Role of Adiponectin in Coronary Heart Disease Risk A Mendelian Randomization Study

Maria Carolina Borges, Debbie A. Lawlor, Cesar de Oliveira, Jon White, Bernardo Lessa Horta, Aluísio J.D. Barros

Rationale: **Hypoadiponectinemia correlates with several coronary heart disease (CHD) risk factors. However, it is unknown whether adiponectin is causally implicated in CHD pathogenesis.**

Objective: **We aimed to investigate the causal effect of adiponectin on CHD risk.**

Methods and Results: **We undertook a Mendelian randomization study using data from genome-wide association studies consortia. We used the ADIPOGen consortium to identify genetic variants that could be used as instrumental variables for the effect of adiponectin. Data on the association of these genetic variants with CHD risk were obtained from CARDIoGRAM (22 233 CHD cases and 64 762 controls of European ancestry) and from CARDIoGRAMplusC4D Metabochip (63746 cases and 130681 controls; ≈ 91% of European ancestry) consortia. Data on the association of genetic variants with adiponectin levels and with CHD were combined to estimate the influence of blood adiponectin on CHD risk. In the conservative approach (restricted to using variants within the adiponectin gene as instrumental variables), each 1 U increase in log blood adiponectin concentration was associated with an odds ratio for CHD of 0.83 (95% confidence interval, 0.68–1.01) in CARDIoGRAM and 0.97 (95% confidence interval, 0.84–1.12) in CARDIoGRAMplusC4D Metabochip. Findings from the liberal approach (including variants in any locus across the genome) indicated a protective effect of adiponectin that was attenuated to the null after adjustment for known CHD predictors.**

Conclusions: **Overall, our findings do not support a causal role of adiponectin levels in CHD pathogenesis. (***Circ Res***. 2016;119:491-499. DOI: 10.1161/CIRCRESAHA.116.308716.)**

> **Key Words:** adiponectin ■ cardiovascular disease ■ coronary artery disease ■ mendelian randomization analysis ■ obesity

Adiponectin, a 30 kDa protein produced mainly by mature
Aadipocytes, has been implicated in a wide spectrum of biological pathways related to peripheral insulin sensitivity, $\frac{1}{2}$ inflammatory response,^{1,2} and atherogenesis.² In contrast to most adipokines, adiponectin secretion is downregulated in obese individuals.3 Observational epidemiological studies support that hypoadiponectinemia is associated with cardiovascular risk factors^{4,5} (eg, insulin resistance and dyslipidemia) and type 2 diabetes mellitus risk⁶; inconsistent findings have been observed on coronary heart disease $(CHD)^{7-10}$ and stroke risk.^{9,11}

Mendelian randomization studies make use of genetic variants as instrumental variables to investigate the effect of environmental exposures and biomarkers on outcomes. Because alleles are randomly allocated during gametogenesis and genotype is a fixed exposure, Mendelian randomization studies are not as vulnerable to confounding and reverse causality and can substantially improve causal inference from observational data.12 Mendelian randomization is regarded as nature's analogue of randomized controlled trials and has successfully been used in cardiovascular research to investigate potential etiologic mechanisms,13 validate and prioritize novel drug targets,¹⁴ and increase understanding of current therapies.15

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There is evidence of a shared allelic architecture of circulating adiponectin with CHD risk and carotid intima–media thickness^{16,17}; however, it remains unanswered if these findings implicate a causal effect of adiponectin on CHD risk or merely shared pleiotropic factors. Our aim was to investigate the causal effect of adiponectin on CHD risk using Mendelian randomization.

Methods

Study Design

We performed a 2-sample Mendelian randomization analysis using summary data from genome-wide association studies (GWAS) consortia. Single-nucleotide polymorphisms (SNPs), previously reported to be associated with blood adiponectin levels, were used as instrumental variables for testing the causal effect of adiponectin on CHD risk. Data on the association of SNPs with (1) adiponectin levels (first sample) and (2) CHD risk (second samples) were combined to estimate the influence of blood adiponectin on CHD risk. To investigate the presence of potential bias (horizontal pleiotropy) or mediation of the effect of adiponectin on CHD via other CHD risk factors (vertical pleiotropy; Online Figure I), we also analyzed data on the association of the selected adiponectinrelated SNPs with a range of CHD risk factors: glycohemoglobin, fasting insulin, high-density lipoprotein cholesterol (HDL-c), low-density lipoprotein cholesterol (LDL-c), triacylglycerols (TAG), body mass index (BMI), and BMI-adjusted waist circumference (WC).

Data Sources

Summary data on the association between SNPs and the phenotypes of interest were extracted from public databases of different consortia: ADIPOGen for adiponectin¹⁸; CARDIoGRAM (Coronary ARtery DIsease Genome-wide Replication And Meta-analysis)¹⁹ and CARDIoGRAMplusC4D Metabochip (CARDIoGRAMplusC4D Metabochip and GWAS meta-analysis)²⁰ for CHD; MAGIC (Meta-Analyses of Glucose and Insulin-Related Traits Consortium) for

glycohemoglobin²¹ and fasting insulin²²; GLGC (Global Lipids Genetics Consortium) for HDL-c, LDL-c, and TAG²³; and GIANT (Genetic Investigation of ANthropometric Traits) for BMI24 and WC.²⁵ Details about each data source are displayed in Online Table I. CARDIoGRAMplusC4D Metabochip includes data from CARDIoGRAM GWAS.

Instrumental Variables

The SNPs for our main instrumental variable analyses (n=17 SNPs) were selected from 145 SNPs strongly (*P*<5×10−8) associated with blood adiponectin levels in the European ancestry GWAS meta-analysis from the ADIPOGen consortium.¹⁸ Independent SNPs were previously selected by Dastani et al¹⁶ by linkage disequilibrium pruning of the genome-wide significant SNPs, retaining SNPs that explained most variance in adiponectin levels in each linkage disequilibrium block (linkage disequilibrium threshold: *R*² <0.05 in HapMap CEU population [Utah residents with Northern and Western European ancestry]; Table 1).

We used 2 sets of instruments (Figure 1):

- 1. A conservative instrumental variable analysis, in which only SNPs within the *ADIPOQ* locus (±50 kb) were considered eligible (n=4 SNPs; C4). *ADIPOQ* is mainly expressed in adipose tissue and encodes adiponectin. We considered this approach unlikely to be biased by horizontal pleiotropy given the functional relationship of *ADIPOQ* to adiponectin levels.
- 2. A liberal analysis, in which independent SNPs from any locus that had reached a genome-wide significant association (*P*<5*10−8) with adiponectin levels in the ADIPOGen consortia GWAS (n=17 SNPs), were included (L17), as previously reported by Dastani et al.¹⁶ These 17 SNPs included the four SNPs within the *ADIPOQ* locus.

Ten of the 17 selected SNPs could be found in CARDIoGRAMplusC4D Metabochip data, 3 of which were proxy SNPs $(R^2 > 0.95$ for CEU population). For the remaining 7 SNPs, data from CARDIoGRAM GWAS was used. As the SNP rs1108842 could not be found in GLGC data, a proxy SNP (rs13083798) in perfect linkage disequilibrium $(R^2=1.0$ for CEU population) was used instead.

Validation of Instrumental Variable Assumptions

Validity of Mendelian randomization analyses results can be compromised if the instrumental variable assumptions are violated. In Online Table II, we described the 3 core assumptions of instrumental variable analysis and the strategies used to address these.

Estimation of Causal Effect

For both liberal and conservative approaches, the β coefficient (log odds ratio of CHD per one natural log greater adiponectin level) and its SE were calculated using the inverse-variance weighted (IVW) method as described by Burgess et al.26 (See Online Data Supplement).

For the liberal approach, we also used the IVW method to estimate the combined effect of adiponectin levels on cardiovascular risk factors (glycohemoglobin, fasting insulin levels, HDL-c, LDL-c, TAG, BMI, and WC). Where we found evidence of an effect of the SNPs on these risk factors, estimates of the association between adiponectin and CHD were adjusted for these risk factors to reduce the possibility that horizontal pleiotropy biased our findings²⁷ (See Online Data Supplement).

Sensitivity Analyses

Assuming that all valid instrumental variables identify the same causal parameter, substantial heterogeneity would be suggestive of pleiotropic SNPs. We evaluated heterogeneity in our IVW estimates using standard tools from the meta-analysis literature: forest plot of per SNP ratio estimate, Cochran Q test, and I^2 values.²⁸⁻³⁰ In addition, to identify overly influential SNPs, additional meta-analyses were performed by removing 1 SNP at a time and recalculating the overall instrumental variable estimates.

SNP	Chr	Position*	Closest Gene	EA	NEA	EAF†	C ₄	L17
rs1415293	1	219730006	ZC3H11B	T	A	0.25	\cdots	$\sqrt{}$
rs1108842	3	52720080	GNL3	C 0.49 A		.	$\sqrt{}$	
rs6810075	3	186548565	ADIPOQ	T	C	0.61	√	$\sqrt{}$
rs16861209	3	186563114	ADIPOQ	A	C	0.08	$\sqrt{}$	$\sqrt{}$
rs17366568	3	186570453	ADIPOQ-AS1, ADIPOQ	G	A	0.93	$\sqrt{}$	$\sqrt{}$
rs3774261	3	186571559	ADIPOQ-AS1, ADIPOQ	A	G	0.50	√	$\sqrt{}$
rs998584	6	43757896	VEGFA	C	A	0.54	\cdots	$\sqrt{}$
rs2980880	8	126480972	TRIB1	A	G	0.71	\cdots	$\sqrt{}$
rs7955516	12	20498036	PDE3A	C	A	0.28	\cdots	$\sqrt{}$
rs601339	12	123174743	HCAR ₂	G	A	0.25	.	$\sqrt{}$
rs6488898	12	124203832	ATP6V0A2	A	G	0.98	\cdots	$\sqrt{}$
rs7978610	12	124468572	ZNF664, FAM101A	C	G	0.27	\cdots	$\sqrt{}$
rs2925979	16	81534790	CMIP	C	T	0.71	\cdots	$\sqrt{}$
rs7200895	16	82644606	CDH ₁₃	T	C	0.69	\cdots	$\sqrt{}$
rs8047711	16	82667671	CDH ₁₃	G	A	0.92	\cdots	$\sqrt{}$
rs12929479	16	82997853	CDH ₁₃	G	A	0.42	.	$\sqrt{}$
rs731839	19	33899065	PEPD	A	G	0.54	.	$\sqrt{}$

Table 1. Characteristics of SNPs Selected for Each Analytic Approach

C4 indicates the 4 SNPs used in the conservative analyses; Chr indicates chromosome; EA, effect allele; EAF, effect allele frequency; L17, 17 SNPs used in the liberal analyses (SNPs selected on the basis of reaching genome-wide significant levels in association with adiponectin, $P < 5 \times 10^{-8}$); and NEA, noneffect allele.

*Genome Reference Consortium Human Build 37.

†1000 Genomes.

Even after adjusting for cardiovascular risk factors associated with our instrument, the liberal approach estimates could still be biased by unknown horizontal pleiotropic pathways that link the adiponectin genetic instrumental variable to CHD independently of path through adiponectin. To explore the presence of this possible bias, the MR-Egger regression method was used.³¹ See Online Data Supplement for a description of this method.

We also undertook a positive control analysis that consisted of a Mendelian randomization analysis in which LDL-c was the biomarker of interest and CHD risk was the outcome (using the IVW and MR-Egger method) because of its established causal role in CHD development (see Online Data Supplement).

Results

Association of the Genetic Instrument With Adiponectin and CHD Risk

Figure 2 shows the associations of SNPs, used as instrumental variables in the conservative (n=4 SNPs within *ADIPOQ* gene) and liberal analyses (n=17 SNPs across the genome), with adiponectin levels and CHD risk. For the conservative approach, each adiponectin-increasing allele was associated with 2.3% reduction in CHD risk (95% confidence interval [CI], −4.1 to −0.4) in CARDIoGRAM data and 0.6% reduction in CHD risk (95% CI, −1.9 to 1.0) in CARDIoGRAMplusC4D Metabochip. For the liberal approach, each adiponectin-increasing allele was associated with 2.3% reduction in CHD risk (95% CI, -3.2 to -1.5) in CARDIoGRAM data and 1.7% reduction in CHD risk (95% CI, −2.3 to −1.1%) in CARDIoGRAMplusC4D Metabochip. Of the 17 SNPs, there was some evidence of heterogeneity (*P*<0.05) between studies

that contributed to each consortium for 3 SNPs: 2 SNPs in CARDIoGRAM (rs1108842 and rs6488898) and 1 SNP in CARDIoGRAMplusC4D Metabochip (rs3774261).

Association of the Genetic Instruments With CHD Risk Factors

More than 50% of individual SNPs were associated with one or more CHD risk factor (glycohemoglobin, fasting insulin levels, HDL-c, LDL-c, TAG, BMI, and WC), and none of these SNPs were located within *ADIPOQ* gene (±50 kb; Table 2). In general, adiponectin-increasing variants were not associated with CHD risk factors in the conservative approach but were related to lower fasting insulin, higher HDL-c, lower TAG, lower WC, and higher BMI in the liberal approach (Figure 3).

Effect of Blood Adiponectin Concentration on CHD Risk

Figure 4 shows the results of all Mendelian randomization analyses assessing the association of genetically predicted adiponectin with CHD risk. Using the conservative approach (including only the 4 SNPs within *ADIPOQ* gene), each unit increase in log adiponectin concentration was associated with an odds ratio for CHD of 0.83 (95% CI, 0.68– 1.01) in CARDIoGRAM and 0.97 (95% CI, 0.84–1.12) in CARDIoGRAMplusC4D Metabochip data set. Using the liberal approach (including 17 SNPs), the odds ratio (OR) for the effect of each unit increase in log adiponectin concentration on CHD was 0.76 (95% CI, 0.65–0.89) in CARDIoGRAM and 0.83 (95% CI, 0.74–0.93) in CARDIoGRAMplusC4D

Figure 1. Analysis plan. Summary data from the association of single-nucleotide polymorphism (SNP) with phenotypes were extracted from genome-wide association study (GWAS) consortia data sets (ADIPOGen, CARDIoGRAM, C4D, MAGIC, GLGC, and GIANT). The effect of adiponectin on CHD was estimated using a conservative Mendelian randomization approach (instrumental variable: SNPs within *ADIPOQ* locus [±50 kb]) and a liberal approach (instrumental variable: SNPs in any locus). For the conservative approach, inversevariance weighted (IVW) method was used. For the liberal approach, IVW method was used in both crude and adjusted analysis for known pleiotropic factors and MR-Egger regression in the analysis accounting for hidden pleiotropy (sensitivity analysis). BMI indicates body mass index; CARDIoGRAM, Coronary Artery Disease Genome-wide Replication and Meta-analysis; CARDIoGRAMplusC4D Metabochip, CARDIoGRAMplusC4D Metabochip meta-analysis; GIANT, genetic investigation of anthropometric traits; GLGC, Global Lipids Genetics Consortium; Hb_{A1c}: glycohemoglobin; HDL, high-density lipoprotein; IV, instrumental variable; LDL, low-density lipoprotein; MAGIC, Meta-Analyses of Glucose and Insulin-related traits Consortium; MR, Mendelian randomization; SNP, singlenucleotide polymorphism; TAG, triacylglycerol; and WC, waist circumference.

Metabochip. When we adjusted these liberal approach results for the CHD risk factors associated with the genetic instrument (fasting insulin, HDL-c, TAG, WC, and BMI), the OR was 0.88 (95% CI, 0.75–1.03) in CARDIoGRAM and 1.00 (95% CI, 0.90–1.12) in CARDIoGRAMplusC4D Metabochip.

Sensitivity Analyses

There was substantial heterogeneity in IVW estimates among the 17 SNPs from the liberal approach in both CARDIoGRAM (*I* 2 =65.2; *P*=1×10−4) and CARDIoGRAMplusC4D Metabochip (*I*²=72.4; *P*=2×10⁻⁶) data (Online Figure II). The effect of removing one SNP at a time on the overall estimate showed that no SNP could explain the observed protective effect in the liberal analysis. The inclusion of the SNPs rs17366568 and rs8047711 slightly underestimated findings from the IVW method in CARDIoGRAM data set (Online Figure III).

By using the MR-Egger method with our liberal instrument, we observed further evidence of directional pleiotropy, that is, the instrument was associated with a decreased log odds of CHD independently of its effect on adiponectin in CARDIoGRAM (log OR, −0.03; 95% CI: −0.05 to −0.02 for the intercept) and in CARDIoGRAMplusC4D Metabochip (log OR, −0.03; 95% CI, −0.05 to −0.02 for the intercept; Online Figure IV). According to Mendelian randomization estimates using the MR-Egger method, each unit increase in log adiponectin concentration was associated with an OR for CHD of 1.25 (95% CI, 0.96–1.63) in CARDIoGRAM and 1.30 (95% CI, 1.06–1.58) in CARDIoGRAMplusC4D Metabochip data set (Figure 4). In the influence meta-analysis, in which we removed 1 of the 17 SNPs at a time from the pooled estimates, all of the results for the remaining 16 SNPs were in the same (positive) direction, but the magnitude of this varied somewhat (Online Figure III).

To investigate any differences between CARDIoGRAM and CARDIoGRAMplusC4D Metabochip, we compared Mendelian randomization results of the effect of LDL-c on CHD risk (positive control analysis). The OR for CHD for each standardized unit increase in LDL-c was 1.70 (95% CI, 1.54–1.88) in CARDIoGRAM and 1.57 (95% CI, 1.47–1.67) in CARDIoGRAMplusC4D Metabochip. After accounting for unknown horizontal pleiotropy (MR-Egger method), estimates were 1.96 (95% CI, 1.59–2.33) for CARDIoGRAM and 1.92 (95% CI, 1.65–2.17) for CARDIoGRAMplusC4D Metabochip.

Discussion

Taken together, our results are not supportive of a protective causal effect of adiponectin on CHD risk. First, we found no consistent evidence that genetic predisposition to elevated blood adiponectin levels is associated to reduced risk of CHD in the analysis restricted to *ADIPOQ* SNPs (conservative approach). Second, in the more liberal analysis, using variants associated with adiponectin across the genome, there was evidence of a protective effect, but this was because of horizontal pleiotropy. This conclusion regarding horizontal pleiotropy resulting in a biased apparent protective effect with our liberal approach is supported by both multivariable Mendelian randomization and MR-Egger. Some of the variants strongly associated with circulating adiponectin, in our liberal analysis,

Figure 2. Forest plots of mean difference in log adiponectin levels and odds ratio of coronary heart disease per allele of singlenucleotide polymorphism (SNP) according to the conservative (A) and liberal (B) approaches. A, Conservative approach including 4 SNPs within *ADIPOQ* gene associated with adiponectin at genome-wide significant levels (*P*<5×10−8; C4). B, Liberal approach including 17 SNPs across the genome associated with adiponectin at genome-wide significant levels (*P*<5×10−8; L17). CHD indicates coronary heart disease; Chr, chromosome; and OR, odds ratio. Results for log adiponectin included 29347 individuals from ADIPOGen consortium and for CHD risk included 86995 individuals (22233 CHD cases) from CARDIoGRAM and 194427 individuals (63746 CHD cases) from CARDIoGRAMplusC4D Metabochip consortium.

are related to loci of potential importance for LDL-c signaling in endothelial cells (*CDH13*) and for vascular biology (eg, *TRIB1* and *VEGFA*), which might explain their pleiotropic effects regarding CHD pathogenesis.18 Last, our results are strengthened by the consistent strong positive associations of LDL-c with CHD when we use the same methods used for adiponectin to test this known causal effect.

Few previous studies have conducted Mendelian randomization analysis to investigate the effect of adiponectin on metabolic diseases. Two smaller studies found evidence that genetically raised adiponectin levels were positively associated with insulin sensitivity.^{32,33} However, a larger study did not provide evidence of a causal role of adiponectin in insulin resistance or type 2 diabetes mellitus 34 but found that genetically raised insulin levels are associated with lower adiponectin

levels, suggesting that the association was possibly because higher insulin levels caused lower adiponectin, rather than the other way round.

We have undertaken the first large Mendelian randomization study of the causal effect of adiponectin on cardiovascular disease risk using GWAS consortia data from CARDIoGRAM (22 233 CHD cases and 64 762 controls) and CARDIoGRAMplusC4D Metabochip (63 746 cases and 130681 controls) with detailed phenotyping of coronary artery disease, myocardial infarction, or both. We applied a rigorous analyses plan to assess the validity and consistency of our findings. This included (1) adopting a systematic prespecified approach to selecting SNPs for our instrumental variables; (2) exploring different scenarios from the plausibly valid (but less well powered) conservative MR approach

	$\mathsf{Hb}_{\mathsf{A1c}}$		Insulin		HDL-c		$LDL-c$		TAG		WC		BMI	
	β	P	β	P	β	P	β	P	β	P	β	P	β	P
rs1415293	0.004	0.513	-0.016	1×10^{-4}	0.015	0.009	-0.012	0.063	-0.014	0.019	-0.006	0.200	0.004	0.283
rs1108842	-0.003	0.586	-0.009	0.027	0.008	0.077	0.010	0.024	-0.009	0.037	-0.020	2×10^{-8}	0.011	0.001
rs6810075*	-0.011	0.123	-0.005	0.214	-0.003	0.813	0.005	0.562	-0.004	0.350	0.003	0.530	-0.001	0.884
rs16861209*	-0.002	0.888	-0.009	0.321	0.000	0.813	0.009	0.245	-0.001	0.610	-0.004	0.570	0.000	0.995
rs17366568*	0.000	0.984	-0.005	0.478	0.009	0.292	0.012	0.343	-0.004	0.587	-0.007	0.400	0.000	1.000
rs3774261*	0.001	0.932	-0.001	0.781	-0.006	0.108	-0.001	0.863	0.001	0.707	0.000	0.910	-0.005	0.100
rs998584	-0.014	0.045	-0.002	0.657	0.026	2×10^{-11}	-0.001	0.936	-0.029	3×10^{-15}	-0.029	6×10^{-15}	0.017	9×10^{-7}
rs2980880	0.014	0.030	0.000	0.967	0.043	1×10^{-26}	-0.040	6×10^{-22}	-0.067	2×10^{-82}	0.001	0.790	0.007	0.026
rs7955516	-0.013	0.051	0.001	0.910	0.019	0.001	-0.003	0.650	-0.007	0.096	0.006	0.210	0.007	0.069
rs601339	0.005	0.528	-0.011	0.036	0.030	3×10^{-6}	0.007	0.284	-0.016	0.013	-0.017	0.003	0.004	0.414
rs6488898	0.027	0.050	-0.004	0.642	0.026	0.007	0.016	0.198	-0.023	0.048	-0.007	0.430	0.021	0.005
rs7978610	0.000	0.959	-0.003	0.432	0.032	2×10^{-9}	-0.020	0.001	-0.029	2×10^{-8}	-0.021	3×10^{-6}	0.013	0.002
rs2925979	-0.001	0.853	-0.005	0.236	0.035	1×10^{-19}	0.003	0.630	-0.021	2×10^{-7}	-0.011	0.003	0.001	0.721
rs7200895	0.000	0.966	0.005	0.347	0.006	0.278	-0.002	0.985	0.005	0.720	-0.001	0.850	-0.002	0.697
rs8047711	-0.020	0.329	-0.005	0.649	0.010	0.887	-0.011	0.482	-0.001	0.678	-0.006	0.700	0.000	0.982
rs12929479	0.002	0.764	-0.006	0.141	-0.008	0.530	-0.010	0.088	-0.004	0.380	-0.011	0.013	-0.016	1×10^{-4}
rs731839	-0.007	0.288	-0.011	0.009	0.022	3×10^{-9}	0.002	0.517	-0.022	3×10^{-9}	0.007	0.059	0.007	0.038

Table 2. Standardized Mean Difference (and *P* **values) of Cardiovascular Risk Factors Per Allele of SNPs Used in Mendelian Randomization Analyses**

SNPs within ADIPOQ gene (±50 kb) are identified by an asterisk (*). After Bonferroni correction, only P values lower than 4.2×10⁻⁴ (0.05÷17 SNPs÷7 phenotypes) were considered statistically significant. BMI indicates body mass index; Hb_{A10}, glycohemoglobin; HDL-c, high-density lipoprotein cholesterol; LDL-c, low-density lipoprotein cholesterol; TAG, triacylglycerols; and WC, waist circumference.

(restricted to SNPs within adiponectin locus) to the wellpowered (but vulnerable to horizontal pleiotropy) liberal MR approach (using SNPs across the genome); (3) extensively investigating the presence of bias because of horizontal pleiotropy by using data from other CHD-related phenotypes (eg, glycemic and lipid and anthropometric traits) and methods to account for it (adjusted IVW method and MR-Egger method); (4) testing our hypotheses in 2 data sets (CARDIoGRAM and CARDIoGRAMplusC4D Metabochip); (5) using a very large sample size that provides us with 100% power to detect an odds ratio of 0.80 and 81% to detect and odds ratio of 0.90 with a 0.05% type 1 error rate (Online Table III); (6) checking

Figure 3. Standardized mean difference (and 95% confidence interval [CI]) in cardiovascular risk biomarkers per 1 U increase in genetically instrumented log adiponectin levels. A, Conservative approach including 4 SNPs within *ADIPOQ* gene associated with adiponectin at genome-wide significant levels (*P*<5×10−8; C4). B, Liberal approach including 17 SNPs across the genome associated with adiponectin at genome-wide significant levels (*P<*5×10⁻⁸; L17). BMI indicates body mass index; Hb_{Ate}, glycohemoglobin; HDL-c, highdensity lipoprotein cholesterol; LDL-c, low-density lipoprotein cholesterol; SNP, single-nucleotide polymorphism; TAG, triacylglicerols; and WC, waist circumference.

the consistency of our findings by performing influence metaanalysis and a positive control analysis; and (7) using 2-sample Mendelian randomization to avoid statistical overfitting in comparison to Mendelian randomization where all analyses are conducted in the same participants³⁵ (in a 1-sample setting, results could be biased in the presence of weak instruments because of genetic variants correlating with confounders by chance).

Some limitations of this study should be considered. First, we were not able to test for effect modification by sex, age, or previous disease because of the use of summary data only. In observational studies, the association between adiponectin levels and CHD outcomes is modified by factors such as the type of event (incident versus prevalent)¹⁰ and age of the participant.36 Surprisingly, we did find a positive association between circulating adiponectin and CHD risk in the MR-Egger analysis with CARDIoGRAMplusC4D Metabochip data set, which is likely to be reflecting a false-positive finding because it was generally inconsistent with results from the conservative approach. We aimed to estimate the causal effect of total adiponectin concentrations, but high-molecular-weight adiponectin is thought to be the biologically active fraction, and we are not able to specifically assess its effect. Although we have explored possible violation of the assumptions of Mendelian randomization (Online Table II), we cannot rule out bias because of possible compensatory mechanisms, known as canalization (eg, counter-regulation of adiponectin receptors expression because of variations in blood adiponectin concentration). That said, we are not aware of any evidence that this might be the case.

The 2-sample Mendelian randomization assumes that both samples come from comparable populations. For our discovery analyses, this was the case, whereas in CARDIoGRAMplusC4D Metabochip, although the majority of the participants were of European ancestry (the same as in ADIPOGen), 9% were from other ethnic backgrounds. However, we think it is unlikely that this will have resulted in a major source of bias. First, double genomic control for ethnicity was undertaken in CARDIoGRAMplusC4D Metabochip to control for confounding by population stratification. Second, we found little evidence of heterogeneity in the association of SNPs with CHD in the 2 consortia, which suggests that (strong) effect modification by genomic ancestry is unlikely. Last, in a positive control study, we showed that 2-sample Mendelian randomization produced similar evidence for the expected positive causal effect of LDL-c on CHD.

Adiponectin concentration in the blood ranges from 1 to 30 ng/mL in healthy adults, which is $\approx 10^{3}$ - to 10⁶-folds higher than the concentration of many hormones and cytokines.³⁷ Blood adiponectin concentration is a modifiable risk factor that can be efficiently targeted by lifestyle modifications, mainly weight loss and dietary changes.³⁸ Our results reinforce that Mendelian randomization studies can be helpful in prioritizing potential drug or lifestyle targets, which could substantially reduce the high costs associated with the development and evaluation of large numbers of compounds or lifestyle changes that fail along the development process.

Overall, our findings are not supportive of a protective role of adiponectin in CHD and indicate that the association of genetically increased adiponectin levels and lower risk of CHD is mainly driven by horizontal pleiotropy.

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the data used are publicly available (Online Table I). Those people acknowledged here and who have made their genome-wide data available to scientist may not necessarily agree with comments made in this article, and the authors take full responsibility for the contents of this article.

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Author Contributions: M.C. Borges, D.A. Lawlor, C. de Oliveira, B.L. Horta, and A.J.D. Barros designed the study. M.C. Borges, D.A. Lawlor, J. White, and A.J.D. Barros conceived the analysis plan. M.C. Borges and C. de Oliveira assisted in data acquisition (from public data basis). M.C. Borges performed analyses. M.C. Borges wrote first draft of article. D.A. Lawlor, C. de Oliveira, J. White, B.L. Horta, and A.J.D. Barros were responsible for critical comments and contributions to final writing of article.

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Disclosures

References

- 1. Yamauchi T, Nio Y, Maki T, *et al*. Targeted disruption of AdipoR1 and AdipoR2 causes abrogation of adiponectin binding and metabolic actions. *Nat Med*. 2007;13:332–339. doi: 10.1038/nm1557.
- 2. Ouchi N, Kihara S, Arita Y, *et al*. Adipocyte-derived plasma protein, adiponectin, suppresses lipid accumulation and class A scavenger receptor expression in human monocyte-derived macrophages. *Circulation*. 2001;103:1057–1063.
- 3. Turer AT, Scherer PE. Adiponectin: mechanistic insights and clinical implications. *Diabetologia*. 2012;55:2319–2326. doi: 10.1007/ s00125-012-2598-x.
- 4. Wildman RP, Mancuso P, Wang C, Kim M, Scherer PE, Sowers MR. Adipocytokine and ghrelin levels in relation to cardiovascular disease risk factors in women at midlife: longitudinal associations. *Int J Obes (Lond)*. 2008;32:740–748. doi: 10.1038/sj.ijo.0803782.
- 5. Yamamoto Y, Hirose H, Saito I, Nishikai K, Saruta T. Adiponectin, an adipocyte-derived protein, predicts future insulin resistance: twoyear follow-up study in Japanese population. *J Clin Endocrinol Metab*. 2004;89:87–90. doi: 10.1210/jc.2003-031163.
- 6. Li S, Shin HJ, Ding EL, van Dam RM. Adiponectin levels and risk of type 2 diabetes: a systematic review and meta-analysis. *JAMA*. 2009;302:179– 188. doi: 10.1001/jama.2009.976.
- 7. Zhang H, Mo X, Hao Y, Huang J, Lu X, Cao J, Gu D. Adiponectin levels and risk of coronary heart disease: a meta-analysis of prospective studies. *Am J Med Sci*. 2013;345:455–461. doi: 10.1097/MAJ.0b013e318262dbef.
- 8. Zhang BC, Liu WJ, Che WL, Xu YW. Serum total adiponectin level and risk of cardiovascular disease in Han Chinese populations: a meta-analysis of 17 case-control studies. *Clin Endocrinol (Oxf)*. 2012;77:370–378. doi: 10.1111/j.1365-2265.2011.04260.x.
- 9. Hao G, Li W, Guo R, Yang JG, Wang Y, Tian Y, Liu MY, Peng YG, Wang ZW. Serum total adiponectin level and the risk of cardiovascular disease in general population: a meta-analysis of 17 prospective studies. *Atherosclerosis*. 2013;228:29–35. doi: 10.1016/j.atherosclerosis.2013.02.018.
- 10. Sook Lee E, Park SS, Kim E, Sook Yoon Y, Ahn HY, Park CY, Ho Yun Y, Woo Oh S. Association between adiponectin levels and coronary heart disease and mortality: a systematic review and meta-analysis. *Int J Epidemiol*. 2013;42:1029–1039. doi: 10.1093/ije/dyt087.
- 11. Kanhai DA, Kranendonk ME, Uiterwaal CS, van der Graaf Y, Kappelle LJ, Visseren FL. Adiponectin and incident coronary heart disease and stroke.

A systematic review and meta-analysis of prospective studies. *Obes Rev*. 2013;14:555–567. doi: 10.1111/obr.12027.

- 12. Smith GD, Ebrahim S. 'Mendelian randomization': can genetic epidemiology contribute to understanding environmental determinants of disease? *Int J Epidemiol*. 2003;32:1–22.
- 13. Collaboration CRPCHDG, Wensley F, Gao P et al. Association between c reactive protein and coronary heart disease: Mendelian randomisation analysis based on individual participant data. *BMJ*. 2011;342:d548
- 14. Sarwar N, Butterworth AS, Freitag DF, *et al*.; IL6R Genetics Consortium Emerging Risk Factors Collaboration. Interleukin-6 receptor pathways in coronary heart disease: a collaborative meta-analysis of 82 studies. *Lancet*. 2012;379:1205–1213. doi: 10.1016/S0140-6736(11)61931-4.
- 15. Swerdlow DI, Preiss D, Kuchenbaecker KB, *et al*.; DIAGRAM Consortium; MAGIC Consortium; InterAct Consortium. HMG-coenzyme A reductase inhibition, type 2 diabetes, and bodyweight: evidence from genetic analysis and randomised trials. *Lancet*. 2015;385:351–361. doi: 10.1016/S0140-6736(14)61183-1.
- 16. Dastani Z, Johnson T, Kronenberg F, Nelson CP, Assimes TL, März W, Richards JB; CARDIoGRAM Consortium; ADIPOGen Consortium. The shared allelic architecture of adiponectin levels and coronary artery disease. *Atherosclerosis*. 2013;229:145–148. doi: 10.1016/j.atherosclerosis.2013.03.034.
- 17. Persson J, Strawbridge RJ, McLeod O, *et al*.; IMPROVE Study Group. Sex-Specific Effects of Adiponectin on Carotid Intima-Media Thickness and Incident Cardiovascular Disease. *J Am Heart Assoc*. 2015;4:e001853. doi: 10.1161/JAHA.115.001853.
- 18. Dastani Z, Hivert MF, Timpson N, *et al*.; DIAGRAM+ Consortium; MAGIC Consortium; GLGC Investigators; MuTHER Consortium; DIAGRAM Consortium; GIANT Consortium; Global B Pgen Consortium; Procardis Consortium; MAGIC investigators; GLGC Consortium. Novel loci for adiponectin levels and their influence on type 2 diabetes and metabolic traits: a multi-ethnic meta-analysis of 45,891 individuals. *PLoS Genet*. 2012;8:e1002607. doi: 10.1371/journal.pgen.1002607.
- 19. Schunkert H, König IR, Kathiresan S, *et al*.; Cardiogenics; CARDIoGRAM Consortium. Large-scale association analysis identifies 13 new susceptibility loci for coronary artery disease. *Nat Genet*. 2011;43:333–338. doi: 10.1038/ng.784.
- 20. Deloukas P, Kanoni S, Willenborg C, et al. Large-scale association analysis identifies new risk loci for coronary artery disease. *Nat Genet*. 2013;45:25–33
- 21. Soranzo N, Sanna S, Wheeler E, *et al*.; WTCCC. Common variants at 10 genomic loci influence hemoglobin A_{1c} levels via glycemic and nonglycemic pathways. *Diabetes*. 2010;59:3229–3239. doi: 10.2337/db10-0502.
- 22. Dupuis J, Langenberg C, Prokopenko I, *et al*.; DIAGRAM Consortium; GIANT Consortium; Global BPgen Consortium; Anders Hamsten on behalf of Procardis Consortium; MAGIC investigators. New genetic loci implicated in fasting glucose homeostasis and their impact on type 2 diabetes risk. *Nat Genet*. 2010;42:105–116. doi: 10.1038/ng.520.
- 23. Willer CJ, Schmidt EM, Sengupta S, *et al*.; Global Lipids Genetics Consortium. Discovery and refinement of loci associated with lipid levels. *Nat Genet*. 2013;45:1274–1283. doi: 10.1038/ng.2797.
- 24. Locke AE, Kahali B, Berndt SI, *et al*.; LifeLines Cohort Study; ADIPOGen Consortium; AGEN-BMI Working Group; CARDIOGRAMplusC4D Consortium; CKDGen Consortium; GLGC; ICBP; MAGIC Investigators; MuTHER Consortium; MIGen Consortium; PAGE Consortium; ReproGen Consortium; GENIE Consortium; International Endogene Consortium. Genetic studies of body mass index yield new insights for obesity biology. *Nature*. 2015;518:197–206. doi: 10.1038/nature14177.
- 25. Shungin D, Winkler TW, Croteau-Chonka DC, *et al*.; ADIPOGen Consortium; CARDIOGRAMplusC4D Consortium; CKDGen Consortium; GEFOS Consortium; GENIE Consortium; GLGC; ICBP; International Endogene Consortium; LifeLines Cohort Study; MAGIC Investigators; MuTHER Consortium; PAGE Consortium; ReproGen Consortium. New genetic loci link adipose and insulin biology to body fat distribution. *Nature*. 2015;518:187–196. doi: 10.1038/nature14132.
- 26. Burgess S, Butterworth A, Thompson SG. Mendelian randomization analysis with multiple genetic variants using summarized data. *Genet Epidemiol*. 2013;37:658–665. doi: 10.1002/gepi.21758.
- 27. Burgess S, Dudbridge F, Thompson SG. Re: "Multivariable Mendelian randomization: the use of pleiotropic genetic variants to estimate causal effects". *Am J Epidemiol*. 2015;181:290–291. doi: 10.1093/aje/kwv017.
- 28. Burgess S, Bowden J, Fall T, Ingelsson E, Thompson SG. Sensitivity analyses for robust causal inference from mendelian randomization analyses with multiple genetic variants. *Epidemiology (in press)*. 2016
- 29. Greco M FD, Minelli C, Sheehan NA, Thompson JR. Detecting pleiotropy in Mendelian randomisation studies with summary data and a continuous outcome. *Stat Med*. 2015;34:2926–2940. doi: 10.1002/sim.6522.
- 30. Egger M, Davey Smith G, Schneider M, Minder C. Bias in meta-analysis detected by a simple, graphical test. *BMJ*. 1997;315:629–634.
- 31. Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments: effect estimation and bias detection through Egger regression. *Int J Epidemiol*. 2015;44:512–525. doi: 10.1093/ije/dyv080.
- 32. Gao H, Fall T, van Dam RM, Flyvbjerg A, Zethelius B, Ingelsson E, Hägg S. Evidence of a causal relationship between adiponectin levels and insulin sensitivity: a Mendelian randomization study. *Diabetes*. 2013;62:1338– 1344. doi: 10.2337/db12-0935.
- 33. Mente A, Meyre D, Lanktree MB, Heydarpour M, Davis AD, Miller R, Gerstein H, Hegele RA, Yusuf S, Anand SS; SHARE Investigators; SHARE-AP Investigators. Causal relationship between adiponectin and metabolic traits: a Mendelian randomization study in a multiethnic population. *PLoS One*. 2013;8:e66808. doi: 10.1371/journal.pone.0066808.
- 34. Yaghootkar H, Lamina C, Scott RA, *et al*.; GENESIS Consortium; RISC Consortium. Mendelian randomization studies do not support a causal role

for reduced circulating adiponectin levels in insulin resistance and type 2 diabetes. *Diabetes*. 2013;62:3589–3598. doi: 10.2337/db13-0128.

- 35. Burgess S, Scott RA, Timpson NJ, Davey Smith G, Thompson SG; EPIC- InterAct Consortium. Using published data in Mendelian randomization: a blueprint for efficient identification of causal risk factors. *Eur J Epidemiol*. 2015;30:543–552. doi: 10.1007/s10654- 015-0011-z.
- 36. Wannamethee SG, Welsh P, Whincup PH, Sawar N, Thomas MC, Gudnarsson V, Sattar N. High adiponectin and increased risk of cardiovascular disease and mortality in asymptomatic older men: does NTproBNP help to explain this association? *Eur J Cardiovasc Prev Rehabil*. 2011;18:65–71. doi: 10.1097/HJR.0b013e32833b09d9.
- 37. Arita Y, Kihara S, Ouchi N, *et al*. Paradoxical decrease of an adiposespecific protein, adiponectin, in obesity. *Biochem Biophys Res Commun*. 1999;257:79–83.
- 38. Esposito K, Pontillo A, Di Palo C, Giugliano G, Masella M, Marfella R, Giugliano D. Effect of weight loss and lifestyle changes on vascular inflammatory markers in obese women: a randomized trial. *JAMA*. 2003;289:1799–1804. doi: 10.1001/jama.289.14.1799.

Novelty and Significance

What Is Known?

- • Adiponectin is a protein produced mainly by mature adipose cells.
- Higher circulating adiponectin levels are associated with lower cardiometabolic risk.
- Some genetic variants are associated with both circulating adiponectin and coronary heart disease risk.

What New Information Does This Article Contribute?

- • Our findings do not support a causal effect of circulating adiponectin levels on the risk of coronary heart disease (CHD).
- Genetic variants that are associated with both circulating adiponectin levels and CHD have pleiotropic effects and do not reflect a direct role of circulating adiponectin in CHD development.

Higher circulating adiponectin levels are associated with better cardiometabolic profile; however, it is unknown whether this association is causal or merely correlative because of confounding factors. We used genetic variants associated with circulating adiponectin levels to test whether adiponectin is causally involved in CHD development, a technique known as Mendelian randomization. Overall, our findings do not support a causal effect of adiponectin on CHD risk, indicating that primary perturbation of circulating adiponectin is unlikely to be a major cause of CHD. Interventions targeting total circulating adiponectin might not be appropriate therapeutic strategies for primary CHD prevention.

and Aluísio J.D. Barros Maria Carolina Borges, Debbie A. Lawlor, Cesar de Oliveira, Jon White, Bernardo Lessa Horta **Role of Adiponectin in Coronary Heart Disease Risk: A Mendelian Randomization Study**

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SUPPLEMENTAL MATERIAL

Supplemental Methods

Inverse variance weighted (IVW) method

For unadjusted and adjusted Mendelian randomization analyses, the inverse variance weighted (IVW) method was used to derive the beta coefficient (log odds ratio of CHD per 1 natural log greater adiponectin level) and its standard error by using the following formulas:

$$
\hat{\beta}_{\text{IVW}} = \frac{\sum_{k=1}^{K} X_k Y_k \sigma_{yk}^{-2}}{\sum_{k=1}^{K} X_k^2 \sigma_{yk}^{-2}} \qquad \qquad SE_{\hat{\beta}_{\text{IVW}}} = \sqrt{\frac{1}{\sum_{k=1}^{K} X_k^2 \sigma_{yk}^{-2}}}
$$

Where X_k is the mean change in log adiponectin level per additional effect allele of SNP k and Y_k is the mean change in log odds of CHD per additional effect allele of SNP k with standard error σ_{Yk} .

The IVW method was also used to estimate the effect of adiponectin on cardiovascular risk factors (HbA1c, fasting insulin levels, HDL-c, LDL-c, TAG, BMI, and WC) $(X_k: \text{mean change in log adiponectin level per additional effect allele of })$ SNP k, Y_k : mean change in the risk factor per additional effect allele of SNP k; σ_{Yk} : standard error of Y_k).

To estimate the association of genetically raised adiponectin and CHD in the model adjusted for cardiovascular risk factors, we used betas for SNP-CHD association as the dependent variable, betas for SNP-adiponectin and SNPcardiovascular risk factors as independent variables and inverse variance weights (with no intercept). This method is equivalent to IVW method when there is only one independent variable 1 .

MR-Egger regression method

The Egger regression has been used for almost two decades to detect small study bias (which may be due to publication bias) in meta-analyses of randomized clinical trials 2 . In this method, the ratio of the effect estimate by its standard error is regressed against the estimate's precision (the inverse of the standard error). Bowden et al. 3 recently proposed an adaptation of the Egger regression to test for bias from horizontal pleiotropy in Mendelian randomization studies.

While the IVW estimate is equivalent to the slope of the best fitting line through the observations that pass through the origin, the MR-Egger estimate would be the best fitting line through the observations in a model that allows the intercept to vary. In this method, the intercept will reflect the average pleiotropic effect across genetic variants (e.g. log odds CHD per allele when difference in adiponectin per allele is zero) and the slope coefficient will provide an estimate of the causal effect provided that the InSIDE (Instrument Strength Independent of Direct Effect) assumption holds, which requires that there is no correlation between SNP-exposure association and direct effects of SNP on outcome.

Bootstrapping (10,000 iterations) was used to derive corrected 95% confidence intervals for MR-Egger intercept and slope using the percentile method 3 .

Positive control analysis

The positive control consisted of a Mendelian randomization analysis in which LDL-c was the biomarker of interest and CHD risk was the outcome. 58

SNPs were reported as strongly associated with LDL-c in GLGC consortium 4 . Of there, 38 could be found in both Cardiogram and CARDIoGRAMplusC4D Metabochip dataset and, thus, were used as the instrumental variable for LDL-c. The crude IVW method was used to estimate the association of LDL-c with CHD risk. Since many SNPs were also associated with other lipid traits (ex: HDL-c, TAG and total cholesterol), the MR-Egger method was also used.

Online Tables

Online Table I. Characteristics of the data sources

Online Table I. Characteristics of the data sources (continued)

† Quality control criteria varied across studies; * NR: not reported in the main consortium publication. § CHD was defined as the presence of coronary artery disease or myocardial infarction. Detailed criteria for CHD definition for each study can be found in the Supplementary material of the main publications ⁵⁻⁷. QC: quality control; GWAS: genome-wide association study; CEU: Centre d'Etude du Polymorphisme Humain collected in Utah; CHD: coronary heart disease; MAF: minor allele frequency; SNP: single nucleotide polymorphism; BMI: body mass index; WC: waist circumference; NR: not reported; NA: not applicable. CARDIoGRAM: Coronary ARtery DIsease Genome-wide Replication And Meta-analysis; MAGIC: Meta-Analyses of Glucose and Insulin-Related Traits Consortium; GLGC: Global Lipids Genetics Consortium; GIANT: Genetic Investigation of ANthropometric Traits.

Online Table II. Core instrumental variable assumptions and strategies to address them

IV: instrumental variable; E: exposure; O: outcome; U: unknown confounder; X: other phenotype: G: other genetic variant in LD; LD: linkage disequilibrium; CHD: coronary heart disease. A dashed arrow was used to indicate weak association between IV and E. *Adapted from Vanderweele ⁹.

Data source	Sample size	Proportion of cases	Type-I error rate (α)	Original OR.	Equivalent standardized OR	R_{x-z}^2	Power
CARDIoGRAM	86,995	25.6%	0.05	0.70	0.80	5%	100%
CARDIOGRAM	86,995	25.6%	0.05	0.80	0.87	5%	97%
CARDIOGRAM	86,995	25.6%	0.05	0.90	0.94	5%	42%
CARDIoGRAMplusC4D Metabochip	194,427	32.8%	0.05	0.70	0.80	5%	100%
CARDIoGRAMplusC4D Metabochip	194,427	32.8%	0.05	0.80	0.87	5%	100%
CARDIoGRAMplusC4D Metabochip	194,427	32.8%	0.05	0.90	0.94	5%	81%

Online Table III. Power simulations for the Mendelian randomization analyses

OR: Assumed true odds ratio of CHD risk per standard deviation of the exposure variable

Conversion of original (per log adiponectin) to equivalent standardized OR (per standard unit of log adiponectin) was made using an external source of individual level data (1982 Pelotas Birth Cohort) with similar adiponectin distribution (adiponectin levels in ADIPOGen consortium: mean = 9.8 µg/ml (SD = 5.6); adiponectin levels in 1982 Pelotas Birth Cohort: mean = 9.3 μ g/ml (SD = 5.7)).

 R^2_{x-z} : proportion of variance explained for the association between the genetic instrument (Z) and adiponectin levels (X). Values approximate findings from Dastani et al 10 and Yaghootkar et al 11 .

Calculations were performed in the power calculator for Mendelian Randomization studies, available at [http://cnsgenomics.com/shiny/mRnd/,](http://cnsgenomics.com/shiny/mRnd/) based on the publication by Brion et al 2013¹².

Online Figures

Online Figure I. Graphical illustration of scenarios of (A) vertical pleiotropy (mediation) and (B) horizontal pleiotropy (bias) by CHD risk factors in the relation among SNPs, adiponectin levels and CHD risk. CHD: coronary heart disease; SNP: single nucleotide polymorphism.

Online Figure II. Meta-analysis and heterogeneity analysis of Mendelian randomization estimates of each SNP for the association of blood adiponectin levels with CHD risk. CHD: coronary heart disease; SNP: single nucleotide polymorphism, Chr: chromosome.

Online Figure III. Pooled odds ratio (and 95%CI) of coronary heart disease risk per unit increase in log adiponectin levels omitting one SNP at a time (influence metaanalysis) estimated by the IVW method and by the MR-Egger method. CHD: coronary heart disease; IVW: inverse-variance weighted method; MR-Egger: Mendelian randomization-Egger method; OR: odds ratio; SNP: single nucleotide polymorphism.

Online Figure IV. Log odds ratio of coronary heart disease and mean increase in log adiponectin levels per adiponectin raising allele in CARDIoGRAM and CARDIoGRAMplusC4D Metabochip. Each data point represents betas for SNP-log OR CHD (Y axis) and SNP-adiponectin (X axis) association (N = 17 SNPs). CHD: coronary heart disease; IVW: inverse-variance weighted method; MR-Egger: Mendelian randomization-Egger method; OR: odds ratio; SNP: single nucleotide polymorphism

Online Figure I.

Online Figure II.

Online Figure III.

Online Figure IV.

Suplemental References

- 1. Burgess S, Dudbridge F, Thompson SG. Re: "Multivariable mendelian randomization: The use of pleiotropic genetic variants to estimate causal effects". *Am. J. Epidemiol.* 2015;181:290-291
- 2. Egger M, Davey Smith G, Schneider M, Minder C. Bias in meta-analysis detected by a simple, graphical test. *BMJ*. 1997;315:629-634
- 3. Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments: Effect estimation and bias detection through egger regression. *Int. J. Epidemiol.* 2015;44:512-525
- 4. Global Lipids Genetics C, Willer CJ, Schmidt EM, Sengupta S, Peloso GM, Gustafsson S, Kanoni S, Ganna A, Chen J, Buchkovich ML, Mora S, Beckmann JS, Bragg-Gresham JL, Chang HY, Demirkan A, Den Hertog HM, Do R, Donnelly LA, Ehret GB, Esko T, Feitosa MF, Ferreira T, Fischer K, Fontanillas P, Fraser RM, Freitag DF, Gurdasani D, Heikkila K, Hypponen E, Isaacs A, Jackson AU, Johansson A, Johnson T, Kaakinen M, Kettunen J, Kleber ME, Li X, Luan J, Lyytikainen LP, Magnusson PK, Mangino M, Mihailov E, Montasser ME, Muller-Nurasyid M, Nolte IM, O'Connell JR, Palmer CD, Perola M, Petersen AK, Sanna S, Saxena R, Service SK, Shah S, Shungin D, Sidore C, Song C, Strawbridge RJ, Surakka I, Tanaka T, Teslovich TM, Thorleifsson G, Van den Herik EG, Voight BF, Volcik KA, Waite LL, Wong A, Wu Y, Zhang W, Absher D, Asiki G, Barroso I, Been LF, Bolton JL, Bonnycastle LL, Brambilla P, Burnett MS, Cesana G, Dimitriou M, Doney AS, Doring A, Elliott P, Epstein SE, Eyjolfsson GI, Gigante B, Goodarzi MO, Grallert H, Gravito ML, Groves CJ, Hallmans G, Hartikainen AL, Hayward C, Hernandez D, Hicks AA, Holm H, Hung YJ, Illig T, Jones MR, Kaleebu P, Kastelein JJ, Khaw KT, Kim E, Klopp N, Komulainen P, Kumari M, Langenberg C, Lehtimaki T, Lin SY, Lindstrom J, Loos RJ, Mach F, McArdle WL, Meisinger C, Mitchell BD, Muller G, Nagaraja R, Narisu N, Nieminen TV, Nsubuga RN, Olafsson I, Ong KK, Palotie A, Papamarkou T, Pomilla C, Pouta A, Rader DJ, Reilly MP, Ridker PM, Rivadeneira F, Rudan I, Ruokonen A, Samani N, Scharnagl H, Seeley J, Silander K, Stancakova A, Stirrups K, Swift AJ, Tiret L, Uitterlinden AG, van Pelt LJ, Vedantam S, Wainwright N, Wijmenga C, Wild SH, Willemsen G, Wilsgaard T, Wilson JF, Young EH, Zhao JH, Adair LS, Arveiler D, Assimes TL, Bandinelli S, Bennett F, Bochud M, Boehm BO, Boomsma DI, Borecki IB, Bornstein SR, Bovet P, Burnier M, Campbell H, Chakravarti A, Chambers JC, Chen YD, Collins FS, Cooper RS, Danesh J, Dedoussis G, de Faire U, Feranil AB, Ferrieres J, Ferrucci L, Freimer NB, Gieger C, Groop LC, Gudnason V, Gyllensten U, Hamsten A, Harris TB, Hingorani A, Hirschhorn JN, Hofman A, Hovingh GK, Hsiung CA, Humphries SE, Hunt SC, Hveem K, Iribarren C, Jarvelin MR, Jula A, Kahonen M, Kaprio J, Kesaniemi A, Kivimaki M, Kooner JS, Koudstaal PJ, Krauss RM, Kuh D, Kuusisto J, Kyvik KO, Laakso M, Lakka TA, Lind L, Lindgren CM, Martin NG, Marz W, McCarthy MI, McKenzie CA, Meneton P, Metspalu A, Moilanen L, Morris AD, Munroe PB, Njolstad I, Pedersen NL, Power C, Pramstaller PP, Price JF, Psaty BM, Quertermous T, Rauramaa R, Saleheen D, Salomaa V, Sanghera DK, Saramies J, Schwarz PE, Sheu WH, Shuldiner AR, Siegbahn A, Spector TD, Stefansson K, Strachan DP, Tayo BO, Tremoli E, Tuomilehto J, Uusitupa M, van Duijn CM, Vollenweider P, Wallentin L, Wareham NJ, Whitfield JB, Wolffenbuttel BH, Ordovas JM, Boerwinkle E, Palmer CN, Thorsteinsdottir U, Chasman DI, Rotter JI, Franks PW, Ripatti S, Cupples LA, Sandhu MS, Rich SS, Boehnke M, Deloukas P, Kathiresan S, Mohlke KL, Ingelsson E, Abecasis GR. Discovery and refinement of loci associated with lipid levels. *Nat. Genet.* 2013;45:1274-1283 5. Preuss M, Koenig IR, Thompson JR, Erdmann J, Absher D, Assimes TL, Blankenberg S, Boerwinkle E, Chen L, Cupples LA, Hall AS, Halperin E, Hengstenberg C, Holm H, Laaksonen R, Li M, Maerz W, McPherson R, Musunuru K, Nelson CP, Burnett MS, Epstein SE, O'Donnell CJ, Quertermous T, Rader DJ,

Roberts R, Schillert A, Stefansson K, Stewart AFR, Thorleifsson G, Voight BF, Wells GA, Ziegler A, Kathiresan S, Reilly MP, Samani NJ, Schunkert H, Consortium CA. Design of the coronary artery disease genome-wide replication and meta-analysis (cardiogram) study a genome-wide association meta-analysis involving more than 22 000 cases and 60 000 controls. *Circulation-Cardiovascular Genetics*. 2010;3:475-U186

- 6. Deloukas P, Kanoni S, Willenborg C, Farrall M, Assimes TL, Thompson JR, Ingelsson E, Saleheen D, Erdmann J, Goldstein BA, Stirrups K, König IR, Cazier JB, Johansson A, Hall AS, Lee JY, Willer CJ, Chambers JC, Esko T, Folkersen L, Goel A, Grundberg E, Havulinna AS, Ho WK, Hopewell JC, Eriksson N, Kleber ME, Kristiansson K, Lundmark P, Lyytikäinen LP, Rafelt S, Shungin D, Strawbridge RJ, Thorleifsson G, Tikkanen E, Van Zuydam N, Voight BF, Waite LL, Zhang W, Ziegler A, Absher D, Altshuler D, Balmforth AJ, Barroso I, Braund PS, Burgdorf C, Claudi-Boehm S, Cox D, Dimitriou M, Do R, Doney AS, El Mokhtari N, Eriksson P, Fischer K, Fontanillas P, Franco-Cereceda A, Gigante B, Groop L, Gustafsson S, Hager J, Hallmans G, Han BG, Hunt SE, Kang HM, Illig T, Kessler T, Knowles JW, Kolovou G, Kuusisto J, Langenberg C, Langford C, Leander K, Lokki ML, Lundmark A, McCarthy MI, Meisinger C, Melander O, Mihailov E, Maouche S, Morris AD, Müller-Nurasyid M, Nikus K, Peden JF, Rayner NW, Rasheed A, Rosinger S, Rubin D, Rumpf MP, Schäfer A, Sivananthan M, Song C, Stewart AF, Tan ST, Thorgeirsson G, van der Schoot CE, Wagner PJ, Wells GA, Wild PS, Yang TP, Amouyel P, Arveiler D, Basart H, Boehnke M, Boerwinkle E, Brambilla P, Cambien F, Cupples AL, de Faire U, Dehghan A, Diemert P, Epstein SE, Evans A, Ferrario MM, Ferrières J, Gauguier D, Go AS, Goodall AH, Gudnason V, Hazen SL, Holm H, Iribarren C, Jang Y, Kähönen M, Kee F, Kim HS, Klopp N, Koenig W, Kratzer W, Kuulasmaa K, Laakso M, Laaksonen R, Lind L, Ouwehand WH, Parish S, Park JE, Pedersen NL, Peters A, Quertermous T, Rader DJ, Salomaa V, Schadt E, Shah SH, Sinisalo J, Stark K, Stefansson K, Trégouët DA, Virtamo J, Wallentin L, Wareham N, Zimmermann ME, Nieminen MS, Hengstenberg C, Sandhu MS, Pastinen T, Syvänen AC, Hovingh GK, Dedoussis G, Franks PW, Lehtimäki T, Metspalu A, Zalloua PA, Siegbahn A, Schreiber S, Ripatti S, Blankenberg SS, Perola M, Clarke R, Boehm BO, O'Donnell C, Reilly MP, März W, Collins R, Kathiresan S, Hamsten A, Kooner JS, Thorsteinsdottir U, Danesh J, Palmer CN, Roberts R, Watkins H, Schunkert H, Samani NJ, Consortium CD, Consortium D, Consortium C, Consortium M, Consortium WTCC. Large-scale association analysis identifies new risk loci for coronary artery disease. *Nat Genet*. 2013;45:25-33
- 7. Schunkert H, König IR, Kathiresan S, Reilly MP, Assimes TL, Holm H, Preuss M, Stewart AF, Barbalic M, Gieger C, Absher D, Aherrahrou Z, Allayee H, Altshuler D, Anand SS, Andersen K, Anderson JL, Ardissino D, Ball SG, Balmforth AJ, Barnes TA, Becker DM, Becker LC, Berger K, Bis JC, Boekholdt SM, Boerwinkle E, Braund PS, Brown MJ, Burnett MS, Buysschaert I, Carlquist JF, Chen L, Cichon S, Codd V, Davies RW, Dedoussis G, Dehghan A, Demissie S, Devaney JM, Diemert P, Do R, Doering A, Eifert S, Mokhtari NE, Ellis SG, Elosua R, Engert JC, Epstein SE, de Faire U, Fischer M, Folsom AR, Freyer J, Gigante B, Girelli D, Gretarsdottir S, Gudnason V, Gulcher JR, Halperin E, Hammond N, Hazen SL, Hofman A, Horne BD, Illig T, Iribarren C, Jones GT, Jukema JW, Kaiser MA, Kaplan LM, Kastelein JJ, Khaw KT, Knowles JW, Kolovou G, Kong A, Laaksonen R, Lambrechts D, Leander K, Lettre G, Li M, Lieb W, Loley C, Lotery AJ, Mannucci PM, Maouche S, Martinelli N, McKeown PP, Meisinger C, Meitinger T, Melander O, Merlini PA, Mooser V, Morgan T, Mühleisen TW, Muhlestein JB, Münzel T, Musunuru K, Nahrstaedt J, Nelson CP, Nöthen MM, Olivieri O, Patel RS, Patterson CC, Peters A, Peyvandi F, Qu L, Quyyumi AA, Rader DJ, Rallidis LS, Rice C, Rosendaal FR, Rubin D, Salomaa V, Sampietro ML, Sandhu MS, Schadt E, Schäfer A, Schillert A, Schreiber S, Schrezenmeir J, Schwartz SM, Siscovick

DS, Sivananthan M, Sivapalaratnam S, Smith A, Smith TB, Snoep JD, Soranzo N, Spertus JA, Stark K, Stirrups K, Stoll M, Tang WH, Tennstedt S, Thorgeirsson G, Thorleifsson G, Tomaszewski M, Uitterlinden AG, van Rij AM, Voight BF, Wareham NJ, Wells GA, Wichmann HE, Wild PS, Willenborg C, Witteman JC, Wright BJ, Ye S, Zeller T, Ziegler A, Cambien F, Goodall AH, Cupples LA, Quertermous T, März W, Hengstenberg C, Blankenberg S, Ouwehand WH, Hall AS, Deloukas P, Thompson JR, Stefansson K, Roberts R, Thorsteinsdottir U, O'Donnell CJ, McPherson R, Erdmann J, Samani NJ, Cardiogenics, Consortium C. Large-scale association analysis identifies 13 new susceptibility loci for coronary artery disease. *Nat Genet*. 2011;43:333-338

8. Smith GD, Lawlor DA, Harbord R, Timpson N, Day I, Ebrahim S. Clustered environments and randomized genes: A fundamental distinction between conventional and genetic epidemiology. *PLoS Med.* 2007;4:e352

9. VanderWeele TJ, Tchetgen Tchetgen EJ, Cornelis M, Kraft P. Methodological challenges in mendelian randomization. *Epidemiology*. 2014;25:427-435

10. Dastani Z, Hivert MF, Timpson N, Perry JR, Yuan X, Scott RA, Henneman P, Heid IM, Kizer JR, Lyytikäinen LP, Fuchsberger C, Tanaka T, Morris AP, Small K, Isaacs A, Beekman M, Coassin S, Lohman K, Qi L, Kanoni S, Pankow JS, Uh HW, Wu Y, Bidulescu A, Rasmussen-Torvik LJ, Greenwood CM, Ladouceur M, Grimsby J, Manning AK, Liu CT, Kooner J, Mooser VE, Vollenweider P, Kapur KA, Chambers J, Wareham NJ, Langenberg C, Frants R, Willems-Vandijk K, Oostra BA, Willems SM, Lamina C, Winkler TW, Psaty BM, Tracy RP, Brody J, Chen I, Viikari J, Kähönen M, Pramstaller PP, Evans DM, St Pourcain B, Sattar N, Wood AR, Bandinelli S, Carlson OD, Egan JM, Böhringer S, van Heemst D, Kedenko L, Kristiansson K, Nuotio ML, Loo BM, Harris T, Garcia M, Kanaya A, Haun M, Klopp N, Wichmann HE, Deloukas P, Katsareli E, Couper DJ, Duncan BB, Kloppenburg M, Adair LS, Borja JB, Wilson JG, Musani S, Guo X, Johnson T, Semple R, Teslovich TM, Allison MA, Redline S, Buxbaum SG, Mohlke KL, Meulenbelt I, Ballantyne CM, Dedoussis GV, Hu FB, Liu Y, Paulweber B, Spector TD, Slagboom PE, Ferrucci L, Jula A, Perola M, Raitakari O, Florez JC, Salomaa V, Eriksson JG, Frayling TM, Hicks AA, Lehtimäki T, Smith GD, Siscovick DS, Kronenberg F, van Duijn C, Loos RJ, Waterworth DM, Meigs JB, Dupuis J, Richards JB, Voight BF, Scott LJ, Steinthorsdottir V, Dina C, Welch RP, Zeggini E, Huth C, Aulchenko YS, Thorleifsson G, McCulloch LJ, Ferreira T, Grallert H, Amin N, Wu G, Willer CJ, Raychaudhuri S, McCarroll SA, Hofmann OM, Segrè AV, van Hoek M, Navarro P, Ardlie K, Balkau B, Benediktsson R, Bennett AJ, Blagieva R, Boerwinkle E, Bonnycastle LL, Boström KB, Bravenboer B, Bumpstead S, Burtt NP, Charpentier G, Chines PS, Cornelis M, Crawford G, Doney AS, Elliott KS, Elliott AL, Erdos MR, Fox CS, Franklin CS, Ganser M, Gieger C, Grarup N, Green T, Griffin S, Groves CJ, Guiducci C, Hadjadj S, Hassanali N, Herder C, Isomaa B, Jackson AU, Johnson PR, Jørgensen T, Kao WH, Kong A, Kraft P, Kuusisto J, Lauritzen T, Li M, Lieverse A, Lindgren CM, Lyssenko V, Marre M, Meitinger T, Midthjell K, Morken MA, Narisu N, Nilsson P, Owen KR, Payne F, Petersen AK, Platou C, Proença C, Prokopenko I, Rathmann W, Rayner NW, Robertson NR, Rocheleau G, Roden M, Sampson MJ, Saxena R, Shields BM, Shrader P, Sigurdsson G, Sparsø T, Strassburger K, Stringham HM, Sun Q, Swift AJ, Thorand B, Tichet J, Tuomi T, van Dam RM, van Haeften TW, van Herpt T, van Vliet-Ostaptchouk JV, Walters GB, Weedon MN, Wijmenga C, Witteman J, Bergman RN, Cauchi S, Collins FS, Gloyn AL, Gyllensten U, Hansen T, Hide WA, Hitman GA, Hofman A, Hunter DJ, Hveem K, Laakso M, Morris AD, Palmer CN, Rudan I, Sijbrands E, Stein LD, Tuomilehto J, Uitterlinden A, Walker M, Watanabe RM, Abecasis GR, Boehm BO, Campbell H, Daly MJ, Hattersley AT, Pedersen O, Barroso I, Groop L, Sladek R, Thorsteinsdottir U, Wilson JF, Illig T, Froguel P, van Duijn CM, Stefansson K, Altshuler D, Boehnke M, McCarthy MI, Soranzo N, Wheeler E, Glazer NL, Bouatia-Naji N, Mägi R, Randall J, Elliott P, Rybin D, Dehghan A,

Hottenga JJ, Song K, Goel A, Lajunen T, Doney A, Cavalcanti-Proença C, Kumari M, Timpson NJ, Zabena C, Ingelsson E, An P, O'Connell J, Luan J, Elliott A, Roccasecca RM, Pattou F, Sethupathy P, Ariyurek Y, Barter P, Beilby JP, Ben-Shlomo Y, Bergmann S, Bochud M, Bonnefond A, Borch-Johnsen K, Böttcher Y, Brunner E, Bumpstead SJ, Chen YD, Chines P, Clarke R, Coin LJ, Cooper MN, Crisponi L, Day IN, de Geus EJ, Delplanque J, Fedson AC, Fischer-Rosinsky A, Forouhi NG, Franzosi MG, Galan P, Goodarzi MO, Graessler J, Grundy S, Gwilliam R, Hallmans G, Hammond N, Han X, Hartikainen AL, Hayward C, Heath SC, Hercberg S, Hillman DR, Hingorani AD, Hui J, Hung J, Kaakinen M, Kaprio J, Kesaniemi YA, Kivimaki M, Knight B, Koskinen S, Kovacs P, Kyvik KO, Lathrop GM, Lawlor DA, Le Bacquer O, Lecoeur C, Li Y, Mahley R, Mangino M, Martínez-Larrad MT, McAteer JB, McPherson R, Meisinger C, Melzer D, Meyre D, Mitchell BD, Mukherjee S, Naitza S, Neville MJ, Orrù M, Pakyz R, Paolisso G, Pattaro C, Pearson D, Peden JF, Pedersen NL, Pfeiffer AF, Pichler I, Polasek O, Posthuma D, Potter SC, Pouta A, Province MA, Rice K, Ripatti S, Rivadeneira F, Rolandsson O, Sandbaek A, Sandhu M, Sanna S, Sayer AA, Scheet P, Seedorf U, Sharp SJ, Shields B, Sigurðsson G, Sijbrands EJ, Silveira A, Simpson L, Singleton A, Smith NL, Sovio U, Swift A, Syddall H, Syvänen AC, Tönjes A, Uitterlinden AG, van Dijk KW, Varma D, Visvikis-Siest S, Vitart V, Vogelzangs N, Waeber G, Wagner PJ, Walley A, Ward KL, Watkins H, Wild SH, Willemsen G, Witteman JC, Yarnell JW, Zelenika D, Zethelius B, Zhai G, Zhao JH, Zillikens MC, Borecki IB, Meneton P, Magnusson PK, Nathan DM, Williams GH, Silander K, Bornstein SR, Schwarz P, Spranger J, Karpe F, Shuldiner AR, Cooper C, Serrano-Ríos M, Lind L, Palmer LJ, Franks PW, Ebrahim S, Marmot M, Wright AF, Stumvoll M, Hamsten A, Buchanan TA, Valle TT, Rotter JI, Penninx BW, Boomsma DI, Cao A, Scuteri A, Schlessinger D, Uda M, Ruokonen A, Jarvelin MR, Peltonen L, Mooser V, Musunuru K, Smith AV, Edmondson AC, Stylianou IM, Koseki M, Pirruccello JP, Chasman DI, Johansen CT, Fouchier SW, Peloso GM, Barbalic M, Ricketts SL, Bis JC, Feitosa MF, Orho-Melander M, Melander O, Li X, Cho YS, Go MJ, Kim YJ, Lee JY, Park T, Kim K, Sim X, Ong RT, Croteau-Chonka DC, Lange LA, Smith JD, Ziegler A, Zhang W, Zee RY, Whitfield JB, Thompson JR, Surakka I, Smit JH, Sinisalo J, Scott J, Saharinen J, Sabatti C, Rose LM, Roberts R, Rieder M, Parker AN, Pare G, O'Donnell CJ, Nieminen MS, Nickerson DA, Montgomery GW, McArdle W, Masson D, Martin NG, Marroni F, Lucas G, Luben R, Lokki ML, Lettre G, Launer LJ, Lakatta EG, Laaksonen R, König IR, Khaw KT, Kaplan LM, Johansson Å, Janssens AC, Igl W, Hovingh GK, Hengstenberg C, Havulinna AS, Hastie ND, Harris TB, Haritunians T, Hall AS, Groop LC, Gonzalez E, Freimer NB, Erdmann J, Ejebe KG, Döring A, Dominiczak AF, Demissie S, de Faire U, Caulfield MJ, Boekholdt SM, Assimes TL, Quertermous T, Seielstad M, Wong TY, Tai ES, Feranil AB, Kuzawa CW, Taylor HA, Gabriel SB, Holm H, Gudnason V, Krauss RM, Ordovas JM, Munroe PB, Kooner JS, Tall AR, Hegele RA, Kastelein JJ, Schadt EE, Strachan DP, Reilly MP, Samani NJ, Schunkert H, Cupples LA, Sandhu MS, Ridker PM, Rader DJ, Kathiresan S, Consortium D, Consortium M, Investigators G, Consortium M, Consortium D, Consortium G, Consortium GBP, Consortium P, investigators M, Consortium G. Novel loci for adiponectin levels and their influence on type 2 diabetes and metabolic traits: A multi-ethnic metaanalysis of 45,891 individuals. *PLoS Genet*. 2012;8:e1002607

11. Yaghootkar H, Lamina C, Scott RA, Dastani Z, Hivert MF, Warren LL, Stancáková A, Buxbaum SG, Lyytikäinen LP, Henneman P, Wu Y, Cheung CY, Pankow JS, Jackson AU, Gustafsson S, Zhao JH, Ballantyne CM, Xie W, Bergman RN, Boehnke M, el Bouazzaoui F, Collins FS, Dunn SH, Dupuis J, Forouhi NG, Gillson C, Hattersley AT, Hong J, Kähönen M, Kuusisto J, Kedenko L, Kronenberg F, Doria A, Assimes TL, Ferrannini E, Hansen T, Hao K, Häring H, Knowles JW, Lindgren CM, Nolan JJ, Paananen J, Pedersen O, Quertermous T, Smith U, Lehtimäki T, Liu CT, Loos RJ, McCarthy MI, Morris AD, Vasan RS, Spector TD,

Teslovich TM, Tuomilehto J, van Dijk KW, Viikari JS, Zhu N, Langenberg C, Ingelsson E, Semple RK, Sinaiko AR, Palmer CN, Walker M, Lam KS, Paulweber B, Mohlke KL, van Duijn C, Raitakari OT, Bidulescu A, Wareham NJ, Laakso M, Waterworth DM, Lawlor DA, Meigs JB, Richards JB, Frayling TM, Consortium G, Consortium R. Mendelian randomization studies do not support a causal role for reduced circulating adiponectin levels in insulin resistance and type 2 diabetes. *Diabetes*. 2013;62:3589-3598

12. Brion MJ, Shakhbazov K, Visscher PM. Calculating statistical power in mendelian randomization studies. *Int J Epidemiol*. 2013;42:1497-1501