A preliminary psychometric evaluation of Music in Dementia Assessment Scales (MiDAS)

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Abstract

Background

Music in Dementia Assessment Scales (MiDAS), an observational outcome measure for music therapy with people with moderate to severe dementia, was developed from qualitative data of focus groups and interviews. Expert and peer consultations were conducted at each stage of the scale development to maximize its content validity. This study aimed to evaluate the psychometric properties of MiDAS.

Methods

Care home residents with dementia attended weekly group music therapy for up to 10 sessions. Music therapists and care home staff were requested to complete weekly MiDAS ratings. The Quality of life scale (QoL-AD) was completed at three-time points.

Results

A total of 629 (Staff=306, Therapist=323) MiDAS forms were completed. The statistical analysis revealed that MiDAS has high therapist inter-rater reliability, low staff inter-rater reliability, adequate staff test-retest reliability, adequate concurrent validity and good construct validity. High factor loadings between the five MiDAS Visual Analogue Scale items: levels of *Interest*, *Response*, *Initiation*, *Involvement* and *Enjoyment* were found.

Conclusion

This study indicates that MiDAS has good psychometric properties despite the small sample size. Future research with a larger sample size could provide a more in depth psychometric evaluation including further exploration of underlying factors. MiDAS provides a measure of engagement with the musical experience and offers insight into who is likely to benefit on other outcomes such as quality of life or reduction in psychiatric symptoms.

Keywords

Running title

Evaluation of the psychometric properties of MiDAS

Introduction

Music-based interventions are popular in clinical practice with people with dementia. Systematic reviews on psychosocial treatments of behavior and psychological symptoms in dementia have found sufficient evidence to support the effectiveness of individually tailored music interventions (O'Connor *et al.*, 2009a; 2009b). A number of systematic reviews on music therapy have also found consistent evidence for short-term reductions in behavioral disturbance and improved mood in people with dementia (e.g. Kodger *et al*., 1999; Vink *et al*., 2013; Ueda *et al*., 2013). On the other hand, music therapy studies have been criticized for the lack of evidence of long-term benefits (Livingston *et al.*, 2005) and for the weaknesses in study designs (Vink *et al*., 2013; McDermott *et al*., 2013). All the reviews agree that the potential benefits of music-based interventions for the wellbeing of people with dementia are promising but the effects of the interventions require further investigations.

There is also concern whether the outcome measures used in the current music therapy studies are most appropriate to evaluate the intervention. Our recent systematic review (McDermott *et al*., 2013) found that the Neuropsychiatric Inventory (NPI) (Cummings *et al*., 1994), the Cohen-Mansfield Agitation Inventory (CMAI) (Cohen-Mansfield, 1986) and various depression scales are most frequently used to evaluate the impact of music therapy on behavioral and psychological symptoms. On the other hand, there is consistent anecdotal clinical evidence that people with dementia benefit

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from 'positive experiences' of music therapy, for instance opportunities for self-expression, increased self-awareness and social interactions with fellow participants. This raises a question if music therapy should only be seen as a 'treatment' to fix neuropsychiatric symptoms.

Music therapy is particularly disadvantaged by the lack of client specific, clinically appropriate and psychometrically evaluated outcome measures. York (1994) developed the Residual Music Skills Test for people with probable Alzheimer's Disease. A test-retest showed a high correlation ($r=0.916$, $p<.000$) (York, 2000). However, testing the musical skills of clients may not be the most appropriate outcome measure to evaluate the intervention. Similarly, an assessment tool to evaluate behavioral responses of people while listening to 'taped music' (Glynn, 1992) is not the most appropriate measure to evaluate the benefits of music therapy for people with dementia, even though the study reported a high inter-rater reliability ($r=0.970-0.999$). In some studies (e.g. Ashida, 2000; Brotons, 2003), researchers used their owns scales but these were not rigorously developed or evaluated for reliability and validity.

Music in Dementia Assessment Scales (MiDAS) (McDermott *et al*., unpublished work), an observational outcome measure for music therapy with people with moderate to severe dementia, was developed from qualitative data of focus groups and interviews conducted with people with dementia, family carers, care home staff and music therapists. Explicit qualitative data analyses methods utilizing the general inductive approach (Thomas, 2006) and the long-table approach (Krueger and Casey, 2000) ensured the rigor and transparency of the scale development. Expert and peer consultations were conducted at each stage of the scale development to maximize the content validity of MiDAS. The pilot MiDAS was field-tested by a music therapist and care home staff for six weeks. Feedback from the therapist and staff and further comments from external music therapists was collated and the revised MiDAS form was produced.

The use of qualitative data for the MiDAS development was to ensure its clinical relevance. The aim of this study was to evaluate the psychometric properties of MiDAS.

Methods

Design

Care home residents with moderate to severe dementia attended weekly group music therapy. Music therapists and care home staff were requested to complete weekly MiDAS ratings. QoL-AD (Quality of Life in Alzheimer's disease, Logsdon *et al*., 1999) was completed at baseline, mid-treatment and end-treatment by care home staff. Statistical analysis was conducted with SPSS version 21. Inter-rater reliability, test-retest reliability, internal consistency, concurrent validity and construct validity of MiDAS were evaluated.

Music in Dementia Assessment Scales (MiDAS)

The MiDAS consists of five Visual Analogue Scales (VAS) items: *Interest*, *Response*, *Initiation*, *Involvement* and *Enjoyment*. Each VAS comprises a 100mm line without intervals, with the two extremes of the scale labeled as 'None at all' and 'Highest'. Examples of behavior and mood of the person that a rater should look for before scoring each VAS are provided on the MiDAS form (Table 1). The MiDAS form also includes a supplementary checklist of six notable reactions from the person (*agitation/aggression*, *withdrawn/low in mood*, *restless/anxious*, *relaxed mood*, *attentive/interested*, and *cheerful/smiling*) and a space for a rater's comment to provide further information when necessary. The checklist of other major reactions should be used to aid clinical interpretations of the VAS items when MiDAS is used for routine clinical evaluation or in a qualitative study. The checklist items are independent from the MiDAS scores. The maximum score for each VAS is 100, thus the maximum MiDAS score per rating is 500.

MiDAS rating procedure

Music therapists and care home staff were requested to complete weekly MiDAS ratings. Before a resident attended his/her music therapy, a staff member with sufficient knowledge of the resident was asked to complete a *Before* rating based on the average presentation of the resident on the day. An *After* form was completed by the same staff member a few hours after the resident's music therapy on the same day before staff finished their shift work. The knowledge about individual residents and the experiences of working in Home A and B varied greatly amongst staff members. Permanent staff members, particularly registered nurses and health care assistants who provided regular daily care to the residents were sought as MiDAS raters whenever possible. Music therapists were asked to complete a *Beginning* form and a *During* form immediately after each session. The *Beginning* rating was based on the therapist's observation of the resident during the first five minutes of the session, whilst *During* rating was based on the observation of the most clinically significant five minutes ('the best five minutes') of that session. It was left to the therapist's clinical judgment to determine what should be considered as 'the most clinically significant five minutes' of that session, since this would vary depending on each participant's clinical goals. For some, this may mean an increased awareness of the therapist and a longer two-way musical interaction. For others, the 'best five minutes' may imply willingness to collaborate with fellow group members and share their ideas. Thus, four MiDAS ratings were to be completed per resident per session in order to assess possible changes in their mood and behavior following music therapy.

Study Sample

Residents who met DSM-IV criteria for dementia (American Psychiatric Association, 1994) in two NHS care homes (Home A and Home B) were recruited for weekly music therapy group for up to 10 sessions. Residents with dementia who also had other psychiatric conditions (e.g. schizophrenia) were included. The music therapy group consisted of between 4 and 6 residents, and each session was guided by the clinical judgments of the two music therapists (RB and MR) with over ten years of experience with this client group. Both therapists used an active music therapy model where clients

are encouraged to explore musical instruments and/or their voice freely. The focus of the sessions included encouragement of musical and verbal interactions between the group members as well as supporting the development of self-awareness of an individual. The use of pre-composed songs combined with exploratory improvisation was frequently used. Ethics approval from the National Research Ethics Service (NRES Committee London East, REC reference 11/LO/0596) was granted in June 2011.

Reliability

Inter-rater reliability examines the degree of objectivity by correlating two sets of ratings by two raters completed at any given time. To assess therapist inter-rater reliability, OM attended RB's music therapy group in Home B as a participant observer for six weeks. OM and RB completed therapist MiDAS ratings independently immediately after each session. For staff inter-rater reliability, an associate health practitioner and an activity co-ordinator in Home A were assigned as staff inter-raters by the unit manager. They were encouraged to rate the residents independently on the day of their music therapy as regularly as their work patterns permitted.

Test-retest reliability measures the repeatability of the test and involves administrating the test twice to the same population group. Staff test-retest reliability was evaluated through permanent members of Home B staff completing before ratings twice on the same residents one week apart on eight separate occasions.

Intraclass Correlation Coefficients (ICC) Case 2 two-way random effects model with absolute agreement (McGraw and Wong, 1996) was utilized, and single scores were quoted for inter-rater reliability and test-retest reliability results. This model was chosen on the basis that the use of two-way random effects ANOVA and single scores allowed random subjects being assessed by random judges at any single occasion, thus maximizing the generalizability of MiDAS.

Internal consistency measures the degree of correlations between the items on the same test and was examined using Cronbach's alpha.

Validity

Face and content validity has already been established during the MiDAS development through the use of consensus methods, expert and peer consultations, and by collating feedback from therapists and care home staff during the refinement stage of the pilot MiDAS.

To evaluate concurrent validity of MiDAS, the QoL-AD (Logsdon *et al.*, 1999) was administered at three time-points (base-line, mid-treatment, end-treatment). QoL-AD, Alzheimer's disease-related quality of life, is a Likert scale consisting of 13 items including: physical health, mood, family, ability to do chores around the house and life as a whole. QoL-AD can be administrated as a self-rating scale or as a proxy measure by a carer. QoL-AD was chosen for three reasons. Firstly, there was no 'gold-standard' music therapy outcome measure to compare with MiDAS and the nature of Quality of Life measures that assess the overall wellbeing of a person was closer to the underlying concept of MiDAS than the nature of measures that assess specific neuropsychiatric symptoms. Secondly, QoL-AD has been utilized successfully in previous studies with people with moderate to severe dementia (e.g. Hoe *et al.*, 2005; Spector and Orrell, 2006). Thirdly, QoL-AD was recommended in a European consensus on outcome measures for psychosocial intervention research in dementia care (Moniz-Cook *et al.*, 2008). Spearman's correlation coefficient was chosen to examine concurrent validity.

Construct validity involves measurement of psychological attributes (Nunnally and Berstein, 1994), including an investigation of the empirical relationships between the measures and an interpretation of the empirical evidence of how it clarifies the construct validity of the particular measure (Carmines and Zeller, 1969). Construct validity of MiDAS was evaluated using exploratory factor analysis and

by examining inter-relationships between the MiDAS items. Although it is widely accepted that a larger sample size would provide more trustworthy results for factor analysis, there is a lack of agreement on a recommended sample size. Hogarty *et al*. (2005) argued other elements such as subjects-to-variables ratio and communality of variance should also be taken into consideration. As this study had a small sample size, squared multiple correlation (SMC) for each variable to estimate communality values as well as Kaiser-Meyer-Olkin Measure of Sampling Adequacy and the Bartlett's test of sphericity (Bartlett, 1954) were evaluated for the appropriateness of factor analysis before conducting Principal Component Analysis (PCA).

Statistical analysis

Statistical analysis was conducted with SPSS version 21. Due to the data characteristics, square root transformation of the original data was conducted to evaluate ICC for inter-rater reliability and test-retest reliability. Tests on non-transformed data (original data) were also conducted to assess if there were any major differences between the results. For the same reason, both Principal Axis Factoring (PAF): a descriptive method suitable when the assumption of normality cannot be established, and PCA were conducted for the evaluation of concurrent validity to compare the results. Because of the small sample size, repeated observations of the same residents at different times were treated as unique observations when evaluating inter-rater reliability, concurrent validity and construct validity (factor analysis Test 1). This method was discussed with two independent statisticians and deemed suitable for the purpose of this study. The second factor analysis was conducted on the mean MiDAS scores of each participant (factor analysis Test 2). Before conducting Test 2, squared multiple correlation (R square) was used as SMC to determine the communality and appropriateness for factor analysis, as suggested by Field (2009).

Results

A total of 19 residents (5 male, 14 female) in the two care homes attended weekly group music therapy for up to 10 sessions (mean attendance 7 sessions). The residents were White British/White European (17) or Black British (2) with a mean age of 81 (SD 7.7). They were either single (6), married (3), separated/divorced (3), or widowed (7). All had diagnosis of dementia: Alzheimer's Disease (10), Vascular dementia (5), and other dementias (4). The mean months in the care homes were 20.1 (SD 14.2).

A total of 629 (Staff=306, Therapist=323) MiDAS forms were completed and included in the psychometric evaluation of MiDAS. The 306 MiDAS forms were completed by a total of 31 nursing staff: consisting of qualified nurses, health care assistants, and an activity co-ordinator. The mean of each five VAS item was: *Interest* 50.17 (SD=26.96), *Response* 51.57 (SD=28.32), *Initiation* 44.34 (SD=30.10), *Involvement* 46.63 (SD=29.58), and *Enjoyment* 46.17 (SD=29.95). The mean total score was 238.87 (SD=136.38).

Inter-rater reliability (therapist =52 paired forms, staff =27) and test-retest reliability (staff =31)
Due to the large SD, square root transformation, rather than log transformation, was chosen because all the VAS items (*Interest*, *Response*, *Initiation*, *Involvement* and *Enjoyment*) had scores of 0 (13/629 MiDAS forms, 11/629, 38/629, 18/629, 15/629 respectively), and these scores would have been excluded from analysis if log transformation was utilized. Skewness and kurtosis of the five VAS items were normalized with square root transformation. The differences between the results of the transformed data and the original data were not significant (ICC differences: range -0.024 to 0.003, mean -0.008, SD 0.011 for Therapist inter-rater reliability) hence the original scores were retained. The results of inter-rater reliability and test-retest reliability are shown on Table 2.

Inter-rater reliability (therapist): A total of 10 residents attended during therapist inter-rater sessions: 7 attended all the three sessions OM observed, 2 residents attended two out of three sessions, and 1

resident attend one of three sessions. In total, 52 paired forms for ten residents were completed. ICC (two-way random model with absolute agreement) revealed high correlations (range: 0.768-0.820).

Inter-rater reliability (staff) and *test-retest reliability*: 27 paired forms were completed by two staff inter-raters at CH A. Very low ICC was found: ranging from 0.127 to 0.362. In order to determine whether the low score was due to the low reliability of MiDAS itself or whether due to the difference in staff roles, we conducted test-retest on staff *Before* forms that were completed by any staff member over any consecutive weeks. 15 paired-forms were found among the 290 forms completed by staff. In addition, OM asked permanent staff members in Home B to complete further MiDAS test-retest (one week apart) on selected residents they knew well on eight separate occasions. In total, 31 paired-forms were examined. Acceptable correlations were found, ICC ranged from 0.498 to 0.609.

Internal consistency

Cronbach's alpha showed high correlations (α =0.967) between the five VAS items (n=629).

Concurrent validity (n=36)

QoL-AD was administered at baseline, mid-treatment, and end-treatment. The three results were combined (n=36) to evaluate an overall correlation, resulting in Spearman's rho=0.480 (*p*=0.003). Due to missing data, analysis was conducted on 10 residents (base-line), 14 residents (mid-treatment), and 12 residents (end-treatment) out of the 19 residents who participated in music therapy. The outcome of Spearman's rank correlation coefficient for each time point was acceptable: 0.524 (base-line), 0.469 (mid-treatment), and 0.474. (end-treatment).

Construct validity

Factor analysis (Test 1: 629 MiDAS ratings)

Individual MiDAS ratings were treated as unique observations and all the 629 forms were included for factor analysis Test 1. The Kaiser–Meyer–Olkin (KMO) measure of sampling adequacy and the Bartlett's test of sphericity (Bartlett, 1954) were carried out. The KMO value was 0.895, which exceeds the value of 0.6 suggested by Pallant (2005). The Bartlett's test of sphericity was statistically significant ($p<0.001$), supporting the factorability of the correlation matrix for Principal Component Analysis (PCA). A Principal Axis Factoring (PAF) was also conducted. The outcomes of KMO, Barlett's Test and the factor loadings were identical for a PCA and a PAF. Eigenvalues greater than 1 were used to select factors. Five components were extracted for the initial eigenvalues but only one component was retained for both PCA and PAF, with a single underlying factor which explained 88.5% of the variance. The one component was interpreted as engagement with music. Correlation between *Response* and *Involvement* was the highest (0.921) and the lowest correlation was observed for *Initiation* and *Enjoyment* was 0.791. (Table 3 here)

Factor analysis (Test 2: 19 residents)

R Square scores on the five VAS items ranged from 0.568 to 0.946 (mean 0.793), exceeding the recommended mean level of communality above 0.7 (MacCallum *et al*., 1999). The KMO value was 0.824, and the Bartlett's test of Sphericity was statistically significant ($p<0.001$). These results suggested factor analysis was acceptable for this study despite its small sample size. Five components were extracted for the initial eigenvalues but only one component was retained, with a single underlying factor which explained 91.1% of the variance. The pattern of the correlations was very similar to that of Test 1, but higher correlations between the items were observed in all but two of the paired correlations (*Initiation* and *Enjoyment*, *Involvement* and *Enjoyment*). (Table 4 here)

Discussion

This study evaluated the psychometric properties of Music in Dementia Assessment Scales (MiDAS). MiDAS is unique in the sense that it is an observational, dementia specific Visual Analogue Scales

developed to measure musical experiences of a person with dementia. Rigorous development of MiDAS using consensus methods, peer and expert consultations ensured its content validity. Staff and therapists who participated in the main study confirmed the clinical relevance of MiDAS, and the preliminary psychometric evaluation revealed that MiDAS has acceptable to good psychometric properties. At the same time, some aspects of the results may need to be interpreted cautiously.

Methodological problems

Although a large number of MiDAS forms were completed for this study, the statistical analysis method was affected by the small sample size (n=19) and missing data (e.g. uncompleted QoL-AD forms). RB and OM completed 52 MiDAS forms each for the purpose of therapist inter-rater reliability. However, just seven out of ten group members attended all three sessions that OM observed, resulting in only seven subjects eligible for ICC analysis. Since the main purpose of the inter-rater was to assess the reliability of the MiDAS as an instrument, not to evaluate the clinical changes in the participants, it was deemed acceptable to treat the 52 paired MiDAS forms as individual cases for this study, but further evaluation is required to confirm the results. Two sets of factor analysis were conducted to accommodate the limitation of using repeated observations of the same residents as independent observations. However, there is a possibility that the small sample size had an impact on the reduced observable variance, affecting the initial eigenvalues and the number of components extracted. Further evaluation with a larger sample size is required. Finally, all the QoL-AD forms were completed by staff, as the residents could not complete the self-rating forms. Concurrent validity of MiDAS should perhaps be tested further with people with mild to moderate dementia who can complete the self-rating QoL-AD.

Challenges of using MiDAS in clinical settings

The reason for the low staff inter-rater reliability may be partially due to the differences in the professional roles of the two raters (an associate health practitioner and an activity co-ordinator). The

health practitioner worked on shifts and mainly worked with the residents on an individual basis whilst the activity co-ordinator provided therapeutic group activities. The health practitioner explained she spoke to the residents before completing MiDAS forms if they appeared asleep or had been showing little response to external stimuli on the day. The activity co-ordinator acknowledged the lack of time spent with individual residents to assess subtle awareness or interest from the residents. This may explain the higher inter-rater reliability for *Initiation* and *Involvement* that residents' reactions were possibly more visible, whilst evaluating *Interest* and *Response* of less active residents required more staff time and possibly a deeper knowledge of each individual resident. Nurses and health care assistants who help with the residents' daily care were considered as the key staff MiDAS raters but involving permanent staff members with in-depth knowledge of the residents was sometimes not possible due to their job responsibilities and time constraints.

Another major challenge of using MiDAS relates to the use of Visual Analogue Scales (VAS) as an observational outcome measure. MiDAS is a relative scale; for instance, one resident's highest level of interest will differ from that of another resident, and the person's optimum levels will shift as dementia progresses. Judging each resident's optimum levels of function and scoring MiDAS relies on a rater's observational skills, his/her relationship with the resident, and their willingness to stop and reflect. It was particularly difficult to have consistent staff raters in Home B. The fact that some residents were rated by four or five different staff members over ten weeks may have had some impact on the consistency of their scores. The therapist RB expressed her concern with regard to the validity of staff scores after observing how some staff members were reluctant to complete the forms as they saw this process as an additional task they were asked to complete. This highlights the importance of training new MiDAS raters to clarify the purpose and procedure of ratings.

Limitations of MiDAS

Factor loadings revealed high correlations between the five VAS items. The high internal consistency was reflected in Cronbach's alpha (0.967). These results may have been influenced by the small number of items (n=5), and may also be an indication that the five VAS items are too similar to discriminate their characteristics from each other. It is also possible that the understanding of each of the items varied greatly between the raters, resulting in averaged mid-range scores for all items and hence high correlations. Since the main aim of MiDAS is to measure what people with dementia value in music, it can be argued having one construct (engagement with music) and using the total score of the five VAS as the main outcome does not devalue the clinical relevance of MiDAS. Nevertheless, each VAS was chosen carefully following the rigorous qualitative analysis. Further exploration of the underlying concepts for each VAS through peer and expert consultations may be useful.

Future recommendations

Having different staff raters every week reduces the reliability of the changes in each resident's MiDAS scores over the time. Nevertheless, this will be the reality in most care homes where shift working is the norm. The psychometric properties of MiDAS need to be evaluated further with a randomized controlled study with a larger sample size in several different care settings where the ethos of care and the levels of staff training are different.

It may be relevant to develop and evaluate a self-rating MiDAS. Likewise, it may also be feasible to develop a MiDAS family carer version that can be used in a music intervention study for people with dementia living in the community. The comparison of self-rating MiDAS scores and proxy MiDAS scores by family carers may clarify what observable clinical changes are and highlight the differences between internal musical experiences for people with dementia and perception of observers. Both the family carer and self-rating version should include the views of people with dementia and their families living in the community.

Conclusion

MiDAS was rigorously developed using qualitative methods and consensus approaches. This study indicates that MiDAS has adequate psychometric properties on a range of attributes even though the sample size was small. This preliminary evaluation revealed MiDAS has high therapist inter-rater reliability, low staff inter-rater reliability, adequate staff test-retest reliability, adequate concurrent validity and good construct validity. The high correlations between the five VAS items highlighted the need for ensuring the clarification of the differences between the items particularly with new MiDAS raters, but it also suggests that adding together the MiDAS items scores produces a meaningful summary score. Future research with a larger sample size could provide a more in depth psychometric evaluation including further exploration of potential underlying factors. The majority of music therapy studies in dementia focus on the reduction of neuropsychiatric symptoms and not the increased wellbeing of people with dementia. MiDAS challenges this culture by aiming to capture some of the aspects of musical experiences that people with dementia themselves would say are meaningful to measure. Much of the research on music in dementia is inconclusive and has failed to offer an explanation of who may or may not benefit. MiDAS provides a measure of engagement with the musical experience and therefore may provide a unique insight into who is likely to benefit in terms of improvement on other outcomes such as quality of life or reduction in psychiatric symptoms.

Conflict of interest

None.

Description of authors' roles

O. McDermott was the main researcher responsible for designing and conducting the study, data collection, the data analysis and writing the article. V. Orgeta was involved in the data analysis and provided statistical support. H.M. Ridder was involved in the development of the MiDAS and

provided clinical expert opinion. M. Orrell supervised the designing and conducting of the study and provided expert advise during the data analysis and writing up. All authors have approved the final draft of this article.

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Table 1. Examples of behaviour and mood for VAS items provided on the MiDAS form

Interest	Did he/she show his/her interest in an activity or other people around him/her? Did his/her posture or facial expression change if activities or music catch his/her attention? Did he/she become animated if activities or music catch his/her attention?
Response	Did his/her facial expression or body-movements indicate his/her awareness of staff or therapist? Did he/she make eye-contact with staff, therapist or other group members? Did he/she join in conversation, music making or make vocal sound?
Initiation	Did he/she try to communicate with staff, therapist or other group members? Did he/she start conversation, start music making, or initiate vocalisation? Did he/she talk about his/her life experiences (reminiscence) or mention music meaningful to them?
Involvement	Did he/she become engaged in conversation, music making, or any forms of communication? Did he/she show his/her enthusiasm in activities that interest him/her?
Enjoyment	Smiling, laughing, brighter mood Playfulness, sense of humour Relaxed mood

Table 2. Therapist and Staff inter-rater reliability and test-retest reliability

	Inter-rater (therapist) ICC (n=52) (CI)	Inter-rater (staff) ICC (n=27) (CI)	Test-retest (staff) ICC (n=31) (CI)
Interest	.791* (.656, .876)	.127 (-.275, .484)	.571* (.288, .766)
Response	.789* (.659, .873)	.189 (-.215, .531)	.498** (.181, .721)
Initiation	.768* (.626, .860)	.351** (-.037, .643)	.665* (.414, .822)
Involvement	.820* (.707, .893)	.362** (-.023, .651)	.614* (.333, .794)
Enjoyment	.787* (.643, .875)	.205 (-.185, .538)	.609* (.327, .791)

ICC Case 2 Model: Two-way random ANOVA, Absolute agreement, Single Measures, 95% Confidence Interval (CI).
* $p < .001$, ** $p < .05$

Table 3. Factor loadings (Principal Component Analysis) Test 1 (n=629)

	Interest	Response	Initiation	Involvement	Enjoyment
Interest	1.00				
Response	.921*	1.00			
Initiation	.821*	.857*	1.00		
Involvement	.859*	.888*	.891*	1.00	
Enjoyment	.843*	.848*	.791*	.845*	1.00

Note: Identical results for PCA and PAF
*$p < .001$

Table 4. Factor loadings (Principal Component Analysis) Test 2 (n=19)

	Interest	Response	Initiation	Involvement	Enjoyment
Interest	1.00				
Response	.947*	1.00			
Initiation	.869*	.928*	1.00		
Involvement	.901*	.955*	.973*	1.00	
Enjoyment	.844*	.885*	.754*	.828*	1.00

*p < .001