

Title

Body mass index trajectories in early childhood in relation to cardiometabolic risk profile and body composition at 5 years of age

Author names

Rasmus Wibaek, Dorte Vistisen, Tsinuel Girma, Bitiya Admassu, Mubarek Abera, Alemseged Abdissa, Kissi Mudie, Pernille Kæstel, Marit E. Jørgensen, Jonathan C.K. Wells, Kim F. Michaelsen, Henrik Friis, Gregers S. Andersen

Author affiliations

Department of Nutrition, Exercise and Sports, University of Copenhagen, Copenhagen, Denmark (RW, BA, MA, PK, KFM, HF); Clinical Epidemiology, Steno Diabetes Center Copenhagen, Gentofte, Denmark (RW, DV, MEJ, GSA); Department of Pediatrics and Child Health, Jimma University, Jimma, Ethiopia (TG); JUCAN Research Center, Jimma University, Jimma, Ethiopia (TG, BA, MA, AA); Department of Population and Family Health, Jimma University, Jimma, Ethiopia (BA); Department of Psychiatry, Jimma University, Jimma, Ethiopia (MA); Department of Laboratory Sciences and Pathology, Jimma University, Jimma, Ethiopia (AA); Ethiopian Public Health Institute, Jimma, Ethiopia (KM); Childhood Nutrition Research Centre, UCL Great Ormond Street Institute of Child Health, London, United Kingdom (JCKW); National Institute of Public Health, Southern Denmark University, Copenhagen, Denmark (MEJ).

Corresponding Author

Rasmus Wibaek, Clinical Epidemiology, Steno Diabetes Center Copenhagen, Niels Steensens

Vej 2, 2820 Gentofte, Denmark. Telephone: +45 28302381. E-mail:

rasmus.wibaek.christensen@regionh.dk

Authors' last names

Wibaek, Vistisen, Girma, Admassu, Abera, Abdissa, Mudie, Kæstel, Jørgensen, Wells,

Michaelsen, Friis, Andersen.

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Short running title

BMI trajectories and metabolic status at 5 years

Abbreviations

Body composition (BC), fat mass (FM), fat mass index (FMI), fat-free mass (FFM), fat-free mass index (FFMI), International Wealth Index (IWI), latent class trajectory (LCT), standard deviation (SD), World Health Organization (WHO).

Clinical Trial Registry

The birth cohort is registered in ISRCTN (<https://www.isrctn.com/>): identifier

ISRCTN46718296.

1 **Abstract**

2 **Background:** Both impaired and accelerated postnatal growth have been associated with
3 adult risk of obesity and cardiometabolic diseases like type-2-diabetes and cardiovascular
4 disease. However, the timing of the onset of cardiometabolic changes and specific growth
5 trajectories linking early growth with later disease risk are not well understood.

6 **Objective:** To identify distinct trajectories of body mass index (BMI) growth from 0-5 years
7 and examine their associations with markers of cardiometabolic risk at age 5 years.

8 **Design:** In a prospective birth cohort study of 453 healthy and term Ethiopian children with
9 BMI assessed a median of 9 times during follow-up, we identified subgroups of distinct BMI
10 trajectories in early childhood using latent class trajectory modelling. Associations of the
11 identified growth trajectories with cardiometabolic markers and body composition at 5 years
12 were analyzed using multiple linear regression analysis in four adjustment models for each
13 outcome.

14 **Results:** Four heterogeneous BMI growth trajectories were identified: “stable low BMI”
15 (19.2%), “normal BMI” (48.8%), “rapid catch-up to high BMI” (17.9%), and “slow catch-up to
16 high BMI” (14.1%). Compared with the “normal BMI” trajectory, children in the “rapid catch-
17 up to high BMI” trajectory had higher triglycerides (range of β -coefficients in model 1-4: 19-
18 21%), C-peptide (23-25%), fat mass (0.48-0.60 kg) and fat-free mass (0.50-0.77 kg) across the
19 four adjustment models. Children in the “stable low BMI” class had lower low-density
20 lipoprotein cholesterol (0.14-0.17 mmol/L), high density lipoprotein cholesterol (0.05-0.09
21 mmol/L), fat mass (0.60-0.64 kg), fat-free mass (0.35-0.49 kg), but higher triglycerides (11-
22 13%).

23 **Conclusions:** The development of obesity and cardiometabolic risk may be established
24 already in early childhood and thus provides further basis for timely interventions targeted at
25 young children from low-income countries with unfavorable growth patterns.

26

27 **Key words**

28 body composition; cohort study; child; developmental origins of health and disease; growth;
29 latent class trajectory modelling; non-communicable diseases; Sub-Saharan Africa.

30 **Introduction**

31 The prevalence of childhood obesity is a major threat to public health worldwide (1, 2). While
32 some progress has been seen in high-income countries, many low-income countries
33 experience a dual burden of malnutrition where rising prevalence of childhood overweight
34 coincides with high pre- and postnatal under-nutrition (3, 4). Both ends of this spectrum of
35 malnutrition have been associated with adult risk of obesity, type 2 diabetes and
36 cardiovascular disease (4, 5), and studies have shown that growth patterns in early childhood
37 is associated with these outcomes (6-8).

38 Most of these studies have examined associations of variability in early-life nutrition with
39 adult risk of disease and are therefore not designed to address the timing of onset of the
40 cardiometabolic adaptations that may occur already in childhood (9-11). Another issue is that
41 early-life growth is typically assessed at one or few timepoints in childhood. However, as
42 childhood growth is a dynamic process and diverging longitudinal growth patterns are likely
43 to associate differently with later cardiometabolic outcomes, studies with detailed and
44 repeated assessment of body size are needed. Additionally, very few of the studies that have
45 explored longitudinal growth patterns and their associations with adiposity and
46 cardiometabolic risk were conducted in low-income populations. These populations are
47 nonetheless likely to be particularly affected as a result of the dual burden of malnutrition
48 seen in many countries currently undergoing rapid economic and nutritional transition (12).
49 Stratification of children into different trajectories of growth associated with different levels
50 of adiposity and cardiometabolic risk profiles may therefore provide an opportunity to
51 identify early and targeted interventions in children from low-income populations.

52 Thus, we aimed to identify distinct trajectories of body mass index (BMI) growth from birth to
53 5 years and to estimate patterns of fat mass (FM) and fat-free mass (FFM) growth in infancy
54 for each of the identified BMI trajectories. Furthermore, we examined the relationships of the
55 identified BMI trajectories with cardiometabolic markers and BC at 5 years of age in a low-
56 income urban population.

57 **Subjects and methods**

58 **Study setting and participants**

59 We used data from the Infant Anthropometry and Body Composition (iABC) birth cohort
60 study, as described elsewhere (13, 14). Briefly, mother-child pairs were recruited within 48
61 hours after delivery at Jimma University Specialized Hospital, Jimma, Ethiopia between
62 December 2008 and October 2012. Eligible mothers were residing in Jimma town and had
63 given birth to a term (gestation >37 weeks) and apparently healthy child above 1500g without
64 congenital malformations. From birth to 60 months of age, we planned a total of 12 visits (0,
65 1.5, 2.5, 3.5, 4.5, 6, 12, 18, 24, 36, 48, 60 months). The modelling of BMI trajectories from 0-5
66 years included children with BMI assessed at the birth visit and at least one time in each of
67 the periods between the 1.5-6- and 12-60-months visits. For the subsequent regression
68 analysis with the 5-year outcomes, the analyses were restricted to children with a valid
69 measurement of the specific outcome in question and full covariate information. As some
70 participants refused blood sampling or were unable to deliver enough blood for analysis of all
71 biomarkers, the number of children included in the regression analyses differed by the
72 specific outcome.

73

74 **Data collection**

75 *Anthropometry and BC*

76 Length was measured to the nearest 0.1 cm in recumbent position using a SECA 416
77 Infantometer for children below 2 years and in standing position using a SECA 213 portable
78 height measurer for children above 2 years (SECA, Hamburg, Germany). Waist circumference
79 was measured to the nearest 0.1 cm in standing position with feet together midway between

80 the iliac crest and lowest costal margin using a non-stretchable measuring tape. Weight from
81 birth to 6 months was measured to the nearest 0.1 g using the built-in electronic scale of the
82 infant air displacement plethysmography (ADP) instrumentation (PEA POD, COSMED, Rome,
83 Italy), from 1-3 years to the nearest 0.1 kg using an electronic UNICEF scale (SECA, Hamburg,
84 Germany) and from 4-5 years to the nearest 1 g using the attached electronic scale of the
85 child/adult ADP instrumentation (BOD POD, COSMED, Rome, Italy). FM and FFM from birth to
86 6 months and at 5 years were measured using ADP in the PEA POD and the BOD POD with a
87 pediatric chair insert, respectively. Both ADP systems are accurate, precise, feasible and safe
88 methods for assessment of BC in infants and children (15, 16), and in the present cohort, the
89 PEA POD has previously been validated against a 3-component model with deuterium dilution
90 (13). In brief, these BC systems estimate total body density from weight and volume
91 measurements, and assuming known values of the density of FM and FFM, calculate the
92 proportion of FM in body weight. The theory and methods behind the PEA POD and BOD POD
93 are described in detail elsewhere (17, 18). A full BC assessment lasted 5-10 minutes and
94 during the volume measurements, the child was placed on a plastic bed (PEA POD) or in a
95 pediatric chair insert (BOD POD) in an enclosed test chamber, not wearing any clothes besides
96 a swim cap (PEA POD and BOD POD) or tight fitted underpants (BOD POD). Calibration of the
97 BC equipment was performed each morning. All the calculations were performed by the built-
98 in computer (PEA POD software version 3.3.0 and BOD POD software version 5.2.0). BMI, FM
99 index (FMI), and FFM index (FFMI) were calculated by dividing weight, FM, and FFM with the
100 squared height in meter, respectively.

101

102 *Blood pressure at 5 years*

103 After relaxing for 5 minutes, systolic and diastolic blood pressure were measured in sitting
104 position using a blood pressure device with age-appropriate cuffs (Pressostabil model, Welch
105 Allyn Inc., Skaneateles Falls, USA). Measurements were done in duplicate and averaged.

106

107 *Cardiometabolic markers at 5 years*

108 A 2 ml venous blood sample was drawn from the antecubital fossa as the last element of
109 assessments of the child following a minimum of 3 hours of fasting. Glucose concentrations
110 were measured in whole blood using the HemoCue Glucose 201 RT System (HemoCue,
111 Ängelholm, Sweden). Glycosylated haemoglobin (HbA1c, mmol/mol) was measured on whole
112 blood using a DCCT aligned Quo-Test® A1c Analyzer (EKF Diagnostics, Cardiff, Wales).
113 Subsequently, serum was obtained by centrifuging the whole blood sample, aliquoted in
114 3x0.4 mL and frozen at -80°C until analyzed. The serum samples were analyzed at the
115 Ethiopian Public Health Institute, using the module c501 of the COBAS 6000 analyzer (Roche
116 Diagnostics International Ltd, Rotkreuz, Switzerland) for total-, LDL-, and HDL cholesterol, and
117 triglyceride concentrations (all lipids in mmol/L), and the module e601 for insulin ($\mu\text{U}/\text{mL}$),
118 and C-peptide (ng/mL). The homeostasis model assessment of insulin resistance index
119 (HOMA-IR) was calculated as $\text{insulin} \times \text{glucose} / 22.5$ (19).

120

121 *Covariates*

122 Information on birth order, child's sex, gestational age, maternal age, educational level, and
123 socioeconomic status of the family was obtained at the birth visit. Maternal postpartum
124 height was measured to the nearest 0.1 cm using a Seca 214 Stadiometer (SECA, Hamburg,

125 Germany). Birth order was self-reported as the number of previous pregnancies. Gestational
126 age of the new-born was assessed by trained research nurses using the New Ballard Score test
127 (20). The International Wealth Index (IWI) was used to assess socioeconomic status of the
128 family. The IWI assesses the material well-being of households in low- and middle-income
129 countries (21), and includes information of 12 material well-being dimensions, including
130 seven household assets, access to two public services and three characteristics of the house.
131 The IWI ranges from 0 to 100 (highest wealth). Information on breastfeeding status was
132 assessed at 4 to 6 months post-partum and divided into four categories: exclusive (no other
133 foods given), almost exclusive (no other foods given except water), predominant (breast milk
134 as primary food) and partial/no (breast milk not the primary food/not breast feeding) (22).

135

136 Ethics

137 Ethical approval was granted by the Jimma University Ethical Review Committee (Ref. no.
138 RPGC/279/2013). Written informed consent was obtained from parents or caregivers of all
139 eligible children. Children with any medical condition observed by the research nurses were
140 referred in accordance with local clinical guidelines.

141

142 Statistical methods

143 All descriptive data are presented as mean (SD) or median (interquartile range) for continuous
144 variables and percentages for categorical variables. P-values <0.05 were considered
145 statistically significant. All analyses were carried out in R version 3.4.1 (The R foundation for
146 Statistical Computing).

147

148 *Identification of latent BMI trajectories in childhood*

149 Heterogeneity in repeated measures of BMI was analyzed using latent class trajectory (LCT)
150 modelling to identify distinct subgroups of children with similar trajectories of BMI growth
151 from 0-5 years (23, 24). We ran a series of LCT models with various specifications of BMI as a
152 function of age and number of subgroups (classes). As described in detail in the **Supplemental**
153 **Methods**, the best fitting model according to our “a priori” criteria was obtained with a four
154 class model specified with natural cubic splines with knot points at 0, 3, 6, 24, 48, and 60
155 months. Using the class assignments from the LCT analysis, we re-estimated the BMI
156 trajectories from 0-60 months for girls and boys separately to assess if there were any sex
157 differences in BMI growth for each of the four assigned classes.

158

159 *FMI and FFMI growth in infancy*

160 For each of the identified BMI trajectory classes, we applied mixed-effects modelling to
161 estimate the corresponding mean growth in FMI and FFMI from 0-6 months of age. We
162 required children to have a minimum of three FMI and FFMI measurements during the first 0-
163 6 months to be included in the modelling. The FMI and FFMI as a function of age were
164 modelled separately and fitted with natural cubic splines with knot points at ages 0, 3 and 6
165 months.

166

167 *Associations of BMI trajectories with BC and cardiometabolic markers at 5 years*

168 We analyzed the relationship of the identified BMI trajectory classes with BC and
169 cardiometabolic markers at 5 years using multiple linear regression analysis. The “normal
170 BMI” trajectory class was used as reference. In all analyses, we log transformed outcomes

171 where the corresponding model residuals were not normally distributed, which resulted in
172 normally distributed model residuals. The resulting estimates were back transformed and
173 presented on the relative scale as percentwise change. We ran four separate models for each
174 outcome. Model 1 was adjusted for sex, birth order, and gestational age. Model 2 was
175 additionally adjusted for the child's exact age at the 5-year visit, maternal age at delivery,
176 maternal postpartum height, maternal educational status, and family socioeconomic status
177 (IWI). Model 3 was additionally adjusted for child birth weight. Model 4 was additionally
178 adjusted for child BMI at the 5-year visit. Since BMI comprises both FM and FFM, the analyses
179 of the outcomes FM and waist circumference were adjusted for the lean component of BMI
180 (FFM and height at 5 years) instead of BMI in model 4. Similarly, the analysis of the outcome
181 FFM was adjusted for FM and height at 5 years (the fat component of BMI). We compared the
182 estimated associations across the four models for a given outcome and exposure using a
183 complete case approach, limiting the analyses to data with complete information on all
184 covariates in model 4. In further analyses, we accounted for multiple testing using the
185 Benjamini-Hochberg approach (25), where the number of tests was set to 45 (15 outcomes
186 for three exposure groups) (**Supplemental Figure 1**). To account for breastfeeding, we ran
187 sensitivity analyses on a smaller sample with available information on breastfeeding status at
188 4 to 6 months post-partum (**Supplemental Figure 2**).

189 **Results**

190 At birth, the mothers were on average 24.6 years and 49% were primiparous. Half of the
191 mothers had completed primary school or higher and 94% were breast feeding either
192 exclusively, almost exclusively or predominantly during the first 4-6 months after birth
193 **(Supplemental Table 1)**. Low birth weight was seen in 9.4% of the children. At 5 years, the
194 children were on average more than 1.2 SD scores below the World Health Organization
195 (WHO) international growth standards for height, 15.3% were stunted, and 5.9% were
196 overweight or obese. The average BMI of 15.0 kg/m² at 5 years was similar to the average of
197 the WHO international growth standards **(Supplemental Table 2)** (26).

198

199 **Latent BMI trajectories in early childhood**

200 A total of 453 children were included in the modelling of the BMI trajectories **(Figure 1)**. The
201 children had their BMI assessed a median (interquartile range) of 9 (8-10) times during follow-
202 up period, contributing a total of 3952 observations to LCT modelling. We identified four
203 distinct BMI trajectory classes from birth to 5 years **(Figure 2 and Supplemental Figure 3)**:
204 Trajectory class 1: “stable low BMI” (19.2%, n=87), 2: “normal BMI” (48.8%, n=221), 3: “rapid
205 catch-up to high BMI” (17.9%, n=81), and 4: “slow catch-up to high BMI” (14.1%, n=64). The
206 ability of the LCT modelling to discriminate the identified classes was acceptable with high
207 median posterior probabilities of assigned class membership above 85% for all four classes
208 **(Supplemental Figure 4)** (27). Children in the “stable low BMI” trajectory class presented on
209 average slow initial BMI gains plateauing after 4 months and remained at a relatively low level
210 in infancy with a small catch-up from around 8 to 27 months. The “normal BMI” trajectory
211 was very similar to the WHO international growth standards with the infancy peak at around

212 5 months (26). The “rapid catch-up to high BMI” trajectory presented on average accelerated
213 initial BMI gains peaking at around 9 months and steadily declining towards a normal BMI
214 level at 60 months. The “slow catch-up to high BMI” trajectory presented on average slow
215 BMI gains from birth until peaking at around 17 months.

216

217 Characteristics of the latent BMI trajectories

218 Background characteristics, BC measures, and cardiometabolic markers of the mother-child
219 pairs according to the BMI trajectory class memberships are shown in **Table 1** and **Table 2**. At
220 birth, we did not observe any statistically significant differences between the four trajectories
221 besides for gestational age. The highest proportions of low birth weight (13-18%) were seen
222 in the two catch-up trajectory classes. At 5 years, children in all groups were on average
223 lighter, shorter, and thinner, compared with the WHO international growth standards (26). At
224 5 years, children in the “stable low BMI” class presented the highest proportions of stunting
225 and underweight, and the lowest proportions of overweight, while children in the “rapid
226 catch-up to high BMI” class presented the lowest proportions of stunting and underweight
227 and had on average 1.2 kg more FM than the “stable low BMI” class.

228

229 FMI and FFMI growth in infancy

230 Each of the four BMI trajectory classes differed in terms of the average velocity of both FMI
231 and FFMI growth from 0-6 months (**Figure 3**). Children in the “rapid catch-up to high BMI”
232 class had the lowest FMI at birth but gained FMI at an accelerated rate through infancy
233 resulting in the highest average FMI at 6 months. While growth in FMI largely reflected the
234 patterns of the BMI classes the first 3-4 months (Figure 3 and Supplemental Figure 3), the

235 accretion patterns of FFMI were slightly concave with much smaller variation between the
236 classes. At 6 months, the difference between the highest and lowest growth trajectory was
237 2.68 kg/m² for FMI and 1.32 kg/m² for FFMI.

238

239 **Associations of BMI trajectories with BC and cardiometabolic markers at 5 years**

240 Compared with children classified in the “normal BMI” trajectory (reference group), children
241 in the “rapid catch-up to high BMI” class had higher triglycerides (range of β -coefficients in
242 model 1-4: 19-21%) and C-peptide (23-25%) across the four models (**Figure 4** and

243 **Supplemental Table 4**). Tendencies of positive associations were also seen for insulin, HbA1c,

244 and HOMA-IR, although confidence bands were wide and included zero. Conversely, children

245 classified in the “stable low BMI” class had lower LDL cholesterol (0.14-0.17 mmol/L), HDL

246 cholesterol (0.05-0.09 mmol/L), but higher triglycerides (11-13%) across the four models.

247 Tendencies of inverse associations were also seen for total cholesterol, and systolic blood

248 pressure, although confidence bands were wide and included zero.

249 In relation to the BC and anthropometry measures, children in the “rapid catch-up to high

250 BMI” class were taller (1.3-1.8 cm), had larger waist circumference (0.9-1.6 cm), higher FM

251 (0.48-0.60 kg) and FFM (0.50-0.77 kg), while children in the “stable low BMI” class had lower

252 height (1.2-1.5 cm), smaller waist circumference (1.0-1.3 cm), and less FM (0.60-0.64 kg) and

253 FFM (0.35-0.49 kg) across the four models. Moreover, children in the “slow catch-up to high

254 BMI” class had higher FM (0.46-0.58 kg), but not higher FFM. When accounting for multiple

255 testing in the results presented in the fully adjusted model (model 4), the associations for the

256 markers of lipid metabolism were no longer significant. However, the inverse associations of

257 “stable low BMI” with waist circumference and FM, and the positive associations of “rapid

258 catch-up to high BMI" with height, FM and FFM remained significant (Supplemental Figure 1).
259 Adjusting all the models for breastfeeding status at 4 to 6 months postpartum did not alter
260 the associations markedly (Supplemental Figure 2).

261 Discussion

262 This is the first study from a sub-Saharan African population to examine how distinct
263 trajectories of adiposity-related BMI the first years of life have different implications on
264 markers of cardiometabolic risk and BC in early childhood. We found that accelerated BMI
265 growth in infancy was associated with higher concentrations of triglycerides, C-peptide,
266 height, waist circumference, FM, and FFM, while low BMI growth was associated with lower
267 LDL, HDL, height, waist circumference, FM, and FFM, but higher triglycerides at 5 years in
268 Ethiopian children. The effect estimates for the cardiometabolic markers did not change
269 markedly across the different adjustment models. Thus, the associations were independent of
270 maternal characteristics, socioeconomic status, and birth weight and were not mediated
271 through size at 5 years. Since this investigation was exploratory rather than confirmatory, we
272 did not account for multiple testing in the primary analysis.

273 Studies examining the health effects of variability in early-life growth have typically used
274 predefined cut-offs to define low, normal or accelerated growth between two or more
275 timepoints (7, 28). This *a priori* classification forces observations into specific categories that
276 may not fully reflect the complex and dynamic trajectories of child growth. Using an
277 exploratory data-driven approach, we were able to identify trajectories of low, normal and
278 accelerated growth associated with markers of adiposity, insulin- and lipid metabolism
279 without *a priori* categorization of growth variability. The generalizability of these trajectories
280 was supported by their similarity to BMI trajectories identified in studies from high-income
281 countries using similar data-driven modelling and overlapping age-periods (29-32). In line
282 with the present findings, these studies also reported associations of accelerated or stable
283 high BMI trajectories in early childhood with risk of obesity (29, 30, 32), elevated levels of

284 markers of adiposity (11, 30), and changes in cardiometabolic status (11, 31) in childhood or
285 early adolescence.

286 The present findings and existing evidence are consistent with the capacity-load model of
287 cardiometabolic risk (33), with early low BMI growth constraining metabolic capacity (e.g.
288 lean mass deficits) while accelerated BMI growth increases metabolic load (e.g. excess FM
289 and waist circumference). Both low metabolic capacity and high metabolic load challenge
290 metabolic homeostasis. We found that FMI accretion from 0-6 months largely reflected the
291 patterns of BMI growth. Children with accelerated BMI growth therefore had the highest FM
292 accretion and on average $\frac{1}{2}$ kg and 1 kg more FM at 5 years than the children with normal and
293 low BMI growth, respectively. Compared with UK children, the Ethiopian children also had
294 higher average FM (boys: 3.11 vs. 4.19 kg, and girls: 3.97 vs. 4.14 kg) and markedly lower FFM
295 (boys: 16.35 vs. 12.27 kg, and girls: 14.60 vs. 12.04 kg) at 5 years (34). As the children with
296 accelerated BMI growth were larger overall, they also presented the highest FFMI accretion
297 and FFM at 5 years. However, the effect estimates for the association with FM and FFM were
298 of almost similar size, despite that the FFM at 5 years was an almost 3 times larger body
299 compartment than FM. Moreover, we have recently shown that FM accretion the first 4
300 months of life was positively associated with FMI at 4 years, but not FFMI (35), and also that a
301 lower birth weight associates with a FM but not a FFM catch-up growth pattern (36).

302 Altogether, it is therefore possible that accelerated BMI growth during a critical window the
303 first 6 months of life induces disproportionately high accretion of FM in relation to FFM, that
304 may result in unfavorable cardiometabolic changes already in childhood. Moreover, children
305 with accelerated BMI growth may be particularly vulnerable to the effects of a high FM as
306 they had the lowest weight at birth, indicating constrained development of metabolic

307 capacity during fetal life, and highest waist circumference at 5 years, related to increased
308 metabolic load. Slow growth during infancy may continue to constrain the development of
309 metabolic capacity, which is supported by a high proportion of stunting at 5 years in children
310 with low BMI growth. However, these children did not appear to have a particularly
311 unfavorable cardiometabolic status as they presented lower cholesterol concentrations and
312 FM compared with the children with a normal BMI growth. These findings are similar to a
313 study of survivors of severe-acute malnutrition in Malawi, where cardio-metabolic risk
314 markers were not elevated despite several indices of low metabolic capacity (37). The likely
315 explanation is that the accompanying lower FM is not sufficient to challenge the low
316 metabolic capacity in these groups of children.

317 Continued follow-up of the present cohort will confirm whether these proposed relationships
318 will persist in the longer term, but it is possible that children with either low or accelerated
319 early growth are at highest risk of developing obesity and cardiometabolic diseases later in
320 life. These new insights may help identify potential early-life targets for interventions that
321 promote FFM accretion, linear growth, and normal birth weight (i.e. metabolic capacity)
322 without increasing excess BMI growth and FM accretion (i.e. metabolic load). This is
323 particularly relevant for public health professionals working in low-income countries where
324 both under-nutrition and excess fat accumulation in early life are pertinent and discernible
325 issues associated with an ongoing nutritional transition (12).

326 The strengths of this study include the 5-year longitudinal design with an average of nine
327 repeated BMI assessments per child, detailed assessment of changes in BMI and BC in a
328 critical window of development from birth to 6 months, and the assessment of multiple
329 cardiometabolic markers and BC indices at 5 years of age. Another strength is the ability of

330 the LCT modelling to capture the complex and dynamic patterns of child growth and let the
331 data speak for themselves by *a posteriori* identification of distinctive subgroups of BMI
332 growth. Finally, over a 5-year period and 12 assessment waves, it was possible to retain 79%
333 and 61% of the cohort of children included in the follow-up study for the growth modelling
334 and regression analysis of the 5-year outcomes, respectively.

335 However, the study also had some limitations. First, longitudinal attrition inevitably makes
336 the LCT modelling of the BMI trajectories more uncertain. Second, it cannot be excluded that
337 children not able to participate in the 5-year follow-up visit have caused some selection bias
338 in the estimated associations, although these mother-child pairs were largely similar to those
339 included in the analysis (**Supplemental Table 3**). Third, the data-driven nature of the LCT
340 modelling may limit generalizability of the present findings and future studies should assess
341 whether the classification of the distinct BMI trajectories commonly apply to BMI growth in
342 this age range. Moreover, as the LCT analysis is not forcing children into groups using
343 predefined cut-offs, the size of each identified class may vary substantially, which may limit
344 the power in the subsequent regression analyses. Although we obtained relatively large class
345 sizes, the effect estimates presented in the regression analysis had relatively large confidence
346 bands which may have resulted in type-2-errors. Fourth, we did not stratify the LCT analysis
347 by sex, as this would have resulted in small class sizes and limit the power in the subsequent
348 regression analyses. However, when re-estimating the BMI trajectories separately for boys
349 and girls using the four identified classes from the whole study sample, we found no clinically
350 meaningful sex differences in BMI growth for any of the four classes (**Supplemental Figure 5**).
351 Fifth, as BMI comprises both weight and height, we were not able to assess how growth
352 trajectories of these individual components was associated with later BC and cardiometabolic

353 risk markers. Finally, the observational nature of this study does not allow us to imply any
354 causal effects. We cannot rule out that important unmeasured covariates such as pre-
355 pregnancy maternal nutritional status, gestational weight gain, paternal BMI, fetal growth
356 trajectories, comprehensive dietary assessment in infancy and childhood, and duration of
357 breastfeeding have confounded our results. However, in a sensitivity analysis we adjusted our
358 results for a crude measure of breastfeeding status at 4-6 months post-partum, which did not
359 affect our results noticeably.

360 In a birth cohort of Ethiopian children, we identified considerable heterogeneity in BMI
361 growth in early childhood, which was associated with FM accretion in early infancy and
362 markers of cardiometabolic status and indices of BC at 5 years. Collectively, our findings offer
363 an important contribution to understand pathways from early growth to later
364 cardiometabolic health by suggesting that distinct trajectories of BMI growth in early life may
365 be a key factor in the complex etiology of obesity and cardiometabolic risk development
366 already from an early age. This is highly relevant for public health professionals in low-income
367 countries, where the dual burden of malnutrition is a pertinent problem, as it offers potential
368 early-life targets for interventions by identifying subgroups of children that may present
369 unfavorable growth trajectories.

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377 performed statistical analysis; RW wrote paper; RW and GSA had primary responsibility for
378 final content. All authors revised and approved the final manuscript.

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Tables

Table 1 Description of the mother-child pairs attending the 5-year follow-up visit and included in the trajectory modelling according to the BMI trajectory class membership (n= 352) ¹

	Trajectory group				p-value ²	Missing, n
	1: Stable low BMI (n=69)	2: Normal BMI (n=173)	3: Rapid catch-up to high BMI (n=60)	4: Slow catch-up to high BMI (n=50)		
Maternal characteristics						
Age at birth (years)	24.8 (4.8)	24.7 (4.6)	24.1 (4.6)	24.4 (5.5)	0.825	0
Postpartum height (cm)	156.6 (5.8)	157.5 (6.3)	157.7 (5.5)	155.6 (6.3)	0.187	2
Postpartum body mass index (kg/m ²)	21.91 (3.06)	22.09 (3.65)	22.33 (3.39)	23.02 (3.78)	0.352	6
Birth order of current child						
First	44.9	44.5	66.7	54.0		
Second	27.5	29.5	16.7	26.0		
Third or above	27.5	26.0	16.7	20.0	0.122	0
Breastfeeding status at 4 to 6 months post-partum						
Exclusive	8.8	15.1	15.5	4.3		
Almost exclusive (water given)	20.6	20.1	24.1	23.4		
Predominant	66.2	57.9	56.9	63.8		
Partial or no	4.4	6.9	3.4	8.5	0.581	20
Maternal education						
No school	5.8	8.7	1.7	10.0		
Some primary school	47.8	47.4	41.7	34.0		
Completed primary school	11.6	14.5	20.0	20.0		
Completed secondary school	18.8	17.9	20.0	26.0		
Higher education	15.9	11.6	16.7	10.0	0.496	0
Socioeconomic status (International Wealth Index)	45.9 (17.2)	44.6 (16.7)	48.2 (18.6)	45.4 (16.5)	0.565	0
Child characteristics at birth						
Sex (boys)	47.8	50.3	51.7	52.0	0.965	0
Gestational age (weeks)	39.2 (0.9)	39.1 (1.0)	39.0 (1.0)	38.7 (0.8)	0.020	0
Weight (kg)	3.10 (0.40)	3.07 (0.38)	2.95 (0.43)	2.98 (0.47)	0.076	0
Length (cm)	49.2 (2.1)	49.3 (1.9)	49.0 (1.9)	48.7 (2.1)	0.291	0
Fat mass (kg)	0.25 (0.20)	0.23 (0.15)	0.18 (0.15)	0.22 (0.16)	0.076	2
Fat-free mass (kg)	2.85 (0.29)	2.85 (0.32)	2.78 (0.35)	2.78 (0.35)	0.322	2
Low birth weight (%) ³	5.8	6.9	13.3	18.0	0.061	0
Child characteristics at 5 years						
Age at 5-year visit (months)	59.8 (1.7)	60.0 (1.4)	60.1 (1.5)	59.8 (1.4)	0.702	0
Weight (kg)	15.14 (1.72)	16.25 (1.88)	17.49 (2.32)	16.76 (2.05)	<0.001	0
Length (cm)	102.8 (4.8)	104.2 (4.2)	106.0 (4.5)	103.8 (3.8)	<0.001	0
Weight for age SDS ⁴	-1.40 (0.81)	-0.90 (0.80)	-0.39 (0.89)	-0.67 (0.87)	<0.001	0
Height for age SDS	-1.44 (0.98)	-1.15 (0.86)	-0.78 (0.95)	-1.25 (0.79)	<0.001	0
BMI for age SDS	-0.73 (0.91)	-0.27 (0.79)	0.14 (0.79)	0.15 (0.84)	<0.001	0
Underweight ⁵	21.7	8.7	0.0	6.0	<0.001	0
Stunted ⁶	26.1	12.1	11.7	16.0	0.043	0
Wasted by BMI (Thinness) ⁷	7.2	2.3	0.0	2.0	0.090	0
Overweight ⁸	1.4	1.7	10	14.0	0.001	0
Obese ⁹	0.0	0.6	3.3	2.0	0.164	0

¹ Data are mean (SD) (for continues normally distributed variables) and percentages (for categorical variables). ² Differences between trajectory groups were calculated by One-way ANOVA F-test (for continuous variables), Pearson's Chi-Square test of independence (for categorical variables with expected counts above 5 in all cells) and Fisher's exact test of independence (for categorical variables with expected counts in any cell below 5). ³ Low birth weight is defined as birth weight <2500 g. ⁴ Standard deviation scores (SDS) are derived using the 2006 (aged <61 months) and 2007 (aged ≥61 months) World Health Organization (WHO) child growth standards. ⁵ Weight for age more than 2 SDS below the age- and sex-specific median of the WHO child growth standards. ⁶ Height for age more than 2 SDS below the age- and sex-specific median of the WHO child growth standards. ⁷ BMI for age more than 2 SDS below the age- and sex-specific median of the WHO child growth standards. ⁸ BMI-for-age more than 1 to 2 SDS above the sex-specific median of the WHO child growth standards. ⁹ BMI-for-age more than 2 SDS above the sex-specific median of the WHO child growth standards.

Table 2 Cardiometabolic markers and body composition at 5 years of age in the children attending the 5-year follow-up visit and included in the trajectory modelling according to the BMI trajectory class membership (n= 352) ¹

	Trajectory group				p-value ²	Missing, n
	1: Stable low BMI (n=69)	2: Normal BMI (n=173)	3: Rapid catch-up to high BMI (n=60)	4: Slow catch-up to high BMI (n=50)		
Glucose metabolism						
Glucose (mmol/L)	5.84 (0.83)	5.90 (0.83)	5.98 (1.08)	5.89 (0.63)	0.844	26
HbA1c (mmol/mol)	37 (5)	37 (4)	38 (4)	38 (5)	0.736	83
Insulin (μ U/mL) ³	5.47 (2.62, 11.58)	5.89 (3.15, 11.08)	8.29 (4.12, 11.61)	6.39 (3.04, 9.43)	0.166	34
C-peptide (ng/mL) ³	0.99 (0.57, 1.24)	1.05 (0.59, 1.54)	1.31 (0.77, 1.84)	1.01 (0.70, 1.35)	0.094	39
HOMA-IR ^{3,4}	1.12 (0.55, 2.53)	1.21 (0.63, 2.33)	1.86 (0.86, 2.58)	1.31 (0.65, 2.07)	0.196	34
Lipids						
Total cholesterol (mmol/L)	3.32 (0.63)	3.43 (0.58)	3.49 (0.71)	3.37 (0.55)	0.430	30
LDL (mmol/L)	1.54 (0.52)	1.68 (0.55)	1.70 (0.68)	1.66 (0.49)	0.322	31
HDL (mmol/L)	0.76 (0.23)	0.81 (0.26)	0.77 (0.27)	0.78 (0.24)	0.434	35
Triglycerides (mmol/L) ³	1.00 (0.79, 1.29)	0.91 (0.68, 1.27)	1.04 (0.83, 1.53)	0.90 (0.72, 1.15)	0.042	35
Blood pressure						
Systolic (mmHg)	86.1 (6.7)	88.0 (7.2)	89.3 (8.0)	87.5 (7.5)	0.092	2
Diastolic (mmHg)	53.6 (6.8)	54.2 (8.2)	55.4 (10.3)	54.3 (9.3)	0.673	2
Anthropometry and body composition						
Body mass index (kg/m ²)	14.32 (1.16)	14.92 (1.10)	15.51 (1.21)	15.52 (1.20)	<0.001	0
Waist circumference (cm)	50.09 (2.69)	51.38 (2.89)	52.77 (3.21)	51.99 (2.85)	<0.001	1
Fat mass (kg)	3.50 (1.22)	4.13 (1.15)	4.68 (1.34)	4.64 (1.24)	<0.001	16
Fat-free mass (kg)	11.64 (1.14)	12.14 (1.33)	12.86 (1.69)	12.10 (1.43)	<0.001	16
Fat mass index (kg/m ²)	3.30 (1.10)	3.79 (0.99)	4.14 (1.05)	4.29 (1.03)	<0.001	16
Fat-free mass index (kg/m ²)	11.02 (0.78)	11.17 (0.80)	11.41 (1.01)	11.26 (0.96)	0.079	16

¹ Data are mean (SD) for continuous variables that are normally distributed and median (interquartile range) for continuous variables that are not following a normal distribution. ² Differences between trajectory groups were calculated by One-way ANOVA F-test for continuous normally distributed variables. Variables found not to follow a normal distribution were log transformed prior to the tests of group differences. ³ Nonnormally distributed. ⁴ Homeostasis model assessment of insulin resistance (HOMA-IR) was calculated as insulin (μ U/mL) \times glucose (mmol/l) / 22.5.

Legends for figures

Figure 1. Flow diagram of the included children and number of BMI observations at each follow-up visit from birth to 60 months.

Figure 2. Distinct body mass index (BMI) trajectories from 0-5 years of children in the iABC birth cohort derived from a latent class trajectory analysis. Solid lines display the class-specific estimated average BMI as a function of age. The dashed lines show the estimated 95% confidence limits. The shaded areas indicate the reference in standard deviation scores (SDS) from the median BMI-for-age according to the international growth standards developed by the World Health Organization. Normal BMI (white) is defined as a BMI-for-age SDS from -1 to 1, mild thinness as ≥ -2 to < -1 SDS (light grey), thinness as < -2 SDS (grey), overweight as > 1 to ≤ 2 SDS (light grey), and obese as > 2 SDS (grey). The density of BMI observations is shown as a rug plot along the x-axis.

Figure 3. Estimated changes in fat mass index and fat-free mass index from 0-6 month for each of the identified latent body mass index trajectories classes.

Figure 4. The coefficients (95% CIs) displayed in the forest plots were derived from multiple linear regression models and represent the mean difference in concentrations of cardiometabolic markers and body composition indices between the reference trajectory 2: normal BMI, and the 3 BMI trajectory categories 1: stable low BMI, 3: rapid BMI catch-up and 4: slow BMI catch-up, respectively. The 4 distinct BMI trajectories (exposure variable) were

derived from a latent class trajectory analysis. The outcome variables insulin, C-peptide, HOMA-IR and triglycerides were log transformed prior to analyses. The resulting effect estimates were back-transformed and presented as percentwise change. Model 1 (the leftmost circle) was adjusted for sex, birth order and gestational age. Model 2 was additionally adjusted for child age at the 5-year visit, maternal age at delivery, maternal postpartum height, maternal educational status and family socioeconomic status (International Wealth Index). Model 3 was additionally adjusted for child birth weight. Model 4 (the rightmost circle) was additionally adjusted for child BMI at the 5-year visit. In model 4, the analyses of FM and waist circumference were adjusted for FFM and height at 5 years instead of BMI, and the analysis of FFM was adjusted for FM and height at 5 years instead of BMI. * $P \leq 0.05$, ** $P \leq 0.01$, *** $P \leq 0.001$.







