago, Dunedin, New Zealand

Thomas Kwapil

Department of Psychology, University of North Carolina at Greensboro, Greensboro, NC, USA

Neus Barrantes-Vidal

Department of Clinical and Health Psychology, Universitat Autònoma de Barcelona, Barcelona, Spain

Alex Cohen, PhD

Department of Psychology, Louisiana State University, Louisiana, LA, USA

Adrian Raine, PhD

Departments of Criminology, Psychiatry, and Psychology, University of Pennsylvania, Philadelphia, PA, USA

Michael T. Compton, M.D.

Department of Psychiatry, Lenox Hill Hospital, New York, NY, USA

Erin B. Tone, PhD

Department of Psychology, Georgia State University, Atlanta, GA, USA

Julie Suhr, PhD Department of Psychology, Ohio University, Athens, OH, USA

Julio Bobes

Department of Psychiatry, University of Oviedo, Oviedo, Spain Center for Biomedical Research in the Mental Health Network (CIBERSAM), Oviedo, Spain

Axit Fumero, PhD

Department of Psychology, University of La Laguna, Santa Cruz de Tenerife, Spain

Stella Giakoumaki, PhD

Department of Psychology, University of Crete, Rethymno, Greece

Ioannis Tsaousis, PhD

Department of Psychology, University of Crete, Rethymno, Greece

Antonio Preti, PhD

Genneruxi Medical Center, Cagliari, Italy

Michael Chmielewski, PhD

Department of Psychology, Southern Methodist University, Dallas TX, USA

Julien Laloyaux, PhD

Department of Biological and Medical Psychology, University of Bergen, Bergen, Norway NORMENT – Norwegian Center of Excellence for Mental Disorders Research, University of Oslo, Oslo, Norway

Psychology and Neuroscience of Cognition Research Unit, University of Liège, Liège, Belgium

Anwar Mechri, PhD

Psychiatry Department, University Hospital of Monastir, Monastir, Tunisia

Mohamed Aymen Lahmar, PhD

Psychiatry Department, University Hospital of Monastir, Monastir, Tunisia

Viviana Wuthrich, PhD

Centre for Emotional Health, Department of Psychology, Macquarie University, Sydney, Australia

Frank Larøi

Department of Biological and Medical Psychology, University of Bergen, Bergen, Norway NORMENT – Norwegian Center of Excellence for Mental Disorders Research, University of Oslo, Oslo, Norway

Psychology and Neuroscience of Cognition Research Unit, University of Liège, Liège, Belgium

Johanna C. Badcock

Centre for Clinical Research in Neuropsychiatry, Division of Psychiatry, Faculty of Health and Medical Sciences, University of Western Australia, Perth, Australia

Assen Jablensky

Centre for Clinical Research in Neuropsychiatry, School of Psychiatry and Clinical Neurosciences, University of Western Australia, Perth, Australia

David Barron

Centre for Psychological Medicine, Perdana University, Serdang, Malaysia

Viren Swami

Department of Psychology, Anglia Ruskin University, Cambridge, UK Centre for Psychological Medicine, Perdana University, Serdang, Malaysia

Ulrich S. Tran

Department of Basic Psychological Research and Research Methods, School of Psychology, University of Vienna, Vienna, Austria

Martin Voracek

Department of Basic Psychological Research and Research Methods, School of Psychology, University of Vienna, Vienna, Austria

Corresponding Author: Eduardo Fonseca-Pedrero University of La Rioja C/ Luis de Ulloa, 2, Edificio VIVES; C.P: 26004, Logroño, La Rioja, Spain

Tel:+34 941 299 031 Fax: +34 941 299 333

e-mail: eduardo.fonseca@unirioja.es

Abstract

The Schizotypal Personality Questionnaire-Brief (SPQ-B) was developed with the

aim of examining variations in healthy trait schizotypy, as well as latent vulnerability to

psychotic-spectrum disorders. No previous study has studied the cross-cultural validity

of the SPQ-B in a large cross-national sample. The main goal of the present study was to

analyze the reliability and the internal structure of SPQ-B scores in a multinational sample

of 28,426 participants recruited from 14 countries. The mean age was 22.63 years (SD =

7.08; range 16-68 years), 37.7% (n = 10,711) were men. The omega coefficients were

high, ranging from 0.86 to 0.92 for the total sample. Confirmatory factor analysis revealed

that SPQ-B items were grouped either in a theoretical structure of three first-order factors

(Cognitive-Perceptual, Interpersonal, and Disorganized) or in a bifactor model (three

first-order factors plus a general factor of schizotypal personality). In addition, the results

supported configural but not strong measurement invariance of SPQ-B scores across

samples. These findings provide new information about the factor structure of schizotypal

personality, and support the validity and utility of the SPQ-B, a brief and easy tool for

assessing self-reported schizotypal traits, in cross-national research. Theoretical and

clinical implications for diagnostic systems, psychosis models, and cross-national mental

health strategies are derived from these results.

Keywords: Schizotypy; Schizotypal personality; Psychosis; Cross-cultural; SPQ-B;

Psychosis risk

1. Introduction

In the past two decades, the early and reliable identification of individuals potentially at-risk for psychotic-spectrum disorders, based on psychometric indices, has become a focus of extensive and expanding research and debate (Addington et al., 2015; Fonseca-Pedrero et al., 2016b; Fusar-Poli et al., 2014; Kline and Schiffman, 2014; Mason, 2015). The identification of specific subgroups of individuals at high risk for psychotic-spectrum disorders may help us to elucidate risks and protective factors, as well as etiological mechanisms and developmental pathways that mitigate, delay, or even prevent the onset of clinically significant psychotic disorders (Barrantes-Vidal et al., 2015).

Schizotypal traits are considered a phenotypic-indicator of schizotypy (Meehl, 1962), a latent personality organization reflecting a putative liability for schizophrenia-spectrum disorders (Barrantes-Vidal et al., 2015; Fonseca Pedrero and Debbané, 2017; Lenzenweger, 2010). Schizotypal traits encompass anomalies and deficits across cognitive (e.g., paranoid ideation, ideas of reference), social/emotional (e.g., anhedonia, no close friends), and behavioural (e.g., odd behaviour and language) systems (Cohen et al., 2015; Fonseca-Pedrero et al., 2017). Previous findings support the notion of assumed phenomenological, temporal, and etiological continuity between the subclinical and clinical psychosis phenotype and lend validity to the concept of schizotypal traits (Cohen et al., 2015; Ettinger et al., 2014; Linscott and van Os, 2013).

Several measurement instruments allow clinicians and researchers to document the presence, frequency, and severity of schizotypal traits (Fonseca-Pedrero et al., 2016b; Mason, 2015). These tools have been developed with the aim of examining variation in healthy trait schizotypy as well as latent vulnerability to psychotic-spectrum disorders in both clinical and non-clinical population (e.g., general population, clinical, and genetic high risk samples). The Schizotypal Personality Questionnaire (SPQ) (Raine, 1991), in its brief version (SPQ-B) (Raine and Benishay, 1995), or its brief revised version (SPQ-BR) (Cohen et al., 2010), measure a broad range of psychotic-like traits—originally nine identified subordinate traits based on the operational definition of Schizotypal Personality Disorder (SPD) (American Psychiatric Association, 1987), and is among the more widely-used measured of this type.

The SPQ-B has been used with patients and relatives of patients with schizophrenia-spectrum disorders (Compton et al., 2007; Moreno-Izco et al., 2015), adolescents (Fonseca-Pedrero et al., 2009), twins (Ericson et al., 2011), outpatients

(Axelrod et al., 2001), and college students (Compton et al., 2009a; Fonseca-Pedrero et al., 2011; Mata et al., 2005; Raine and Benishay, 1995). The psychometric properties of the SPQ-B have been examined previously. For instance, the reliability of scores and several sources of evidence of validity have been demonstrated (e.g., Fonseca-Pedrero et al., 2016b; Mason, 2015). Moreover, translations of the measure have been validated in several countries (e.g., France, China, Spain, Turkey, Switzerland, etc.) (e.g., Aycicegi et al., 2005; Ma et al., 2015; Ortuño-Sierra et al., 2013).

Examination of the SPQ-B factor structure has yielded factorial solutions of two (Aycicegi et al., 2005), three (Compton et al., 2009a; Fonseca-Pedrero et al., 2011, 2009; Ma et al., 2015; Mata et al., 2005; Ortuño-Sierra et al., 2013; Tran et al., 2015), and four factors (Cohen et al., 2010; Fonseca-Pedrero et al., 2010). The three-factor model characterized by Cognitive-Perceptual (e.g., hallucinations, ideas of reference, magical thinking or paranoid ideation), Interpersonal (e.g., blunted affect, social anxiety or lack of close friends), and Disorganized (e.g., odd behavior and speech) dimensions has been widely replicated across studies. However, although the underlying structure of schizotypal personality, as assessed via the SPQ-B, has been analyzed, previous research has produced some contradictory results. These mixed findings are partially explained by variations in sampling method (random, convenience), sample characteristics (clinical, non-clinical, and country), and the data-analytic approach employed (exploratory vs. confirmatory factor analysis).

To the best of our knowledge, no previous studies has validated the psychometric quality of SPQ-B scores across multiple countries. For instance, we have little information about the factorial structure of SPQ-B scores and its possible variation across countries, particularly non-Western countries. Moreover, as previous studies have demonstrated with the SPQ, alternative models (e.g., Barron et al., 2017; Preti et al., 2015) may better explain the latent structure of SPQ-B scores. Thus, it is important to gather new information about the validity of this tool through cross-cultural research and collaborative multinational studies. Furthermore, and despite the globalization of psychosis research, no previous study has analyzed the psychometric quality of psychosis risk screeners in multinational samples.

The purpose of the present study was to analyze the psychometric properties of SPQ-B scores in a large sample recruited from 14 countries. Derived from this main goal are the following specific objectives: a) to estimate the reliability of SPQ-B scores across countries; b) to study the internal structure of SPQ-B scores across countries; and c) to

analyze the measurement invariance of SPQ-B scores across countries. We hypothesized that the three-factor model of the SPQ-B would have adequate goodness-of-fit indices across samples. Moreover, we hypothesized that new measurement models, such as a bifactor model, would fit adequately. In addition, we further hypothesized that SPQ-B scores would show configural measurement invariance across samples.

2. Method

2.1. Participants

Participants were gathered from 24 sites across 14 countries (Australia, Austria, Belgium, Canada, China, Germany, Greek, Italy, Mauritius, New Zealand, Spain, Tunisia, United States of America, and United Kingdom). Data from the present study, focused on reporting of full SPQ scores, has been published elsewhere (Fonseca-Pedrero et al., 2017) and the present study focused specifically on the SPQ-B. The overall sample consisted of 28,426 participants. The mean age was 22.63 years (SD = 7.08; range 16-68 years). A total of 14.5% (n = 4,113) of participants did not provide age. Participant were 10,711 males (37.7%) and 17,208 females (60.5%); 507 (1.8%) did not specify sex. Thus, 27,919 (98.2%) participants reported sex and 22,888 (80.52%) reported age. In this study, we considered information at the country, and not research site, level. Information about the age, sex, and other participant characteristics are reported in Table 1. Information about sampling procedures and demographic characteristics of the samples across sites are presented in the Supplementary Materials.

-----Insert Table 1-----

2.2. Instrument

2.2.1. The Schizotypal Personality Questionnaire-Brief (SPQ-B)

The SPQ-B provides a common index of schizotypal traits across all countries. The SPQ-B is a 22-item (*True/False*) self-report scale based on the SPQ (Raine, 1991) for the assessment of SPD traits as defined by *DSM-III-R* diagnostic criteria (American Psychiatric Association, 1987). The SPQ-B includes items that fall within three domains: Cognitive-Perceptual (ideas of reference, paranoid ideation, magical thinking, and unusual perceptual experiences), Interpersonal (social anxiety, no close friends, blunted affect, and paranoid ideation), and Disorganized (odd speech and behavior). In the present study, the items of the brief version were extracted from the original SPQ validated for

each country. Item selection was based on the original brief SPQ: English (Raine, 1991), Spanish (Fonseca-Pedrero et al., 2014b), Italian (Fossati et al., 2003), Chinese (Chen et al., 1997), Arabic (Lahmar et al., 2014), French (Dumas et al., 2000), Creole (Reynolds et al., 2000), Greek (Tsaousis et al., 2015), and German version (Klein et al., 1997).

2.3. Procedure

Conventions for obtaining informed consent required by each investigator's research institution, as well as IRB or ethical committees were followed. All participants provided written informed consent prior to participation. The study was conducted in accordance with the guidelines of the Declaration of Helsinki (World Medical Association, 2013). In the present study the SPQ-B scores being reported are derived from the administration of the full 74 item SPQ (see Fonseca-Pedrero et al., 2017). Similarly, the SPQ was sometimes administered in the context of larger studies (see Supplemental Material for further information).

2.4. Data analyses

Descriptive statistics for the items of the SPQ-B items were calculated as the first step. In order to test the reliability of SPQ-B scores, and due to the limitations of Cronbach's α (Dunn et al., 2014), coefficient ω was estimated (Zinbarg et al., 2005). Next, in order to analyse the internal structure of SPQ-B scores, and based on previous literature, several confirmatory factor analyses (CFAs) were conducted at the item level. Considering the categorical nature of the data, we used the robust mean-adjusted weighted least square method (WLSMV) for parameter estimation (Muthén and Muthén, 1998-2012). The following goodness-of-fit indices were used: Chi-square (χ^2), Comparative Fit Index (CFI), Tucker-Lewis Index (TLI), Root Mean Square Error of Approximation (RMSEA), and Weighted Root Mean Square Residual (WRMR). CFI and TLI values greater than .95 are preferred and those close to .90 are considered acceptable; RMSEA values should be under .08 for a reasonable fit, and under .05 for a good fit, whereas WRMR values less than .08 are considered evidence of good model (Brown, 2006; Hu and Bentler, 1999).

Taking into account previous studies, different measurement models were tested: a) a unidimensional model; b) a bidimensional solution with a Cognitive-Perceptual, and a Negative factor (Siever and Gunderson, 1983); c) the Raine et al. (1994) model that includes Cognitive-Perceptual, Interpersonal, and Disorganized dimensions with Items 7,

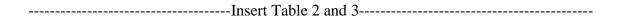
9, 14, and 17 overlapping (i.e., cross-loading) in both the Cognitive-Perceptual and Interpersonal dimensions; d) the Raine and Benishay (1995) three-factor solution with no item cross-loadings allowed, and; e) a bifactor model that includes a general factor of schizotypal personality and three first order factors (Cognitive-Perceptual,, Interpersonal, and Disorganized). Correlations among error terms were not permitted. Finally, and with the aim of studying measurement invariance across countries, we conducted successive multi-group CFAs models (MGCFAs models) for categorical outcomes (Muthén and Asparouhov, 2002).

The relatively few missing values in the data were replaced by regression-based estimates, to which an error component was added, based on the SPSS Missing Value Analysis module. SPSS 22.0 (IBM Corp Released, 2013), Mplus 7.4 (Muthén and Muthén, 1998-2012), FACTOR 10.5 (Ferrando and Lorenzo-seva, 2017), and R (R Development Core Team, 2011) were used for the data analyses.

3. Results

3.1 Descriptive statistics and internal consistency of the SPQ-B scores

Means and standard deviations for the SPQ-B items for all countries are shown in Table 2. Internal consistency values for SPQ-B scores in the total sample and by country are shown in Table 3. Omega coefficients were adequate for data from all participating countries. Values for the total sample were 0.86, 0.91, 0.89, and 0.92 for the Cognitive-Perceptual, Interpersonal, and Disorganized subscales, and the Total score, respectively. Across countries, values ranged from 0.77 (Cognitive-Perceptual for China) to 0.94 (total score for the United States, Interpersonal and Disorganization for Germany).



3.2. Internal structure of schizotypal traits

Goodness-of-fit indices for the analyzed models are presented in Table 4. As can be seen, the models that showed the best fit in all the countries were the bifactor and Raine et al. (1994) models (models c and e). The bifactor model displayed better goodness-of-fit indices, but, as explained below, the factor loadings in this solution revealed some inconsistencies. It is worth noting that, in several countries, some of the goodness-of-fit indices such as CFI and TLI were close to the standard cut-off values, but still inadequate.

In particular, values of CFI lower than .90 were observed in both models, especially in the model of Raine et al. Nonetheless, RMSEA values in both factorial solutions were good for all of the countries analyzed. As noted by Yu (2002), the RMSEA index may be preferred for analysis with the WLSMV estimator and ordered categorical variables. Thus, by this standard, the goodness-of-fit indices for the analyzed models could be considered adequate.

Tables 5 and 6 show the factor loadings for each of the 22 items for the Raine et al. (1994) and the bifactor models, respectively. In addition, the means and range of the factor loadings for the SPQ-B items in the two models are presented. In the case of the Raine et al. (1994) model, correlations among the latent variables were calculated, with averages of 0.561 (Cognitive-Perceptual-Disorganized), 0.286 (Positive-Interpersonal), and 0.593 for the total sample. As can be seen, some factor loadings on the latent factors of the bifactor model were negative and nominally not significant, thus suggesting that this model could be further improved. Factor loadings for the Raine et al. (1994) model were all adequate and statistically significant.

-----Insert Table 5 and 6-----

3.3. Measurement invariance of the SPQ-B scores across countries

Measurement invariance across all participating countries was studied for the two models that displayed best fit, namely the Raine et al. (1994) model (χ^2 = 19973.89; df = 2828; CFI = 0.912; TLI = 0.90; RMSEA = 0.055, with 95% CI: 0.054-0.055; WRMR = 8.62) and the bifactor models (χ^2 = 14564.89; df = 2618; CFI = 0.938; TLI = 0.924; RMSEA = 0.047 with 95% CI: 0.047-0.048; WRMR = 7.01). The configural invariance model, in which no equality constraints were imposed, showed an adequate fit to the data for both models. Next, a strong invariance model was tested with the item thresholds and factor loadings constrained to equality across groups. The Δ CFI between the constrained and the unconstrained models was over 0.01, indicating that strong invariance was not supported in the case of the bifactor model (χ^2 = 23498.71; df = 3086; CFI = 0.895; TLI = 0.890; RMSEA = 0.057 with 95% CI: 0.056-0.058; WRMR = 9.80). For the Raine et al. (1994) model, no convergence was found and the program did not allow us to calculate strong invariance parameters. The Δ CFI between the constrained and the unconstrained models was over 0.01, indicating that strong invariance was not supported. Hence, the

results support configural invariance, whereas strong measurement invariance of the SPQ-B across the 14 countries studied was not tenable.

4. Discussion

The psychometric assessment of schizotypal traits offers distinctive benefits, such as being relatively inexpensive, non-invasive, and useful for screening large samples of the general population, as well as for identifying participants at increased risk for psychosis (e.g., Fonseca-Pedrero et al., 2016b; Lenzenweger, 2010; Mason, 2015). For these purposes, and in tandem with global mental health research strategies, there is a clear need for psychometrically sound tools for both psychosis risk and schizotypal screening, which are validated across countries, to use in international research studies and diverse cultural settings. To date, no study has attempted to validate the SPQ-B in a cross-national sample. Furthermore, it remains unclear whether the factorial structure underlying SPQ-B scores is invariant across multiple countries. Thus, the main goal of the present study was to analyse the reliability, internal structure and measurement invariance by country of SPQ-B scores in a multinational sample of participants recruited from 14 countries.

Our analyses highlighted several important findings. First, SPQ-B scores showed adequate levels of internal consistency across countries. The reliability of SPQ-B scores, estimated with coefficient omega, was generally above 0.8. This research provides further support for the reliability of the SPQ-B scores, extending previous findings to non-clinical samples from different countries and variable study contexts. Thus, the SPQ-B could be used as a screening instrument to identify individuals who may be at increased risk for psychosis-spectrum disorders as well as to examine variations in healthy trait schizotypy in cross-cultural studies.

Second, examination of the factorial structure underlying the SPQ-B scores indicated that schizotypal traits have a multidimensional, rather than unidimensional, structure. SPQ-B items were grouped, in the present analysis, in a theoretical structure of three first-order factors (i.e., Cognitive-Perceptual, Interpersonal, and Disorganization dimensions) as well as in a bifactor model (three first-order factors plus general factor of schizotypal personality). In fact, this is the first study to show that it is possible to derive a total score for the SPQ-B and to obtain distinct subscores for the three classic schizotypal dimensions. Schizotypal personality is a multifaceted construct phenotypically similar to that found in patients with psychosis (e.g., Liddle, 1987). Just

as the manifestation of schizophrenia is heterogeneous – encompassing a broad range of emotional, cognitive, perceptual, social and behavioral functions – schizotypy involves a diverse set of traits. Numerous studies, using the SPQ-B or its brief versions, have obtained evidence of such a three-factor structure for schizotypal personality (Compton et al., 2009a; Fonseca-Pedrero et al., 2011, 2009; Ma et al., 2015; Mata et al., 2005; Ortuño-Sierra et al., 2013; Tran et al., 2015), consistent with the Raine et al. (1994) model. Furthermore, the present results corroborate those found when comparing SPQ scores across samples (e.g., Bora and Arabaci, 2009; Compton et al., 2009b; Fonseca-Pedrero et al., 2016a; Fonseca-Pedrero et al., 2017; Fossati et al., 2003; Raine et al., 1994; Reynolds et al., 2000). Futhermore, this factorial structure is similar to those found in the new measure of schizotypy named the Multidimensional Schizotypy Scale (MSS) (Kwapil et al., in press).

Third, multigroup CFA showed that the SPQ-B three-factor model had configural, but not strong measurement invariance, across countries. Similar results have been found in prior research using the SPQ and its brief versions, as well as other schizotypy tools (e.g., the short form of the Oxford-Liverpool Inventory of Feelings and Experiences and Chapman's scales of psychosis proneness) (Cicero, 2015; Fonseca-Pedrero et al., 2015, 2014a; Kwapil et al., 2012; Ortuño-Sierra et al., 2013). For instance, Ortuño-Sierra et al. (2013), when comparing the factorial equivalence of the SPQ-B between Spanish and Swiss adolescents, found that SPQ-B scores had configural and partial strong invariance across the two samples. In addition, the present results demonstrated that several items showed differential functioning by country. To date, differential item functioning (DIF) for psychosis risk or schizotypy measures has yet to be thoroughly addressed. In crosscultural research, it is vital to test whether varied groups show differing probabilities of success on (or likelihood of endorsing) an item after matching on the underlying construct (e.g., schizotypy) that the item is intended to measure (Byrne et al., 2009; Zumbo, 2007). DIF is of particular importance in international, comparative, and cross-cultural research particularly in efforts to ensure fairness and equity in testing (Zumbo, 2007). The present findings suggest that some schizotypal traits reflecting emotion, behavior, and cognition may differ across countries, at least those that were included in the present study. In fact, schizotypal traits assessed in different cultures have the potential to provide us with information about cultural variations in social and affective functioning (Cohen et al., 2015). Similar results have been found when psychotic symptoms or psychotic-like experiences are analyzed in samples recruited around the world (Largi et al., 2014; Nuevo et al., 2012; Woods et al., 2014). The finding of configural measurement equivalence across cultures provides essential evidence of construct validity for the schizotypal dimensions, as well as evidence of the cross-cultural validity of SPQ-B scores; however, examination of DIF by sex, age, and language will be an important next step in future studies.

The results of the present study should be considered in light of the following limitations. First, there is an inherent problem in the use of self-reports as indirect indicators of schizotypal traits. Second, the nature of the sample, composed of a majority of college students, precludes the generalization of the results to other populations of interest. Third, the fact that not all the samples employed the infrequency response to detect those participants who displayed random or pseudo-random patterns of responses may undermine the validity and generalizability of the results found in the present crossnational study. Finally, in the present study, the items of the SPQ-B were extracted from the original full version of the SPQ.

5. Conclusions

We have provided the first comprehensive validation study of the SPQ-B using a large, multinational sample from 14 countries. These results provides new information about the brief assessment of schizotypal traits using the same psychometric tool and analytic procedures to compare results obtained in different countries and linguistic groups. In addition, our results demonstrated that schizotypal personality is composed, at a minimum, of three dimensions (i.e., Cognitive-Perceptual, Interpersonal, and Disorganized), and is perhaps encompassed by a general schizotypal factor. The results derived from this cross-national study have theoretical and clinical implications for diagnostic systems, psychosis models, and cross-national mental health strategies.

References

- Addington, J., Stowkowy, J., Weiser, M., 2015. Screening tools for clinical high risk for psychosis. Early Interv. Psychiatry. 9, 345-356.
- American Psychiatric Association, 1987. Diagnostic and statistical manual of mental disorders (3rd ed. revised) (DSM-III-R), Washington, DC: APA.
- Axelrod, S.R., Grilo, M.C., Sanislow, C., McGlashan, T.H., 2001. Schizotypal Personality Questionnaire-Brief: Factor structure and convergent validity in inpatient adolescent. J. Pers. Disord. 15, 168–179.
- Aycicegi, A., Dinn, W.M., Harris, C.L., 2005. Validation of Turkish and English versions of the Schizotypal Personality Questionnaire-B. Eur. J. Psychol. Assess. 21, 34–43.
- Barrantes-Vidal, N., Grant, P., Kwapil, T., 2015. The role of schizotypy in the study of the etiology of schizophrenia spectrum disorders. Schizophr. Bull. 41, S408-416.
- Barron, D., Morgan, K. D., Towell, T., Jaafar, J. L., Swami, V. 2017. Psychometric properties of the Malay Schizotypal Personality Questionnaire: Measurement invariance and latent mean comparisons of Malay and Chinese adults. Asia-Pacific Psychiatry, in press.
- Bora, E., Arabaci, L.E., 2009. Confirmatory Factor Analysis of Schizotypal Personality Traits in University Students. Turkish J. Psychiatry. 20, 339–345.
- Brown, T.A., 2006. Confirmatory factor analysis for applied research. Guilford Press, New York.
- Byrne, B.M., Oakland, T., Leong, F.T., van de Vijver, F.J., Hambleton, R., Cheung, F.M., Bartram, D., 2009. A critical analysis of cross-cultural research and testing practices: Implications for improved education and training in psychology. Train. Educ. Prof. Psychol. 3, 29–105.
- Chen, W.J., Hsiao, C.K., Lin, C.C.H., 1997. Schizotypy in community samples: The three-factor structure and correlation with sustained attention. J. Abnorm. Psychol. 106, 649–654.
- Cicero, D.C., 2015. Measurement Invariance of the Schizotypal Personality

 Questionnaire in Asian, Pacific Islander, White, and Multiethnic Populations.

 Psychol. Assess.
- Cohen, A., Mohr, C., Ettinger, U., Chan, R.C.K., Park, S., 2015. Schizotypy as an organizing framework for social and affective sciences. Schizophr. Bull. 41, S427-

- Cohen, A.S., Matthews, R.A., Najolia, G.M., Brown, L.A., 2010. Toward a more psychometrically sound brief measure of schizotypal traits: introducing the SPQ-Brief Revised. J. Pers. Disord. 24, 516–537.
- Compton, M.T., Chien, V.H., Bollini, A., 2007. Psychometric properties of the Brief version of the Schizotypal Personality Questionnaire in relatives with schizophrenia-spectrum disorders and non-psychotic control. Schizophr. Res. 91, 122–131.
- Compton, M.T., Goulding, S.M., Bakeman, R., McClure-Tone, E.B., 2009a. An examination of the factorial structure of the Schizotypal Personality Questionnaire-Brief (SPQ-B) among undergraduate students. Schizophr. Res. 115, 286–289.
- Compton, M.T., Goulding, S.M., Bakeman, R., McClure-Tone, E.B., 2009b.

 Confirmation of a four-factor structure of the Schizotypal Personality

 Questionnaire among undergraduate students. Schizophr. Res. 111, 46–52.
- Dumas, P., Bouafia, S., Gutknecht, C., Saoud, M., Dalery, J., d'Amato, T., 2000. [Validation of the French version of the Raine Schizotypal Personality Disorder Questionnaire--categorial and dimensional approach to schizotypal personality traits in a normal student population]. Encephale 26, 23–29.
- Dunn, T.J., Baguley, T., Brunsden, V., 2014. From alpha to omega: a practical solution to the pervasive problem of internal consistency estimation. Br. J. Psychol. 105, 399–412.
- Ericson, M., Tuvblad, C., Raine, A., Young-Wolff, K., Baker, L.A., 2011. Heritability and longitudinal stability of schizotypal traits during adolescence. Behav. Genet. 41, 499–511.
- Ettinger, U., Meyhöfer, I., Steffens, M., Wagner, M., Koutsouleris, N., 2014. Genetics, cognition, and neurobiology of schizotypal personality: a review of the overlap with schizophrenia. Front. Psychiatry 5, 18.
- Ferrando, P.J., Lorenzo-seva, U. 2017. Program FACTOR at 10: Origins, development and future directions. Psicothema. 29, 236–240.
- Fonseca-Pedrero, E., Compton, M., Tone, E.B., Ortuño-Sierra, J., Paino, M., Fumero A., Lemos-Girádez, S., 2014a. Cross-cultural invariance of the factor structure of the Schizotypal Personality Questionnaire across Spanish and American college students. Psychiatry Res. 30, 1071–1076.
- Fonseca-Pedrero, E., Debbané, M, Ortuño-Sierra, J., Chan, R.C.K., Cicero, D.C.,

- Zhang, L.C., Brenner, C., Barkus, E., Linscott, R.J., Kwapil, T., Barrantes-Vidal, N., Cohen, A., Raine, A., Compton, M.T., Tone, E.B., Suhr, J., Muñiz, J., Fumero, A., Giakoumaki, S., Tsaousis, I., Preti, A., Chmielewski, M., Laloyaux, J., Mechri, A., Lahmar, M.A., Wuthrich, V., Larøi, F., Badcock, J.C., Jablensky, A., 2017. The structure of schizotypal personality traits: a cross-national study. Psychol. Med. 17:1-12. doi: 10.1017/S0033291717001829.
- Fonseca-Pedrero, E., Debbané, M., Schneider, M., Badoud, D., Eliez, S., 2016a. Schizotypal traits in adolescents with 22q11.2 deletion syndrome: validity, reliability and risk for psychosis. Psychol. Med. 46, 1005–1013.
- Fonseca-Pedrero, E., Fumero, A., Paino, M., de Miguel, A., Ortuño-Sierra, J., Lemos Giraldez, S., Muñiz, J., 2014b. Schizotypal Personality Questionnaire: New sources of validity evidence in college students. Psychiatry Res. 219, 214–220.
- Fonseca-Pedrero, E., Gooding, D., Debbané, M., Muñiz, J., 2016b. Psychopathology: Psychosis assessment and high-risk paradigms, in: Bartram, D., Cheung, F., Geisinger, K.F., and Iliescu, D. (Eds). The ITC International Handbook of Testing and Assessment. Oxford University Press, UK, pp. 147-170.
- Fonseca-Pedrero, E., Lemos-Giráldez, S., Paíno-Piñeiro, M., Villazón-García, U., Muñiz, J., 2010. Schizotypal traits, obsessive-compulsive symptoms, and social functioning in adolescents. Compr. Psychiatry 51, 71–77.
- Fonseca-Pedrero, E., Lemos-Giráldez, S., Paino, M., Villazón-García, U., Muñiz, J., 2009. Validation of the Schizotypal Personality Questionnaire Brief form in adolescents. Schizophr. Res. 111, 53–60.
- Fonseca-Pedrero, E., Ortuño-Sierra, J., Sierro, G., Daniel, C., Cella, M., Preti, A., Mohr,C., Mason, O., 2015. The measurement invariance of schizotypy in Europe. Eur.Psychiatry 30, 837–844.
- Fonseca-Pedrero, E., Paino, M., Lemos-Giráldez, S., Sierra-Baigrie, S., Muñiz, J., 2011. Measurement invariance of the Schizotypal Personality Questionnaire-Brief across gender and age. Psychiatry Res. 190, 309–315.
- Fonseca Pedrero, E., Debbané, M., 2017. Schizotypal traits and psychotic-like experiences during adolescence: An update. Psicothema. 29, 5-17.
- Fossati, A., Raine, A., Carretta, I., Leonardi, B., Maffei, C., 2003. The three-factor model of schizotypal personality: Invariance across age and gender. Pers. Individ. Dif. 35, 1007–1019.
- Fusar-Poli, P., Carpenter, W.T., Woods, S.W., McGlashan, T.H., 2014. Attenuated

- Psychosis Syndrome: Ready for DSM-5.1? Annu. Rev. Clin. Psychol. 10, 155–192.
- Hu, L.-T., Bentler, P.M., 1999. Cut off criteria for fit indexes in covariance structure analysis: Conventional criteria versus new alternatives. Struct. Equ. Model. 6, 1–55.
- IBM Corp Released, 2013. IBM SPSS Statistics for Windows, Version 22.0. IBM Corp, Armonk, NY.
- Klein, C., Andresen, B., Jahn, T., 1997. Erfassung der schizotypen Persönlichkeit nach DSM-III-R: Psychometrische Eigenschaften einer autorisierten deutschsprachigen Übersetzung des" Schizotypal Personality Questionnaire"(SPQ) von Raine. Diagnostica. 43, 347-369.
- Kline, E., Schiffman, J., 2014. Psychosis risk screening: A systematic review. Schizophr. Res. 158, 11–18.
- Kwapil, T.R., Gross, G.M., Silvia, P.J., Raulin, M.L., Barrantes-Vidal, N. in press. Development and psychometric properties of the Multidimensional Schizotypy Scale: A new measure for assessing positive, negative, and disorganized schizotypy. Schizophr Res. 2017 Jul 20. pii: S0920-9964(17)30402-4. doi: 10.1016/j.schres.2017.07.001.
- Kwapil, T.R., Ros-Morente, A., Silvia, P.J., Barrantes-Vidal, N., 2012. Factor invariance of psychometric schizotypy in Spanish and American samples. J. Psychopathol. Behav. Assess. 34, 145–152.
- Lahmar, M.L., Gassab, L., Beltaief, F., Mechri, A., 2014. Psychometric properties of the arabic version of the schizotypal personality questionnaire in Tunisian university students. Tunis Med 92, 318–322.
- Larøi, F., Luhrmann, T.M., Bell, V., Christian, W.A.J., Deshpande, S., Fernyhough, C., Jenkins, J., Woods, A., 2014. Culture and hallucinations: overview and future directions. Schizophr. Bull. 40, S213-220.
- Lenzenweger, M.F., 2010. Schizotypy and schizophrenia: The view from experimental psychopathology. Guilford Press, New York.
- Liddle, P., 1997 The symptoms of chronic schizophrenia: A re-examination of the positive-negative dichotomy. Br J Psychiatry. 151, 145-151.
- Linscott, R.J., van Os, J., 2013. An updated and conservative systematic review and meta-analysis of epidemiological evidence on psychotic experiences in children and adults: on the pathway from proneness to persistence to dimensional

- expression across mental disorders. Psychol. Med. 43, 1133–1149.
- Ma, W.-F., Lane, H.-Y., Chiang, L.-C., Wu, P.-L., Yang, S.-J., Tsai, G.E., 2015.

 Assessing the construct validity of the chinese-version schizotypal personality questionnaire-brief on male and female undergraduate students. J. Nurs. Res. 23, 162–166.
- Mason, O., 2015. The assessment of schizotypy and its clinical relevance. Schizophr. Bull. 41, S374-85.
- Mata, I., mataix-Cols, D., Peralta, V., 2005. Schizotypal Personality Questionnaire-Brief: factor structure and influence of sex and age in a nonclinical population. Pers. Individ. Dif. 38, 1183–1192.
- Meehl, P.E., 1962. Schizotaxia, schizotypy, schizophrenia. Am. Psychol. 17, 827–838.
- Moreno-Izco, L., Sánchez-Torres, A.M., Lorente-Omeñaca, R., Fañanás, L., Rosa, A., Salvatore, P., Peralta, V., Cuesta, M.J., 2015. Ten-year stability of self-reported schizotypal personality features in patients with psychosis and their healthy siblings. Psychiatry Res. 227, 283–289.
- Muthén, B.O., Asparouhov, T., 2002. Latent variable analysis with categorical outcomes: Multiple-group and growth modeling in Mplus. Mplus Web Note No. 4, at http://www.statmodel.com/mplus/examples/webnote.html.
- Muthén, L.K., Muthén, B.O., Mplus User's Guide. Seventh Edition. Muthén & Muthén, Los Angeles, CA. 1998-2012.
- Nuevo, R., Chatterji, S., Verdes, E., Naidoo, N., Arango, C., Ayuso-Mateos, J.L., 2012. The continuum of psychotic symptoms in the general population: A cross-national study. Schizophr. Bull. 38, 475–485.
- Ortuño-Sierra, J., Badoud, D., Knecht, F., Paino, M., Eliez, S., Fonseca-Pedrero, E., Debbané, M., 2013. Testing measurement invariance of the schizotypal personality questionnaire-brief scores across Spanish and Swiss adolescents. PLoS One 8. doi:10.1371/journal.pone.0082041
- Preti, A., Siddi, S., Vellante, M., Scanu, R., Muratore, T., Gabrielli, M., Tronci, D., Masala, C., Petretto, D.R., 2015. Bifactor structure of the Schizotypal Personality Questionnaire (SPQ). Psychiatry Res. 230, 940–950.
- R Development Core Team, 2011. R: A Language and Environment for Statistical Computing. R Foundation for Statistical Computing. Retrieved from http://www.R-project.org, Vienna, Austria.
- Raine, A., 1991. The SPQ: A scale for the assessment of schizotypal personality based

- on DSM-III-R criteria. Schizophr. Bull. 17, 555-564.
- Raine, A., Benishay, D., 1995. The SPQ-B: a brief screening instrument for schizotypal personality disorder. J. Pers. Disord. 9, 346–355.
- Raine, A., Reynolds, C., Lencz, T., Scerbo, A., Triphon, N., Kim, D., 1994. Cognitive-perceptual, interpersonal, and disorganized features of schizotypal personality. Schizophr. Bull. 20, 191–201.
- Reynolds, C.A., Raine, A., Mellingen, K., Venables, P.H., Mednick, S.A., 2000. Three-factor model of schizotypal personality: invariance across culture, gender, religious affiliation, family adversity, and psychopathology. Schizophr. Bull. 26, 603–618.
- Siever, L.J., Gunderson, J.G., 1983. The search for a schizotypal personality: historical origins and current status. Compr. Psychiatry 24, 199–212.
- Tran, U. S., Stieger, S., Voracek, M. 2015. Mixed-footedness is a more relevant predictor of schizotypy than mixed-handedness. Psychiatry Res. 225, 446-451.
- Tsaousis, I., Zouraraki, C., Karamaouna, P., Karagiannopoulou, L., Giakoumaki, S.G., 2015. The validity of the Schizotypal Personality Questionnaire in a Greek sample: Tests of measurement invariance and latent mean differences. Compr. Psychiatry 62, 51–62.
- Woods, A., Jones, N., Bernini, M., Callard, F., Alderson-Day, B., Badcock, J.C., Bell,
 V., Cook, C.C.H., Csordas, T., Humpston, C., Krueger, J., Laroi, F., McCarthy-Jones, S., Moseley, P., Powell, H., Raballo, A., Smailes, D., Fernyhough, C., 2014.
 Interdisciplinary approaches to the phenomenology of auditory verbal hallucinations. Schizophr. Bull. 40. 40 Suppl 4:S246-54.
- World Medical Association, 2013. World Medical Association Declaration of Helsinki: Ethical principles for medical research involving human subjects. J. Am. Med. Assoc. 310, 2191–2194.
- Yu, C.Y., 2002. Evaluating cutoff criteria of model fit indices for latent variable models with binary and continuous outcomes. Doctoral dissertation. Los Angeles: University of California.
- Zinbarg, R.E., Revelle, W., Yovel, I., Li, W., 2005. Cronbach's α, Revelle's β, and McDonald's ωH: Their relations with each other and two alternative conceptualizations of reliability. Psychometrika. 70, 123–133.
- Zumbo, B. D., 2007. Three generations of DIF analyses: Considering where it has been, where it is now, and where it is going. Lang Assess Q. 4, 223-233.

Table 1

Demographic characteristics of the sample

	Cour	ntry	Ge	nder		Age	
	n	%	Male	Females	M	SD	Range
US	10,477	36.9	3,162	7,212	22.0	6.7	16-55
Spain	1,123	4.0	224	899	20.2	2.0	18-29
New Zealand	1,698	6.0	515	1,183	20.1	3.0	17-51
Italy	649	2.3	305	344	24.3	3.5	19-38
Australia	1,931	6.8	634	1,294	28.5	11.2	17-55
Belgium	893	3.1	245	648	24.9	9.1	17-55
UK	1,199	4.2	404	795	22.8	6.5	16-68
Tunisia	458	1.6	137	321	20.4	1.4	18-29
China	4,907	17.3	2,973	1,533	19.7	1.0	17-24
Canada	1,849	6.5	562	1,287	20.8	2.9	18-53
Greece	1,041	3.7	390	651	32.4	9.9	17-55
Mauritius	1,201	4.2	688	513	23.4	1.2	21-27
Austria	611	1.4	294	317	33.2	12.6	19-66
Germany	389	2.1	178	211	32.7	13.2	19-66
Total	28,426	100	10,711	17,208	22.63	7.08	16-68

Table 2. Descriptive statistics for the SPQ-B across countries and total sample

	(n = 1)	SA 0,477)	Sp. (n = 1	,123)	Ne Zeal (n = 1	and ,698)	Ita (n =	649)	Austr (n = 1	,931)	Belg	893)	(n =	JK 1,199)	Tun (n = -	458)	(n = 4)		Can (n = 1	,849)	Gre (n = 1	,041)	Maur (n = 1	,201)	Aus (n = 1	390)	(n =	many 610)	San (N = 2	
Items	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD	Mean	SD
1	0.32	0.47	0.46	0.50	0.23	0.42			0.42	0.49	0.47	0.50	0.42	0.49	0.53	0.50	0.31		0.30	0.46		0.50	0.41	0.49	0.29	0.45	0.26	0.44	0.34	0.48
2	0.4	0.49	0.30		0.38		0.26	0.44					0.48	0.50	0.36				0.29		0.23		0.26	0.44			0.27	0.45	0.38	0.49
3	0.31	0.46	0.30		0.28		0.31	0.46		0.44		0.47	0.34	0.47	0.54		0.31		0.27		0.28		0.47	0.50			0.15	0.36	0.31	0.46
	0.24	0.43	0.34	0.48	0.25	0.44	0.27	0.45	0.33	0.47	0.28	0.45	0.31	0.46	0.37	0.48	0.69	0.46	0.24	0.43	0.35	0.48	0.19	0.39	0.13	0.34	0.11	0.31	0.34	0.47
	0.36	0.48	0.46	0.50	0.25	0.43	0.32	0.47	0.39	0.49	0.30	0.46	0.33	0.47	0.68	0.47	0.85	0.36	0.30	0.46	0.35	0.48	0.37	0.48	0.25	0.43	0.23	0.42	0.44	0.50
	0.16	0.37	0.42	0.49	0.14	0.34	0.09	0.29	0.17	0.38	0.10	0.30	0.26	0.44	0.12	0.33	0.13	0.34	0.12	0.33	0.07	0.26	0.18	0.38	0.12	0.32	0.12	0.33	0.16	0.37
	0.24	0.43	0.17	0.37	0.17	0.38	0.29	0.45	0.15	0.35	0.25	0.44	0.24	0.43	0.57	0.5	0.09	0.29	0.23	0.42	0.38	0.49	0.61	0.49	0.13	0.34	0.07	0.26	0.22	0.42
	0.19	0.39	0.13	0.34	0.13	0.34	0.18	0.38	0.14	0.35	0.28	0.45	0.23	0.42	0.48	0.5	0.37	0.48	0.17	0.37	0.15	0.36	0.28	0.45	0.19	0.4	0.16	0.37	0.22	0.41
	0.3	0.46	0.18	0.38	0.31	0.46	0.12	0.33	0.28	0.45	0.22	0.41	0.38	0.49	0.27	0.44	0.16	0.36	0.27	0.44	0.17	0.37	0.34	0.48	0.15	0.35	0.12	0.33	0.25	0.44
)	0.35	0.48	0.22	0.42	0.34	0.48	0.12	0.33	0.34	0.47	0.13	0.34	0.36	0.48	0.24	0.43	0.21	0.41	0.32	0.47	0.13	0.33	0.36	0.48	0.38	0.49	0.31	0.46	0.30	0.46
	0.38	0.49	0.34	0.48	0.33	0.47	0.15	0.35	0.45	0.50	0.30	0.46	0.36	0.48	0.50	0.5	0.29	0.45	0.30	0.46	0.22	0.42	0.58	0.49	0.18	0.38	0.15	0.36	0.35	0.48
	0.15	0.35	0.10	0.30	0.14	0.35	0.09	0.28	0.17	0.37	0.17	0.37	0.25	0.43	0.16	0.37	0.19	0.39	0.08	0.27	0.15	0.35	0.18	0.39	0.15	0.36	0.20	0.40	0.15	0.36
, ,	0.37	0.48	0.41	0.49	0.42	0.49	0.25	0.44	0.35	0.48	0.46	0.50	0.46	0.50	0.45	0.5	0.43	0.50	0.32	0.47	0.27	0.44	0.36	0.48	0.28	0.45	0.30	0.46	0.38	0.47
	0.45	0.50	0.47	0.50	0.38	0.49	0.59	0.49	0.29	0.45	0.42	0.49	0.50	0.50	0.68	0.47	0.25	0.43	0.42	0.49	0.56	0.50	0.73	0.45	0.38	0.49	0.38	0.49	0.42	0.49
;	0.35	0.48	0.41	0.49	0.28	0.45	0.16	0.36	0.25	0.44	0.34	0.48	0.38	0.49	0.38	0.49	0.53	0.50	0.35	0.48	0.47	0.50	0.26	0.44	0.34	0.47	0.32	0.47	0.37	0.48
	0.30	0.46	0.36	0.48	0.30	0.46	0.24	0.43	0.34	0.47	0.31	0.46	0.40	0.49	0.28	0.45	0.21	0.41	0.25	0.44	0.26	0.44	0.43	0.50	0.18	0.39	0.16	0.37	0.29	0.45
,	0.30	0.46	0.24	0.43	0.24	0.43	0.19	0.40	0.21	0.41	0.44	0.50	0.36	0.48	0.72	0.45	0.22	0.41	0.23	0.42	0.46	0.50	0.55	0.50	0.19	0.39	0.11	0.32	0.29	0.45
	0.23	0.42	0.12	0.33	0.17	0.38	0.06	0.24	0.15	0.36	0.16	0.36	0.26	0.44	0.22	0.42	0.12	0.33	0.22	0.41	0.14	0.34	0.37	0.48	0.26	0.44	0.22	0.42	0.20	0.39
3	0.27	0.44	0.13	0.33	0.19	0.39	0.11	0.31	0.17	0.37	0.36	0.48	0.29	0.46	0.31	0.46	0.07	0.26	0.20	0.40	0.11	0.32	0.23	0.42	0.15	0.36	0.14	0.35	0.20	0.40
	0.24	0.43	0.17	0.38	0.14	0.35	0.14	0.35	0.17	0.38	0.29	0.45	0.29	0.45	0.39	0.49	0.31	0.46	0.20	0.40	0.21	0.41	0.33	0.47	0.17	0.38	0.17	0.37	0.24	0.43
)	0.34	0.47	0.35	0.48	0.31	0.46	0.18	0.39	0.25	0.43	0.28	0.45	0.33	0.47	0.38	0.49	0.16	0.36	0.28	0.45	0.22	0.42	0.38	0.49	0.15	0.36	0.12	0.33	0.28	0.45
1 2	0.48	0.5	0.54	0.50	0.43	0.50	0.47	0.50	0.37	0.48	0.59	0.49	0.51	0.50	0.66	0.47	0.04	0.20	0.48	0.50	0.48	0.50	0.52	0.50	0.46	0.50	0.42	0.49	0.40	0.49

Subscales																														
POS	2.41	2.07	2.21	1.79	2.21	1.88	1.62	1.71	2.57	1.92	2.18	1.84	2.86	2.09	3.09	1.81	2.97	1.55	1.99	1.87	2.10	1.80	2.69	1.90	1.71	1.84	1.52	1.62	2.44	1.93
INT	2.79	2.42	2.86	2.09	2.29	2.24	2.28	1.79	2.32	1.99	2.82	2.20	2.99	2.36	3.92	2.11	1.79	1.63	2.58	2.28	2.90	2.23	3.86	2.08	2.18	2.13	1.95	1.94	2.58	2.35
DIS	1.55	1.70	1.57	1.41	1.30	1.46	1.08	1.38	1.26	1.51	1.82	1.53	1.88	1.80	2.30	1.62	1.62	1.41	1.28	1.53	1.10	1.34	1.84	1.68	1.06	1.47	1.03	1.44	1.51	1.59
Total score	6.74	4.99	6.64	3.91	5.80	4.27	4.98	3.83	6.15	4.04	6.82	4.29	7.73	4.89	9.31	4.16	6.37	3.46	5.85	4.43	6.09	4.17	8.39	4.65	4.95	4.28	4.50	3.79	6.54	4.50

 $Note.\ SD = Standard\ Deviation;\ POS = Positive;\ INT = Interpersonal;\ DIS = Disorganized$

Table 3

Omega coefficients for the SPQ-B scores across countries and total sample

SPQ-B	US	Spain	NZ	Italy	Australia	Belgium	UK	Tunisia	China	Canada	Greece	Mauritus	Austria	Germany	Total
Positive	0.88	0.84	0.87	0.87	0.84	0.85	0.86	0.79	0.77	0.87	0.88	0.83	0.88	0.91	0.86
Interpersonal	0.93	0.90	0.93	0.89	0.88	0.91	0.92	0.85	0.88	0.92	0.92	0.86	0.92	0.94	0.91
Disorganization	0.91	0.85	0.92	0.91	0.91	0.84	0.93	0.85	0.88	0.90	0.90	0.86	0.92	0.94	0.89
Total Score	0.94	0.89	0.93	0.93	0.91	0.91	0.93	0.88	0.89	0.93	0.92	0.91	0.92	0.93	0.92

Note. NZ= New Zealand

Table 4

Goodness-of-fit indices of the models tested in the confirmatory factor analysis

	χ^2	df	CFI	TLI	RMSEA (90% CI)	WRMR
Model a: Unidimensional						
US	13644.01	209	.820	.801	.085 (.084081)	6.825
Spain	5375.73	209	.700	.668	.090 (.088092)	4.469
New Zealand	2717.42	209	.775	.751	.084 (.081-087)	3.178
Italy	743.17	209	.826	.808	.063 (.058068)	1.709
Australia	2730.01	209	.740	.713	.079 (.076082)	3.220
Belgium	1505.39	209	.748	.722	.083 (.079087)	2.423
UK	2654.50	209	.761	.736	.099 (.095010)	3.188
Tunisia	598.52	209	.783	.754	.064 (.061066)	1.523
China	4309.68	209	.772	.751	.064 (.062067)	3.904
Canada	3036.70	209	.785	.762	.086 (.083088)	3.371
Greece	1578.64	209	.793	.774	.080 (.079082)	2.475
Mauritus	741.28	209	.921	.912	.046 (.042050)	1.564
Austria	721.908	209	.803	.782	.079 (.073086)	1.727
Germany	971.177	209	.749	.723	.077 (.072082)	1.998
Total sample	42494.65	209	.768	.743	.084 (.084085)	12.104
Model b:Bidimensional						
US	14069.65	208	.855	.839	.080 (.079081)	6.960
Spain	1479.03	208	.742	.713	.074 (.070077)	2.423
New Zealand	2285.08	208	.814	.793	.077 (.074080)	2.921
Italy	667.05	208	.850	.834	.058 (.053063)	1.611
Australia	2430.40	208	.774	.748	.072 (.068074)	3.042
Belgium	1357.85	208	.783	.749	.078 (.074081)	2.305
UK	2.293.71	208	.796	.774	.091 (.088095)	2.968
Tunisia	525.48	208	.817	.796	.058 (.052064)	1.415
China	3870.22	208	.796	.773	.060 (.058062)	3.703
Canada	2456.85	208	.829	.810	.076 (.074079)	3.035
Greece	1205.26	208	.853	.838	.073 (.068075)	2.164
Mauritus	608.31	208	.940	.934	.040 (.036044)	1.412
Austria	580.94	208	.856	.841	.069 (.061074)	1.531
Germany	801.94	208	.805	.783	.068 (.063073)	1.814
Total sample	37064.26	208	.797	.775	.079 (.078080)	11.325
Model c: Three factor model						
US	8297.27	202	.915	.903	.062 (.061063)	5.184
Spain	990.75	202	.840	.820	.059 (.055063)	1.943
New Zealand	1336.89	202	.900	.880	.058 (.055060)	2.186
Italy	414.88	202	.931	.921	.040 (.035046)	1.211
Australia	1180.56	202	.899	.885	.050 (.047053)	2.054
Belgium	897.01	202	.865	.846	.062 (.058066)	1.820
UK	1444.63	202	.897	.861	.072 (.068075)	2.285
Tunisia	396.64	202	.871	.871	.046 (.039053)	1.195
China	2847.80	202	.852	.831	.052 (.050053)	3.170

Greece 872.69 202 .899 .884 .056 (.053060) 1.790 Mauritus 521.96 202 .952 .945 .036 (.033040) 1.292 Austria 374.84 202 .933 .924 .047 (.039054) 1.154 Germany 482.39 202 .908 .895 .048 (.042053) 1.342 Total sample 22683.56 202 .876 .859 .048 (.042053) 1.342 Model d:Three factor model (no overlay) V V V .876 .852 .068 (.067069) 5.860 Spain 1245.25 206 .789 .763 .067 (.063071) 2.208 New Zealand 1675.86 206 .868 .852 .065 (.062068) 2.476 Italy 510.998 206 .869 .853 .056 (.062068) 2.476 Italy 510.998 206 .869 .853 .056 (.062068) 2.431 UK 1.656.99 206 .85	Canada	1482.74	202	.903	.889	.059 (.056061)	2.291
Austria 374.84 202 .933 .924 .047 (.039054) 1.134 Germany 482.39 202 .908 .895 .048 (.042053) 1.342 Total sample 22683.56 202 .876 .859 .063 (.062063) 8.727 Model di:Three factor model (no verlap) US 10267.63 206 .889 .882 .068 (.067069) 5.860 Spain 1245.25 206 .789 .763 .067 (.063071) 2.208 New Zealand 1675.86 206 .868 .852 .065 (.062068) 2.476 Italy 510.998 206 .891 .889 .048 (.043053) 1.383 Australia 1474.20 206 .869 .853 .056 (.054059) 2.333 Belgium 1020.36 206 .842 .823 .067 (.062071) 1.971 UK 1.565.99 206 .858 .841 .077 (.041054) 2.246 China 3552	Greece	872.69	202	.899	.884	.056 (.053060)	1.790
Germany 482.39 202 .908 .895 .048 (.042053) 8.727 Total sample 22683.56 202 .876 .859 .063 (.062063) 8.727 Model d:Three factor model (no overlap) US 10267.63 206 .895 .882 .068 (.067069) 5.860 Spain 1245.25 206 .789 .763 .067 (.063071) 2.026 New Zealand 1675.86 206 .868 .852 .065 (.062068) 2.476 Italy 510.998 206 .809 .853 .048 (.043053) 1.383 Australia 1474.20 206 .869 .853 .056 (.054059) 2.333 Belgium 1020.36 206 .842 .823 .067 (.062071) 1.971 UK 1.656.99 206 .858 .841 .077 (.071073080) 2.484 Tunisia 418.60 206 .878 .863 .065 (.062069) 2.572 Greace <t< td=""><td>Mauritus</td><td>521.96</td><td>202</td><td>.952</td><td>.945</td><td>.036 (.033040)</td><td>1.292</td></t<>	Mauritus	521.96	202	.952	.945	.036 (.033040)	1.292
Total sample 22683.56 202 .876 .859 .063 (.062063) 8.72 Model d:Three factor model (no overlap) US 10267.63 206 .895 .882 .068 (.067069) 5.860 Spain 1245.25 206 .789 .763 .067 (.063071) 2.208 New Zealand 1675.86 206 .868 .852 .065 (.062068) 2.476 Italy 510.998 206 .901 .889 .048 (.043053) 1.383 Australia 1474.20 206 .869 .853 .056 (.062068) 2.433 Belgium 1020.36 206 .842 .823 .067 (.062071) 1.971 UK 1.656.99 206 .858 .841 .077 (.073080) 2.484 Tunisia 418.60 206 .877 .862 .047 (.041054) 1.246 China 3552.65 206 .813 .791 .058 (.056069) 3.541 Canada 180.923 206	Austria	374.84	202	.933	.924	.047 (.039054)	1.154
Model d:Three factor model (no overlay) US 10267.63 206 .895 .882 .068 (.067069) 5.860 Spain 1245.25 206 .789 .763 .067 (.063071) 2.208 New Zealand 1675.86 206 .868 .852 .065 (.062068) 2.476 Italy 510.998 206 .901 .889 .048 (.043053) 1.383 Australia 1474.20 206 .869 .853 .056 (.054059) 2.333 Belgium 1020.36 206 .842 .823 .056 (.062071) 1.971 UK 1.656.99 206 .874 .862 .047 (.041054) 1.246 China 3552.65 206 .813 .791 .058 (.056059) 3.541 China 3552.65 206 .877 .862 .047 (.041054) 1.246 China 1809.23 206 .878 .863 .055 (.062069) 2.063 Mauritus 614.40	Germany	482.39	202	.908	.895	.048 (.042053)	1.342
US 10267.63 206 .895 .882 .068 (.067069) 5.80 Spain 1245.25 206 .789 .763 .067 (.063071) 2.208 New Zealand 1675.86 206 .868 .852 .065 (.062068) 2.476 Italy 510.998 206 .901 .889 .048 (.043053) 1.383 Australia 1474.20 206 .869 .853 .056 (.054059) 2.333 Belgium 1020.36 206 .842 .823 .067 (.062071) 1.971 UK 1.656.99 206 .858 .884 .077 (.073080) 2.484 Tunisia 418.60 206 .878 .862 .047 (.041054) 1.246 China 3552.65 206 .813 .791 .058 (.062069) 3.541 Canada 1809.23 206 .878 .863 .065 (.062069) 2.063 Mauritus 614.40 206 .893 .880 .059 (.05	Total sample	22683.56	202	.876	.859	.063 (.062063)	8.727
Spain 1245.25 206 .789 .763 .067 (.063071) 2.208 New Zealand 1675.86 206 .868 .852 .065 (.062068) 2.476 Italy 510.998 206 .901 .889 .048 (.043053) 1.383 Australia 1474.20 206 .869 .853 .056 (.054059) 2.333 Belgium 1020.36 206 .842 .823 .067 (.062071) 1.971 UK 1.656.99 206 .858 .841 .077 (.073080) 2.484 Tunisia 418.60 206 .877 .862 .047 (.041054) 1.246 China 3552.65 206 .813 .791 .058 (.056059) 3.541 Canada 1809.23 206 .878 .863 .065 (.062068) 2.572 Greece 1124.98 206 .861 .845 .065 (.062068) 2.572 Germany 701.291 206 .893 .880 .059	Model d:Three factor model	(no overlap)					
New Zealand 1675.86 206 .868 .852 .065 (.062068) 2.47 Italy 510.998 206 .901 .889 .048 (.043053) 1.383 Australia 1474.20 206 .869 .853 .056 (.054059) 2.333 Belgium 1020.36 206 .842 .823 .067 (.062071) 1.971 UK 1.656.99 206 .858 .841 .077 (.073080) 2.484 Tunisia 418.60 206 .877 .862 .047 (.041054) 1.246 China 3552.65 206 .813 .791 .058 (.056059) 3.541 Canada 1809.23 206 .878 .863 .065 (.062068) 2.572 Greece 1124.98 206 .861 .845 .065 (.062069) 2.063 Mauritus 614.40 206 .893 .880 .059 (.055062) .041 (.037-044) 1.414 Austria 484.997 206 .837	US	10267.63	206	.895	.882	.068 (.067069)	5.860
Italy 510.998 206 .901 .889 .048 (.043053) 1.383 Australia 1474.20 206 .869 .853 .056 (.054059) 2.333 Belgium 1020.36 206 .842 .823 .067 (.062071) 1.971 UK 1.656.99 206 .858 .841 .077 (.073080) 2.484 Tunisia 418.60 206 .877 .862 .047 (.041054) 1.246 China 3552.65 206 .813 .791 .058 (.056059) 3.541 Canada 1809.23 206 .878 .863 .065 (.062068) 2.572 Greece 1124.98 206 .861 .845 .065 (.062069) 2.063 Mauritus 614.40 206 .939 .932 .041 (.037044) 1.414 Austria 484.997 206 .893 .880 .059 (.052066) 1.362 Germany 701.291 206 .837 .817 .063 (.	Spain	1245.25	206	.789	.763	.067 (.063071)	2.208
Australia 1474.20 206 .869 .853 .056 (.054059) 2.333 Belgium 1020.36 206 .842 .823 .067 (.062071) 1.971 UK 1.656.99 206 .858 .841 .077 (.073080) 2.484 Tunisia 418.60 206 .877 .862 .047 (.041054) 1.246 China 3552.65 206 .813 .791 .058 (.056059) 3.541 Canada 1809.23 206 .878 .863 .065 (.062068) 2.572 Greece 1124.98 206 .861 .845 .065 (.062069) 2.063 Mauritus 614.40 206 .939 .932 .041 (.037044) 1.414 Austria 484.997 206 .893 .880 .059 (.052066) 1.362 Germany 701.291 206 .837 .817 .063 (.058068) 1.671 Total sample 28597.38 206 .844 .825 <td< td=""><td>New Zealand</td><td>1675.86</td><td>206</td><td>.868</td><td>.852</td><td>.065 (.062068)</td><td>2.476</td></td<>	New Zealand	1675.86	206	.868	.852	.065 (.062068)	2.476
Belgium 1020.36 206 .842 .823 .067 (.062071) 1.971 UK 1.656.99 206 .858 .841 .077 (.073080) 2.484 Tunisia 418.60 206 .877 .862 .047 (.041054) 1.246 China 3552.65 206 .813 .791 .058 (.056059) 3.541 Canada 1809.23 206 .878 .863 .065 (.062068) 2.572 Greece 1124.98 206 .861 .845 .065 (.062069) 2.063 Mauritus 614.40 206 .939 .932 .041 (.037044) 1.414 Austria 484.997 206 .893 .880 .059 (.052066) 1.362 Germany 701.291 206 .837 .817 .063 (.058068) 1.671 Total sample 28597.38 206 .844 .825 .070 (.069070) 9.878 Model e: bifactor US 5847.31 187 .941	Italy	510.998	206	.901	.889	.048 (.043053)	1.383
UK 1.656.99 206 .858 .841 .077 (.073080) 2.484 Tunisia 418.60 206 .877 .862 .047 (.041054) 1.246 China 3552.65 206 .813 .791 .058 (.056059) 3.541 Canada 1809.23 206 .878 .863 .065 (.062068) 2.572 Greece 1124.98 206 .861 .845 .065 (.062069) 2.063 Mauritus 614.40 206 .939 .932 .041 (.037044) 1.414 Austria 484.997 206 .893 .880 .059 (.052066) 1.362 Germany 701.291 206 .837 .817 .063 (.058068) 1.671 Total sample 28597.38 206 .844 .825 .070 (.069070) 9.878 Model e: bifactor US 5847.31 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .941	Australia	1474.20	206	.869	.853	.056 (.054059)	2.333
Tunisia 418.60 206 .877 .862 .047 (.041054) 1.246 China 3552.65 206 .813 .791 .058 (.056059) 3.541 Canada 1809.23 206 .878 .863 .065 (.062068) 2.572 Greece 1124.98 206 .861 .845 .065 (.062069) 2.063 Mauritus 614.40 206 .939 .932 .041 (.037044) 1.414 Austria 484.997 206 .893 .880 .059 (.052066) 1.362 Germany 701.291 206 .837 .817 .063 (.058068) 1.671 Total sample 28597.38 206 .844 .825 .070 (.069070) 9.878 Model e: bifactor 1 .847 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .946 .921 <t< td=""><td>Belgium</td><td>1020.36</td><td>206</td><td>.842</td><td>.823</td><td>.067 (.062071)</td><td>1.971</td></t<>	Belgium	1020.36	206	.842	.823	.067 (.062071)	1.971
China 3552.65 206 .813 .791 .058 (.056059) 3.541 Canada 1809.23 206 .878 .863 .065 (.062068) 2.572 Greece 1124.98 206 .861 .845 .065 (.062069) 2.063 Mauritus 614.40 206 .939 .932 .041 (.037044) 1.414 Austria 484.997 206 .893 .880 .059 (.052066) 1.362 Germany 701.291 206 .837 .817 .063 (.058068) 1.671 Total sample 28597.38 206 .844 .825 .070 (.069070) 9.878 Model e: bifactor US 5847.31 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .994 .875 .049 (.045053) 1.544 New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950	UK	1.656.99	206	.858	.841	.077 (.073080)	2.484
Canada 1809.23 206 .878 .863 .065 (.062068) 2.572 Greece 1124.98 206 .861 .845 .065 (.062069) 2.063 Mauritus 614.40 206 .939 .932 .041 (.037044) 1.414 Austria 484.997 206 .893 .880 .059 (.052066) 1.362 Germany 701.291 206 .837 .817 .063 (.058068) 1.671 Total sample 28597.38 206 .844 .825 .070 (.069070) 9.878 Model e: bifactor US 5847.31 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .898 .875 .049 (.045053) 1.544 New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .991	Tunisia	418.60	206	.877	.862	.047 (.041054)	1.246
Greece 1124.98 206 .861 .845 .065 (.062069) 2.064 Mauritus 614.40 206 .939 .932 .041 (.037044) 1.414 Austria 484.997 206 .893 .880 .059 (.052066) 1.362 Germany 701.291 206 .837 .817 .063 (.058068) 1.671 Total sample 28597.38 206 .844 .825 .070 (.069070) 9.878 Model e: bifactor 5847.31 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .898 .875 .049 (.045053) 1.544 New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .912 .892 .049 (.046051) 1.830 Belgium 695.55 187 .901 .878 <td>China</td> <td>3552.65</td> <td>206</td> <td>.813</td> <td>.791</td> <td>.058 (.056059)</td> <td>3.541</td>	China	3552.65	206	.813	.791	.058 (.056059)	3.541
Mauritus 614.40 206 .939 .932 .041 (.037044) 1.414 Austria 484.997 206 .893 .880 .059 (.052066) 1.362 Germany 701.291 206 .837 .817 .063 (.058068) 1.671 Total sample 28597.38 206 .844 .825 .070 (.069070) 9.878 Model e: bifactor US 5847.31 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .898 .875 .049 (.045053) 1.544 New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .991 .878 .055 (.051060) 1.532 UK 957.491 187 .992 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .991	Canada	1809.23	206	.878	.863	.065 (.062068)	2.572
Austria 484.997 206 .893 .880 .059 (.052066) 1.362 Germany 701.291 206 .837 .817 .063 (.058068) 1.671 Total sample 28597.38 206 .844 .825 .070 (.069070) 9.878 Model e: bifactor US 5847.31 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .898 .875 .049 (.045053) 1.544 New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .912 .892 .049 (.046051) 1.830 Belgium 695.55 187 .901 .878 .055 (.051060) 1.532 UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912	Greece	1124.98	206	.861	.845	.065 (.062069)	2.063
Germany 701.291 206 .837 .817 .063 (.058068) 1.671 Total sample 28597.38 206 .844 .825 .070 (.069070) 9.878 Model e: bifactor US 5847.31 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .898 .875 .049 (.045053) 1.544 New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .912 .892 .049 (.046051) 1.830 Belgium 695.55 187 .901 .878 .055 (.051060) 1.532 UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187	Mauritus	614.40	206	.939	.932	.041 (.037044)	1.414
Total sample 28597.38 206 .844 .825 .070 (.069070) 9.878 Model e: bifactor US 5847.31 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .898 .875 .049 (.045053) 1.544 New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .912 .892 .049 (.046051) 1.830 Belgium 695.55 187 .901 .878 .055 (.051060) 1.532 UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187 .992 .866 .046 (.044048) 2.640 Canada 1006.38 187 <t< td=""><td>Austria</td><td>484.997</td><td>206</td><td>.893</td><td>.880</td><td>.059 (.052066)</td><td>1.362</td></t<>	Austria	484.997	206	.893	.880	.059 (.052066)	1.362
Model e: bifactor US 5847.31 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .898 .875 .049 (.045053) 1.544 New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .912 .892 .049 (.046051) 1.830 Belgium 695.55 187 .901 .878 .055 (.051060) 1.532 UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187 .892 .866 .046 (.044048) 2.640 Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187	Germany	701.291	206	.837	.817	.063 (.058068)	1.671
US 5847.31 187 .941 .927 .054 (.053055) 4.123 Spain 687.21 187 .898 .875 .049 (.045053) 1.544 New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .912 .892 .049 (.046051) 1.830 Belgium 695.55 187 .901 .878 .055 (.051060) 1.532 UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187 .892 .866 .046 (.044048) 2.640 Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187 .921 .903 .052 (.048056)	Total sample	28597.38	206	.844	.825	.070 (.069070)	9.878
Spain 687.21 187 .898 .875 .049 (.045053) 1.544 New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .912 .892 .049 (.046051) 1.830 Belgium 695.55 187 .901 .878 .055 (.051060) 1.532 UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187 .892 .866 .046 (.044048) 2.640 Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187 .921 .903 .052 (.048056) 1.547 Mauritus 415.24 187 .966 .958 .032 (.028-	Model e: bifactor						
New Zealand 902.85 187 .936 .921 .047 (.044051) 1.695 Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .912 .892 .049 (.046051) 1.830 Belgium 695.55 187 .901 .878 .055 (.051060) 1.532 UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187 .892 .866 .046 (.044048) 2.640 Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187 .921 .903 .052 (.048056) 1.547 Mauritus 415.24 187 .966 .958 .032 (.028036) 1.127 Austria 299.357 187 .957 .947 .039 (.0	US	5847.31	187	.941	.927	.054 (.053055)	4.123
Italy 338.92 187 .950 .939 .035 (.029041) 1.051 Australia 1036.82 187 .912 .892 .049 (.046051) 1.830 Belgium 695.55 187 .901 .878 .055 (.051060) 1.532 UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187 .892 .866 .046 (.044048) 2.640 Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187 .921 .903 .052 (.048056) 1.547 Mauritus 415.24 187 .966 .958 .032 (.028036) 1.127 Austria 299.357 187 .957 .947 .039 (.031047) .956 Germany 373.595 187 .939 .924 .040 (.034046) 1.102	Spain	687.21	187	.898	.875	.049 (.045053	1.544
Australia 1036.82 187 .912 .892 .049 (.046051) 1.830 Belgium 695.55 187 .901 .878 .055 (.051060) 1.532 UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187 .892 .866 .046 (.044048) 2.640 Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187 .921 .903 .052 (.048056) 1.547 Mauritus 415.24 187 .966 .958 .032 (.028036) 1.127 Austria 299.357 187 .957 .947 .039 (.031047) .956 Germany 373.595 187 .939 .924 .040 (.034046) 1.102	New Zealand	902.85	187	.936	.921	.047 (.044051)	1.695
Belgium 695.55 187 .901 .878 .055 (.051060) 1.532 UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187 .892 .866 .046 (.044048) 2.640 Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187 .921 .903 .052 (.048056) 1.547 Mauritus 415.24 187 .966 .958 .032 (.028036) 1.127 Austria 299.357 187 .957 .947 .039 (.031047) .956 Germany 373.595 187 .939 .924 .040 (.034046) 1.102	Italy	338.92	187	.950	.939	.035 (.029041)	1.051
UK 957.491 187 .925 .907 .059 (.055062) 1.749 Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187 .892 .866 .046 (.044048) 2.640 Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187 .921 .903 .052 (.048056) 1.547 Mauritus 415.24 187 .966 .958 .032 (.028036) 1.127 Austria 299.357 187 .957 .947 .039 (.031047) .956 Germany 373.595 187 .939 .924 .040 (.034046) 1.102	Australia	1036.82	187	.912	.892	.049 (.046051)	1.830
Tunisia 339.87 187 .912 .891 .042 (.035049) 1.072 China 2124.12 187 .892 .866 .046 (.044048) 2.640 Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187 .921 .903 .052 (.048056) 1.547 Mauritus 415.24 187 .966 .958 .032 (.028036) 1.127 Austria 299.357 187 .957 .947 .039 (.031047) .956 Germany 373.595 187 .939 .924 .040 (.034046) 1.102	Belgium	695.55	187	.901	.878	.055 (.051060)	1.532
China 2124.12 187 .892 .866 .046 (.044048) 2.640 Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187 .921 .903 .052 (.048056) 1.547 Mauritus 415.24 187 .966 .958 .032 (.028036) 1.127 Austria 299.357 187 .957 .947 .039 (.031047) .956 Germany 373.595 187 .939 .924 .040 (.034046) 1.102	UK	957.491	187	.925	.907	.059 (.055062)	1.749
Canada 1006.38 187 .938 .923 .049 (.046052) 1.780 Greece 709.26 187 .921 .903 .052 (.048056) 1.547 Mauritus 415.24 187 .966 .958 .032 (.028036) 1.127 Austria 299.357 187 .957 .947 .039 (.031047) .956 Germany 373.595 187 .939 .924 .040 (.034046) 1.102	Tunisia	339.87	187	.912	.891	.042 (.035049)	1.072
Greece 709.26 187 .921 .903 .052 (.048056) 1.547 Mauritus 415.24 187 .966 .958 .032 (.028036) 1.127 Austria 299.357 187 .957 .947 .039 (.031047) .956 Germany 373.595 187 .939 .924 .040 (.034046) 1.102	China	2124.12	187	.892	.866	.046 (.044048)	2.640
Mauritus 415.24 187 .966 .958 .032 (.028036) 1.127 Austria 299.357 187 .957 .947 .039 (.031047) .956 Germany 373.595 187 .939 .924 .040 (.034046) 1.102	Canada	1006.38	187	.938	.923	.049 (.046052)	1.780
Austria 299.357 187 .957 .947 .039 (.031047) .956 Germany 373.595 187 .939 .924 .040 (.034046) 1.102	Greece	709.26	187	.921	.903	.052 (.048056)	1.547
Germany 373.595 187 .939 .924 .040 (.034046) 1.102	Mauritus	415.24	187	.966	.958	.032 (.028036)	1.127
	Austria	299.357	187	.957	.947	.039 (.031047)	.956
Total sample 17695.42 187 .904 .881 .057 (.057058) 7.357	Germany	373.595	187	.939	.924	.040 (.034046)	1.102
` /	Total sample	17695.42	187	.904	.881	.057 (.057058)	7.357

Note. χ^2 = Chi square; df = degrees of freedom; CFI = Comparative Fit Index; TLI = Tucker-Lewis Index; RMSEA = Root Mean Square Error of Approximation; CI = Confidence Interval; WRMR= Weighted Root Mean Square Residual.

Table 5

Factor loadings for the Bifactor model.

	US	Spain	NZ	Italy	Australia	Belgium	UK	Tunisia	China	Canada	Greece	Mauritus	Austria	Germany	Total sample	Acros	ss samples
General factor																Mean	Range
1	0.68	0.47	0.73	0.39	0.37	0.57	0.64	0.42	0.53	0.61	0.54	0.66	0.66	0.54	0.58	0.55	0.37-0.73
2	0.32	0.18	0.18	0.33	0.23	0.19	0.27	0.21	0.20	0.24	0.22	0.51	0.51	0.30	0.27	0.25	0.18-0.50
3	0.57	0.37	0.41	0.67	0.53	0.56	0.50	0.44	0.54	0.47	0.49	0.63	0.63	0.49	0.54	0.51	0.37-0.67
4	0.38	0.22	0.33	0.49	0.37	0.27	0.30	0.18	0.07	0.30	0.15	0.33	0.33	0.49	0.20	0.28	0.07-0.49
5	0.34	0.23	0.30	0.47	0.39	0.32	0.33	0.31	0.12	0.31	0.29	0.31	0.31	0.45	0.22	0.31	0.12-0.47
6	0.60	0.31	0.50	0.63	0.50	0.77	0.57	0.51	0.73	0.51	0.45	0.72	0.72	0.52	0.57	0.56	0.31-0.77
7	0.71	0.70	0.74	0.59	0.68	0.63	0.66	0.33	0.51	0.73	0.74	0.57	0.57	0.89	0.65	0.63	0.33-0.74
8	0.78	0.57	0.70	0.73	0.62	0.55	0.75	0.57	0.62	0.72	0.58	0.70	0.70	0.64	0.67	0.66	0.55-0.73
9	0.56	0.54	0.56	0.71	0.57	0.57	0.46	0.56	0.67	0.55	0.54	0.59	0.59	0.84	0.59	0.57	0.43-0.71
10	0.46	0.35	0.33	0.57	0.33	0.43	0.46	0.37	0.37	0.41	0.45	0.47	0.47	0.29	0.44	0.42	0.33-0.57
11	0.51	0.33	0.42	0.44	0.23	0.36	0.47	0.39	0.38	0.42	0.60	0.42	0.42	0.59	0.44	0.42	0.23-0.59
12	0.31	0.17	0.17	0.33	0.23	0.16	0.14	0.30	0.17	0.27	0.21	0.32	0.32	0.36	0.25	0.22	0.05-0.33
13	0.57	0.31	0.40	0.63	0.50	0.39	0.47	0.61	0.24	0.51	0.51	0.54	0.54	0.43	0.48	0.47	0.31-0.63
14	0.67	0.66	0.69	0.43	0.54	0.64	0.66	0.35	0.43	0.66	0.68	0.53	0.53	0.58	0.59	0.58	0.35-0.69
15	0.53	0.43	0.53	0.43	0.25	0.48	0.41	0.47	0.21	0.43	0.50	0.42	0.42	0.48	0.40	0.42	0.21-0.50
16	0.52	0.34	0.39	0.45	0.42	0.47	0.52	0.35	0.38	0.50	0.45	0.50	0.50	0.58	0.50	0.44	0.38-0.50
17	0.64	0.64	0.56	0.72	0.57	0.47	0.60	0.28	0.45	0.59	0.63	0.51	0.51	0.63	0.62	0.56	0.45-0.72
18	0.69	0.61	0.65	0.34	0.53	0.55	0.67	0.59	0.70	0.67	0.66	0.34	0.34	0.63	0.63	0.59	0.34-0.73
19	0.66	0.46	0.61	0.59	0.59	0.56	0.61	0.69	0.74	0.63	0.52	0.65	0.65	0.58	0.65	0.61	0.46-0.74
20	0.78	0.65	0.75	0.75	0.70	0.78	0.76	0.73	0.73	0.76	0.74	0.67	0.67	0.72	0.75	0.73	0.65-0.78
21	0.54	0.38	0.43	0.50	0.38	0.37	0.45	0.46	0.47	0.47	0.60	0.53	0.53	0.69	0.47	0.47	0.37-0.60
22	0.55	0.57	0.57	0.18	0.35	0.45	0.58	0.15	0.59	0.57	0.60	0.46	0.46	0.61	0.47	0.47	0.15-0.62

Latent factors																	
Positive																	
2	0.64	0.65	0.69	0.71	0.42	0.57	0.70	0.73	0.60	0.61	0.76	0.39	0.81	0.92	0.58	0.66	0.39-0.92
4	0.45	0.40	0.41	0.27	0.44	0.42	0.41	0.42	0.47	0.46	0.32	0.39	0.51	0.51	0.55	0.42	0.27-0.51
5	0.63	0.61	0.59	0.48	0.63	0.66	0.61	0.50	0.60	0.65	0.67	0.41	0.60	0.68	0.70	0.59	0.41-0.68
9	0.04	0.26	0.30	0.12	0.24	0.25	0.29	0.29	0.09	0.38	0.35	0.28	0.18	0.11	0.19	0.23	0.04-0.38
10	0.32	0.32	0.28	0.17	0.43	0.24	0.27	0.23	0.22	0.39	0.31	0.24	0.10	0.18	0.23	0.26	0.10-0.39
12	0.51	0.59	0.64	0.65	0.35	0.70	0.58	0.45	0.30	0.47	0.61	0.35	0.84	0.69	0.49	0.55	0.30-0.84
16	0.42	0.43	0.40	0.30	0.41	0.31	0.35	0.41	0.37	0.41	0.37	0.11	0.27	0.40	0.31	0.35	0.11-0.43
17	0.21	0.18	0.22	0.09	0.13	0.57	0.24	0.26	0.08	0.24	0.18	0.19	0.20	0.08	0.11	0.21	0.08-0.57
Interpersonal																	
1	0.18	0.34	0.22	0.42	0.25	0.33	0.07	0.31	0.28	0.18	0.22	-0.03	0.38	0.31	0.22	0.25	0.01-0.33
7	0.13	-0.12	0.16	0.09	0.31	0.16	0.22	0.18	-0.07	0.04	-0.10	0.03	-0.10	-0.11	0.19	0.06	0.03-0.24
11	0.68	0.75	0.76	0.83	0.47	0.79	0.78	0.65	0.60	0.77	0.69	0.56	0.70	0.56	0.67	0.68	0.47-0.83
14	0.22	0.07	0.17	0.32	0.45	0.20	0.14	0.36	0.19	0.17	-0.05	0.19	0.14	0.16	0.30	0.19	-0.05-0.45
15	0.56	0.50	0.52	0.58	0.62	0.50	0.56	0.53	0.49	0.61	0.45	0.33	0.52	0.56	0.46	0.52	0.33-0.62
18	0.27	0.28	0.37	0.30	0.53	0.28	0.29	0.27	0.39	0.30	0.36	0.25	0.39	0.37	0.35	0.33	0.25-0.53
21	0.72	0.85	0.76	0.68	0.59	0.86	0.84	0.66	0.62	0.78	0.72	0.71	0.58	0.54	0.75	0.71	0.59-0.86
22	0.39	0.17	0.36	0.45	0.63	0.35	0.29	0.50	0.53	0.38	0.19	0.31	0.23	0.34	0.44	0.37	0.17-0.63
Disorganized																	
3	0.44	0.40	0.64	0.35	0.52	0.31	0.53	0.08	0.26	0.56	0.57	-0.09	0.54	0.51	0.42	0.40	0.08-0.64
6	0.70	0.74	0.70	0.54	0.77	0.62	0.74	0.44	0.44	0.70	0.78	0.05	0.74	0.75	0.65	0.62	0.05-0.78
8	0.04	0.09	0.04	-0.25	0.18	-0.15	0.08	0.06	-0.27	0.09	0.20	0.05	0.36	0.21	-0.02	0.05	0.04-0.27
13	0.36	0.28	0.38	0.10	0.33	-0.20	0.46	-0.25	0.17	0.39	0.33	0.12	0.45	0.44	0.29	0.24	0.09-0.52
19	0.54	0.79	0.59	0.58	0.65	0.42	0.54	0.49	0.40	0.57	0.55	0.07	0.70	0.57	0.49	0.53	0.07-0.79
20	-0.11	-0.02	-0.07	-0.20	-0.08	-0.32	-0.07	-0.38	-0.54	-0.07	-0.10	0.95	0.13	0.17	-0.16	0.05	0.01 -0.95

Table 6

Factor loadings for the Raine et al. (1994) model

т.															Total		
Items	Us	Spain	NZ	Italy	Australia	Belgium	UK	Tunisia	China	Canada	greece	Mauritus	Austria	Germany	sample	acro	oss samples
																Mean	Range
Positive																	
2	0.60	0.53	0.54	0.54	0.43	0.47	0.58	0.57	0.49	0.53	0.60	0.63	0.72	0.74	0.55	0.57	0.43-0.74
4	0.60	0.45	0.57	0.61	0.59	0.50	0.52	0.40	0.31	0.53	0.32	0.43	0.73	0.69	0.46	0.52	0.31-0.73
5	0.62	0.56	0.63	0.64	0.68	0.64	0.61	0.58	0.41	0.61	0.64	0.42	0.76	0.78	0.54	0.61	0.41-0.78
7	0.31	0.46	0.29	0.42	0.34	0.31	0.26	0.21	0.37	0.38	0.36	0.52	0.49	0.35	0.23	0.36	0.20-0.52
9	0.54	0.56	0.50	0.58	0.53	0.55	0.48	0.52	0.40	0.59	0.59	0.55	0.51	0.31	0.46	0.51	0.39-0.59
10	0.63	0.53	0.52	0.66	0.54	0.57	0.62	0.51	0.54	0.61	0.64	0.56	0.35	0.32	0.59	0.54	0.51-0.64
12	0.54	0.49	0.50	0.55	0.41	0.50	0.41	0.55	0.38	0.49	0.53	0.41	0.81	0.61	0.49	0.51	0.38-0.81
14	0.22	0.25	0.27	0.18	0.14	0.27	0.28	0.08	0.12	0.29	0.31	0.30	0.06	0.17	0.13	0.21	0.06-0.31
16	0.75	0.59	0.64	0.56	0.62	0.65	0.73	0.59	0.63	0.72	0.70	0.56	0.73	0.70	0.71	0.66	0.56-0.75
17	0.46	0.52	0.45	0.58	0.38	0.47	0.46	0.38	0.34	0.52	0.45	0.50	0.42	0.30	0.357	0.44	0.38-0.58
Interperso	nal																
1	0.74	0.61	0.78	0.60	0.48	0.70	0.66	0.56	0.65	0.67	0.61	0.70	0.67	0.67	0.66	0.65	0.48-0.78
7	0.53	0.35	0.61	0.29	0.59	0.49	0.57	0.28	0.29	0.47	0.48	0.11	0.50	0.60	0.57	0.44	0.11-0.61
9	0.52	0.21	0.29	0.28	0.23	0.23	0.17	0.26	0.47	0.19	0.21	0.15	0.50	0.57	0.31	0.31	0.15-0.57
11	0.79	0.73	0.76	0.88	0.47	0.81	0.87	0.70	0.60	0.79	0.87	0.60	0.84	0.77	0.73	0.75	0.47-0.87
14	0.59	0.52	0.58	0.45	0.67	0.54	0.52	0.46	0.45	0.52	0.48	0.32	0.58	0.52	0.63	0.52	0.32-0.57
15	0.73	0.65	0.73	0.73	0.59	0.72	0.64	0.72	0.40	0.69	0.65	0.53	0.65	0.68	0.59	0.65	0.40-0.73
17	0.35	0.33	0.31	0.28	0.65	0.20	0.32	0.08	0.28	0.25	0.38	0.10	0.36	0.50	0.41	0.31	0.08-0.65
18	0.79	0.71	0.78	0.49	0.80	0.66	0.77	0.70	0.87	0.78	0.77	0.43	0.75	0.73	0.77	0.72	0.43-0.87
21	0.82	0.80	0.77	0.87	0.68	0.84	0.88	0.77	0.71	0.84	0.88	0.73	0.90	0.82	0.79	0.81	0.68-0.90
22	0.69	0.63	0.71	0.41	0.68	0.61	0.68	0.40	0.79	0.71	0.67	0.58	0.69	0.63	0.65	0.64	0.40-0.79

D	Disorganize	ed																
3		0.67	0.54	0.61	0.71	0.71	0.60	0.67	0.44	0.55	0.63	0.64	0.62	0.70	0.68	0.64	0.63	0.44-0.71
6		0.77	0.59	0.74	0.71	0.81	0.75	0.80	0.52	0.76	0.72	0.68	0.73	0.83	0.79	0.71	0.73	0.59-0.83
8		0.80	0.63	0.74	0.71	0.66	0.56	0.78	0.57	0.64	0.76	0.66	0.72	0.78	0.79	0.68	0.70	0.57-0.79
13	3	0.65	0.44	0.52	0.64	0.61	0.39	0.60	0.60	0.26	0.62	0.61	0.56	0.60	0.53	0.55	0.55	0.39-0.62
19	9	0.79	0.77	0.80	0.68	0.85	0.60	0.79	0.67	0.78	0.79	0.69	0.67	0.87	0.80	0.76	0.76	0.60-0.87
20	0	0.77	0.67	0.77	0.74	0.65	0.76	0.75	0.70	0.70	0.76	0.76	0.71	0.79	0.85	0.73	0.74	0.65-0.85
F	actor Corr	elations																
F.	2-F1	0.64	0.04	0.52	0.77	0.56	0.67	0.60	0.60	0.59	0.60	0.553	0.79	0.50	0.42	0.62	0.58	0.04-0.79
F.	3-F1	0.38	0.05	0.25	0.32	0.22	0.23	0.30	0.27	0.20	0.30	0.264	0.59	0.40	0.21	0.29	0.29	0.05-0.59
F.	3-F2	0.70	0.04	0.61	0.57	0.47	0.62	0.59	0.68	0.73	0.62	0.656	0.80	0.61	0.67	0.65	0.59	0.04-0.80