- 1 **A Dynamic Framework for the Study of Optimal Birth Intervals Reveals the Importance of**
- 2 **Sibling Competition and Mortality Risks**
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- 4 **Short title:** A Dynamic Framework for Optimal Birth Spacing
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22 **Abstract**

23 Human reproductive patterns have been well studied but the mechanisms by which physiology, 24 ecology and existing kin interact to affect the life history need quantification. Here, we create a 25 model to investigate how age-specific interbirth intervals adapt to environmental and intrinsic 26 mortality, and how birth patterns can be shaped by competition and help between siblings. The 27 model provides a flexible framework for studying the processes underlying human reproductive scheduling. We developed a state-based optimality model to determine age-dependent and family-28 29 dependent sets of reproductive strategies, including the state of the mother and her offspring. We 30 parameterised the model with realistic mortality curves derived from five human populations. 31 Overall, optimal birth intervals increase until the age of 30 after which they remain relatively 32 constant until the end of the reproductive lifespan. Offspring helping each other does not have much 33 effect on birth intervals. Increasing infant and senescent mortality in different populations decreases 34 interbirth intervals. We show that sibling competition and infant mortality interact to lengthen 35 interbirth intervals. In lower-mortality populations, intense sibling competition pushes births further 36 apart. Varying the adult risk of mortality alone has no effect on birth intervals between populations; 37 competition between offspring drives the differences in birth intervals only when infant mortality is 38 Iow. These results are relevant to understanding the demographic transition, because our model 39 predicts that sibling competition becomes an important determinant of optimal interbirth intervals 40 only when mortality is low, as in post-transition societies. We do not predict that these effects alone 41 can select for menopause.

Keywords: interbirth intervals; humans; state-dependent optimality modelling; life history evolution;42 43 sibling competition.

44

45 **Introduction**

On attaining sexual maturity, humans have substantial reproductive potential and populations are46 47 capable of rapid expansion. This feature of the human life history may have contributed to the successful migration and colonisation that has been a characteristic of our species. Present day48 49 populations exhibiting natural fertility have a typical interbirth interval (IBI) in the range of 3-5 years 50 (Sear & Mace, 2008). Shorter first birth intervals are associated with increased lifetime reproductive 51 success (Nenko et al., 2013). Moreover, IBIs increase with age until reproduction is physiologically no 52 longer possible after the age of menopause, although the age at last reproduction tends to occur 53 well before this (Sievert, 2006). Explanations of the (proximate) mechanisms underlying these 54 patterns have so far met with mixed success. 55 Here, we construct a flexible framework in which factors relating to individual human reproductive 56 success are analysed from an evolutionary perspective. Our model explores how reproductive 57 schedules adapt to mortality risks (both intrinsic and environmental) and kin effects, potentially 58 explaining the variation in human life history across the world. 59 At birth, human children are particularly altricial compared to other great apes and require intensive 60 and protracted maternal investment. While mothers are breastfeeding, fertility is usually suppressed (but see Short et al. 1991). This can act as a natural contraceptive, protecting both the mother and61 62 existing children from too close birth spacing (Ellison et al., 1993). Nevertheless, human infants are 63 weaned early compared to other great apes. This increases the fertility of the mother and may 64 require alloparents (usually kin) to help in providing for the child (Bogin, 1997; Hawkes et al., 1997). 65 Although young children are capable of foraging to some degree, they remain nutritionally 66 dependent on others for many years (Kaplan, 1996). The age of puberty depends on rates of growth 67 and development, which in turn depend on the levels of nutrition received during infancy and 68 childhood.

Human females suffer an unusually high hazard during childbirth, which increases with age (Grimes,69 1994; Abitbol, 1996). At older maternal ages, there is a general age-related increase in IBIs and70 71 offspring are weaned later, as in many other primates (Caro et al., 1995). Younger offspring are particularly vulnerable if their mother dies (Willführ & Gagnon, 2013).72

There is inevitable competition between siblings for maternal attention and resources. Newborns 74 are likely to divert maternal attention from existing children and the youngest child must usually be 75 weaned before the mother is again fertile. Young children with many young siblings may therefore 76 be exposed to higher mortality risks than if they are the sole recipient of the mother's provisioning. In order to combat these risks, the World Health Organization recommends a minimum birth interval 78 of two years (WHO, 2006).

79 Older children can share some of the burden of care for the young with their mother, helping to

80 reduce the mortality risk of their younger siblings. This has been observed in some farming

81 populations such as the Mandinka in Gambia, the Maya in Mexico, and the Chewa in Malawi, as well

82 as in a 17th Century Québécoise population (Sear et al., 2002; Beise, 2005; Sear, 2008; Kramer, 2010).

83 The timing of births can have important consequences for reproductive success. The risks of adverse

84 outcomes due to short IBIs are well documented. However, there is evidence to suggest that

85 extended spacing between births (longer than 50 months) is also linked to events such as preterm

86 birth and low birth weight (Conde-Agudelo et al., 2006). This may be due to phenotypic correlations

whereby a female may already be experiencing low fertility or poor nutritional status. Thus,87

88 understanding how mechanisms such as sibling competition can affect birth spacing might be

89 important for understanding patterns of infant mortality.

90 **Models of optimal reproductive scheduling**

91 In foraging populations, women have no option but to carry infants, which poses a considerable

92 energetic burden. The !Kung San -- a foraging people of the Kalahari desert -- are largely dependent

93 on mongongo nuts as a food source (Howell, 1979). They frequently move foraging site and must 94 carry both food and young offspring. Mathematical models quantifying the load of food that can be efficiently carried, along with the demands of young offspring who must be provisioned and also95 carried, predicted an optimal IBI of approximately 4 years, which is typical of this population (Blurton96 97 Jones & Sibly, 1978; Blurton Jones, 1986; Anderies, 1996). However, other related groups of San 98 people who have a different local ecology, which does not expose them to the same reproductive 99 constraint imposed by the need to carry food, have similar IBIs of 4 years (Hill & Hurtado, 1996).

100 A more complex model considered the influence of a female's age and stochasticity in her foraging success on the survival of her children and her optimal reproductive strategy (Anderies, 1996), 102 where older females were assumed to forage less efficiently. A female's probability of survival 103 depended on her age and, if she gave birth, on the risk of mortality in childbirth. Through maximising lifetime reproductive success, an IBI of 4 years was a robust response to all realistic conditions, and104 105 showed only a small increase in the optimal interval with age of the mother. The predicted optimal 106 IBIs matched observations of !Kung reproductive decisions (Anderies, 1996).

107 Mace (1998) used the same framework to show how reproductive decisions are sensitive to inherited wealth, when parental resources are required for the next generation to marry and 109 reproduce; the more parental resources are needed, the smaller the optimal family size. When this 110 was the case, higher mortality risk in the environment caused increased fertility through 111 'replacement' births even though the overall family size of surviving offspring was not much altered.

Here, we apply the well-developed technique of state-dependent optimality modelling (Houston & 113 McNamara, 1999) to investigate reproductive decisions in human life history. We develop a general 114 but comprehensive dynamic model that offers the flexibility to examine optimal age-related 115 reproductive strategies across a variety of contexts relevant to human physiology, ecology and social 116 organisation. A dynamic modelling framework can add greater realism to models of reproductive

117 behaviour. This allows maternal decisions to be evaluated in terms of their long-term fitness 118 consequences; crucially, decisions depend on the mother's state.

119 The aim is to identify the key determinants of the age-related increase in IBI given exposure to 120 mortality hazards from the mother's socioecological environment. Rather than generating quantitative predictions for observed birth intervals, this model is intended to understand the 122 factors driving human life history variation. This is not explicitly a model of menopause, as it does not include a third generation with which to explore grandmother effects. However in one set of 124 experiments we extend the possible reproductive span to the end of life in order to investigate 125 whether maternal mortality hazards and offspring effects can select for reproductive cessation.

126 **Materials and Methods**

127 **The Model**

128 The purpose of the model is to determine the optimal IBIs over the course of an individual female

129 life cycle. A woman can produce a child once every two or more years. However, there are

130 considerable risks associated with reproduction both for the mother and her existing family. First,

the mother is exposed to the risks associated with childbirth, which increase with age (Grimes, 1994;

132 Blanc et al., 2013). Second, offspring spaced too closely encounter competition for maternal

133 provisions; for example, the youngest child must be weaned before the next is born.

134 For each existing child in the model, a newborn sibling diverts attention from the mother that would

135 otherwise be directed towards them. A newborn child can therefore bring an associated reduction in

136 survival for all siblings. Finally, even in the absence of a newborn child, existing siblings have a

137 detrimental effect on one another. The model examines the interaction of these parameters in

138 determining an optimal schedule of births for a female.

- 139 In order to determine the optimal birth decisions, the model can be characterised as a discrete-time 140 Markov Decision Process (MDP) and solved by stochastic dynamic programming. The MDP contains 141 the following elements:
- The finite set of states is described by mother's age *x* and family structure *C*, discussed below.142
- 143 $\mathcal U$ is the set of actions {reproduce, do not reproduce}.
- 144 $P_u(x)$ is the mother's probability of surviving, given her age, *x*, and her birth decision, *u*.
- $Q_{\mu}(x, C, C')$ calculates the survival probabilities of each of the children in family structure *C*, which
- becomes family structure *C'* the following year, given their mother's age, *x*, and her birth decision, *u*.146
- 147 This accounts for all combinations of child survival, including where all the children die, as well as the
- 148 effects of sibling competition and juvenile help.
- 149 $R_{\nu}(x, C, C')$ is one half of the expected number of offspring that mature next year, given the
- 150 mother's age, x, her birth decision *u* and the effects of sibling competition or help on the maturing
- child's survival as family structure *C* transitions to *C'* (Houston & McNamara, 1999). *C'* is the family151
- 152 structure corresponding to **C** with children ageing one year and newborns being present (or not)
- 153 according to birth decision *u* and the mortality risks for the mother and her children. This element is
- 154 half the total expected offspring since the model tracks only females.

155 **State Variables**

Females in the model make an annual decision (*u*) whether to give birth or not, depending on their156 157 age and the structure of their existing family. The state variables are the mother's age and the age 158 and number of children in her existing family. A female is assumed to mature at 15 years. The model 159 tracks her birth decisions from sexual maturity until the age of 50. Twinning is excluded from the 160 model so she can only give birth to a single child and the minimum birth spacing is set at two years, 161 to allow a reasonable period of lactational amenorrhea while remaining computationally tractable. 162 Given these constraints, a mother can have 987 possible family compositions. (A family of children 163 aged between 1 and 14 can be represented as a 14-bit binary string where the presence of a child is

164 marked with a 1. For example, a family with a two-year-old and a 14-year old would be 165 01000000000001. Neighbouring binary digits cannot both be 1; hence, there are 987 possible 166 combinations.) Since the sexually mature lifespan is 35 years, the model will optimise birth decisions 167 over $35 \times 987 = 34,545$ states. The state space is (x, C) where:

- 168 1. *x* is the set of maternal ages between 15 years and 50 years.
- 2. *C* is the family structure (i.e. mother's offspring): a set of child ages between 1 year and 14 170 years for up to 7 children, including no offspring. There will always be a minimum spacing of
- 171 2 years between children.

172 **State Transitions**

173 The model considers all possible combinations of family in each year that can result in the case of

174 none, any or all children surviving. One of the strengths of state-dependent optimality modelling as a

175 methodology is its ability to account for a range of future states. The probability of each permutation

176 is calculated from mortality data that, in turn, depend on the structure of the family, the mother's

177 age and whether or not she gives birth.

178 **Mortality**

179 The mother's mortality rate is comprised of age-specific senescent and maternal components, and

180 an age-independent extrinsic term (equation (1)); child mortality is a decreasing exponential

181 function of age (equation (2)) (Siler, 1979). In order to situate the model in a real-world context, we

182 parameterise the mortality model using cross-cultural data (see Supplementary Table S1 and Fig.

183 1a).

$$
\mu_{\text{adult}}(x) = \mu_{\text{extrinsic}} + \mu_{\text{senescent}}(x) + \mu_{\text{material}}(x) \tag{1}
$$

where:184

185 $\mu_{\text{extrinsic}} = a_2$

186 $\mu_{\text{sensor}}(x) = a_3 e^{x b_3}$

187
$$
\mu_{\text{material}}(x) = \begin{cases} \alpha_{\text{birth}} x^2 - x \beta_{\text{birth}} + \gamma_{\text{birth}} \\ \alpha_{\text{birth}} e^{(x - x_{\text{maturity}}) \beta_{\text{birth}}} + (\mu_{\text{extrinsic}} - \alpha_{\text{birth}}) \end{cases}
$$

188 Here, *x* is the mother's age; a_2 , a_3 and b_3 are population-specific mortality parameters; $α_{birth} β_{birth}$ 189 and γ_{birth} are maternal mortality parameters. The two maternal mortality functions are discussed 190 below.

- The sources of mortality are considered to be independent and can therefore simply be added together to obtain total mortality. The annual probability of survival is $\exp{-\mu_{\text{adult}}(x)}$.
- 193 According to how these parameters have been estimated in the published literature, the hunter-
- 194 gatherer populations (Ache and Tsimane) have the lowest infant mortality rates for newborns but
- 195 eventually have the highest infant, extrinsic and senescent mortality rates. Hunter-gatherers and the
- 196 Taiwanese pastoralists have the greatest increases in senescent mortality while modern Swedes
- have the lowest; the Gambia data provide an intermediate case. We also parameterised the model
- 198 with artificially low and high mortality curves to ensure our results are not confounded by these
- 199 counterintuitive published mortality parameters (results not shown).
- 200 Infant mortality is characterised by two age-related curves, describing mortality in the presence and
- 201 absence of the mother. If the mother dies, her child is exposed to a ten-fold increase in mortality risk
- (Shanley et al., 2007). Children under the age of two will die if their mothers die.202
- 203 Maternal mortality is either a J-shaped or exponential function (equation 1). Parameters for the J-
- shaped function were calculated from data presented in Blanc et al. (2013) fitted to a second-degree
- polynomial; the exponential function was fitted to data in Grimes (1994). See Supplementary Table
- 206 S2 for parameter values and Fig. 1b for a visual representation of maternal mortality.
- 207 The model is run under different assumptions concerning the relative importance of the inter-
- 208 relationships of children with each other and their mothers, firstly with each factor in isolation and
- then in combination. As the state variables are the mother's and her children's ages and, as the

210 decision of whether to give birth is annual, these ages are simply incremented by one year. A 14 year 211 old child matures and becomes independent of the mother the following year. Mortality introduces a stochastic element into the model, as there is a finite probability that the mother and any one (or 213 even all) the children may not survive to the following year. For example, for a 30 year old woman 214 with a 3 year old child who gives birth, there are 8 different states that need to be considered in the following year (see Supplementary Table S3 for an example calculation).215

216 **Sibling Competition**

Siblings compete for maternal resources and thus have detrimental effects on each other's survival 218 (Hill & Hurtado, 1996; Rutstein, 2005; Bøhler & Bergström, 2008). To model this, we calculate a 219 weighting factor for each child that increases or decreases her mortality risk, depending on the ages 220 of her siblings. In the absence of quantitative models of human sibling competition in the literature, 221 we assume a linear, additive effect for four levels of competition: none, low, medium and high (Fig. 222 $2a$).

223 A high weighting results in a large effect on mortality; conversely a low weighting results in a negligible effect on mortality. For a child aged *y* with siblings in family structure *C*, the total mortality224 225 rate for the child, $\mu_{child}(y)$, is given by her intrinsic mortality, $\mu_{childintrinsic}(y)$, modified by the sum 226 of these weightings:

$$
\mu_{\text{child}}(y) = \mu_{\text{children}}(y) \times (1 + \sum_{c} \text{weightings})
$$
 (2)

where:

$$
\mu_{\text{children}}(y) = a_1 e^{-yb_1}
$$

227

228 Here, a_1 , b_1 are population-specific mortality parameters. The sum of weights due to family structure, *C*, exclude the weight of the focal child age *y*. The child's annual probability of survival is229 230 exp $\{-\mu_{\text{child}}(y)\}.$

231 **Juvenile Help**

- 232 In some models we assume children over the age of 10 can have beneficial effects in the family by
- 233 decreasing their siblings' risk of dying. As for sibling competition, quantitative models of age-based
- 234 levels of help are absent from the literature. Thus, we model help as a linear, additive effect which
- 235 decreases the detrimental effect of the weighting described above for four different intensities of
- 236 help: none, low, medium and high (Fig. 2b).
- 1037 Juvenile help, as modelled here, has a weaker effect than sibling competition. In order to investigate
- 238 the extent to which this assumption affects our results, we also conducted a sensitivity analysis
- 239 where we varied the weightings of help relative to competition.

240 **The Dynamic Programming Equation**

- For each birth decision (action *u*) taken by an adult female of a certain age (*x*) with family structure241
- (set of children) C , we calculate the number of offspring in the following year from:
- 243 1. The adult female's probability of surviving to the next year.
- 244 2. For each possible family structure next year, the probability the mother is in the new state
- (age and family structure), given her survival and the survival of her offspring.245
- 246 3. For each possible family structure next year, the probability that a new child is born and 247 survives.
- 248 The decision of whether or not to give birth is taken in view of the risk of childbirth and the burden
- 249 of having a dependent child the following year, if it survives. Children that are 15 years old are
- 250 considered independent of the mother and, assuming female demographic dominance, only adult
- 251 females are included in the calculations (Charlesworth, 1994).
- 252 Given the mother's age and present family structure, the optimal birth strategy is determined by the
- fitness of the strategy, i.e. maximising the maximum eigenvalue of the projection matrix (Houston &253
- 254 McNamara, 1999). We define $f_t(x, C)$ as the expected number of descendants left *t* years in the

future by a female in state (x, C) . Initially $f_0(x, C) = 1$ for all ages x and family structures *C* except 256 $f_0(x_{\text{dead}},\emptyset) = 0$ (i.e. there are no fitness benefits to dying without children). From f_0 , we can 257 calculate f_1, f_2 , etc. from the dynamic programming equation:

$$
f_{t+1}(x, \mathcal{C}) := \max_{u} \sum_{\mathcal{C}'} \{ [P_u(x)Q_u(x, \mathcal{C}, \mathcal{C}')]f_t(x+1, \mathcal{C}') \} + \left[\left(1 - P_u(x) \right) Q_u(x_{\text{dead}}, \mathcal{C}, \mathcal{C}') f_t(x_{\text{dead}}, \mathcal{C}') \right] + \left[R_u(x, \mathcal{C}, \mathcal{C}') f_t(15, \emptyset) \right] \}
$$
\n(3)

258 where:

259 i) The census time is prior to the reproductive decision, therefore 15 year olds have only just 260 matured.

261 ii) The probability of a 14 year old surviving to become mature in the next year is not affected 262 by her mother's survival. However, the maturing child's survival can depend on the

presence of siblings, including babies born under birth decision *u* given the mother's age263 and current family structure, *C*.264

265 iii) Mature males are assumed to have the same reproductive value as females and an even 266 Sex ratio is assumed.

267 iv) The minimum IBI is two years but in the event of a newborn not surviving to the next time 268 **interval, the focal female can reproduce again.**

269 The growth rate of a population following the optimal strategy is given by the ratio $\lambda_{t+1} =$

270 $f_{t+1}(s_0)/f_t(s_0)$ for a reference state, s_0 (McNamara, 1991). The iteration process was judged to

271 have converged on an optimal strategy when $\lambda_{t+1} \cong \lambda_t$, to seven decimal places.

272 **The Simulated Population**

273 The optimal IBI is determined as a function of all possible states. Stochasticity is inherent in the 274 model as there can be a number of states in the next time interval with a calculated probability 275 depending on the probability of survival of children and mother. The population is simulated by 276 modelling population growth forward in time using the state-dependent optimal strategy. The

annual population growth rate at the stable age distribution has the same value as the relative 278 fitness determined in the dynamic optimisation procedure outlined above.

279 In the results that follow, the population is described in terms of the average IBIs. There are a 280 number of different ways to define IBI, such as an average of all birth spacings at a given age, or the 281 interval between a newborn and the next child. For example, a 35 year old female with 3 children of 5, 9 and 12 years old who gives birth has an average birth spacing of 4 years or alternatively a birth 283 interval at 35 of 5 years. An additional problem in defining IBI is how to include children who have 284 died. For example in the previous example the 35 year old female may have given birth in the previous year, in which case the IBI at 34 was 4 years, but the baby died. In the work that follows,285 286 the IBI relates to the spacing between a newborn baby and the next youngest child, unless stated 287 otherwise.

288 **Probabilistic Age at First Birth**

10 15 289 In order to involve fewer degrees of freedom, the model fixes the age of first birth at 15 years and 290 does not impose menopause. Although this paper is concerned with reproductive schedules 291 throughout the lifespan rather than the initial decision to reproduce, we ran a set of experiments 292 where age at first birth was probabilistic. Females were still assumed to mature at age 15 but gave 293 birth for the first time with a probability calculated from the function $y = 0.25 + 0.15x$, where x 294 is the age between 15 and 20. Thus, newly mature females have a probability = 0.25 of giving birth at 295 age 15, linearly increasing such that first birth is guaranteed by age 20.

296 The code is freely available; see Supplementary Information for download instructions.

297 **Results**

298 IBIs increase from first reproduction until age 30 in the Ache, Sweden and Taiwan populations (Fig. 299 3; red, green and blue lines, respectively), after which they remain relatively constant until the end

-
- of the reproductive span at age 50. Birth intervals in the Tsimane and Gambian populations (Fig. 3;

301 purple and yellow lines) decrease slightly from the age of 20 and again remain constant until aged 50. Fig. 3 shows these effects for the cases where there is no risk of dying in childbirth, averaged across all sibling effects (competition and juvenile help). The average IBI hovers in the range 2.05-303-303-304 2.72 years across populations.

305 **Sibling competition and juvenile help**

306 Length of the optimal IBI is sensitive to how severely children compete for maternal resources as 307 well as to mortality risks in the population (Fig. 4). In the Taiwanese population, for example, 308 increasing the intensity of sibling competition from 'none' to 'high' causes the median IBI to increase 309 by 1.24 years. When there are higher levels of environmental mortality, such as in the Gambian and Tsimane populations, birth intervals are less affected by the level of sibling competition. In 'easier'310 environments, such as Sweden, birth intervals increase with the intensity of sibling competition.311

Juvenile help, on the other hand, has a small effect on birth spacing, which only becomes apparent

313 after the age of 30 (Fig. 5 and Supplementary Fig. S3). The highest level of help decreases the IBI only

by a maximum of 0.15 years (in the Taiwan population with 'medium' sibling competition).314

Supplementary Table S5 shows the extent to which sibling competition and juvenile help can extend

316 or contract birth intervals. In order to understand the effect that our assumption of weaker levels of

at help compared to competition, we varied the strength of juvenile help. Even when help has the

318 same, but opposite, weighting as sibling competition, IBIs are not strongly affected except when help

319 is 'high' intensity but competition is 'low' or absent (Supplementary Fig. S4).

In order to tease apart the independent effects of infant and senescent mortality, we ran the model 321 holding each of these two factors constant in turn. When children were not exposed to any mortality hazards -- but the rate of senescent and extrinsic mortality could vary across all populations -- birth 323 intervals remained at the minimum of 2 years, regardless of the levels of sibling competition or juvenile help (results not shown but follow the same pattern as the red lines in Fig. 4). This is324

325 unsurprising, since sibling effects cannot occur when there is no infant mortality.

Increasing the intensity of sibling competition lengthens the birth intervals when infant mortality 327 occurred but the rate of adult mortality was held constant across populations (results not shown but 328 are the same as in Fig. 4). Thus, infant mortality, in the presence of sibling competition, appears to 329 drive increases in IBIs.

330 **Menopause can be favoured under extreme age-dependent maternal mortality**

331 To explore the circumstances that might select for menopause, we increased the potential

332 reproductive span to a maximum age of 90. An age-related risk of dying during childbirth has a

and inegligible effect on birth spacing when the mortality function is J-shaped; females continued to

334 reproduce until death (Supplementary Fig. S2, panel A). Reproductive cessation only becomes

adaptive under extreme levels of maternal mortality risk that increase exponentially with age

(Supplementary Fig. S2, panel B). It should be noted that effects in old age, such as menopause,336

337 would be more realistic had the model included grandmaternal effects on child survival (which this

338 model does not attempt to do; see Discussion).

339 **Probabilistic age at first birth does not affect birth decisions later in life**

340 When age at first birth was probabilistic rather than fixed at 15, the female experienced an initial spike in birth intervals where they increased to a maximum of 3.95 years (Supplementary Fig. S5)3141 Before dropping, at age 22, to the minimum of 2 years. After this point, optimal birth intervals follow 343 the same pattern as shown in Fig. 4.

344 **Discussion**

345 The model uses a comprehensive description of the mother and her family structure to obtain the 346 optimal birth strategy that maximises the number of offspring who survive to sexual maturity, a key 347 component of fitness. Alongside this, the model takes into account the stochastic year to year 348 changes that can occur in the family across a set of realistic mortality hazards derived from five 349 human populations. Optimal reproductive decisions are based on the complex interaction of family

350 members and the environment. Although not explicitly included, the strength of these interactions is 351 likely to be determined by resource availability.

352 The dynamic, state-dependent framework presented here shows how mortality hazards and sibling 353 competition interact to produce a range of life history strategies. IBIs increase with age in three of 354 the five simulated populations until the age of 30, after which birth spacing remains constant (Fig. 3). 355 In low mortality environments (e.g. modern Sweden), increasing the intensity of sibling competition 356 results in longer IBIs compared to high mortality environments (e.g. Tsimane; Fig. 4). Even at young ages, mothers reproduce below their maximum potential level of reproductive output in order to 358 enhance the survival prospects of existing children (Figs. 3 and 4). Siblings providing help to each 359 other did little to reduce optimal IBIs (Fig. 5, Supplementary Table S5 and Fig. S4). These effects 360 alone do not induce menopause; it is only in the presence of extreme and exponentially increasing age-related risks of dying during childbirth that reproductive cessation becomes adaptive 362 (Supplementary Fig. S2).

363 Our results predict many aspects of observed life history patterns. Among Ache hunter-gatherers, 364 the initial birth interval for women giving birth at age 15 was \sim 2.5 years; the median IBI of Ache 365 women is 3 years and remains relatively constant throughout her life (Hill & Hurtado, 1996). Under 366 high levels of sibling competition, the Ache IBI in our model reached a maximum of 2.96 years (Fig. 4, 367 top-left panel).

In the absence of published empirical data, we modelled sibling competition and juvenile help as368 369 linearly increasing or decreasing (respectively) the mortality risks of other children in the family. The abled of a particular child depended on her age and affected all siblings equally. A more realistic 371 implementation of this might include the ages of siblings in the effects. Newborns could have a more 372 deleterious effect on young siblings rather than older ones who are capable of provisioning 373 themselves, although this will depend of ecology; among the Ache, for example, children older than 10 years who were raised with more competing juvenile siblings suffered higher mortality (Hill &

375 Hurtado, 1996). In order to understand kin effects independently of environment-specific mortality, 376 the patterns of kin effects were assumed to be the same across all five modelled populations. Future 377 work could also tailor the levels of help and competition to the mortality rates in different ecologies. 378 We expect that altering the dynamics of kin effects in these ways would lead to greater divergence 379 in reproductive schedules between populations but less variation within a population. Less intense 380 sibling competition brought about by 'easier' environments might lead to shorter birth intervals, all 381 else equal. Introducing other allocarers, such as grandparents (see below), into the model could also 382 alleviate the effects of sibling competition.

383 We assumed that sibling competition occurs over maternal resources that are directly invested in 384 one offspring at the expense of others, with effects that diminish with age, although sibling 385 competition for parental resources can continue into adulthood (Mace, 2013). Social institutions 386 such as arranged marriages can also affect sibling competition depending on birth order, the 387 presence of same-sex siblings and local demography. In South Asia, for example, the presence of 388 older sisters can increase a girl's education by allowing her to remain in school rather than marry 389 (Vogl, 2013).

390 Our modelling framework also assumes that all offspring have equal quality. However, the reproductive value – and sex – of the youngest child can affect a mother's IBIs. For example,391 392 firstborn boys of high reproductive value often receive additional care with an associated delay to 393 the next child (Mace & Sear, 1997). Other primates, and indeed other mammals, also have a delayed 394 interval following the birth of a male offspring (Bercovitch & Berard, 1993; Birgersson, 1998). 395 Children in Tanzania were more likely to be weaned later when they were later-born or heavier at 396 birth, while socioeconomic status also played a role: high-status females and low-status males 397 received less parental investment in the form of breastfeeding (Wander & Mattison, 2013). 398 Maternal quality could be modelled by introducing a probability of birth depending on her fertility.

Fertility is variable in terms of ecological conditions and physiological status of women, as shown by

the seasonality of birth, response to food supply, and the effect of lactational amenorrhoea (Ellison400 et al., 1993; Kaplan, 1996). Juveniles will inevitably vary according to the quality of care they have received with a corresponding effect on growth and age of reproductive maturity (McNamara &402 403 Houston, 1996).

In addition to the sex, reproductive value, and birth order of offspring, birth spacing decisions404 405 respond to other circumstances, such as the constraint that carrying food imposes on the number of 406 young offspring (Blurton Jones & Sibly, 1978; Blurton Jones, 1986; Anderies, 1996). In farming and About the populations, the heritable resources needed in adult life to go on to reproduce also 408 constrain reproductive schedules (Mace, 1998). The current version of the model only tracks female offspring. It would be interesting to include males in order to test hypotheses about birth order and 410 the sex ratios of offspring (following, e.g. Trivers & Willard, 1973 and Leimar, 1996).

All The model tracks individual females and their children under the age of 15 years, from which IBIs are 412 calculated. Once a female offspring is beyond the age of 15, it is no longer possible to know if her 413 mother is alive or how old she is. In some model scenarios, this leads to a sudden increase in reproduction in late life, because the risk of death no longer has any cost once children are 15414 415 (Supplementary Fig. S1). In reality, grandmothers can continue to enhance the fitness of older 416 offspring (Sear & Mace, 2008), so such late life peaks in fertility are an artefact of the model 417 structure being limited to two generations. Clearly, to examine more closely the effect of All grandmaternal care, children must be followed beyond independence and the influence of the maternal grandmother can then be modelled explicitly. In a similar fashion, we expect that inclusion419 420 of a grandmaternal generation would allow menopause to evolve in the model without assuming 421 extreme maternal mortality rates.

In summary, we have developed a comprehensive, dynamic framework for the study of optimal IBIs422 423 and explored how sibling behaviour affects maternal reproductive success in different ecologies. At 424 high levels of infant mortality, sibling effects become less important, presumably as the risk of

- 425 mortality exceeds the risks associated with competition should each child survive (Fig. 4). This 426 suggests a stronger role for sibling competition only when mortality is low, as has been noted in 427 some modern populations (Lawson & Mace, 2009).
- As mentioned above, extensions of this framework could take into account a third, grandmothering
- 429 generation in order to test hypotheses about the evolution of menopause. As it stands, our model
- 430 treats the nuclear family as an 'island', unaffected by the lives and strategies of others. Future work
- 431 might investigate the effects of other family members (e.g. grandparents, spouses, in-laws,
- 432 stepparents) on optimal reproductive scheduling..

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439 **References**

- 440 Abitbol, M. 1996. *Birth and human evolution: anatomical and obstetrical mechanics in* 441 *primates*. Bergin & Garvey.
- 442 Anderies, J. M. 1996. An adaptive model for predicting !Kung reproductive performance: A 443 stochastic dynamic programming approach. *Ethol. Sociobiol.* **17**: 221–245.
- 444 Beise, J. 2005. The helping grandmother and the helpful grandmother: The role of maternal 445 and paternal grandmothers in child mortality in the 17th and 18th century population 446 of French settlers in Quebec, Canada. In: *Gd. Evol. Significance Second Half Female Life* 447 (E. Voland, A. Chasiotis, & W. Schiefenhövel, eds), p. 215−238. Rutgers University Press, 448 New Brunswick.
- Bercovitch, F. B., & Berard, J. D. 1993. Life history costs and consequences of rapid reproductive maturation in female rhesus macaques. *Behav. Ecol. Sociobiol.* **32**: 103– 109.
- Birgersson, B. 1998. Male-biased maternal expenditure and associated costs in fallow deer. *Behav. Ecol. Sociobiol.* **43**: 87–93.
- Blanc, A. K., Winfrey, W., & Ross, J. 2013. New findings for maternal mortality age patterns: aggregated results for 38 countries. *PLoS One* **8**: e59864.
- Blurton Jones, N. G. 1986. Bushman birth spacing: A test for optimal interbirth intervals. *Ethol. Sociobiol.* **7**: 91–105.
- Blurton Jones, N. G., & Sibly, R. M. 1978. Testing adaptiveness of culturally determined behavior: do bushman women maximize their reproductive success by spacing births widely and foraging seldom? In: *Hum. Behav. Adapt.* (V. Reynolds & N. G. Blurton Jones, eds), pp. 135–157. Taylor and Francis, London.
- Bogin, B. 1997. Evolutionary hypotheses for human childhood. *Am. J. Phys. Anthropol.* **104**: 63–89.
- Bøhler, E., & Bergström, S. 2008. Subsequent pregnancy affects morbidity of previous child. *J. Biosoc. Sci.* **27**: 431–442.
- Caro, T. M., Sellen, D. W., Parish, a., Frank, R., Brown, D. M., Voland, E., & Mulder, M. B. 1995. Termination of reproduction in nonhuman and human female primates. *Int. J. Primatol.* **16**: 205–220.
- Charlesworth, B. 1994. *Evolution in Age-Structured Populations*. Cambridge University Press.
- Conde-Agudelo, A., Rosas-Bermudez, A., & Kafury-Goeta, A. C. 2006. Birth Spacing and Risk of Adverse Perinatal Outcomes. *JAMA* **295**: 1809–1823.
- Ellison, P. T., Panter-Brick, C., Lipson, S. F., & O'Rourke, M. T. 1993. The ecological context of human ovarian function. *Hum. Reprod.* **8**: 2248–2258.
- Grimes, D. A. 1994. The morbidity and mortality of pregnancy Still risky business. *Am. J. Obstet. Gynecol.* **170**: 1489–1494.
- Hawkes, K., O'Connell, J. F., & Blurton Jones, N. G. 1997. Hadza women's time allocation, offspring provisioning, and the evolution of long postmenopausal life spans. *Curr. Anthropol.* **38**: 551–577.
- Hill, K., & Hurtado, A. M. 1996. *Aché Life History: The Ecology and Demography of a Foraging People*. Aldine de Gruyter, New York.
- Houston, A. I., & McNamara, J. M. 1999. *Models of Adaptive Behaviour: An Approach Based on State*. Cambridge University Press.
- Howell, N. 1979. *Demography of the Dobe Kung*. Academic Press, New York.
- Kaplan, H. 1996. A theory of fertility and parental investment in traditional and modern human societies. *Yearb. Phys. Anthropol.* **39**: 91–135.
- Kramer, K. L. 2010. Cooperative Breeding and its Significance to the Demographic Success of Humans. *Annu. Rev. Anthropol.* **39**: 417–436.
- Lawson, D. W., & Mace, R. 2009. Trade-offs in modern parenting: a longitudinal study of sibling competition for parental care. *Evol. Hum. Behav.* **30**: 170–183.
- Leimar, O. 1996. Life-history analysis of the Trivers and Willard sex-ratio problem. *Behav. Ecol.* **7**: 316–325.
- Mace, R. 1998. The coevolution of human fertility and wealth inheritance strategies. *Philos. Trans. R. Soc. B Biol. Sci.* **353**: 389–397.
- Mace, R. 2013. Cooperation and conflict between women in the family. *Evol. Anthropol.* **22**: 251–258.
- Mace, R., & Sear, R. 1997. Birth interval and the sex of children in a traditional African population: an evolutionary analysis. *J. Biosoc. Sci.* **29**: 499–507.
- McNamara, J. M. 1991. Optimal life histories: A generalisation of the Perron-Frobenius theorem. *Theor. Popul. Biol.* **40**: 230–245.
- McNamara, J. M., & Houston, A. I. 1996. State-dependent life histories. *Nature* **380**: 215– 221.
- Nenko, I., Hayward, A. D., & Lummaa, V. 2013. The effect of socio-economic status and food availability on first birth interval in a pre-industrial human population. *Proc. R. Soc. B Biol. Sci.* **281**: 1–9.
- Rutstein, S. O. 2005. Effects of preceding birth intervals on neonatal, infant and under-five years mortality and nutritional status in developing countries: evidence from the demographic and health surveys. *Int. J. Gynaecol. Obstet.* **89 Suppl 1**: S7–24.
- Sear, R. 2008. Kin and Child Survival in Rural Malawi. *Hum. Nat.* **19**: 277–293.
- Sear, R., & Mace, R. 2008. Who keeps children alive? A review of the effects of kin on child survival. *Evol. Hum. Behav.* **29**: 1–18.
- Sear, R., Steele, F., McGregor, I. A., & Mace, R. 2002. The effects of kin on child mortality in rural Gambia. *Demography* **39**: 43–63.
- Shanley, D. P., Sear, R., Mace, R., & Kirkwood, T. B. L. 2007. Testing evolutionary theories of menopause. *Proc. R. Soc. B Biol. Sci.* **274**: 2943–2949.
- Short, R. V., Lewis, P. R., Renfree, M. B., & Shaw, G. 1991. Contraceptive effects of extended lactational amenorrhoea: beyond the Bellagio Consensus. *Lancet* **337**: 715–717.
- Sievert, L. L. 2006. *Menopause: A Biocultural Perspective*. Rutgers University Press, Piscataway, New Jersey.
- Siler, W. 1979. A competing-risk model for animal mortality. *Ecology* **60**: 750–757.
- Trivers, R. L., & Willard, D. E. 1973. Natural Selection of Parental Ability to Vary the Sex Ratio of Offspring. *Sci.* **179** : 90–92.
- Vogl, T. S. 2013. Marriage institutions and sibling competition: Evidence from South Asia. *Q. J. Econ.* 1017–1072.
- Wander, K., & Mattison, S. M. 2013. The evolutionary ecology of early weaning in Kilimanjaro, Tanzania. *Proc. Biol. Sci.* **280**: 20131359.
- WHO. 2006. *Report of a WHO Technical Consultation on Birth Spacing*.
- Willführ, K. P., & Gagnon, A. 2013. Are stepparents always evil? Parental death, remarriage,
- and child survival in demographically saturated Krummhörn (1720-1859) and
- expanding Québec (1670-1750). *Biodemography Soc. Biol.* **59**: 191–211.

531 **Figures**

Figure 1: (a) Age-specific mortality risk in each of the five modelled populations: Ache (red), Gambia 533 (yellow), Sweden (green), Taiwan (blue), Tsimane (purple) from birth until a maximum lifespan of 90 534 (although note that reproductive spans last until age 50). (b) Maternal mortality hazards during the 535 reproductive span (ages 15-50): Solid lines show fitted probability functions for a J-shaped mortality 536 function derived from Blanc et al. (2013) with three levels: low (blue); medium (green); high (red). 537 Dashed lines show an exponential maternal mortality function derived from Grimes (1994) with 538 three levels: low (brown), medium (green) and high (red). See Supplementary Tables S1 and S2 for 539 parameters and references.

Figure 2: Age-specific sibling effects. Higher weights have stronger effects on sibling survival. (a) 541 Sibling competition (b) Sibling help. Intensities of sibling effects: none (red), low (yellow), medium 542 (blue) and high (purple).

Figure 3: Age-specific interbirth intervals (IBIs) for each of the five modelled populations: Ache (red), 544 Gambia (yellow), Sweden (green), Taiwan (blue), Tsimane (purple). Infant, adult and extrinsic 545 mortality vary according to the population parameters (see Table S1). Here, maternal mortality was 546 set to 'none', meaning the focal female did not face any increase in mortality due to giving birth. 547 Data points are the mean IBI values across the range of sibling competition and juvenile help 548 parameters. Values presented here do not include IBIs after the death of children, so IBIs are 549 independent of any replacement effect.

Figure 4: The effects of sibling competition on interbirth interval for the five modelled populations. 551 Here, each population experiences mortality according to the parameters in Table S1. Each panel 552 shows IBI for the four intensities of sibling competition: none (red); low (yellow); medium (blue); 553 high (purple).

Figure 5: Juvenile help does not have a strong effect on interbirth intervals. The curves are optimal 555 birth intervals for Sweden, for each of the four intensities of sibling competition: none (red); low

- 556 (yellow); medium (blue); high (purple). Panels show, from left to right, increasing levels of sibling
- 557 help (see Fig. 2b). Supplementary Fig. S3 shows the effects of juvenile help across all modelled
- 558 populations and Supplementary Fig. S4 illustrates a sensitivity analysis on our juvenile help
- 559 assumptions.