

PAIN Publish Ahead of Print DOI: 10.1097/j.pain.000000000002082

Bidirectional longitudinal associations between loneliness and pain, and the role of inflammation Anna Loeffler^{a,b}, Andrew Steptoe^{a,*}

^aDepartment of Behavioural Science and Health, University College London, London, UK

^bInstitute for Social Science Research, University of Queensland, Brisbane, Queensland, Australia

*Corresponding author: Department of Behavioural Science and Health, University College London, 1-19 Torrington Place, London WC1E 6BT, UK, a.steptoe@ucl.ac.uk

This is an open access article distributed under the Creative Commons Attribution License 4.0 (CCBY), which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

Abstract

Pain and loneliness are consistently associated, but the direction of the relationship is uncertain. We assessed bidirectional associations over a 4 year period in a sample of 4,906 men and women (mean 65.1 ± 8.72 years) who were participants in the English Longitudinal Study of Ageing. The role of inflammation in these links was also investigated. Pain was defined by reports of being often troubled by pain at a moderate or severe intensity, while loneliness was measured using the shortened UCLA scale. Age, sex, ethnicity, educational attainment, wealth as a marker of socioeconomic resources, marital status, physical activity, and depressive symptoms were included as covariates. We found that baseline loneliness was associated with pain 4 years later after adjusting for baseline pain and other

covariates (odds ratio (OR) = 1.25, 95% CI 1.06-1.47, P = 0.007). Similarly, baseline pain independently predicted loneliness 4 years later (OR = 1.34, 95% CI 1.14-1.58, P = 0.001). Associations remained significant after additional adjustment for baseline mobility impairment. Likelihood of pain on follow-up was heightened when baseline loneliness was accompanied by elevated C-reactive protein concentration (OR = 1.50, 95% CI 1.13-2.00, P = 0.006), while inflammation did not predict future loneliness or contribute to the association between baseline pain and future loneliness. Both pain and loneliness are distressing experiences that impact well-being and quality of life. We conclude that there were bidirectional longitudinal relationships between pain and loneliness in this representative sample of older men and women, but that the mechanisms underlying these processes may differ.

Keywords: Loneliness; chronic pain; Longitudinal: C-reactive protein

1. Introduction

Chronic pain and loneliness are common problems among older adults. Population studies indicate that the prevalence of chronic pain is more than 50% in people age 65 and over, ^{1; 14; 30} while loneliness contributes to increased risk of physical and cognitive functional decline, mental ill-health and to cardiovascular disease.^{3; 32; 39} Loneliness and pain are positively associated,³⁴ and it has been argued that there are common brain mechanisms underlying physical and social pain.^{10; 11; 15} However, the temporal relationship between loneliness and pain is uncertain. Pain may limit activity and social engagement, thereby contributing to loneliness, whereas the stress of loneliness may exacerbate pain.

Longitudinal studies are valuable in identifying the temporal sequence. A study of cancer survivors, benign controls, and older adults showed that loneliness predicted increases in pain up to 4 years later independently of sleep problems, physical exercise and other covariates.¹⁹ By contrast, an

analysis of the Jerusalem Longitudinal Cohort Study found no association between loneliness and chronic pain 7 years later in people aged 70-90 years at baseline.³⁸ Results for the reverse relationship have been inconsistent as well. An investigation of the Health and Retirement Study over a 4 year period indicated that incident loneliness was predicted by pain, but only if pain was present both at baseline and follow-up.¹² Another community study documented cross-sectional associations between pain and loneliness, but found that baseline pain was minimally related to future loneliness.⁴

The potentially bidirectional links between pain and loneliness have typically been investigated in separate studies, making it difficult to determine the relative importance of the two temporal sequences. We therefore tested associations between loneliness and future pain, and pain and future loneliness, in a single study of a nationally representative population sample involving 4,906 men and women aged 52 and older. We also studied the contribution of a number of factors that potentially explain part of the association. These include socioeconomic status which is linked both to pain and loneliness, and physical inactivity.^{5; 29; 35} Another potential confounder is depression, since depressive symptoms are closely linked both with loneliness and with pain.^{7; 30}

The second aim of this study was to evaluate the possible role of inflammation in linking pain and loneliness. Inflammation contributes to the experience of many types of disease-related pain, and there are associations between inflammatory biomarkers and nonspecific pain.²³ Low grade systemic inflammation may induce sensitization of pain pathways and peripheral nociceptors.^{2; 16} There is also evidence that inflammation is related to loneliness, possibly as part of the stress response associated with perceptions of social isolation, Thus loneliness is associated with proinflammatory gene expression,⁸ and with elevated levels of markers of inflammation such as C-reactive protein (CRP) and interleukin-6.^{28; 33} It is therefore conceivable that inflammation constitutes a biological pathway through which loneliness augments risk of future pain, and vice versa. We therefore tested whether heightened inflammation indexed by elevated CRP augmented the association between loneliness and future pain, or the link between pain and future loneliness, 4 years later.

2. Methods

Data were analysed from the English Longitudinal Study of Ageing (ELSA) a nationally-representative population study of men and women age 50 and older living in England.³⁷ The study started in 2002, and data are collected every 2 years using face-to-face computer assisted personal interviews held in participants' homes, coupled with a self-completion questionnaire. The baseline for these analyses was wave 2 (2004) since this was the first wave in which loneliness was assessed. A total of 8,039 participants completed measures of loneliness and pain in Wave 2, of whom 5,183 also completed assessments 4 years later in Wave 4 (2008). Data on covariates were missing for 277 individuals, therefore the analytic sample was 4,906 (2,158 men and 2,748 women). There were no significant differences between individuals included and excluded from the longitudinal analyses in baseline levels of loneliness or pain. The analyses of the role of inflammation were carried out on 3,701 participants who had blood samples obtained during a separate home visit by study nurses. The study was approved through the National Research Ethics Service, and all participants provided informed consent.

2.1. Measures

We measured pain by participant self-report, using the questions, 'Are you often troubled with pain?' and, if so, 'How bad is the pain most of the time?' (with options of mild, moderate, or severe). In line with previous research, we specifically focused on pain that was classed as moderate or severe.^{13; 42} Loneliness was measured with the three-item short form of the Revised UCLA loneliness scale.¹⁸ Each item was scored on a 3-point scale: *hardly ever or never, some of the time*, and *often*. We defined those

individuals who reported they were lonely some of the time on at least two items as lonely. Both pain and loneliness were assessed at baseline (2004) and follow-up (2008).

The covariates in these analyses included sex, age, ethnicity, education, wealth, marital status, physical activity and depressive symptoms, as these factors are potentially associated with chronic pain and loneliness. Age was modelled as a continuous variable. Participants were categorised into white European and other because the number of participants of non-European origin was very small.³⁷ Education was classified according to the person's highest educational qualification into three categories: low (no qualifications), intermediate (qualifications at the end of state-regulated schooling), and higher (high school graduation up to university degree). Wealth is a robust indicator of economic resources among older people,⁹ and was measured with a detailed assessment of financial, housing and physical wealth (such as land, business wealth and jewellery), excluding pension wealth. Respondents were divided into married/cohabiting and not married at baseline. Physical activity was assessed by a series of questions concerning frequency of vigorous, moderate and light activities categorised according to their metabolic equivalent (MET). The vigorous activities corresponded to $MET \ge 6$, moderate to \geq 3.5 and <6, and light \geq 2 to <3.5. They were subsequently classified into 5 levels: Level 0 - No moderate and no vigorous activity; Level 1 – Moderate activity once a week or 1–3 times a month and no vigorous activity; Level 2 - Moderate activity more than once a week and no vigorous activity, or vigorous activity 1-3 times a month and no moderate activity; Level 3 - Moderate activity once a week or more and vigorous activity once a week or 1-3 times a month, or vigorous activity once a week and moderate activity 1–3 times a month or never, or moderate activity 1–3 times a month and vigorous activity 1–3 times a month; Level 4 - Vigorous activity more than once a week, with or without moderate activity. ²⁶ Depressive symptoms were measured using the 8-item Centre for Epidemiologic Studies Depression Scale (CES-D),³⁶ a shortened scale with a Cronbach α of 0.78 in this sample. The item on loneliness was excluded,⁶ and we used a score of ≥ 3 to indicate the presence of significant depressive symptoms.³¹

High sensitivity plasma CRP concentration was assayed from blood samples obtained during study nurse visits to participants' homes at baseline. We classified individuals with values $\geq 3 \text{ mg/L}$ as having raised CRP levels, since this is an established threshold in population and clinical studies. Results were the same when CRP was analysed as a continuous variable. Individuals with values $\geq 20 \text{ mg/L}$ were excluded because high values may indicate the presence of an acute infection or serious acute illness.

In sensitivity analyses we also included mobility impairment as a covariate. Participants were asked about their ability to carry out 10 tasks such as climbing a flight of stairs without resting, or picking up a small coin. They were subsequently classified into those who did or did not experience any mobility impairment at baseline.

2.2. Statistical analysis

Data were analysed using a series of multivariable logistic regressions, and results are presented as odds ratios (OR) with 95% confidence intervals (CI). In longitudinal analysis, we first assessed whether loneliness at baseline was a predictor of chronic pain 4 years later, independently of baseline chronic pain and other covariates. We therefore computed the adjusted odds of reporting chronic pain on follow-up with the low loneliness group as the reference category. Four models were tested. Model 1 included age, sex along with loneliness as determinants of future pain. In model 2, we added baseline pain to the regression. Ethnicity, education, wealth, marital status and physical activity were added in model 3 to discover the extent to which these factors explained associations between loneliness and future pain. Depressive symptoms were added in model 4 to detect the extent to which the association of loneliness with future pain was independent of depression.

A second set of analyses restricted analysis to people who had no pain at baseline, so tested the association between baseline loneliness and incident pain using the same set of models as described above. We also computed the proportion of people with and without loneliness at baseline who developed chronic pain over the follow-up period.

A parallel strategy was used to analyse the association between chronic pain and future loneliness. Model 1 included age, sex, and baseline chronic pain as determinants of loneliness 4 years later. Model 2 added baseline loneliness, while ethnicity, education, wealth, marital status and physical activity were included in model 3. Depressive symptoms were added in model 3. The analysis of incident loneliness tested the association between baseline pain and the development of loneliness among people who were not lonely at baseline.

In sensitivity analyses, the repeated both the regressions on future chronic pain and future loneliness, adding mobility impairment as an additional covariate to the fully adjusted models,

Interactions with inflammation were tested by categorising respondents into groups of low or high baseline loneliness and low or high CRP for the analyses of chronic pain at 4 years. The adjusted ORs of pain for each category were assessed with the low lonely / low CRP group as the reference category. Likewise, we classified people into baseline pain or no pain and low or high CRP for the analyses of loneliness at 4 years, estimating the adjusted OR of loneliness with the no pain / low CRP group as the reference category.

3. Results

The study sample comprised 2,158 men and 2,748 women with an average age of 65.1 ± 8.72 years. The large majority were of white European background with moderate levels of education; less than one third had college education (Table 1). Men tended to be more educated, wealthier and more likely to be married than women. Significant depressive symptoms were present in 11.3% of participants overall, and were more common among women. Physical activity levels were significantly higher in men than women. Around 30% reported moderate or high loneliness, while 24.6% experienced moderate/severe pain at baseline. Over the 4 year follow-up period, 16.5% of respondents with no pain at baseline reported moderate/severe pain, while 37.1% who were in pain at baseline were no longer experiencing pain on follow-up. The incidence of loneliness among people who were not lonely at baseline was 17.4%, whereas 33.5% of lonely people at baseline were no longer lonely 4 years later. There were significant gender differences in loneliness and pain, with women reporting higher levels than men. Overall, 31.9% of participants had CRP concentrations \geq 3.0 mg/L, with higher rates among women. Cross-sectionally, loneliness at baseline was positively associated with moderate/severe pain after adjustment for age, sex, ethnicity, education, wealth, marital status, physical activity and depression (partial r = 0.063, P < 0.001).

3.1 Loneliness as a predictor of future chronic pain

The longitudinal associations between loneliness and pain 4 years later are summarized in Table 2. Loneliness at baseline was related to the presence of moderate/severe pain 4 years later independently of age and sex (model 1), with an adjusted odds ratio of 1.45 (95%CI 1.25-1.69) when baseline pain had also been taken into account (model 2). The association was reduced to 1.35 after ethnicity, education, wealth, marital status and physical activity had been taken into account (model 3), and fell further after the inclusion of depressive symptoms in the regression (model 4). Nonetheless, the association remained significant, with 25% higher adjusted odds of moderate/severe pain per unit increase in loneliness. The full details of regression model 4 are provided in Supplementary Table 1 (available at http://links.lww.com/PAIN/B182). Other independent determinants were of pain after 4 years were baseline pain, female gender, lesser wealth, physical inactivity, being married, and depressive symptoms at baseline. In sensitivity analysis including baseline mobility impairment as a

covariate, the association between loneliness and follow-up pain remained significant (Adjusted OR = 1.19, 95%CI 1.01-1.40, P = 0.039).

The analysis of incident pain among the 3,697 individuals who were free of pain at baseline is summarised in Supplementary Table 2 (available at http://links.lww.com/PAIN/B182). The results are comparable to those in the primary analysis, with an adjusted odds of new pain reports of 1.39 (95%CI 1.13-1.79, P = 0.001) for lonely individuals in the fully adjusted model. In absolute terms, this translated into incident pain in 20.0% of participants who were lonely at baseline, compared with 15.2% in those who were not lonely.

3.2. Moderate/severe pain as a predictor of future loneliness

Baseline moderate/severe pain was associated with loneliness 4 years later, independently of age, sex, baseline loneliness, ethnicity, education, wealth, marital status, physical activity and depressive symptoms (Table 2). The adjusted OR for overall pain was reduced from 1.86 in model 1 to 1.34 in model 4, indicating that baseline loneliness, sociodemographic factors and depressive symptoms explained a substantial proportion of the association between baseline pain and future loneliness. Nonetheless, the OR of 1.34 indicates that among older people, pain does presage greater loneliness in the future. The full regression results from model 4 are detailed in Supplementary Table 3 (available at http://links.lww.com/PAIN/B182) and indicate that baseline loneliness, older age, female gender, lower wealth, unmarried status, and depressive symptoms were independently associated with future loneliness. The sensitivity analysis indicated that association between loneliness and follow-up pain remained significant after baseline mobility impairment had been included as a covariate (Adjusted OR = 1.25, 95%CI 1.05-1.49, P = 0.012).

The analysis of incident loneliness on follow-up among people who were not lonely at baseline is detailed in Supplementary Table 2 (available at http://links.lww.com/PAIN/B182). In the fully

adjusted model, the odds of incident loneliness were 1.58 (1.28-1.95), P < 0.001, slightly greater than in the analysis of the full sample. Numerically, 15.9% of participants with no pain at baseline became lonely on follow-up, compared with 23.1% of those who were in pain at baseline.

3.3. Interactions with inflammation

We tested the involvement of inflammation in the bidirectional links between loneliness and chronic pain in the 3,701 participants for whom high sensitivity CRP was available at baseline. Baseline CRP was associated with both baseline pain (r = 0.13, P < 0.001) and baseline loneliness (r = 0.034, P =0.043). Longitudinally, CRP concentration predicted future pain (OR = 1.22, 95% CI 1.02-1.45, P =0,029), but not future loneliness (OR = 0.97, 95% CI 0.81-1.16, P = 0.76). To test the contribution of inflammation in the relationship between loneliness and future chronic pain, we compared four groups defined by the presence of loneliness and elevated CRP at baseline. As shown in Table 3, 49.2% of participants were in the low lonely / low CRP group, and the smallest proportion were in the high lonely / high CRP group (9.8%). After all covariates had been taken into account, there was a significant interaction between loneliness-inflammation group and future chronic pain. Compared with the low lonely / low CRP group, chronic pain on follow-up was significantly more common in all other groups. Notably, there was a trend across groups (P = 0.003), with the highest odds of chronic pain in the high lonely / high CRP group (OR = 1.50, 95% CI 1.13-2.00, P = 0.006). These results are illustrated in Figure 1 (upper panel), where it is evident that 7.1% more individuals in the high lonely / high CRP group experienced moderate or several chronic pain on follow-up than did the low lonely / low CRP group, and the other two groups showed intermediate levels.

The corresponding analysis that tested the combination of baseline pain and CRP in relation to future loneliness is also summarised in Table 3. In this case, pain was the primary determinant of future loneliness, and CRP concentration had no additional role. Thus the adjusted ORs for baseline pain / low

CRP and baseline pain / high CRP were almost identical (1.34 and 1.35 respectively). As can be seen in Figure 1 (lower panel), future loneliness was more common among participants who reported moderate/severe pain at baseline compared to those with no pain. It appears therefore that inflammation augmented the association of baseline loneliness with future pain, but did not play a role in the link between baseline pain and future loneliness.

4. Discussion

These analyses examined the bidirectional associations between pain and loneliness over a 4-year period in a large population sample of older men and women. We found that loneliness at baseline predicted moderate/intense pain on follow-up, and that pain at baseline was related to later loneliness, confirming that there are two-way associations. The links were explained in part by demographics, physical activity, depressive symptoms, and mobility impairment, but even when these factors had been taken into account, the associations remained robust. Analyses of a subsample in which CRP was measured suggested a different role of inflammation in the two directions of association. Higher levels of inflammation appeared to augment the relationship between baseline loneliness and future pain, such that the combination of loneliness and inflammation conferred a greater risk of pain on follow-up than either factor individually. By contrast, inflammation did not play a role in the association between pain at baseline and future loneliness.

The prevalence of moderate/severe pain increased slightly between baseline when participants were aged 65.1y on average, and 4 years later, with a larger increase among women (3.9%) than men (2.7%). Loneliness increased by only a modest extent on average, as has been observed previously.⁴¹ However, there were shifts in both directions in a substantial minority of respondents, with increases and decreases in pain and loneliness over time. The strongest predictors of future pain and loneliness were baseline levels of the two experiences. There was a 6-fold increase in risk of pain on follow-up

for people who were in pain at baseline, while lonely individuals at baseline were 7 times more likely to be lonely 4 years later. Notably, the strength of associations between baseline loneliness and future pain, and baseline pain and future loneliness, were very similar, with adjusted ORs of 1.25 and 1.34 respectively. The other factors entered into the statistical models also explained similar proportions of the two relationships. Computing the proportion of association explained each set of factors,²⁴ it is apparent that 22% of the association between baseline loneliness and future pain was explained by sociodemographic factors (ethnicity, education, wealth), marital status, and physical activity. Analogously, 24% of the link between baseline pain and later loneliness was explained by these same factors. The pattern reinforces the close interdependence of pain and loneliness in this sample. One influential theory argues that societal exclusion activates brain regions also responsible for processing physical pain.¹⁰ It is suggested that evolutionarily, pain computations responsible for preventing danger were adopted by social attachment systems to trigger in response to social separation in order to avoid harmful consequences,²⁵ creating a centrally driven association between loneliness and physical pain. However, the direct overlap between neural processes implicated in physical and social pain is uncertain, and the association may reflect more fundamental motivational pathways.¹⁵

While the associations between loneliness and later pain parallel previous research,¹⁹ the longitudinal results for pain being related to later loneliness were more robust than those of other population studies. For example, analyses of the Health and Retirement study reported by Emerson et al¹² did not find that baseline pain predicted future loneliness, though the combination of pain both at baseline and follow-up was related to loneliness. Another study of a community sample from Arizona found minimal associations between pain intensity or pain frequency and loneliness measured 6-53 months later.⁴ Differences may relate to the definitions of pain and measures of loneliness used in these investigations, and in the selection of covariates.

There was some overlap between the other variables included in these analyses that were related both with future pain and future loneliness. Female gender, lower socioeconomic status defined by wealth, and depressive symptoms were associated with both directions of the pain – loneliness relationship. This reflects the greater vulnerability of these groups to chronic pain and loneliness. By contrast, physical inactivity was a risk factor for future pain but not loneliness, while older age and being unmarried were associated with subsequent loneliness but not pain. The role of physical activity in protecting against episodes of future pain has been established.³⁵ Loneliness is also correlated with physical inactivity, but there is less evidence that for longitudinal associations at older ages.²¹ Increasing age and not having a marital partner are recognised risk factors for loneliness, while the links between marital status and pain are more complicated, and vary with the nature and quality of marital interactions.²²

Our analyses involving inflammation identified some differences between the two directions of association. These analyses were conducted on a smaller sample (3,701 compared with 4,906) because CRP was obtained on a separate occasion from the main assessments. Not all participants had a study nurse visit, and blood samples were not obtained from everyone because of the presence of clinical factors such as clotting disorders, or refusal. However, these analyses suggest that inflammation may play a role in the relationship between loneliness and future pain, but not in the association between baseline pain and future loneliness. This is apparent in Table 3, where the adjusted OR for people in the high lonely/high CRP group (1.50) was greater than for the high lonely/low CRP or low lonely/high CRP groups (1.36 and 1.27 respectively), with a significant trend across groups. Loneliness is associated with proinflammatory gene expression, with heightened inflammatory responses to stress, and with raised IL-6 concentration.^{8: 17: 33} Not all lonely people express heightened inflammation, but the combination seems relevant to future pain experience. This is not to imply that inflammation is the only or even the key pathway linking loneliness with pain. It has been argued that social isolation

increases pain inference and pain intensity mediated by pain vigilance, the abnormal focus on signals of pain and potential injury.²⁰ Further, loneliness may increases pain sensitivity, thereby amplifying the individual's future pain experience.

By contrast, CRP was not relevant to the association between baseline pain and future loneliness. The increased likelihood of being lonely on follow-up was the same when baseline pain was coupled with lower or higher CRP concentrations. Even though there was a cross-sectional relationship between CRP and loneliness at baseline, CRP concentration on its own was not associated with future loneliness. This indicates that other biological, social or psychological pathways not involving inflammation are relevant to the link between pain and later loneliness. Several factors may be involved. For instance, the experience of moderate/severe pain may curtail social activity, involvement in family events, and reduce cultural engagement, all of which are protective against future pain. ^{13; 27} People with chronic pain can feel misunderstood when loved ones judge or ignore the suffering person's feelings, promoting a sense of isolation.⁴⁰

This study has a number of limitations. Although loneliness was assessed using a standardised questionnaire, pain was measured with a single item concerning pain intensity, and duration of pain experience was indexed by asking participants whether they were 'often' troubled by pain. The study is observational, and we took into account a wide range of sociodemographic, behavioural, and emotional factors that potentially confound the links between pain and loneliness. However, there may be unmeasured factors that explain the association, so causal conclusions cannot be drawn. Only a small proportion of the sample were of non-white European origin, so results may not generalise to other sectors of the population. It should also be noted that neither pain nor loneliness are the primary focus of assessments in ELSA, reducing the possibility of expectations concerning their relationship influencing responses.

In conclusion, this analysis of a large representative population sample confirmed bidirectional associations between loneliness and pain over a 4-year period. The magnitude of associations was similar in the two directions, and remained after adjusting statistically for demographic factors, physical activity levels and depressive symptoms. Nonetheless, we found some distinctions between the mechanisms involved in these relationships, with systemic inflammation playing a more prominent role in the association between loneliness and subsequent pain than in the link between pain and later loneliness. This suggests that the pathways responsible for the associations differ. Both pain and loneliness are distressing experiences that impact well-being and quality of life. Greater understanding of these bidirectional relationships may help identify methods for breaking these links.

Conflict of interest statement

The Authors have no conflicts of interest to declare

Acknowledgements

The English Longitudinal Study of Ageing is administered by a team of researchers based at the University College London, NatCen Social Research, the Institute for Fiscal Studies, and the University of Manchester. Funding is provided by National Institute of Aging (R01AG017644) and by a consortium of UK government departments coordinated by the National Institute of Health Research. ELSA data are available through the UK Data Service.

References

- [1] Andrews P, Steultjens M, Riskowski J. Chronic widespread pain prevalence in the general population: A systematic review. Eur J Pain 2018;22(1):5-18.
- [2] Ansar W, Ghosh S. C-reactive protein and the biology of disease. Immunol Res 2013;56(1):131-142.
- [3] Beutel ME, Klein EM, Brahler E, Reiner I, Junger C, Michal M, Wiltink J, Wild PS, Munzel T, Lackner KJ, Tibubos AN. Loneliness in the general population: prevalence, determinants and relations to mental health. BMC Psychiatry 2017;17(1):97.
- [4] Boggero IA, Sturgeon JA, Arewasikporn A, Castro SA, King CD, Segerstrom SC. Associations of pain intensity and frequency with loneliness, hostility, and social functioning: cross-sectional, longitudinal, and within-person relationships. Int J Behav Med 2019;26(2):217-229.
- [5] Booher L. The impact of low socioeconomic status in adults with chronic pain: an integrative review. Orthop Nurs 2019;38(6):381-389.
- [6] Cacioppo JT, Hawkley LC, Thisted RA. Perceived social isolation makes me sad: 5-year crosslagged analyses of loneliness and depressive symptomatology in the Chicago Health, Aging, and Social Relations Study. Psychol Aging 2010;25(2):453-463.
- [7] Chou KL. Reciprocal relationship between pain and depression in older adults: evidence from the English Longitudinal Study of Ageing. J Affect Disord 2007;102(1-3):115-123.
- [8] Cole SW, Hawkley LC, Arevalo JM, Sung CY, Rose RM, Cacioppo JT. Social regulation of gene expression in human leukocytes. Genome Biol 2007;8(9):R189.
- [9] Demakakos P, Biddulph JP, Bobak M, Marmot MG. Wealth and mortality at older ages: a prospective cohort study. J Epidemiol Community Health 2016;70(4):346-353.

- [10] Eisenberger NI. The pain of social disconnection: examining the shared neural underpinnings of physical and social pain. Nature Rev Neuroscience 2012;13(6):421-434.
- [11] Eisenberger NI, Lieberman MD, Williams KD. Does rejection hurt? An FMRI study of social exclusion. Science 2003;302(5643):290-292.
- [12] Emerson K, Boggero I, Ostir G, Jayawardhana J. Pain as a risk factor for loneliness among older adults. J Aging Health 2018;30(9):1450-1461.
- [13] Fancourt D, Steptoe A. Physical and psychosocial factors in the prevention of chronic pain in older age. J Pain 2018;19(12):1385-1391.
- [14] Fayaz A, Croft P, Langford RM, Donaldson LJ, Jones GT. Prevalence of chronic pain in the UK: a systematic review and meta-analysis of population studies. BMJ open 2016;6(6):e010364.
- [15] Ferris LJ, Jetten J, Hornsey MJ, Bastian B. Feeling hurt: revisiting the relationship between social and physical pain. Rev Gen Psychol 2019;23(3):320-335.
- [16] Gao X, Kim HK, Chung JM, Chung K. Reactive oxygen species (ROS) are involved in enhancement of NMDA-receptor phosphorylation in animal models of pain. Pain 2007;131(3):262-271.
- [17] Hackett RA, Hamer M, Endrighi R, Brydon L, Steptoe A. Loneliness and stress-related inflammatory and neuroendocrine responses in older men and women.
 Psychoneuroendocrinology 2012;37:1801-1809.
- [18] Hughes ME, Waite LJ, Hawkley LC, Cacioppo JT. A short scale for measuring loneliness in large surveys: Results from two population-based studies. Res Aging 2004;26:655-672.
- [19] Jaremka LM, Andridge RR, Fagundes CP, Alfano CM, Povoski SP, Lipari AM, Agnese DM, Arnold MW, Farrar WB, Yee LD, Carson WE, 3rd, Bekaii-Saab T, Martin EW, Jr., Schmidt CR, Kiecolt-Glaser JK. Pain, depression, and fatigue: loneliness as a longitudinal risk factor. Health Psychol 2014;33(9):948-957.

- [20] Karayannis NV, Baumann I, Sturgeon JA, Melloh M, Mackey SC. The impact of social isolation on pain interference: a longitudinal study. Ann Behav Med 2019;53(1):65-74.
- [21] Kobayashi LC, Steptoe A. Social isolation, loneliness, and health behaviors at older ages: longitudinal cohort study. Ann Behav Med 2018;52(7):582-593.
- [22] Leonard MT, Cano A, Johansen AB. Chronic pain in a couples context: a review and integration of theoretical models and empirical evidence. J Pain 2006;7(6):377-390.
- [23] Lim YZ, Wang Y, Cicuttini FM, Hughes HJ, Chou L, Urquhart DM, Ong PX, Hussain SM. Association between inflammatory biomarkers and nonspecific low back pain: a systematic review. Clin J Pain 2020;36(5):379-389.
- [24] Lin DY, Fleming TR, De Gruttola V. Estimating the proportion of treatment effect explained by a surrogate marker. Stat Med 1997;16(13):1515-1527.
- [25] Macdonald G, Leary MR. Why does social exclusion hurt? The relationship between social and physical pain. Psychol Bull 2005;131(2):202-223.
- [26] McMunn A, Hyde M, Janevic M, Kumari M. Health. In: M Marmot, J Banks, R Blundell, C Lessof, J Nazroo, editors. Health, Wealth and Lifestyles of the Older Population in England. London: Institute for Fiscal Studies, 2003. pp. 207-428.
- [27] Moulin DE, Clark AJ, Speechley M, Morley-Forster PK. Chronic pain in Canada--prevalence, treatment, impact and the role of opioid analgesia. Pain Res Manag 2002;7(4):179-184.
- [28] Nersesian PV, Han HR, Yenokyan G, Blumenthal RS, Nolan MT, Hladek MD, Szanton SL. Loneliness in middle age and biomarkers of systemic inflammation: Findings from Midlife in the United States. Soc Sci Med 2018;209:174-181.
- [29] Niedzwiedz CL, Richardson EA, Tunstall H, Shortt NK, Mitchell RJ, Pearce JR. The relationship between wealth and loneliness among older people across Europe: Is social participation protective? Prev Med 2016;91:24-31.

- [30] Patel KV, Guralnik JM, Dansie EJ, Turk DC. Prevalence and impact of pain among older adults in the United States: findings from the 2011 National Health and Aging Trends Study. Pain 2013;154(12):2649-2657.
- [31] Radloff LS. The CES-D Scale: A self-report depression scale for research in the general population. Appl Psychol Meas 1977;1:385-401.
- [32] Shankar A, McMunn A, Demakakos P, Hamer M, Steptoe A. Social isolation and loneliness: Prospective associations with functional status in older adults. Health Psychol 2017;36(2):179-187.
- [33] Smith KJ, Gavey S, NE RI, Kontari P, Victor C. The association between loneliness, social isolation and inflammation: A systematic review and meta-analysis. Neurosci Biobehav Rev 2020;112:519-541.
- [34] Smith TO, Dainty JR, Williamson E, Martin KR. Association between musculoskeletal pain with social isolation and loneliness: analysis of the English Longitudinal Study of Ageing. Br J Pain 2019;13(2):82-90.
- [35] Steffens D, Maher CG, Pereira LS, Stevens ML, Oliveira VC, Chapple M, Teixeira-Salmela LF, Hancock MJ. Prevention of low back pain: a systematic review and meta-analysis. JAMA internal medicine 2016;176(2):199-208.
- [36] Steffick DE. Documentation of Affective Functioning Measures in the Health and Retirement Study. Ann Arbor: Survey Research Center University of Michigan., 2000.
- [37] Steptoe A, Breeze E, Banks J, Nazroo J. Cohort profile: the English Longitudinal Study of Ageing. Int J Epidemiol 2013;42(6):1640-1648.
- [38] Stessman J, Rottenberg Y, Shimshilashvili I, Ein-Mor E, Jacobs JM. Loneliness, health, and longevity. J Gerontol A Biol Sci Med Sci 2014;69(6):744-750.

- [39] Valtorta NK, Kanaan M, Gilbody S, Ronzi S, Hanratty B. Loneliness and social isolation as risk factors for coronary heart disease and stroke: systematic review and meta-analysis of longitudinal observational studies. Heart 2016;102(13):1009-1016.
- [40] Wernicke S, de Witt Huberts J, Wippert PM. The pain of being misunderstood: Invalidation of pain complaints in chronic low back pain patients. J Health Psychol 2017;22(2):135-147.
- [41] Yang K. Longitudinal loneliness and its risk factors among older people in England. Can J Aging 2018;37(1):12-21.
- [42] Yiengprugsawan V, Steptoe A. Impacts of persistent general and site-specific pain on activities of daily living and physical performance: A prospective analysis of the English Longitudinal Study of Ageing. Geriat Gerontol Int 2018;18(7):1051-1057.

Ć

Table 1Participant characteristics

	Total (n = 4905)	Men (n = 2156)	Women (n = 2749)	P difference
Age (y)	65.11 ± 8.72	64.86 ± 8.45	65.31 ± 8.92	0.074
Ethnicity (White European)	4853 (98.9%)	2133 (98.8%)	2720 (99.0%)	0.68
Education				
Lower	1586 (32.3%)	559 (25.9%)	1027 (37.4%)	< 0.001
Intermediate	1950 (39.7%)	815 (37.8%)	1135 (41.3%)	
Higher	1370 (27.9%)	784 (36.3%)	586 (21.3%)	
Wealth:				
Lowest	663 (13.5%)	245 (11.4%)	418 (15.2%)	< 0.001
2	862 (17.6%)	358 (16.6%)	506 (18.4%)	
3	997 (20.3%)	426 (19.7%)	571 (20.8%)	
4	1142 (23.3%)	531 (24.6%)	611 (22.2%)	
Highest	1240 (25.3%)	598 (27.7%)	642 (23.4%)	

	2421 (60 70()	1705 (70.00())	1606 (61 70/)	.0.001
Marital status (Married)	3421 (69.7%)	1725 (79.9%)	1696 (61.7%)	< 0.001
Depression (Significant	552 (11.3%)	169 (7.8%)	383 (13.9%)	< 0.001
	, , , , , , , , , , , , , , , , , , ,	× ,	· · · · ·	
symptoms)				
symptoms				
Physical activity	2.28 ± 1.24	2.46 ± 1.20	2.13 ± 1.25	< 0.001
Loneliness (wave 2)	1448 (29.5%)	544 (25.3%)	903 (32.9%)	< 0.001
Lonenness (wave 2)	1440 (29.3%)	344 (23.3%)	903 (32.970)	<0.001
Loneliness (wave 4)	1563 (31.9%)	574 (26.6%)	989 (36.0%)	< 0.001
Mad/anna main (man 2)	1209 (24 60/)	412 (10 10/)	795 (28.9%)	< 0.001
Mod/severe pain (wave 2)	1208 (24.6%)	413 (19.1%)	795 (28.9%)	<0.001
Mod/severe pain (wave 4)	1370 (27.9%)	470 (21.8%)	900 (32.8%)	< 0.001
		` <i>´</i>	, ,	
	11(2)(21,00())		705 (25 00)	.0.001
C-reactive protein $\geq 3 \text{ mg/L}$	1162 (31.9%)	437 (26.9%)	725 (35.9%)	< 0.001

Table 2Bidirectional associations between loneliness and chronic pain

Adjusted odds ratios (OR) with 95% confidence intervals (CI)

	Loneliness as a predictor of future chronic pain		Chronic pain as a predictor of loneliness	
	Adjusted OR	Р	Adjusted OR	Р
	(95% CI)		(95% CI)	
Model 1	1.74 (1.53-1.99)	<0.001	1.86 (1.62-2.13)	<0.001
Adjusted for age and sex				
Model 2 Additionally adjusted for baseline pain ^a or loneliness ^b	1.45 (1.25-1.69) ^a	<0.001	1.55 (1.32-1.81) ^b	<0.001
Model 3	1.35 (1.16-1.58)	< 0.001	1.42 (1.21-1.67)	<0.001
Additionally adjusted for ethnicity, education, wealth, marital status, and physical activity				

(1.14-1.58)	0.001

	Grouping	%	Adjusted OR (95% CI)	P
Loneliness and CRP	Low lonely / low CRP	49.2%	$1 (ref)^a$	0.028
as predictors of future chronic pain	Low lonely / high CRP High lonely / low CRP	21.9% 19.0% 9.8%	1.27 (1.03 - 1.58) 1.36 (1.07 - 1.71) 1.50 (1.12 - 2.00)	0.028 0.011 0.006
	High lonely / high CRP		1.50 (1.13 – 2.00)	0.006
Chronic pain and CRP as predictors of	No pain / low CRP No pain / high CRP	55.6% 21.3%	1 (ref) ^b 0.93 (0.75-1.15)	0.43
future loneliness	Pain / low CRP Pain / high CRP	12.7% 10.5%	1.34 (1.04-1.73) 1.35 (1.02-1.78)	0.023 0.036

Table 3Loneliness, chronic pain, and C-reactive protein

^{*a*} Adjusted for age, sex, baseline chronic pain, ethnicity, education, wealth, marital status, physical activity, and depressive symptoms

^b Adjusted for age, sex, baseline loneliness, ethnicity, education, wealth, marital status, physical activity, and depressive symptoms

Figure 1 legend

- Upper panel Proportion of individuals experiencing moderate/severe pain at 4 year follow up in low lonely / low CRP, low lonely / high CRP, high lonely / low CRP, and high lonely / high CRP groups, adjusted for age, sex, baseline pain, ethnicity, education, wealth, marital status, physical activity and depressive symptoms. Error bars are standard errors of the mean (s.e.m.).
- Lower panel Proportion of Individuals experiencing moderate/severe pain at 4 year follow up in low pain / low CRP, low pain / high CRP, high pain / low CRP, and high pain / high CRP groups, adjusted for age, sex, baseline loneliness, ethnicity, education, wealth, marital status, physical activity and depressive symptoms. Error bars are standard errors of the mean (s.e.m.).

