Dark Energy Survey Year 3 Results: Galaxy mock catalogs for BAO analysis

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ABSTRACT

The calibration and validation of scientific analysis in simulations is a fundamental tool to ensure unbiased and robust results in observational cosmology. In particular, mock galaxy catalogs are a crucial resource to achieve these goals in the measurement of baryon acoustic oscillation (BAO) in the clustering of galaxies. Here we present a set of 1952 galaxy mock catalogs designed to mimic the Dark Energy Survey Year 3 BAO sample over its full photometric redshift range $0.6 < z_{photo} < 1.1$. The mocks are based upon 488 ICE-COLA fast *N*-body simulations of full-sky light cones and were created by populating halos with galaxies, using a hybrid halo occupation distribution – halo abundance matching model. This model has ten free parameters, which were determined, for the first time, using an automatic likelihood minimization procedure. We also introduced a novel technique to assign photometric redshift for simulated galaxies, following a two-dimensional probability distribution with VIMOS Public Extragalactic Redshift Survey data. The calibration was designed to match the observed abundance of galaxies as a function of photometric redshift errors, and the clustering amplitude on scales smaller than those used for BAO measurements. An exhaustive analysis was done to ensure that the mocks reproduce the input properties. Finally, mocks were tested by comparing the angular correlation function $w(\theta)$, angular power spectrum C_{ℓ} , and projected clustering $\xi_p(r_{\perp})$ to theoretical predictions and data. The impact of volume replication in the estimate of the covariance is also investigated. The success in accurately reproducing the photometric redshift uncertainties and the galaxy clustering as a function of redshift render this mock creation pipeline as a benchmark for future analyses of photometric galaxy surveys.

Key words. catalogs - large-scale structure of Universe - galaxies: distances and redshifts - Galaxy: halo - methods: numerical

1. Introduction

Over recent years, a large international effort has been focused on constraining the dark energy properties, measuring the cosmological parameters with high accuracy, and testing the Lambda cold dark matter (ACDM) paradigm. This led to the development of new techniques and data combinations that allow tighter constraints: cosmic microwave background (CMB); Type Ia supernovae (SNe Ia); galaxy clustering (GC); weak lensing (WL); baryon acoustic oscillation (BAO); etc. Especially in recent years, BAO (Peebles & Yu 1970; Sunyaev & Zeldovich 1970) has become a powerful alternative to building the Hubble diagram, which now allows one estimate cosmological parameters by itself (Percival et al. 2007; Beutler et al. 2011; Alam et al. 2021). Most of the probes mentioned before require the measure-

ment of redshifts with high fidelity, giving more significance to

spectroscopic surveys (e.g., the WFC3 Infrared Spectroscopic Parallel Survey – WISP¹, Baryon Oscillation Spectroscopic Survey – BOSS², *Euclid*³, Wide-Field Infrared Survey Telescope – WFIRST⁴, Dark Energy Spectroscopic Instrument – DESI⁵). However, photometric surveys have some advantages over the spectroscopic ones. In particular, every observed galaxy can be used in the cosmological analysis although in practice is necessary to select a galaxy population that presents a prominent spectral feature that can be captured with broadband filters. Besides, many successful techniques are used to estimate true redshifts given observed photometric redshifts (photo-*z*)

http://wisps.ipac.caltech.edu

² http://www.sdss3.org/surveys/boss.php

³ https://www.euclid-ec.org/

⁴ https://wfirst.gsfc.nasa.gov/

⁵ https://www.desi.lbl.gov/

within a given uncertainty. Then, the statistical power makes imaging surveys almost as competitive as spectroscopic surveys in the measurement of galaxy clustering. One example of these techniques is the Directional Neighbourhood Fitting (DNF, De Vicente et al. 2016), which stands out as one of the most robust and accurate determinations of photo-*z*, and is thus used in several photometric surveys (see e.g., Drlica-Wagner et al. 2018; Sevilla-Noarbe et al. 2021; Euclid Collaboration 2020). The Dark Energy Survey (DES)⁶ is a badge example. DES has mapped the southern sky for six years covering an area of ~5000 deg² and has recorded data from a few hundred million distant galaxies. These numbers will be pushed even beyond by future projects (e.g., the Legacy Survey of Space and Time – Rubin LSST⁷, Spectro-Photometer for the History of the Universe, Epoch of Reionization, and Ices Explorer – SPHEREx⁸).

For the large data sets that these projects produce, the calibration and replication of scientific analysis in simulations previous to unblinding (procedure explained below), is a fundamental tool to ensure unbiased and robust results. This task requires the fulfillment of two requirements: (i) a realistic simulation of the observed cosmological volume with a final galaxy catalog that mimics the data and (ii) a large number of realizations varying the initial conditions that allows a full control of statistical uncertainties. The negative effects of introducing simulated volume replications to achieve the first requisite should not be underestimated. For example, the consequence of over-estimation of the covariances was found in the mocks used for this work. Accomplishing the second requirement by using pure N-body simulations is computationally impossible when the number of needed realizations is hundreds or thousands. Using approximate methods allows to have the desired number of runs with less computational resources (see e.g., Coles & Jones 1991; Scoccimarro & Sheth 2002; Koda et al. 2016; Avila et al. 2015; Chuang et al. 2015a; Izard et al. 2018). All these methods reduce the resolution of the simulation on small scales in exchange for computing speed. But when we focus our study on BAO scales, as the purpose of this work, it has been shown that the accuracy of these approximate methods is more than sufficient for a precise analysis (Chuang et al. 2015b; Lippich et al. 2019; Blot et al. 2021). For example, Izard et al. (2016) demonstrates that the ICE-COLA method yields a matter power spectrum within 1% for $k \leq 1$ h Mpc⁻¹ and a halo mass function within 5% of those in the N-body. Nowadays, all cosmological surveys need to develop their own galaxy mock catalogs in order to properly simulate the characteristics of the data. BOSS, eBOSS and the first year of DES data (DES Y1) for example have designed their own mocks (Manera et al. 2013; Chuang et al. 2015a,b; Kitaura et al. 2016; Avila et al. 2018; Zhao et al. 2021). The BAO analysis using the first three years of DES data (DES Y3) is structured in three papers: Carnero Rosell et al. (2022) presents a systematics analysis of the galaxy sample for DES Y3 BAO measurement, this work describes the simulations used in the analysis and the main DES Y3 BAO paper presents the angular distance constrains and cosmology in DES Collaboration (2021). An analogous work was made for DES Y1: the DES Y1 sample was presented in Crocce et al. (2019), a description of the mocks was shown in Avila et al. (2018) and DES Collaboration (2019) as the main DES Y1 BAO paper including a ~4% precision D_A measurement. In this case, the work was accompanied by several method papers (Chan et al. 2018; Ross et al. 2017; Camacho et al. 2019).

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A common process in the analysis of new Surveys data release is to blind the data in certain ways. The implementation of a rigorous process of unblinding can reduce or eliminate confirmation bias. A strict blinding strategy has been applied to this work. The final set of mocks, presented here, were completed before computing α (BAO shift parameter) on the final data vector, and before plotting the angular two-point correlation function or clustering C_{ℓ} of the DES Y3 sample. Only three pre-unblinding values of the angular clustering on scales lower than one degree (unused for the BAO analysis) were provided to calibrate the clustering amplitude of the mocks. Once the mocks were done and the data passed through the rigorous process of unblinding we could compare the clustering both in configuration and harmonic space of the mocks with the final post-unblinding measurements of the data. We refer the interested reader to DES Collaboration (2021) for more details about the unblinding process of the data.

This paper is arranged as follows. In Sect. 2 we briefly describe the reference sample. Then, Sect. 3 describes the main features of the used dark matter halo catalogs. On the one hand, we present the fast simulation used to perform the mocks for the analysis, and on the other hand, our benchmark pure N-body simulation. In Sect. 4 we detail step by step the process we followed to create our galaxy catalogs, which match the main properties of the data. One of the most important aspects of our pipeline is the automatic calibration, which is exhaustively detailed in Sect. 5. After describing how the mocks are created, in Sect. 6 we compare them with the real data and theoretical models in terms of covariance matrices and clustering measurements, both in angular configuration space and angular harmonic space. Additionally, in Sect. 7 we investigate the effect of replicated structures in the mocks. Finally, we conclude in Sect. 8 with our summary and conclusions.

2. Reference data

The BAO analysis for the DES Y3 data set is based on the Y3 GOLD catalog (Sevilla-Noarbe et al. 2021), which contains nearly 390 million objects, with a depth reaching $S/N \sim 10$ for extended objects up to i < 22.3 (*AB*), and top-of-the-atmosphere photometric accuracy under 3 mmag. This data set has been compiled from the coaddition of nearly 40 000 exposures in the grizY optical and near-infrared bands taken during the first three years of observations. It used the DECam instrument (Flaugher et al. 2015) from the Blanco Telescope in Cerro Tololo (Chile) and covers 5000 deg² of the southern hemisphere.

The catalog includes positional, photometric, and morphological information, using a multi-epoch, multi-band fitting procedure of the object's shape in every exposure where the object is present (the Single Object Fitting, or SOF, method). This is the basis of all the measurements of the object mentioned before. In addition, the Y3 GOLD catalog contains flagging information to assess the quality of the measured object, and ancillary survey information about low-quality regions in the sky, and survey properties in general (seeing, airmass, etc).

From the Y3 GOLD, we select a sample of red galaxies used to measure the BAO scale, with a color selection similar to the DES Y1 analysis presented in Crocce et al. (2019). The DES Y3 selection shows an increase in the number density of galaxies (due to improvements on the DESDM⁹ reduction), which allows us to extend the redshift range of the analysis to photometric

⁶ https://www.darkenergysurvey.org

⁷ https://www.lsst.org

⁸ http://spherex.caltech.edu/

⁹ DES Data Management in National Center for Supercomputing Applications (NCSA, Morganson et al. 2018).

Table 1. Selection process to create the BAO sample in DES Y3.

Keyword	Cut		
Gold	Observations present in the Y3 GOLD catalog		
Quality	FLAGS_GOLD=0		
Footprint	4108.47deg^2		
Color selection	(i-z) + 2.0(r-i) > 1.7		
Completeness cut	<i>i</i> < 22.3		
Flux selection	$17.5 < i < 19 + 3$ Z_MEAN		
Star-galaxy separation	EXTENDED_CLASS_MASH_SOF=3		
DNF photo-z range	[0.6–1.1]		

Notes. We refer to Sevilla-Noarbe et al. (2021) and Carnero Rosell et al. (2022) and online documentation for details about the meaning of the cuts.



Fig. 1. DES Y3 sample footprint, covering 4108.47 deg^2 of the southern sky.

redshift $z_{\text{photo}} < 1.1$ with i < 22.3 (*AB*). The selection applied to create the sample is summarized in Table 1 (details can be found in Carnero Rosell et al. 2022).

Furthermore, the footprint mask is selected accordingly removing regions with depth less than 22.3, plus additional quality cuts explained in the aforementioned references. We use all HEALPix maps with NSIDE=4096 found in the online release¹⁰. In Fig. 1, we show the angular distribution of the BAO footprint, covering 4108.47 deg².

One of the most critical aspects of any photometric analysis is the measurement of the redshift. For the sample, we characterize the N(z) (the true redshift distribution) in each tomographic bin using the "VIMOS Public Extragalactic Redshift Survey" (Guzzo et al. 2014, VIPERS) catalog as a reference, since it is a complete sample from redshift above z = 0.5 up to i = 22.5 (AB). VIPERS observed in two fields, named W1 and W4, both overlapping DES. The total overlap area is 16.324 deg^2 This provides, after several selection processes described in Carnero Rosell et al. (2022), a final sample of 8362 galaxies with spectroscopic redshift, *z*_{VIPERS}, available for redshift calibration. Carnero Rosell et al. (2022) use VIPERS to validate the performance of the photometric redshift (called Z_MEAN in the DES catalogs) and to estimate the true redshift (Z_MC) distribution of the DES Y3 sample. DNF predicts Z_MEAN as the best-value in the fitted hyper-plane and also defines Z_MC as the closest friend. In this work, we use these overlapping galaxies for the

opposite purpose, assigning z_{photo} to the simulated galaxies. In other words, we re-sample the z_{spec} vs z_{photo} diagram found from VIPERS to assign z_{photo} to mock galaxies.

3. Halo light cone catalogs

In this section, we describe the halo catalogs from dark matter simulations used in this paper. We start by describing the ICE-COLA fast simulations and then our benchmark pure *N*-body simulation, MICE Grand Challenge. It is important to make clear that both sets of simulations share the same cosmology, mass resolution, and halos found with a Friends of Friends algorithm.

3.1. ICE-COLA fast simulations

To build a large number of mocks we use a set of 488 fast *N*-body simulations generated with the ICE-COLA code (Izard et al. 2016). The COmoving Lagrangian Acceleration (COLA) method solves for the evolution of the matter density field using second-order Lagrangian Perturbation Theory (2LPT) combined with a Particle-Mesh (PM) solver to integrate the particle orbits at small scales, where 2LPT start to deviate from the full *N*-body solution (Tassev et al. 2013; Koda et al. 2016). The ICE-COLA code extends on this method to produce on-the-fly light cone halo catalogs and weak lensing maps (Izard et al. 2018).

The simulations use 2048³ particles in a box of size of 1536 Mpc h⁻¹ to match the mass resolution of the MICE Grand Challenge simulation (see Sect. 3.2). Here we use the optimal code parameters found in Izard et al. (2016), namely 40 time-steps, a starting redshift of $z_{ini} = 19$ and a PM grid of 27 times the number of particles. Halos are found with a Friends of Friends (FOF) algorithm with linking length b = 0.2. We refer the interested readers to Koda et al. (2016) and Izard et al. (2016) for more complex analysis and thorough validation of the method.

It is important to know the limitations of fast simulations to be able to use them in the range of scales where they agree with *N*-body, but also (moderately) beyond those scales with a systematic error that can be quantified (i.e, under control).

3.2. MICE Grand Challenge simulation

As was mentioned before, it is really important to validate the range of scales where we can trust fast simulations. In our case, we use the MICE Grand Challenge¹¹ (Fosalba et al. 2015b,a; Crocce et al. 2015, MICE hereafter), simulation as the benchmark *N*-body run. MICE is an all-sky light cone *N*-body simulation evolving 4096³ dark-matter particles in a ~29 Gpc³ h⁻³ comoving volume. The assumed cosmology corresponds to the best-fit of WMAP five-year data (Komatsu et al. 2009). This is consistent with a flat Λ CDM model with $\Omega_m = 0.25$, $\Omega_{\Lambda} = 0.75$, $\Omega_b = 0.044$, $n_s = 0.95$, $\sigma_8 = 0.8$ and h = 0.7.

As our reference simulation, its cosmological parameters are also used for our ICE-COLA runs, where only the initial condition changes among the 488 fast simulations. An exhaustive validation of the simulations used here has been done in Izard et al. (2016), finding a matter power spectrum within 1% for $k \leq 1 h \text{ Mpc}^{-1}$ and demonstrating that ICE-COLA fast simulation can perfectly be used for BAO purposes. In Fig. 2, we compare the halo masses and the clustering for redshift z = 0.5(blue) and z = 1 (red). The top panel shows the ratio between the MICE and the ICE-COLA halo mass function, the lowest halo

¹⁰ https://www.darkenergysurvey.org/the-des-project/ data-access/

¹¹ More information is available at http://maia.ice.cat/mice/



Fig. 2. *Top panel*: ratio of the Halo Mass Function between MICE and ICE-COLA. Blue for redshift z = 0.5 and red for z = 1. Shaded areas correspond to the standar deviation of the 488 ICE-COLA runs. *Bottom panel*: ratio between the MICE and the ICE-COLA ACF for a sample of halos with $M_{halo} > 1.46 \times 10^{12} M_{\odot}$.

mass plotted here correspond to the mass limit of the ICE-COLA mock used in this paper, $1.46 \times 10^{12} M_{\odot}$ (50 particles). Both simulations are consistent, as was found in Izard et al. (2016), within an accuracy of ~5%. The bottom panel of Fig. 2 shows the ratio between the MICE and the ICE-COLA angular two-point correlation function (ACF). In this case, the clustering is calculated using halos with more than 50 particles in a full-octant comoving output shell of width 125 Mpc and 166 Mpc for redshift z = 0.5 and z = 1, respectively. This threshold of 50 particles is not set deliberately but is a resulting minimum number of particles of the halos of the mocks presented in this work. For clustering, the accuracy is within ~5% up to scales of one degree. For higher angular distances the error increase, especially for redshift z = 1 where these scales correspond to larger 3D distances.

4. Galaxy light cone catalogs

With the simulations and the corresponding ICE-COLA halo catalogs presented in the previous section, we can start now by describing, step by step, the mechanism used to construct a galaxy mock beginning with a halo catalog. Some of our recipes described below closely follows Carretero et al. (2015) and we use a similar hybrid Halo Occupation Distribution -Halo Abundance Matching model modeling strategy presented by Avila et al. (2018) in the analysis of the DES Y1 data release. Before going into the details, it is important to remark that simulated box replications are needed to have light cones reaching higher redshifts than 0.6 (corresponding limiting redshift if we set the light cone origin at the center of a box-size of 1 536 Mpc h⁻¹) and covering the DES Y3 footprint. Four boxes on each Cartesian direction are needed (a total of 64 simulated boxes) to create a full-sky light cone up to redshift ~1.4. The implications that these replications have on the analysis are discussed in more detail in Sect. 7.

Here, we create a mock of galaxies for BAO analysis from a halo catalog of a fast simulation. However, the described procedure can be applied to any sort of halo catalog to mimic any kind of galaxies samples. A key aspect that makes this pipeline successful is the inclusion of an automatic calibration, as discussed in Sect. 5.

4.1. Halo occupation distribution

The relation between galaxies and halos is not univocal, as one halo can harbor more than one galaxy. Furthermore, the more massive the halo the higher the number of galaxies it has, reaching quantities of hundreds of galaxies in a single halo. The halo occupation distribution (HOD; Jing et al. 1998; Benson et al. 2000; Seljak 2000), describes the relation between halos and galaxies, in terms of several parameters. In other words, the HOD tells us how many galaxies a halo of a given mass has on average, $\langle N|M_{halo} \rangle$.

Two different functions are needed to describe the HOD of a sample of galaxies. One for the central galaxies and another for satellites as they model clustering on different scales in the halo model. Centrals shape large scales (halo-halo correlations) and satellite small scales (intra-halo correlations). The complexity of the function can be as high as desired, in order to match the behavior of the sample with higher accuracy. We focus here on the large-scale structure, leaving aside a complex function that would allow us to model the small scales. Therefore, we assign to each halo, one central galaxy

$$N_{\rm cent} = 1, \tag{1}$$

and a number of satellite galaxies following a Poisson distribution with mean

$$N_{\rm sat} = \frac{M_{\rm halo}}{M_1},\tag{2}$$

where M_{halo} is the mass of the halo, and M_1 is a free HOD parameter. This simple HOD is used to populate all halos with galaxies in the light cone. However, the particular sample selection is a sub sample of this generic HOD assignment. Therefore, the final values of $N_{cent}(M_{halo})$ and $N_{sat}(M_{halo})$ that compose the sample of the mocks differs from the expression defined on Eqs. (1) and (2).

Once N_{cent} and N_{sat} values are determined, the next step is to populate the halos with galaxies following these HOD quantities. One central galaxy is placed in the center of each halo and the velocity is assumed to be equal to its host halo. On the other hand, satellites galaxies are distributed inside the halo following a spherical NFW (Navarro et al. 1996) profile. Concentrations, needed to model the NFW density profile, are taken from Cooray & Sheth (2002) where the inputs are the mass and redshift of halos. We also model the velocity of galaxies with simplistic assumptions, using a simple Gaussian distribution centered at the velocity of the host halo and assuming a standard deviation proportional to the velocity dispersion of it (Sheth & Diaferio 2001; Carretero et al. 2015).

This first free parameter, M_1 , allows to control the clustering: by increasing the parameter, we undersample the most massive halos decreasing the linear bias. It also introduces a 1-halo term that fades away as we increase M_1 . More details can be found in Avila et al. (2018). Figure 3 shows the evolution of M_1 over the five tomographic bins where we assume a linear interpolation among these values. In Sect. 5 is explained in detail how these values are obtained.

At this point, we already have a general galaxy catalog with positions and velocities, made by populating all halos on the light cone. The next step is to introduce a second free parameter by setting pseudo luminosities. This second HOD parameter allows setting a sample by selecting only high luminosity galaxies from the general catalog.



Fig. 3. Evolution of the first HOD free parameter M_1 . One value for each tomographic bin. Dashed line correspond to the interpolation assumed for all the redshift range.

4.2. Pseudo-luminosity assignment

The best tracers for BAO signal are brightest galaxies (see e.g., Comparat et al. 2013) and they represent a few percent of the total number of galaxies. Not all galaxies resulting from the previous step 4.1 will enter into our selection to perform the BAO analysis. Therefore, this step is needed to select that few percent of galaxies.

An efficient way to select the tracers of our mocks is by assigning a pseudo luminosity l_p to all galaxies and then selecting the most luminous ones. To set the luminosities we rely on the halo abundance matching (HAM) techniques (Kravtsov et al. 2004; Conroy et al. 2006; Guo et al. 2010), where it is assumed that the most massive (luminous) galaxy lives in the most massive halo, the second most massive galaxy lives in the second most massive halo, and so on. On top of the mean (deterministic) relation between mass and luminosity assumed in the AM technique, we add some scattering to make this matching closer to observations. We model l_p with a Gaussian scatter around the halo mass M_{halo} in logarithmic scales:

$$\log(l_{\rm p}) = \log(M_{\rm halo}) + \Delta_{\rm LM} R_{\mu=0\,\sigma=1}^{\rm gauss},\tag{3}$$

where Δ_{LM} is our second free parameter which controls the amount of scatter. We note that l_{p} is modeled in arbitrary scales. The purposes of defining a luminosity for galaxies are two:

- 1. It allows one to match the abundance and redshift distribution of data by selecting the most luminous galaxies. More details on Sect. 4.3.
- 2. Its definition, and therefore the introduction of the second free parameter Δ_{LM} , also influences the clustering. As we decrease this value, lower mass halos go out of our selection and higher mass halos enter it, effectively increasing the bias.

Figure 4 shows the evolution of this scatter parameter Δ_{LM} as a function of redshift. In the same way as M_1 , we assume for Δ_{LM} one value for each tomographic bin and applying linear interpolation among these values as a function of redshift. The modeling explained in this subsection follows the same procedure used in Avila et al. (2018). As was mentioned earlier, in Sect. 5 it is explained how these values are obtained.

4.3. Photometric redshifts

This is perhaps one of the most dedicated and challenging steps of the mock creation pipeline. For the ICE-COLA mocks we know the true redshift and we need to model the observed one for each simulated galaxy, contrary to what happens for observations. DES is a photometric survey, therefore the measurement of the redshift has a precision much lower than spectroscopic surveys. For example, Carnero Rosell et al. (2022) show that the dispersion on the photo-*z* for the DES Y3 sample is $\sigma_{68} = 0.054$ on average for the five tomographic bins. σ_{68} is



Fig. 4. Evolution of the second HOD parameter Δ_{LM} . One value for each tomographic bin. Dashed line corresponds to the interpolation assumed for all redshift range.

defined as the value such as 68 per cent of the galaxies have $|z_{\text{photo}} - z_{\text{spec}}|/(1 + z_{\text{spec}}) < \sigma_{68}$. These uncertainties must be modeled in the simulation to have consistent clustering measurements.

Each galaxy in DES has an observed photometric redshift Z_MEAN, derived from the magnitude measured in each filter. In addition, as was explained in Sect.2, there is a small sample of 8362 galaxies for which we also have the true spectroscopic redshifts from VIPERS (z_{VIPERS}). The combination of using DES and VIPERS result in our mocks matching the abundance of DES Y3 BAO galaxies $n(Z_MEAN)$ and the redshift distribution $N(z_{VIPERS})$ on each tomographic bin of these galaxies present in both surveys.

We start by dividing the interval $\mathbb{Z}_{MEAN} = [0.6, 1.1]$ into L thin bins of with $\Delta \mathbb{Z}_{MEAN} = 0.01$. Then, according to the data, we can express the number of galaxies each l bin has as $n(\mathbb{Z}_{MEAN}^l)$. This is the first condition we want to accomplish with the mocks: match the abundance of each l thin photometric bin.

Secondly, we select *M* spectroscopic bins of width $\Delta z_{\text{VIPERS}} = 0.025$, here bins are thicker because of the smaller number of VIPERS galaxies. Then, we can determine the probability of having a galaxy in a given pair of bins (l,m) as $P(Z_\text{MEAN}^l, z_{\text{VIPERS}}^m)$. It is important to remark that this matrix is built only using those DES galaxies which have a z_{VIPERS} . Mocks need to satisfy this 2D probability distribution $P(Z_\text{MEAN}, z_{\text{VIPERS}})$ and, at the same time, match the abundance of galaxies $n(Z_\text{MEAN})$. By combining both, the number of galaxies an ICE-COLA mock should have at a given pair of bins(l,m) can be calculated as

$$A_{l,m} = n(\mathbb{Z}_{MEAN}^{l}) \times P(\mathbb{Z}_{MEAN}^{l}, z_{VIPERS}^{m}).$$
(4)

The assignment of photometric redshifts z_{photo} to galaxies in mocks is then performed in two steps. Firstly, we separate the simulated galaxies into L and M bins and assigning z_{photo} by following the distribution $P(Z_MEAN, z_{VIPERS})$. And finally, we choosing from each (l,m) pair of bins the $A_{l,m}$ most luminous galaxies, given the luminosities defined on Eq. (3). The Fig. 5 shows the resulting $n(z_{\text{photo}})$ for the mocks compared with data. Gray histograms correspond to DES Y3 sample and red points with error bars represent the mocks. The agreement is almost perfect, as expected given that it is done by construction. On the other hand, to achieve a good match on redshift distribution $N(z_{\text{spec}})$ for each tomographic bin is normally not so easy, but with this technique, it is also achieved by construction. This is shown in Fig. 6 where filled green histograms correspond to VIPERS data while the black line denotes distribution for DES Y3. As in Fig. 5 red points represent the average of the mocks and error bars correspond to the maximum and minimum. The goal here was to assign z_{photo} in such a way that it gives a redshift



Fig. 5. Photometric redshift distribution of data and mocks. Gray histogram corresponds to DES Y3 sample. Red points represent the average over the 1952 ICE-COLA mocks and error bars denote the maximum and minimum.

distribution $N(z_{\text{spec}})$ on each tomographic bin matching those of the DES Y3 BAO galaxies present in VIPERS.

Some important quantities must be compared for a correct photoz validation, and are evaluated on each tomographic bin. These are the mean redshift (\bar{z}) , the width of the N(z) (W68, opposed to the dispersion), and the dispersion on the photo- $z(\sigma_{68})$. The performance of these quantities in the mocks are analyzed against VIPERS galaxies (for those we have both Z_MEAN and z_{VIPERS}) in Fig. 7. The top panel shows the difference in \bar{z} between the estimated true redshift for DES Y3 sample (Z_MC, dotted black) and ICE-COLA mocks (red) against VIPERS (ZVIPERS). In the medium and bottom panel, we present the evolution of σ_{68} and W68 as a function of \overline{z} for each sample, respectively. In all three quantities studied here, the agreement between VIPERS and the mocks is very satisfactory, showing a difference within 1%. Of course, our method to assign z_{photo} by construction should show a perfect match, with zero uncertainties. However, the small differences we see in Fig. 7 come from the fact that the sample of VIPERS galaxies (8362) is not fully representative of the 2D space z_{photo} - z_{spec} . We want to stress that the difference between z_{VIPERS} and Z_MC represents an estimation of the uncertainty we have on the redshifts, and hence the precision of the mocks is well below those uncertainties. We refer the interested reader to Carnero Rosell et al. (2022) for details on the performance of using VIPERS as a training sample for photo-z validation and true redshift assignment of the sample.

This step is different from what was done for DES Y1 by Avila et al. (2018). Here we use an "exact" method which may propagate noise while Avila et al. (2018) used an "analytical" procedure (fitting a double skewed Gaussian) which may miss some photo-z features.

4.4. Masking

Finally, our last step is to create galaxy mocks with the same footprint as the data. The angular mask of the DES Y3 sample, described in Sect 2, has an area of 4108.47 deg² and final ICE-COLA mocks must have the same characteristics with the same HEALPix resolution NSIDE=4096 as the data mask. To satisfy this and at the same time be efficient in creating as many catalogs as possible, four masks are placed on each full-sky light cone. This allows us to go from having 488 ICE-COLA runs to having 1952 BAO galaxy mocks at the end, quadrupling the number of simulations. This is illustrated in Fig. 8 where four DES Y3 BAO footprints are placed, without overlapping, in a full-sky light cone. Although this configuration of masks max-



Fig. 6. True redshift z_{spec} distribution in each tomographic bin. Green filled histograms correspond to VIPERS, black lines represent the distribution for Y3 data and red point are the average over the 1952 ICE-COLA mocks, while error bars correspond to the maximum and minimum. Histograms are normalized to have an integral of unity.

imizes the number of mocks, on Sect. 7 we return to this and analyze the negative implications it has on the covariance matrix.

5. Calibration

The pipeline used to generate the BAO mocks contains two free parameters per tomographic bin, M_1 and Δ_{LM} , as detailed in the previous section. This amounts to a total of ten free parameters that should be varied altogether to minimize the difference between the measurements and the mocks. In more detail, we want to minimize the discrepancy between the measured $w(\theta)$ and the one obtained from the mocks. Let us emphasize that " $w(\theta)$ " in this section refers to the measurement of the angular correlation function calculated only for the three pre-unblinding angular apertures $\theta = [0.58, 0.75, 0.92]$. The main problem that we have to address is that we cannot obtain $w(\theta)$ from the mocks without first specifying the values of these parameters, generating the mocks, and measuring the observable quantity. If the parameter space were small we could attempt to vary the parameters one at a time, run a few cases, and try to determine approximately the best values for the parameters. However, a 10-dimensional parameter space and ~500 CPU hours to generate the mocks make it effectively impossible to follow this brute force approach.

In this work, we make use of the novel technique presented in Tutusaus (in prep.), where an automatic calibration procedure is implemented into the pipeline to enable us to sample the parameter space and provide the values giving the best agreement with the data in a fully automatized way. We present the basic idea of this method, while we refer the reader to Tutusaus (in prep.) for all the details.



Fig. 7. Basic metrics for the photo-*z* validation on each tomographic bin. *From top to bottom*: difference between \bar{z} and \bar{z}_{VIPERS} , evolution of σ_{68} and W68 as a function \bar{z} . Black and red curves correspond to DES Y3 and ICE-COLA, respectively, and green lines on *bottom panels* represent VIPERS.



Fig. 8. Configuration of four DES Y3 BAO footprints on a fullsky. Mask are constructed with HEALPix assuming a resolution NSIDE=4096.

The first step of the calibration is the determination of the minimal number of mocks that need to be used to get a statistically representative measurement of $w(\theta)$. This will allow us to calibrate with a subset of the mocks and use the best-fit parameters for all of them. In addition to the number of mocks, we also need to find the optimal area for calibration. If the area considered is too small, the determination of $w(\theta)$ will be affected by cosmic variance and it might not be statistically representative. To find the minimum number of mocks with the smallest area that we need to use for the calibration (based on a maximum feasible computation time), we have compared $w(\theta)$ from the selected mocks to the mean $\overline{w}(\theta)$ of the full set of mocks as a function of the number of mocks for different areas. The results are shown in Fig. 9. As can be seen in the figure, we need a large number of mocks of 300 deg^2 (dotted red line) to obtain a $w(\theta)$ representative of the full sample. However, if we consider mocks with the same angular coverage as DES



Fig. 9. Agreement between $w(\theta)$ of a subset of mocks and the mean of $w(\theta)$ for all mocks as a function of the number of mocks used for the selection. The dotted red line stands for mocks of 300 deg², while the dot-dashed blue line represents mocks of 900 deg² and the dashed green one corresponds to mocks of 1500 deg² The solid black line stands for mocks of 4108.47 deg² with the mask of DES Y3 BAO data. The black open circle denotes the selection of the number of mocks and area used for the calibration in this analysis. The number 15 on the *y*-axis corresponds to the number of clustering measurements averaged (three angular apertures times five redshift bins).

Y3 BAO data we can obtain a good representative of the full $w(\theta)$ by just using five mocks. This is the area and number of mocks used for the calibration and it is represented with a black empty circle in Fig. 9. We choose the combination number of mock-area requiring less computational time to get uncertainties within 2%. Note that this optimization of the area and number of mocks have been performed using a fixed value for the calibration parameters, $M_1 = [13.5, 13.9, 14.5, 13.8, 13.2]$ and $\Delta_{LM} = [1.06, 1.23, 1.97, 3.14, 2.21]$, but our goal here was to determine the size of the representative subset of mocks, not the agreement with the data yet. Therefore, these fixed values do not have a significant impact on the subset of mocks that will be used for the calibration. Moreover, we note that the chosen 2% accuracy is somewhat arbitrary. As it can be seen in Fig. 9, considering more mocks provides a better agreement. However, we have verified that 2% is enough for our purposes, guaranteeing uncertainties within 1σ , and it still allows us to use a reduced number of mocks per point in the calibration.

Once we have determined how many mocks and which area we will use for the calibration, we need to start sampling the parameter space to determine the best-fit parameters. The main idea is to sample a given hypercube in the parameter space. In each point we generate five mocks using the DES Y3 BAO mask, measure $w(\theta)$, and compute the value of the χ^2 of the measured $w(\theta)$ in the mocks to the real measurements:

$$\chi^2 = (w(\theta)_{\text{data}} - w(\theta)_{\text{mocks}})^{\mathrm{T}} C^{-1} (w(\theta)_{\text{data}} - w(\theta)_{\text{mocks}}).$$
(5)

We note that *C*, which enters into the χ^2 , is the standard covariance matrix of the 1952 mocks and takes into account the correlations between the different tomographic bins. To obtain *C* we calculated $w(\theta)$ for the 1952 mocks previously created using the fixed value for the calibration parameters mentioned above.

This approach is not different from a standard Monte Carlo Markov chain. However, it is important to note that each evaluation in a point of the parameter space is extremely expensive in computational time since it implies generating five DES Y3-like mocks and measuring $w(\theta)$ on them. Moreover, we are not interested in the posterior of the calibration parameters M_1 and Δ_{LM} , but rather on their best-fit values, since this is the only quantity needed to generate mocks close to the real



Fig. 10. Agreement between the pre-unblinding $w(\theta)$ measurements of the data (open black circle) and the output of the calibrated mocks for the five tomographic redshift bins (red lines). The error bars from the mock measurements have been obtained as the square root of the diagonal of the covariance matrix using the 1952 mocks created with the fixed values.

measurements. A straightforward approach would be to use a simple χ^2 minimization algorithm to go directly to the minimum of the χ^2 function, but the generation of the mocks contains an intrinsic random component when assigning the position and properties of galaxies. This introduces an important stochastic behavior in our problem and makes unusable the standard minimization algorithms.

In this work, following Tutusaus (in prep.), we have decided to use the differential evolution stochastic minimization algorithm first proposed by Storn & Price (1997). The essential idea of the algorithm is to use a population of candidate solutions. We first initialize the population using a Latin Hypercube sampling; then, iteratively, these candidate solutions are combined to generate a new population and the χ^2 is evaluated at each position. In more detail, the distance between two random candidate solutions is used to displace the best candidate solution so far (minimum χ^2). If the new candidates are better than the previous ones they are accepted and belong to the new population; otherwise they are discarded and the new population is completed with candidates from the old population. Note that, because of this, the size of the population remains constant. The process ends when the standard deviation of the χ^2 values of the population is smaller than a given tolerance times the mean of the χ^2 values. The best candidate of the population at the end of the process is the best-fit used to generate the final mocks.

Once the calibration parameters have been determined and the five mocks generated, we can check the agreement between the $w(\theta)$ from the mocks and the real measurements to verify how accurate the calibration is. The results are shown in Fig. 10. For each one of the tomographic bins, we represent the data with open black circles and the measurements from the five mocks with red lines. The errors have been obtained as the square root of the diagonal of the covariance matrix. The agreement is well within 1 σ for all the bins, giving a goodness of the model $\chi^2_{mod}/d.o.f. = 7.58/5$. Degrees of freedom (d.o.f.) equal five corresponds to the number of θ -bins= 15, three apertures times five tomographic bins, minus the ten free parameters, two (M_1 and Δ_{LM}) per bin.

6. Analysis

In previous sections, we explained the methodology and the procedure used to create 1952 mocks reproducing all the relevant properties of the DES Y3 sample. In this section, we analyze the clustering of these mock catalogs in different spaces. We also include in the analysis a theoretical model prediction for those statistics.

6.1. Theoretical model

The theoretical template is computed using the redshift-space power spectrum

$$P(k,\mu) = (1 + \beta\mu^2)^2 b^2 \{ [P_{\rm lin}(k) - P_{\rm sm}(k)] D_{\rm BAO} + P_{\rm sm}(k) \}, \qquad (6)$$

where μ is the dot product between \hat{k} and the line-of-sight direction, *b* is the linear bias, and $\beta = f/b$ with *f* being the linear growth rate. The power spectrum is built using the linear power spectrum $P_{\rm lin}(k)$ and the linear no-wiggle power spectrum $P_{\rm sm}^{12}$. The nonlinear damping of the BAO feature is modeled by

$$D_{\text{BAO}}(k,\mu) = \exp\{-k^2 [\mu^2 \Sigma_{\parallel}^2 + (1-\mu^2) \Sigma_{\perp}^2 + f\mu^2 (\mu^2 - 1)\delta \Sigma_{\parallel}^2]\},$$
(7)

with $\Sigma_{\parallel} = (1 + f)\Sigma_{\perp}$. The damping scales Σ_{\perp} and $\delta\Sigma$ are computed following Baldauf et al. (2015). In MICE cosmology, $\Sigma_{\perp} = 5.80 \,\mathrm{Mpc}\,\mathrm{h}^{-1}$ and $\delta\Sigma = 3.18 \,\mathrm{Mpc}\,\mathrm{h}^{-1}$ at redshift 0 and they are scaled to higher redshift by the growth factor. See DES Collaboration (2021) for more details about the procedure to obtain these quantities. Once provided with $P(k,\mu)$, we computed the anisotropic redshift-space correlation function $\xi(s,\mu)$ through a Fourier transform (see Chan, in prep.). The angular correlation function is obtained after projecting ξ weighted by the redshift distribution n(z) (normalized to 1),

$$w(\theta) = \int dz_1 \int dz_2 n(z_1) n(z_2) \xi(s(z_1, z_2, \theta), \mu(z_1, z_2, \theta)).$$
(8)

The harmonic power spectrum template C_{ℓ} is derived from *w* by a Legendre transform

$$C_{\ell} = 2\pi \int_{-1}^{1} d\mu \ w(\arccos(\mu)) \mathcal{L}_{\ell}(\mu), \tag{9}$$

where \mathcal{L}_{ℓ} is the Legendre polynomial. For more details on the modeling, see the main DES Y3 BAO paper (DES Collaboration 2021).

6.2. Angular correlation function: $w(\theta)$

Around ten thousand ACF must be calculated for the mocks (1952 mocks times five bins). For this reason, we resort to a code that allows the calculation using pixels, reducing the computational time. It is important to point out that since we only use angular apertures greater than one degree for fitting, any effect from the pixelization should be negligible. We use the public code CUTE (Alonso 2012). CUTE supports the Landy & Szalay estimator (Landy & Szalay 1993):

$$w(\theta) = \frac{\mathrm{DD}(\theta) - 2\mathrm{DR}(\theta) + \mathrm{RR}(\theta)}{\mathrm{RR}(\theta)},\tag{10}$$

where DD, DR and RR represent the total number of Data-Data, Data-Random and Random-Random pairs separated by an angular θ projected distance, respectively. In this case, Data correspond to galaxies in the mocks while Randoms are created by sampling the same volume with random points. The total number of randoms is 20 times the average number of galaxies in the mocks. The same random catalog is used to calculate the clustering for all 1952 mocks. The chosen pixel resolution is $n_{\text{pix-shp}} = 4096$ which yields pixels with an angular θ resolution of 2.1 arcmin.

¹² Defined by following the 1D Gaussian smoothing in log-space described in Appendix A of Vlah et al. (2016).



Fig. 11. Angular two-point correlation function for each tomographic bin. Red lines correspond to the average over all the mocks and shaded light-red bands correspond to the standard deviation. Dashed blue lines indicate the theoretical prediction described on Sect. 6.1 and black points correspond to the pre-unblinding data values used in the calibration showed on Sect. 5. Finally, solid black lines with error bars represent the final post-unblinding measurement of the data.

Figure 11 shows the result of the angular two-point correlation function of 1952 mocks for the 5 tomographic bins. It can be noticed that the difference between the pre-unblinding values used for calibration (open black circles) and the averaged ACF for the mocks (solid red lines) differs from what was obtained on Sect. 5 during the calibration procedure (see Fig. 10). However, these small differences are expected given the degree of representativeness when using only five mocks for the calibration (0.25%) of the total number of mocks). It is important to keep in mind that the number of five mocks has been determined with a fixed calibration. Therefore, combining the error introduced by these fixed parameters and the allowed 2% accuracy, we can expect the final accuracy with the bestfit calibration and all mocks to be slightly above 2%. To be more precise, this number has increased from 2 to 3.1%. However, within the corresponding uncertainties, the agreement is still within 1σ for the five tomographic bins, giving a global $\chi^2_{\rm mod}$ /d.o.f. = 16.72/5. The increase in the $\chi^2_{\rm mod}$ may be due either to the low representativeness of the five mocks used for the calibration or the presence of a strong anti-correlation among the data points. Nevertheless, when we quantify the goodness of the fit between the mocks to the data, using all θ bins in Fig. 11, we find remarkable good values $\chi^2_{m-d}/d.o.f. =$ bin. In conclusion, such accuracy is enough for our purposes and we do not consider rerunning the pipeline with more mocks in each point of the calibration. In more detail, only the first bin shows one pre-unblinding data point out of 1σ from the mocks. In this case, the mean $w(\theta)$ of the mock has changed ~4% from the value found during the calibration. Although this value is double of 2% what was foreseen in the calibration (see Fig. 9), the global change is within the expectation. Blue curves in Fig. 11 denotes the theoretical prediction described in Sect. 6.1, and it is clear from the figure that modeled $w(\theta)$ agree almost perfectly with the measurements on the mocks. Finally, solid black lines with error bars correspond to the final post-unblinding measurements of the data (using a brute force configuration of CUTE).

6.3. Projected Clustering: $\xi_w(s_{\perp})$

In photometric surveys, most of the radial BAO information is lost due to redshift uncertainties. Additionally, the photo-*z* uncertainty causes the BAO scale in the 3D correlation function, $\xi(s)$, to deviate from its true position. However, Ross et al. (2017) demonstrated that when the correlation function is plotted against the transverse scale $s_{\perp} = s\sqrt{1-\mu^2}$, the BAO peak appears where s_{\perp} equals to the true sound horizon scale. Thus, (angular) BAO information can still be retrieved via the 3D correlation analysis.

Following the methodology from Ross et al. (2017) and also described in DES Collaboration (2021), we show in Fig. 12 the 3D wedge correlation function $\xi_w(s_\perp)$ measured from the mocks (due to computational expenses, only 120 mocks are used) and data. The results are also obtained from CUTE using Eq. (10), by replacing $w(\theta)$ with $\xi(s_\perp, s_\parallel)$, and then integrating over s_\parallel for the scales with $\mu < 0.8$. As a comparison, we have also plotted the corresponding theory prediction. While Ross et al. (2017) assumes Gaussian photo-*z* distribution, the prediction makes use of the photo-*z* distribution from the mocks, resulting in better agreement with the numerical measurements. Further details of the comparison between the mock results and the theory will be presented in Chan (in prep.). We find a good agreement for the 3D clustering of data, mocks and theory.



Fig. 12. 3D wedge correlation function for μ range [0,0.8]. With the same color code as previous figures, solid red line corresponds to ICE-COLA mocks and shaded light-red band to its standard deviation. Post-unblinding measurement of data is shown with black points and the theory with a dashed blue line.

6.4. Angular power spectrum: C_{ℓ}

We also measured the clustering signal in the Fourier conjugate space of angular distances on the sphere, the so-called harmonic space, by estimating the angular power spectra of galaxy number counts, C_{ℓ} . Although constructed from the same underlying field, the angular power spectrum and the correlation function present different advantages and downsides. Most notably, the correlation function is relatively straightforward to estimate in the presence of an angular survey mask. Still, its estimates are largely correlated. On the other hand, the power spectrum requires a deconvolution of the angular mask, but the correlation between scales reduces. Taking these pros and cons into account, it is clearly desirable to have complementary information from both statistics. See Giannantonio et al. (2016) for the first implementation of these complementary estimators in the context of DES Y1 data analyses.

We begin by estimating pixelized galaxy overdensity maps, $\delta_g(\hat{n}) = N_g(\hat{n})/\bar{N}_g - 1$, for the ICE-COLA mocks, where \hat{n} is the pixel position in the sphere, $N_g(\hat{n})$ the pixelized galaxy number counts and \bar{N}_g the mean number of galaxies per pixel. We then used the "Pseudo- C_ℓ " method (Hivon et al. 2002) to estimate the angular power spectra. The Pseudo- C_ℓ method deconvolves the incomplete sky coverage mode mixing effect on a set of band power bins using analytical methods and has the advantage of being less computationally expensive, reaching equivalent error estimates than optimal quadratic estimators. Also, Pseudo- C_ℓ estimators are effectively "unbiased" concerning maximum likelihood estimators. In particular, we use the implementation of the NaMASTER code¹³ (Alonso et al. 2019).

The discrete nature of galaxy number counts introduces a shot-noise contribution to the estimated galaxy overdensity maps and, consequently, a bias to the estimated C_{ℓ} . We account for this "noise bias" analytically, following Alonso et al. (2019) and Nicola et al. (2020) by subtracting this Poissonian noise from our power spectrum. For each ICE-COLA mock, we consider the partial sky coverage introduced by its associated mask, as shown in Fig. 8. To optimize the BAO feature detection, we bin the power spectra in bands of $\Delta \ell = 20$ from a minimum multipole $\ell_{\min} = 10$ up to a maximum multipole of 1 000.

Figure 13 shows the mean and standard deviation of the estimated angular power spectra of the 1952 ICE-COLA mocks for the five tomographic bins (solid red lines) together with the theory prediction (dashed blue lines). Black points correspond to the post-unblinding measurements of the data.

7. Effect of replications

Without taking into account the tails on redshift distribution for the tomographic bins on the edges ([0.6-0.7] and [1.0-1.1]), the maximum comoving line-of-sight distance that can be found among galaxies on the mocks is ~1000 Mpc h⁻¹. Even if we consider the tails, the distance is lower than the box of the simulation, 1536 Mpc h⁻¹. In other words, no halo of ICE-COLA simulations is used more than one time along any give line-of-sight. But this problem does occur for different lines-of-sight. The area of the survey is very wide implying that for higher redshift several numbers of simulated boxes are used to equal the volume of the BAO Y3 sample. Inevitably this implies the use of the same halo structures on each mock. These repeated halos leave at different times but originate from the same initial structure they will obviously correlate, to leading order overdensities just grow linearly.

Table 2 shows the upper limit for the percentage of replicated "halos" that can be found among bins. To obtain these values, we randomly sampled a box with particles and then created the light cone by replicating this box. These numbers are representative of all the repeated "halos" in a light cone and not only those used to make up the DES Y3 sample, which are a few percent of the total. This fact, together with the selection efficiency, makes that the values are shown on Table 2 stand for a very conservative bound. However, these numbers refer to the replications present in a single mock but each light cone is used to create four mocks (see Fig. 8). This inevitably will introduce replications among different mocks created with the same light cone in addition to those among the bins of a single mock. Numbers in parentheses in Table 2 correspond to the percentage of repeated random particles among the four mocks made from one light cone. For example, the pair of bins (1,3) has on average (mean among the four mocks) 11.9% repeated random particles. This number increase to 30.2% when considering the four mocks. The difference between the 11.9 and the 30.2% come from the particles which are not repeated in one mock but do in others.

The main impact of this replication problem becomes noticeable when the cross-covariance matrix is analyzed. The repeated structures in different bins introduce a spurious correlation among measured $w(\theta)$ of tomographic bins. This effect is shown in Fig. 14 wherein the top panel we compare the covariance matrix of the ICE-COLA mocks (lower diagonal) with the covariance matrix computed using COSMOLIKE halo model (Krause & Eifler 2017; Fang et al. 2020, upper diagonal). From the former, the high degree of correlation between bins that are not adjacent is visible. This strong correlation can be seen clearly in the bottom panel of Fig. 14 where we show one column of the covariance matrix corresponding to an aperture of $\theta = 2.7$ deg for COSMOLIKE (solid black line) and ICE-COLA (dashed blue line). For simplicity, we are using in this plot a $\Delta \theta = 0.2$.

8. Conclusions

The performance of well-validated mocks for DES Y3 BAO analysis is crucial for obtaining robust scientific results. The analysis of the data collected by the Dark Energy Survey during the first 3 years of the project poses a great scientific challenge in the development of the required mocks. We have created a significant number of mocks, 1952, adequate for statistical analysis

¹³ https://github.com/LSSTDESC/NaMaster



Fig. 13. Measured angular power spectra for the tomographic redshift bins considered. Red lines correspond to the average over all the mocks and black points represent the post-unblinding measurements of the data. The theoretical prediction is shown with dashed blue lines.

Table 2. Percentage of replications among tomographic bins.

BIN	1	2	3	4	5
1	3.9 (22.6)	9.7 (30.6)	11.9 (30.2)	15.9 (30.3)	17.8 (30.8)
2	_	5.9 (27.9)	13.9 (28.7)	15.7 (29.1)	17.7 (28.9)
3	_	-	8.1 (28.7)	16.5 (29.2)	18.5 (28.1)
4	_	-	_	8.9 (30.7)	18.2 (27.4)
5	-	_	-	—	10.1 (29.9)

Notes. Each value must be interpreted as the percentage of random points found duplicated in the corresponding pair of bins only in one mock. Numbers in parentheses represent the same but considering the four mocks done with one light cone.

of galaxy clustering. The mocks have been built by populating the halo catalogs of 488 ICE-COLA fast simulations. This task is made up of many steps that make up our pipeline, those are:

- 1. HALO CATALOGS Run 488 ICE-COLA simulations with different initial conditions. Create light cones halo catalogs on-the-fly by replicating simulated boxes.
- 2. GALAXY CATALOGS
 - (a) HOD: Populate halos with one central galaxy and N_{sat} satellite galaxies following a Poisson distribution, setting the first free parameter per tomographic bin, M_1 .
 - (b) HAM: Assign a pseudo-luminosity l_p to galaxies, setting the second free parameter per tomographic bin, Δ_{LM} .
 - (c) Model photometric redshifts using a highly accurate method which follows a 2D probability distribution relating photometric and spectroscopic redshifts of VIPERS data.
 - (d) Apply four DES Y3 BAO mask over each full-sky light cone and quadrupling the total number of galaxy mocks compared to halo catalogs.
- 3. CALIBRATION Set an automatic calibration procedure using a differential evolution stochastic minimization algorithm to find the best values for the ten free parameters. This step is a loop of step (ii).

FINAL GALAXY CATALOGS – Repeat step (ii) with the final set of parameters.

The automatic calibration and the photometric redshift assignment procedure are key to obtaining final mocks that reproduce with high accuracy the principal properties of the data: (1) observed volume, (2) abundance of galaxies, redshift distribution, redshift uncertainty, and (3) clustering as a function of redshift. Firstly (1), the replications of the simulated box allowed us to achieve the observational volume of the DES Y3 sample. However, the use of even bigger simulations would be beneficial, to avoid the repetition of structures and therefore the introduction of spurious cross-correlations and cross-covariance among not adjacent bins. This should be considered when designing mocks for the calculation of the covariance matrix in future surveys. Secondly (2), the use of overlapping galaxies between DES Y3 and VIPERS gave rise to a very realistic photo-z assignment. In these mocks, the abundance of galaxies, $n(z_{photo})$, the redshift distribution $N(z_{spec})$ and their uncertainties \bar{z} , W68 and σ_{68} are in excellent agreement with VIPERS data, showing discrepancies lower than one per cent. Finally (3), the clustering measured in this set of mocks shows an excellent agreement with the pre-unblinding data used for calibration. The uncertainty on the calibration procedure resulted in a goodness of the



Fig. 14. Covariance matrix of the angular correlation function. *Top*: comparison of the correlation matrix obtained from COSMOLIKE halo model (upper diagonal) and from ICE-COLA mocks (lower diagonal). *Bottom*: columm of the covariance matrix $Cov[\theta_i^N, \theta_j^M]$ with fixed $\theta_j^1 = 2.7$ deg. The *i* and *j* subscripts correspond to the index of tha data vector (from 0 to 124, with θ ranging in [0, 5]deg in $\Delta \theta = 0.2$ deg steps and four redshift bins), and *N*, *M* superscripts represent the index of the tomographic redshift bins.

model χ^2_{mod} /d.o.f. = 7.58/5 while for the final mocks this value increased a bit up to 16.72/5.

DES Collaboration (2021) presents the best fit and uncertainty for the angular scale of the BAO angular distance measurement D_A using the Y3 Dark Energy Survey data release. The mocks presented in this work have been a crucial tool in the procedure of obtaining these results. ICE-COLA mocks have been used to run robustness tests and optimize the methodology. Finally, it is important to remark that the pipeline that created this set of mocks should work perfectly for any future surveys which need realistic mocks for galaxy clustering analysis.

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References

- Alam, S., Aubert, M., Avila, S., et al. 2021, Phys. Rev. D, 103, 083533
- Alonso, D. 2012, ArXiv e-prints [arXiv:1210.1833]
- Alonso, D., Sanchez, J., Slosar, A., & LSST Dark Energy Science Collaboration 2019, MNRAS, 484, 4127
- Avila, S., Murray, S. G., Knebe, A., et al. 2015, MNRAS, 450, 1856
- Avila, S., Crocce, M., Ross, A. J., et al. 2018, MNRAS, 479, 94
- Baldauf, T., Mirbabayi, M., Simonović, M., & Zaldarriaga, M. 2015, Phys. Rev. D, 92, 043514
- Benson, A. J., Cole, S., Frenk, C. S., Baugh, C. M., & Lacey, C. G. 2000, MNRAS, 311, 793
- Beutler, F., Blake, C., Colless, M., et al. 2011, MNRAS, 416, 3017
- Blot, L., Corasaniti, P.-S., Rasera, Y., & Agarwal, S. 2021, MNRAS, 500, 2532
- Camacho, H., Kokron, N., Andrade-Oliveira, F., et al. 2019, MNRAS, 487, 3870
- Carnero Rosell, A. C., Rodriguez-Monroy, M., Crocce, M., et al. 2022, MNRAS, 509, 778
- Carretero, J., Castander, F. J., Gaztañaga, E., Crocce, M., & Fosalba, P. 2015, MNRAS, 447, 646
- Chan, K. C., Crocce, M., Ross, A. J., et al. 2018, MNRAS, 480, 3031
- Chuang, C.-H., Kitaura, F.-S., Prada, F., Zhao, C., & Yepes, G. 2015a, MNRAS, 446, 2621
- Chuang, C.-H., Zhao, C., Prada, F., et al. 2015b, MNRAS, 452, 686
- Coles, P., & Jones, B. 1991, MNRAS, 248, 1
- Comparat, J., Jullo, E., Kneib, J.-P., et al. 2013, MNRAS, 433, 1146
- Conroy, C., Wechsler, R. H., & Kravtsov, A. V. 2006, ApJ, 647, 201
- Cooray, A., & Sheth, R. 2002, Phys. Rep., 372, 1
- Crocce, M., Castander, F. J., Gaztañaga, E., Fosalba, P., & Carretero, J. 2015, MNRAS, 453, 1513

- Crocce, M., Ross, A. J., Sevilla-Noarbe, I., et al. 2019, MNRAS, 482, 2807 DES Collaboration 2019, MNRAS, 483, 4866
- DES Collaboration (Abbott, T. M. C., et al.) 2021, ArXiv e-prints [arXiv:2107.04646]
- De Vicente, J., Sánchez, E., & Sevilla-Noarbe, I. 2016, MNRAS, 459, 3078
- Drlica-Wagner, A., Sevilla-Noarbe, I., Rykoff, E. S., et al. 2018, ApJS, 235, 33 Euclid Collaboration (Desprez, G., et al.) 2020, A&A, 644, A31
- Fang, X., Eifler, T., & Krause, E. 2020, MNRAS, 497, 2699
- Flaugher, B., Diehl, H. T., Honscheid, K., et al. 2015, AJ, 150, 150
- Fosalba, P., Crocce, M., Gaztañaga, E., & Castander, F. J. 2015a, MNRAS, 448, 2987
- Fosalba, P., Gaztañaga, E., Castander, F. J., & Crocce, M. 2015b, MNRAS, 447, 1319
- Giannantonio, T., Fosalba, P., Cawthon, R., et al. 2016, MNRAS, 456, 3213
- Guo, Q., White, S., Li, C., & Boylan-Kolchin, M. 2010, MNRAS, 404, 1111
- Guzzo, L., Scodeggio, M., Garilli, B., et al. 2014, A&A, 566, A108
- Hivon, E., Górski, K. M., Netterfield, C. B., et al. 2002, ApJ, 567, 2
- Izard, A., Crocce, M., & Fosalba, P. 2016, MNRAS, 459, 2327
- Izard, A., Fosalba, P., & Crocce, M. 2018, MNRAS, 473, 3051
- Jing, Y. P., Mo, H. J., & Börner, G. 1998, ApJ, 494, 1
- Kitaura, F.-S., Rodríguez-Torres, S., Chuang, C.-H., et al. 2016, MNRAS, 456, 4156
- Koda, J., Blake, C., Beutler, F., Kazin, E., & Marin, F. 2016, MNRAS, 459, 2118 Komatsu, E., Dunkley, J., Nolta, M. R., et al. 2009, ApJS, 180, 330
- Krause, E., & Eifler, T. 2017, MNRAS, 470, 2100
- Kravtsov, A. V., Berlind, A. A., Wechsler, R. H., et al. 2004, ApJ, 609, 35
- Landy, S. D., & Szalay, A. S. 1993, ApJ, 412, 64
- Lippich, M., Sánchez, A. G., Colavincenzo, M., et al. 2019, MNRAS, 482, 1786
- Manera, M., Scoccimarro, R., Percival, W. J., et al. 2013, MNRAS, 428, 1036
- Morganson, E., Gruendl, R. A., Menanteau, F., et al. 2018, PASP, 130, 074501
- Navarro, J. F., Frenk, C. S., & White, S. D. M. 1996, ApJ, 462, 563
- Nicola, A., Alonso, D., Sánchez, J., et al. 2020, JCAP, 2020, 044
- Peebles, P. J. E., & Yu, J. T. 1970, ApJ, 162, 815
- Percival, W. J., Cole, S., Eisenstein, D. J., et al. 2007, MNRAS, 381, 1053
- Ross, A. J., Banik, N., Avila, S., et al. 2017, MNRAS, 472, 4456
- Scoccimarro, R., & Sheth, R. K. 2002, MNRAS, 329, 629
- Seljak, U. 2000, MNRAS, 318, 203
- Sevilla-Noarbe, I., Bechtol, K., Carrasco Kind, M., et al. 2021, ApJS, 254, 24
- Sheth, R. K., & Diaferio, A. 2001, MNRAS, 322, 901 Storn, R., & Price, K. V. 1997, J. Glob. Optim., 11, 341
- Sunyaev, R. A., & Zeldovich, Y. B. 1970, Ap&SS, 7, 3
- Tassev, S., Zaldarriaga, M., & Eisenstein, D. J. 2013, JCAP, 2013, 036
- Vlah, Z., Seljak, U., Yat Chu, M., & Feng, Y. 2016, JCAP, 2016, 057 Zhao, C., Chuang, C.-H., Bautista, J., et al. 2021, MNRAS, 503, 1149
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