# Diminishing returns drive altruists to help extended family

P. Kennedy<sup>1\*</sup>, S. Sumner<sup>1†</sup>, P. Botha<sup>1</sup>, N. J. Welton<sup>2</sup>, A. D. Higginson<sup>3</sup> & A. N. Radford<sup>1</sup>

Supplementary files: (1) Supplementary Information (modelling detail), (2) dataset and statistical code, (3) simulation code, (4) Extended Data (supplementary figures).

#### Abstract:

Altruism between close relatives can be easily explained. However, paradoxes arise when organisms divert altruism towards more-distantly-related recipients. In some social insects, workers 'drift' extensively between colonies and help raise less-related foreign brood, seemingly reducing inclusive fitness. Since being highlighted by W. D. Hamilton, three hypotheses (bet-hedging, indirect reciprocity, and diminishing returns to cooperation) have been proposed for this surprising behaviour. Here we show using inclusive fitness theory that bet-hedging and indirect reciprocity could only drive cooperative drifting under improbable conditions. However, diminishing returns to cooperation create a simple context in which sharing workers is adaptive. Using a longitudinal dataset comprising over a quarter of a million nest-cell observations, we quantify cooperative payoffs in the Neotropical wasp *Polistes* canadensis, where drifting occurs at high levels. As the worker-to-brood ratio rises in a worker's home colony, the predicted marginal benefit of a worker for expected colony productivity diminishes. Helping on related colonies can allow effort to be focused on related brood that are more in need of care. Finally, we use simulations to show that cooperative drifting evolves under diminishing returns when dispersal is local, allowing altruists to focus their efforts on related recipients. Our results indicate the power of nonlinear fitness effects to shape social organization, and suggest that models of eusocial evolution should be extended to include neglected social interactions within colony networks.

<sup>&</sup>lt;sup>1</sup>School of Biological Sciences, University of Bristol, 24 Tyndall Avenue, Bristol, BS8 1TQ, UK

<sup>&</sup>lt;sup>2</sup>Population Health Sciences, Bristol Medical School, University of Bristol, Canynge Hall, 39 Whatley Road, Bristol, BS8 2PS, UK

<sup>&</sup>lt;sup>3</sup>Centre for Research in Animal Behaviour, University of Exeter, Exeter, EX4 4QG, UK

<sup>&</sup>lt;sup>†</sup>Current address: Centre for Biodiversity and Environment Research (CBER), University College London, Division of Biosciences, Gower Street, London, WC1E 6BT, UK

<sup>\*</sup>Correspondence: patrick.kennedy@bristol.ac.uk

<sup>\*</sup>Corresponding author *ORCID number:* 0000-0002-2524-6192

Altruism – sacrifice of Darwinian fitness to increase that of a recipient – is easily explained when occurring between close relatives<sup>1</sup>, through an increase in the altruist's inclusive fitness (success at propagating copies of its genes in the population)<sup>2–4</sup>. However, paradoxical forms of altruism arise when individuals divert help from the closest relatives and towards more-distantly-related recipients<sup>5,6</sup>. This seemingly irrational behaviour appears to reduce inclusive fitness<sup>7</sup>.

Social insect colonies are often impenetrable 'fortresses'<sup>8</sup>. However, in some (especially primitively-eusocial) species, between-colony movement ('drifting') by both workers<sup>6,7,9–12</sup> and foundresses<sup>13</sup> can be extensive. Often, opportunities for reproductive parasitism<sup>14</sup> or nest inheritance<sup>13</sup> provide clear direct fitness motives. In other cases, drifters lack obvious opportunities for direct fitness<sup>7,11,12</sup>, and help more-distant relatives than recipients on their home colony. In some primitively-eusocial *Polistes* paper wasps, the extreme extent of drifting by nonreproductive workers has become clear: in the tropics (where *Polistes* originated), 56% of workers within a population were detected at multiple colonies<sup>12</sup>. Drifting workers perform standard cooperative tasks<sup>12</sup> (henceforth, 'cooperative drifting'), creating 'extended kin groups'<sup>11</sup> (networks of cooperating colonies). However, despite primitively-eusocial insects being long-standing models for understanding trajectories to complex eusociality<sup>15</sup>, the driver of cooperative drifting has not been identified.

Three hypotheses have been proposed to explain the evolution of cooperative drifting by non-reproductive primitively-eusocial workers. The *bet-hedging* hypothesis<sup>12</sup> suggests that helping multiple related colonies ('1', '2', and '3' in **Fig. 1a**) avoids the risk of investing in a single colony that could succumb to chance failure or predation. By diversifying investments, workers accept reduced expected inclusive fitness for the benefit of reduced variance in inclusive fitness. The *indirect reciprocity* hypothesis<sup>5</sup> suggests that helping in partner colonies ('2' in **Fig. 1b**) leads to other workers ('3' in **Fig. 1b**) – who may be nonrelatives – helping the home colony ('1' in **Fig. 1b**). For this to be beneficial, it is assumed that help exchanged between neighbours is worth more than help from natal workers due to 'social heterosis': a negative relationship between relatedness and ability to improve colony productivity<sup>16</sup>. The *diminishing returns* hypothesis<sup>7,12</sup> suggests that the marginal benefit provided by a worker diminishes as the number of workers tending brood increases<sup>17,18</sup> ('1' in **Fig. 1c**), an effect first highlighted by Michener in 1964<sup>19</sup>, and helping on related colonies allows effort to be redirected towards brood that are more in need of care ('2' in **Fig. 1c**). There has been no formal comparison of these hypotheses, so we begin by using inclusive fitness theory to assess the plausibility of each.

# Results

Selection for cooperative drifting

First, we show that, under normal conditions, bet-hedging<sup>12</sup> cannot select for cooperative drifting. We derive an explicit mean-risk trade-off for bet-hedging traits for an organism maximizing its inclusive fitness<sup>3</sup> (*Methods 1*):

$$\max_{\mathbf{q}} \left\{ \mathbf{q}^{\dagger} \mathbf{\mu} - v \cdot \rho \sqrt{\mathbf{q}^{\dagger} \mathbf{C} \mathbf{q}} \right\}$$
 subject to 
$$\mathbf{q}^{\dagger} \mathbf{1} = 1$$
 (1)

where the organism must choose the optimal weights (q) to place on different investments, balancing the expectation ( $\mathbf{q}^{\dagger}\mathbf{\mu}$ , where  $\dagger$  denotes transpose) and variance ( $\mathbf{q}^{\dagger}\mathbf{C}\mathbf{q}$ ) in absolute inclusive fitness returns, for particular values of risk-aversion (v) and the correlation  $(\rho)$  between the portfolio's return and whole-population average reproductive success ( $\overline{w}$ ). The vector  $\mu$  contains the offspring production per unit invested (weighted by offspring relatedness), C is the variance—covariance matrix for the investment payoff rates, and 1 is a vector of ones. A 'risky' portfolio has high variance. In Methods 1, we show that the value of risk aversion (v) that maximizes inclusive fitness is the coefficient of variation in  $\overline{w}$ . However, under demographic stochasticity – which is generated by random colony failures – meaningful fluctuations in  $\overline{w}$  will only occur if the population is extremely small (Methods 1), so  $v \approx 0$ . Moreover, in reasonably sized populations, fluctuations in  $\overline{w}$  that do happen due to random colony failure occur independently of fluctuations in the portfolio's returns, so  $\rho \approx 0$ . Since  $v \approx 0$  (and  $\rho \approx 0$ ), the second term of (1) is negligible, so workers should maximise only the expectation of inclusive fitness ( $\mathbf{q}^{\dagger}\mathbf{\mu}$ ). The same logic underpins why bet-hedging against randomly occurring clutch failure is an unlikely explanation for birds distributing eggs over multiple nests<sup>20</sup> or parasitoids distributing eggs over multiple hosts<sup>21</sup>. In summary, the bet-hedging hypothesis for cooperative drifting is mathematically coherent (Fig. 1d) only in an extremely small population and/or in cases where drifting would lead to a very small reduction in expected inclusive fitness  $\mathbf{q}^{\dagger}\mathbf{\mu}$ , circumstances that only rarely arise in the natural world. One way, for instance, is if there were near-negligible differences in relatedness towards brood on home and neighbouring colonies, but *Polistes* drifters face falls in relatedness<sup>12</sup>.

Second, we show that indirect reciprocity<sup>5</sup> is unlikely to explain cooperative drifting. Using simulations, Nonacs<sup>5</sup> has argued that drifters will automatically help only at colonies sharing indirect reciprocator alleles because these are the only colonies willing to accept them. However, there is no reason to expect pleiotropy between these two behaviours, so (unlike Ref.<sup>5</sup>) we allow the emission (y) and acceptance (m) of drifters to evolve independently in a kin-selection model ( $Methods\ 2$ ). As in Ref.<sup>5</sup>, we find that when m is prevented from decreasing whilst y increases, drifting can evolve ( $Methods\ 2$ ). But when this unjustified assumption is relaxed, indirect reciprocity collapses due to the invasion of

free-riding: colonies willingly accept foreign workers  $^{10}$  (m=1) whilst none of their own workers drift (y=0) (Methods 2; Fig. 1e).

Third, we show that diminishing returns<sup>7,12</sup> can be a simple driver of cooperative drifting (*Methods 3*): as the effect of helping closely-related recipients declines, diversion of altruism towards more-distantlyrelated recipients can satisfy Hamilton's rule (increase in benefit b more than compensates for the decrease in relatedness r). In a kin-selection model using the framework of Davies  $et\ al.^{22}$ , we let the probability of brood surviving to maturity (K) be determined by a simple diminishing returns function:  $K=1-(1-h)^T$ , where  $0 \le h \le 1$  is help received from workers. Higher values of T result in stronger diminishing returns. Help (h) received by each brood depends on the worker-to-brood ratio  $\psi$ in their colony, which may vary stochastically between and within colonies through time. For simplicity, we consider neighbouring colonies with equal brood numbers, lying at different points on a line of possible worker-to-brood ratios. Worker-to-brood ratios greater than  $\psi=1$  occur only on colonies in terminal decline, so we focus on  $0 < \psi \le 1$ . Drifting reduces a worker's relatedness to the brood it cares for to the (nonzero) proportions  $d_{\mathbb{Q}}$  and  $d_{\mathbb{C}}$  for female and male brood respectively, relative to raising siblings at home. When a home colony has worker-to-brood ratio  $\psi$ , selection favours a small increase in drifting (y) to a colony with a proportion g of the home colony's workforce when the improvement in indirect fitness from escaping diminishing returns (left-hand side) compensates for reductions in recipient relatedness (right-hand side):

$$\frac{\left(1 - \psi x (1 - z)(g + y)\right)^{T - 1}}{\left(1 - \psi x (1 - z)(1 - y)\right)^{T - 1}} > \frac{4}{\left(3d_{\varphi} + d_{\delta}\right)}$$
(2)

where x is the proportion of females that are behaviourally sterile (not reproductive) and z is the proportion of offspring that are male, so x(1-z) is the proportion of offspring that are workers.

To find candidate evolutionarily stable strategies (ESS) for drifting level y, we set equal the left and right sides of Inequality 2 and rearrange for y. Letting consanguinity to male and female brood be devalued by the same amount ( $d_{\mathbb{Q}} = d_{\mathbb{S}} = d$ ), the ESS drifting level  $y^*$  occurs at:

$$y^* = \frac{1}{1 + \sqrt[T-1]{d}} \left[ 1 - \sqrt[T-1]{d}g - \frac{1 - \sqrt[T-1]{d}}{\psi x (1 - z)} \right]$$
(3)

Given the possibility to drift to related colonies at sufficiently steeper points on an inclusive-fitness-returns curve, workers should tolerate a reduction in recipient relatedness ( $Methods\ 3$ ; Fig. 1f; Extended Data 1,2). The ESS drifting level  $y^*$  increases whenever a gap in worker payoff increases between two related colonies:  $y^*$  increases with higher T (Fig. 1f), higher  $\psi$  and lower g, and for scenarios in which there is a larger workforce (lower z and higher x). For example, a worker may be twice as related to

brood on its home colony as to brood on a partner colony (d=%). Assume the home colony (with  $\psi=1$ ) has twice as many workers as the partner colony ( $g=\frac{1}{2}$ ). With diminishing returns factor of T=4, three-quarters of females developing as non-reproductive workers ( $x=\frac{3}{4}$ ), and a sex ratio of 3:1 female eggs to male eggs ( $z=\frac{1}{4}$ ), the expected equilibrium level of cooperative drifting from the home colony would be for 13.2% of worker effort to be divested to the less-related partner colony ( $y^*=0.132$ ). As workers drift and group sizes consequently change, the marginal payoff of staying on the home colony rises and the marginal payoff of helping on the partner colony falls until  $y=y^*$ , where any further drifting would reduce inclusive fitness.

# Diminishing payoffs in the wild

To investigate plausible ranges within which diminishing returns to cooperation may exist, we used longitudinal field tracking of brood development and worker numbers in a Neotropical paper wasp. Polistes canadensis is a model species in which cooperative drifting is common and can reduce relatedness to recipient brood considerably (e.g., r=0.56 at home vs r=0.19 on partner colonies)<sup>12</sup>. Colonies typically consist of a single queen and behaviourally nonreproductive daughter workers and non-nestmate drifters<sup>12</sup>. As workers emerge, colonies can grow in group size from fewer than 10 females to up to 200 females. In dominance hierarchies below the queen, a small number of high-ranking wasps have an opportunity to inherit the nest on the death of the resident queen<sup>23</sup>; most subordinates perform helping behaviours (including foraging, nest hygiene, brood inspection, nest building, and nest defense)<sup>12,24,25</sup>. The lack of a covering nest envelope allows clear observation of individual brood (**Fig. 2a**) and the opportunity to document whole-colony development through time<sup>26</sup>.

We aimed to assess how a colony's success at producing new adults is associated with numbers of workers and brood on the colony. Across 56 days, we made over a quarter-of-a-million observations of more than 20,000 individually tracked brood cells on 91 wild post-worker-emergence colonies. The stage of brood development in every cell in each colony was recorded repeatedly (*Methods 4*; **Extended Data 3**) and treated as a state in a Markov model (**Fig. 2b,c**). The number of adults on the nest at night (when all workers are present) was counted at intervals over the observation period. This resulted in a dataset of 123,116 state-transitions involving live brood on 85 colonies in 471 colony-observations for which worker number can be predicted by interpolation through the night censuses (*Methods 4*). To investigate the extent to which there may be variation in payoffs within networks of colonies, we asked how the number of workers correlated with colony success at different points of larval development.

Using between-colony variation in the number of workers and worker-to-brood ratio, whilst controlling for within-colony variation and colony state (using the extent of brood-cell emptiness as a

proxy), a Bayesian hierarchical model predicts that colonies with higher worker-to-brood ratios and worker numbers are associated with higher brood development pace and lower brood death. The Markov model predicts that the expected time for a single egg-containing brood cell to produce an adult successfully in colonies with different worker and brood numbers ('expected mean first passage time', eMFPT) declines as more workers tend the brood (Fig. 3a,b; Extended Data 4). Multiplying 1/eMFPT by brood number estimates the expected number of new adults that can be produced per day (whole-colony productivity), which is highest in colonies with many workers and many brood (Fig. 3c,d). The slope of whole-colony productivity with respect to worker number (Fig. 3e,f) then provides a prediction of the marginal increase in the daily number of brood successfully raised associated with each additional worker (i.e., plausibility values for the payoff rate). The 'payoff' in Fig. 3e,f represents an empirical estimate of the benefit b in Hamilton's rule from working on a colony of the given size. An investment in a partner colony 2 is in the inclusive-fitness interest of a worker from a home colony 1 if  $r_2b_2 > r_1b_1$ .

The difference in the predicted payoff from the model suggests that — within the main parameter space occupied by colonies — workers are more valuable (lighter colours in Fig. 3e) at colonies with a relatively 'understaffed' workforce faced with large brood-rearing challenges than at colonies with fewer brood to rear. Future models that consider within-colony dynamics may further explain finer-scale variation in brood development rates (see Extended Data 5–8; SI Table S3). However, the prediction of variation in plausible payoff rates between colonies with different workforce sizes suggests a context in which workers can increase indirect fitness by helping at less-closely-related colonies.

# Scales of competition

To explore the demographic conditions that can sustain cooperative drifting, we use agent-based haplodiploid simulations (*Methods 5*) of a large population of monogynous colonies distributed over a square lattice (where each colony has eight neighbours in a 'Moore neighbourhood'). To model simple colony growth, we assume a 'unit' is a group of workers (e.g., 10 workers), and let colonies produce a new unit every three time-steps (e.g., 3 weeks) until reaching a maximum of 10 units. We assume a saturated environment in which each square on the lattice is occupied by one colony. At each time-step, a randomly chosen 10% of colonies die and are replaced. The foundress of the replacement colony is drawn either from lottery competition among the local eight colonies in the Moore neighbourhood (when female philopatry is assumed) or globally from the whole population (when female philopatry is relaxed). Males compete globally for mating with females in lottery competition. The number of

reproductively-destined offspring that colonies produce is a diminishing-returns function of the helper effort h on the colony,  $1-(1-h)^T$  as above. Colonies with more helpers therefore produce more reproductively-destined females and males, and so have a greater chance in competition for nest sites and mating, respectively. We evolve units' propensity to drift for workers on colonies with more than five units, and show the results of competition by introducing mutant alleles to resident populations at an initial frequency of 5%. After 1,000 time-steps, we record the mutant frequency and plot the average change in frequency over three replicate simulations (*Methods 5*).

Under linear returns, drifting does not invade, regardless of demography (**Fig. 4a-c**). We then consider moderate diminishing returns (T=3) under three conditions: female philopatry and altruism directed at local colonies in the Moore neighbourhood (**Fig. 4d**); female philopatry and altruism directed at partner colonies whose queens are parents or daughters of their own queen (**Fig. 4e**); and female global dispersal with altruism directed at local colonies in the Moore neighbourhood (**Fig. 4f**). (Female global dispersal with altruism directed at the genealogically-close partner colonies is not considered, because global female dispersal makes these pairings vanishingly rare.) Cooperative drifting can invade only when we allow female philopatry and kin-directed altruism (**Fig. 4e**): under this scenario, neighbouring nests develop genealogical links (spatial kin structure), and – when drifters are shared preferentially within these links – cooperative drifting occurs at a more local scale than the spatial scale of competition. *Polistes* wasps often form colony clusters within wider aggregations of tens to hundreds of colonies 11,12,27, a context in which cooperative drifting can evolve by spatial selection under diminishing returns.

#### Discussion

Established accounts of the evolution of eusociality assume actors must choose either to stay as helpers or leave as reproductives<sup>8,28</sup>. Our results suggest that diminishing returns may drive altruists to diversify their help across recipients: workers in some primitively-eusocial societies may increase inclusive fitness by providing altruism to recipients beyond their home colony. Under positive kinship, spatial kin clustering, and diminishing returns<sup>17,29</sup>, worker investments can evolve to become diffusible public goods.

Our model predicts the conditions under which we expect cooperative drifting to have evolved (**Equation 2**, **Fig. 1f**). Intuitively, drifting is more likely when there are stronger diminishing returns (higher T), a larger difference in workforce between nests (smaller g), increased total workforce (higher  $\psi$ ), and a greater capacity to target kin (higher  $d_{\mathbb{Q}}$  and  $d_{\mathbb{Q}}$ ). For simplicity in **Equation 2**, we assume that all colonies have the same sex ratio, but between-colony sex-ratio variation suggests an additional

factor: a colony producing mainly brothers has a reduced worker relatedness to the brood, at which point switching colony may be rational for a worker. In short, drifting offers a simple route to boost inclusive fitness when neighbouring colonies differ in parameters that determine the value of a worker. Differences in worker and brood number arise easily among *P. canadensis* colonies (**Fig. 3a**), which are subject to several sources of stochasticity. These include fluctuations in worker number due to the high attrition rate of foraging workers<sup>12</sup>, frequent loss of brood to parasitoids, presumed loss of brood due to disease (based on workers' hygienic removal of larvae), episodes of queen replacement, and so on. Fluctuations in brood cohort size translate into fluctuations in workforce size once the brood pupate.

Since Michener<sup>19</sup> highlighted diminishing returns in hymenopteran societies in 1964, a number of studies across ants, bees, and wasps have revealed declines in per-capita productivity with rising group size (e.g. <sup>18,30–33</sup>). This so-called 'reproductivity effect' has not proved universal (e.g. <sup>34–36</sup>), but its frequent occurrence leads to 'Michener's paradox': why do apparently partly-redundant helpers exist<sup>26,30</sup>? Previous tests of the reproductivity effect have used snapshots of per-capita productivity. By contrast, we provide a prediction of plausible ranges for the payoffs of cooperation in a primitively-eusocial insect using colony dynamics. Diminishing returns exist, but – through cooperative drifting – workers can mitigate redundancy arising from stochastic variation in worker-to-brood ratios between colonies.

The extent of drifting across primitively-eusocial insects remains to be explored<sup>5,12</sup>. However, the relatively high levels of drifting observed in Neotropical species such as *P. canadensis* contrast with, for example, the European wasp *P. dominula*, which also forms dense colony aggregations<sup>13</sup> but shows high aggression towards neighbours. This difference in social organization may be due to differences in the intensity of diminishing returns (e.g., due to food availability or parasite density). Higher stochastic predation of workers in some species may undermine workers' abilities to track need across nests. Alternatively, drifting may be more likely in the tropics: unlike temperate species in which foundresses establish nests synchronously in the spring, tropical species often establish nests throughout the year<sup>25</sup>, and so nests may be more likely to differ in worker-to-brood ratio. Tropical species may also experience less uncertainty in neighbour relatedness, since nests are more commonly founded by local dispersal from parent nests (simulated in **Fig. 4e**), although kin spatial structure can be reestablished in temperate species by natal philopatry of spring foundresses<sup>37</sup>. Direct comparisons between species with and without cooperative drifting are needed.

Cooperative drifting has also emerged among complex eusocial species. Ant 'supercolonies' exist when nests with multiple queens (polygyny) exchange workers (polydomy)<sup>6,38</sup>. Supercoloniality results in remarkably low-relatedness cooperation, and remains a theoretical challenge. The evolution of supercoloniality is likely to involve informational constraints preventing nepotism<sup>6</sup>, although some

positive relatedness may be maintained by cryptic kin structure<sup>39</sup>. Our results are relevant here: the initial drivers of low-relatedness cooperation are unlikely to have been either bet-hedging by risk-spreading at the expense of the expectation of inclusive fitness (**Equation 1**; **Fig. 1d**) or the reciprocity scenario proposed by Ref.<sup>5</sup> (**Fig. 1e**). In principle, diminishing returns may initially have favoured partial diversion of altruism to more-distantly-related colonies. However, supercoloniality and primitively-eusocial cooperative drifting are not completely analogous. Supercoloniality may have been a relatively small step for ants that had already evolved high within-colony polygyny – and consequently reduced relatedness<sup>40</sup> – for other reasons. Unlike primitively-eusocial wasps, the first step to explaining cooperative drifting in ants is explaining polygyny<sup>41</sup>.

Manipulating colony networks by adjusting worker-to-brood ratio ( $\psi$ ) may offer tests of whether wasps make strategic adjustments to investments (y). Empirical studies are needed to identify whether host workers discern cooperative drifters and adjust acceptance thresholds (m) adaptively<sup>42,43</sup> according to need. Future theoretical work could assess the tension between selfish and cooperative drifting in determining the acceptance of foreign workers. Wasps with high resource-holding potential may exploit the relaxation of nest boundaries to drift for direct fitness (e.g., joining dominance hierarchies on multiple nests to maximise chance of nest inheritance). Models of the mechanisms individual workers might use to distribute their effort would be useful, potentially inspired by resource-use models in foraging theory<sup>44</sup>.

Nonlinear payoffs exert strong effects on social evolution: diminishing returns can limit the tragedy of the commons<sup>45</sup>, promote polymorphic equilibria<sup>46</sup>, and increase sharing in reproductive skew games<sup>47</sup>. However, the extent to which diminishing returns shape investment patterns remains little quantified – despite clear theoretical predictions. A tempting explanation for divestment across recipients is that actors help different recipients in proportion to relatedness (an idea known as the 'proportional altruism' model<sup>48</sup>). This argument commits the 'gamblers' fallacy'<sup>49</sup>: instead, it is always best to invest in the single recipient who carries the highest inclusive fitness returns at any one time<sup>50</sup>. To explain altruism towards more-distant relatives, in the 1980s Altmann<sup>49</sup>, Weigel<sup>17</sup>, and Schulman and Rubenstein<sup>29</sup> highlighted diminishing returns to investment by a *single* individual. Here, we have considered diminishing returns to investment by *multiple* individuals. In both cases, diminishing returns provide a simple explanation for helping multiple recipients, which awaits empirical study in many species. Our results indicate the power of nonlinear fitness effects to shape social organisation, and suggest that models of eusocial evolution should be extended to include neglected social interactions within colony networks.

**Author contributions.** PK, SS, and ANR planned field data collection, and PK and PB collected field data. PK and ADH conducted modelling. PK, NJW, and ANR conducted statistical analysis, and PK, SS, and ANR interpreted the results. PK drafted the manuscript and all authors contributed to its development.

**Data availability.** The transitions data for *P. canadensis* is available as supplementary information.

**Code availability.** The statistical code and individual-based simulation code are both available as supplementary information.

**Competing interest statement.** The authors declare no competing interests.

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#### Figure captions

- Fig. 1 | Three adaptive hypotheses have been proposed for cooperative drifting. (a) Bet-hedging. (b) Indirect reciprocity. (c) Diminishing returns. (d) Cooperative drifting is only favoured by bet-hedging (above dashed line) at extremely small population sizes and when recipients differ only slightly in relatedness to the actor. Equilibrium divestment levels shown ( $Methods\ 1$ ). (e) Cooperative drifting cannot be explained by the indirect reciprocity hypothesis because free-riders invade ( $y \to 0, m \to 1$ ). Arrows show direction of selection. Black quarter-circle is the global attractor ( $Methods\ 2$ ). (f) Diminishing returns can select for cooperative drifting. Stronger diminishing returns and higher relatedness favour more cooperative drifting ( $Methods\ 3$ ). Relative relatedness is d in Equation 3 of main text. For illustration, we plot equal reductions in relatedness for female and male brood ( $d_{\circ} = d_{\circ} = d$ ).  $\psi = 1$ ; g = 0.5; z = 0.75; z = 0.25.
- Fig. 2 | Brood development in the Neotropical paper wasp *Polistes canadensis*. (a) *P. canadensis* colonies allow easy observation of individual brood cells. *Photograph: PK*. (b) Wasp development involves states from egg (state 1) through larvae (2 to 6) to pupa (7) and finally a new adult (definitions of states in Extended Data 3). This can be represented using a Markov model. (c) Baseline transition probabilities between developmental states using an intercepts-only model (*Methods 4*). In using developmental rates to produce a measure of colony productivity as a function of worker behaviour, we set transitions to death as transitions back to egg (state 1), in order to exclude confounding effects from between-queen variation in egg-laying rates.
- Fig. 3 | Brood-rearing rates in *Polistes canadensis*. (a) Expected number of days for an idealised brood cell successfully to produce an adult ('expected mean first passage time', eMFPT) predicted by between-colony effects. Brood size denotes number of prepupal brood (states 1–6). Worker group size denotes interpolated group size using night-time censuses. White dots denote colony observations. To illustrate predictions in the main parameter space, we generated the prediction of eMFPT for 1,000 simulated colonies inside the convex hull set by those colony observations with up to 150 workers and over 40 brood (only 10 colony observations, from six colonies, in the dataset of 471 colony observations used to generate the model, lie outside this range). (b) 95% credible intervals corresponding to *a*. (c) Predicted whole-colony productivity (daily expected number of new adults, given the eMFPT values) is highest in colonies with many workers and many brood. (d) 95% credible intervals corresponding to *c*. (e) Slope of predicted whole-colony productivity with respect to worker group size, representing predicted effect of adding a new worker (the 'payoff' provided to the recipients), an empirical estimate of the potential benefit *b* in Hamilton's rule. (f) 95% credible intervals corresponding to *e*. See *Methods 4*.
- Fig. 4 | Evolution of cooperative drifting in a spatially explicit social haplodiploid simulation. Results of competition from agent-based simulations of the invasion of drifting. Red denotes mutants invading from a starting frequency ( $p_{\rm initial}$ ) of 5% of the population; blue denotes mutants declining. White denotes negligible change in frequency. Convergence-stable strategies are marked by asterisks. Conditions were (a–c) no diminishing returns, T=1, and (d–f) moderate diminishing returns, T=3; (a, d) female philopatry and altruism directed at local colonies in the Moore neighbourhood, (b, e) female philopatry and altruism directed at partner colonies with queens who are parents or daughters of their own queen, (c, f) global female dispersal with altruism directed at local colonies in the Moore neighbourhood. Nonlinear returns drive inter-colony cooperation only under both local female dispersal and kin-directed altruism (panel e).

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#### Methods

Tables of notation and supplementary detail on models are available in the **Supplementary Information**, and supplementary figures are provided in the **Extended Data**.

# 1. Bet-hedging

We consider when an inclusive fitness maximising actor should divide its investments between separate recipients to minimise the risk that its investments will be lost<sup>12</sup>. Notation is summarized in **Table S1** (see **Supplementary Information**). Following Grafen<sup>3</sup>, we start with the Price equation under uncertainty and treat the target of selection as an individual maximand. The expected change in allele frequency due to selection  $\Delta \bar{G}$  — where average reproductive success for the population is  $\bar{w}$ , and  $I_i$  is absolute inclusive fitness — is equal to the covariance over individuals i between the expectation of relative inclusive fitness  $\frac{I_i}{\bar{w}}$  and the individual's genotype  $G_i$ 

$$\mathbb{E}_{\omega}[\Delta \bar{G}] = \text{Cov}_i\left(G_i, \mathbb{E}_{\omega}\left[\frac{I_i}{\bar{w}}\right]\right) \tag{1.1}$$

where expectations  $\mathbb{E}$  are taken across possible states of the population  $\omega \in \Omega$  that may occur. The absolute inclusive fitness is the sum of all effects  $b_{ijt}$  of actor i on the absolute fitness of recipient j (including the actor itself) across each role t (i.e.,  $I_i = \sum_{j}^{N} r_{ij} b_{ijt}$ ). For derivation of Equation 1.1, see the first term on the right-hand side of Eq. 6 in Grafen<sup>3</sup>. Under fitness additivity and frequency-independence, the quantity  $\mathbb{E}_{\omega}\left[\frac{I_i}{\overline{w}}\right]$  is considered to represent a maximand of organismal behaviour<sup>3</sup>: it is a target of selection (as it covaries with genotype) that is under the control of the actor since  $I_i$  is composed of the effects of the actor's behaviours.

The expectation of a ratio of random variables can be approximated by the Taylor series expansion. The Taylor series is an acceptable approximation when  $\overline{w}$  does not rise to greater that  $2\mathbb{E}_{\omega}[\overline{w}]^{51}$ . We expand the maximand to the second order as follows:

$$\mathbb{E}_{\omega} \left[ \frac{I_i}{\overline{w}} \right] \approx \frac{1}{\mathbb{E}_{\omega}[\overline{w}]} \left( \mathbb{E}_{\omega}[I_i] - \frac{\sigma_{\omega}[\overline{w}]}{\mathbb{E}_{\omega}[\overline{w}]} \rho \sigma_{\omega}[I_i] \right) \tag{1.2}$$

where  $\sigma_{\omega}[\cdot]$  denotes standard deviation over the states  $\omega \in \Omega$ , and  $\rho$  denotes the correlation between  $I_i$  and  $\overline{w}$ . The division by  $\mathbb{E}_{\omega}[\overline{w}]$  does not affect the optimal decision, and so we focus on the terms in the brackets.

To make more explicit the quantity under a focal organism's control, we can describe as  $\mathbf{q}$  the vector of investment weights (the proportions of its total resource) that the individual chooses to place on different investments that affect the reproductive success of itself and other individuals. The expectation of the focal individual's absolute inclusive fitness ( $\mathbb{E}_{\omega}[I_i]$ ) is given by  $\mathbf{q}^{\dagger}\mathbf{\mu}$ , where  $\mathbf{\mu}$  is the vector of expected inclusive

fitness payoffs from the different potential options (the expectation over the different states  $\omega \in \Omega$ ). By convention,  $\dagger$  denotes transpose. Likewise, the standard deviation of a focal individual's absolute inclusive fitness appearing in Equation 1.2 ( $\sigma_{\omega}[I_i]$ ) is given by  $\sqrt{\mathbf{q}^{\dagger}\mathbf{C}\mathbf{q}}$ , where  $\mathbf{C}$  is the variance-covariance matrix of the payoffs over the states  $\omega \in \Omega$ , and  $\mathbf{q}^{\dagger}\mathbf{C}\mathbf{q}$  gives the variance (over the states  $\omega \in \Omega$ ) of the individual's absolute inclusive fitness.

The aim for the organism can then be expressed using the following maximand (echoing financial portfolios<sup>52</sup>):

$$\max_{\mathbf{q}} \left\{ \mathbf{q}^{\dagger} \mathbf{\mu} - v \cdot \rho \cdot \sqrt{\mathbf{q}^{\dagger} \mathbf{C} \mathbf{q}} \right\}$$
 (1.3)

where the sum of  $\mathbf{q}$  is one. The coefficient of variation in population average reproductive success (v) is not affected by the organism's choice of  $\mathbf{q}$ . Whereas in economics, individuals have subjective risk aversions, 'risk aversion' in biology is imposed by the environment: a higher v makes individuals more averse to having an inclusive fitness investment portfolio that has correlated fluctuations with population average reproductive success. An environment can have high v or low v, and this is imposed on the organism. The bet-hedging hypothesis implies that an optimisation trade-off exists within  $\mathbf{q}$ , balancing the expectation and the variance in inclusive fitness profit (measured in the absolute number of offspring-equivalents produced).

Equation 1.4 highlights that variance in inclusive fitness ( $\mathbf{q}^{\dagger}\mathbf{C}\mathbf{q}$ ) only matters when the level of stochasticity v (i.e., the coefficient of variation in population average reproductive success  $\overline{w}$ ) and the correlation  $\rho$  are non-negligible. This is generally only true when there is environmental stochasticity that makes a genotype's total reproductive success fluctuate in a correlated fashion between environmental states. For instance, there may be some distinct environmental states when one type of colony does worse, such as big colonies fail when there is a drought. However, this would require a genotype starting from a position of overinvesting workers in one type of colony, such that it may then be able to reduce its variance by diverting some effort towards a different type of colony that doesn't fail in that environmental state.

The bet-hedging hypothesis proposed by Sumner et al.<sup>12</sup> is based on hedging against individual risks to colonies: demographic stochasticity. In general, demographic stochasticity can only generate meaningful fluctuations in  $\overline{w}$  when population (or deme) size is very small<sup>53,54</sup>. Next, we illustrate the effect of population size, which we then plot in **Fig. 1d** of the main text.

To illustrate the bet-hedging hypothesis in a specific example, we switch to a neighbour-modulated perspective<sup>55</sup>. Notation is summarised in **Table S2**. We focus on the effects experienced by recipients due to the actions of related actors. Thus, instead of focusing on the expected relative inclusive fitness  $\mathbb{E}_{\omega}\left[\frac{l_i}{\overline{w}}\right]$  of an actor i, we focus on the expected relative fitness  $\mathbb{E}_{\omega}\left[\frac{w_j}{\overline{w}}\right]$  of a recipient j. The Taylor approximation

allows us to write the condition for selection of a trait of interest in a population undergoing pure demographic stochasticity as:

$$\operatorname{Cov}_{j}\left(G_{j}, \frac{1}{\mathbb{E}_{\omega}[\overline{w}]}\left(\mathbb{E}_{\omega}[w_{j}] - \frac{\operatorname{Cov}_{\omega}[w_{j}, \overline{w}]}{\mathbb{E}_{\omega}[\overline{w}]}\right)\right) > 0$$

$$\therefore \operatorname{Cov}_{j}\left(G_{j}, \mathbb{E}_{\omega}[w_{j}] - \frac{\operatorname{Var}_{\omega}[w_{j}]}{N\mathbb{E}_{\omega}[\overline{w}]}\right) > 0$$

$$(1.4)$$

since  $\text{Cov}_{\omega}[w_j, \overline{w}] = \rho_{w_j, \overline{w}} \text{Var}_{\omega}[w_j]$  and under pure demographic stochasticity<sup>56</sup>  $\rho_{w_j, \overline{w}} = \frac{1}{N}$ 

We consider the following scenario. Workers invest in colony reproductive success (where  $w_j$  is the reproductive success of the colony's queen) with linear returns  $w_j = An_j$ , where  $n_j$  is the number of workers helping at queen j's colony and A is a constant. After this investment period, queens are exposed to random catastrophe (such as a predation of the nest) with independent probability  $\theta$  which reduces their reproductive success to a proportion k of its value. Workers must decide during the investment period whether to invest solely in their home colony, where the queen is the closest related queen, or divest some of their investment to neighbour colonies.

We assume that there are two sizes of colony with equal numbers of brood: colony type 1 and colony type 2. Type 1 and type 2 colonies have a high  $(n_1)$  and low  $(n_2)$ , where  $0 < n_2 < n_1$  number of workers, respectively. For each genotype we let half the colonies be type 1 and half be type 2, and pair each type 1 colony with a type 2 colony. For simplicity, we assume a haploid asexual population (i.e., workers share their queen's allele for the trait of interest). We ask whether workers on a type 1 colony should divest part of their investments to a type 2 colony in order to hedge against the risk of their investments on the home nest being lost to random colony failure. If workers divest from a type 1 colony, they are paired with a foreign queen on a type 2 colony identical at the focal locus with probability  $\alpha$  or a random queen (who may or may not be identical at the focal locus) on a type 2 colony with probability  $1 - \alpha$ . A mutant worker from a type 1 colony with divestment level  $\gamma$  will expend a proportion  $\gamma$  of its effort on the neighbouring type 2 colony and a proportion  $1 - \gamma$  on its own colony. We assume a population monomorphic for a resident strategy  $\gamma$ .

The absolute fitness  $w_2$  of a queen on an  $n_2$  colony carrying a mutant allele y if no catastrophe occurs is:

$$w_2 = A\{n_2 + n_1(\alpha y + (1 - \alpha)[Py + (1 - P)\bar{y}])\}$$
(1.5)

where P is the frequency of mutant allele in the population. The absolute fitness  $w_1$  of a queen on an  $n_1$  colony carrying the mutant allele is as follows if no catastrophe occurs:

$$w_1 = An_1(1 - y) (1.6)$$

There are then four outcomes for any given queen:

	Starts with few workers $(n_2)$	Starts with many workers $(n_1)$	
No catastrophe occurs	$w_2$	$w_1$	
Catastrophe occurs	$kw_2$	$kw_1$	

The probability of each outcome is:

	Starts with few workers $(n_2)$	Starts with many workers $(n_1)$	
No catastrophe occurs	$1-\theta$	$1-\theta$	
	2	2	
Catastrophe occurs	$\theta$	$\theta$	
	2	2	

Accordingly, the variance of w over population states  $\omega$  is:

$$\operatorname{Var}_{\omega}[w] = \frac{1}{2} \left( (1 - \theta)((w_2 - \mathbb{E}_{\omega}[w])^2 + (w_1 - \mathbb{E}_{\omega}[w])^2) + \theta((kw_2 - \mathbb{E}_{\omega}[w])^2 + (kw_1 - \mathbb{E}_{\omega}[w])^2) \right)$$
(1.7)

where:

$$\mathbb{E}_{\omega}[w] = \frac{1-\theta}{2}(w_1 + w_2) + \frac{\theta}{2}k(w_1 + w_2) \tag{1.8}$$

To ask if a small increase in the level of altruism divested to relatives  $(y - \bar{y})$  can invade, we evaluate the derivative of the approximation of expected relative fitness (which covaries with genetic value in Equation 1.5). We set this to equal to zero to find the candidate equilibria  $y^*$ :

$$\frac{\partial \left( \mathbb{E}_{\omega} [w_j] - \frac{\operatorname{Var}_{\omega} [w_j]}{N \mathbb{E}_{\omega} [\overline{w}]} \right)}{\partial y} \bigg|_{y = \overline{y}} = 0$$
(1.9)

As  $y^*$  is too complex to give intuition, we plot  $y^*$  for a range of population sizes (N) and differences in relatedness between the home and partner colony in **Fig. 1d** of the main text (i.e., when 'relative relatedness' is one, there is no difference between a worker's relatedness to the home colony and the partner colony). The region in which divestment to multiple relatives is favoured is very narrow.

# 2. Indirect reciprocity

In this section and the diminishing returns section below, we apply the Taylor-Frank method<sup>55</sup> to a haplodiploid population, using the framework of Davies *et al.*<sup>22</sup>, where x is level of altruistic sterility, z is the sex ratio (proportion male), y is probability of drifting, and m is probability of accepting incoming drifters. An overview and further detail is given in Section 2 of the **Supplementary Information**, with notation in **Table S1** and the consanguinity values for haplodiploids in **Table S2**.

We model the indirect reciprocity hypothesis for drifting between unrelated colonies proposed by Ref.<sup>5</sup> in a haplodiploid population. Ref.<sup>5</sup> argues that a colony's 'willingness to accept drifters' (m) is an honest signal of its cooperative willingness to emit drifters, allowing the evolution of drifting in the absence of genealogical kinship. We follow Nonacs<sup>5</sup> in letting m be the 'tag' used passively to determine whether to drift to a target colony: if the colony will let the worker in, it enters. Accordingly, we look for evolutionarily stable strategy,  $y^*$  and  $m^*$ .

We randomly match each colony with a local unrelated recipient colony. Capital letters denote the mean value of the trait in a social group. We assume that the probability with which females develop as sterile helpers is the same on all colonies (x). The following events occur:

- 1. Of helper females on a focal ('home') colony, a proportion  $Y_f$  will be emitted to a recipient colony as drifters (i.e.,  $Y_f$  is the mean value of the drifting trait in the home colony). Of these, a proportion  $M_r$  are on average accepted (since  $M_r$  is the average level of willingness to accept drifters among members of the recipient colony that receives drifters from the focal colony). Any workers that attempt to drift but are rejected by the neighbour colony return to their home colony and work there.
- 2. Of helper females on a donor colony to the 'home' colony, a proportion  $Y_d$  will be emitted to the home colony as drifters. Of these, a proportion  $M_f$  are accepted.

To include 'social heterosis', we let help have the effect u on an unrelated partner colony and the effect a on the home colony. When u > a, workers can make a bigger contribution on an unrelated partner colony (having unrelated genotypes in the workforce amplifies productivity). When u < a, workers can make a bigger contribution on their own home colony (having unrelated genotypes in the workforce impairs productivity).

Brood developing on the home colony receive *h* units of help:

$$h = [a(1 - M_{\rm r}Y_{\rm f}) + uM_{\rm f}Y_{\rm d}]x \tag{2.1}$$

In this section, we assume linear returns to cooperation, such that K = h, where K is a brood member's probability of successfully developing and h is the help received during its development (see **Supplementary Information** for table of notation).

In Case A, we show that the result of Ref.<sup>5</sup> can be recovered in the improbable scenario where workers that accept incoming drifters are unable to stop themselves from drifting in turn. In Case B, we show that cooperative drifting does not invade when this constraint is removed.

# 2.1 Case A: Constrained pleiotropy between emission and acceptance of drifters

Here, we force m and y into pleiotropy, so that an increase in the trait value of one is accompanied by an increase in the trait value of the other. We assume  $M_f = \beta Y_f$ ,  $M_r = \beta Y_r$ , and  $\overline{m} = \beta \overline{y}$ , where  $\beta$  is a constant  $(\beta > 0)$ . Substituted into Equation 2.1:

$$h = [a(1 - \beta Y_r Y_f) + u\beta Y_f Y_d] x \tag{2.2}$$

In the Supplementary Information, we provide the background to the relative fitness functions. The relative fitness of both male and female brood is:

$$W_{\vec{\mathcal{O}}} = W_{\mathcal{P}} = \frac{x}{\overline{K}} \left( \beta u Y_f Y_d + \alpha (1 - \beta Y_r Y_f) \right)$$
 (2.3)

In the absence of kinship between colonies (as assumed by Ref.<sup>5</sup>), selection favours a small increase in drifting *y* when it increases the fitness of the home brood (sisters and brothers on the home colony), which occurs when:

$$c_{\varphi} \frac{\partial W_{\varphi}}{\partial Y_{f}} p_{\text{sis}} + c_{\mathring{\circlearrowleft}} \frac{\partial W_{\mathring{\circlearrowleft}}}{\partial Y_{f}} p_{\text{bro}} > 0$$
 (2.4)

where  $c_{\uparrow}$  is the class reproductive value for females, which for haplodiploids is  $\frac{2}{3}$ , and  $c_{\circlearrowleft}$  is the class reproductive value for males, which for haplodiploids is  $\frac{1}{3}$ . The terms  $p_{sis}$  and  $p_{bro}$  are haplodiploid consanguinities to sisters and brothers, respectively, which are  $\frac{3}{8}$  and  $\frac{1}{4}$ . Evaluating with both  $Y_f$  and  $Y_r$  at the population average drifting level  $\bar{y}$ , the effect on the fitness of female and male brood in the home nest due to a small increase in drifting by workers from the home nest is:

$$\left. \frac{\partial W_{\Diamond}}{\partial Y_{\mathbf{f}}} \right|_{Y_{\mathbf{f}} = \bar{y}} = \left. \frac{\partial W_{\circlearrowleft}}{\partial Y_{\mathbf{f}}} \right|_{Y_{\mathbf{f}} = \bar{y}} = (u - a) \frac{x \beta \bar{y}}{\bar{K}}$$
(2.5)

Substituting the relevant consanguinities ( $p_{sis}$  and  $p_{bro}$ ) and Equation 2.5 into the condition for selection on drifting (Equation 2.4) gives:

$$(u-a)\frac{x\beta\bar{y}}{3\bar{K}} > 0 \tag{2.6}$$

Since  $\frac{x\beta\bar{y}}{3\bar{K}}$  has a positive real value, drifting is favoured in this scenario if:

$$u > a \tag{2.7}$$

Accordingly, cooperative drifting can invade under the hypothesis proposed by Ref.<sup>5</sup> when increasing the emission of drifters (y > 0) to non-relatives is the unavoidable price of increasing the acceptance of drifters (m > 0) from non-relatives. Due to a constraint forcing pleiotropy between the traits  $(\beta)$ , it is worth paying the price of losing home workers when the incoming non-relatives increase the colony's productivity more than home workers (through social heterosis, u > a).

#### 2.2 Case B: Absence of pleiotropy between emission and acceptance of drifters

We now remove the assumption of pleiotropy, so that *m* and *y* are treated independently. Again assuming that drifters would be unrelated to brood they care for on partner nests as in Ref.<sup>5</sup>, the condition for selection to favour a small increase in drifting is the same as Equation 2.4. The relative fitness of male and female brood is now:

$$W_{\vec{o}} = W_{\varphi} = \frac{x}{\overline{K}} \left( M_{\rm f} u Y_{\rm d} + a (1 - M_{\rm r} Y_{\rm f}) \right) \tag{2.8}$$

Evaluating again with both  $Y_f$  and  $Y_r$  at the population average drifting level  $\bar{y}$ , and with  $M_f$  and  $M_r$  at the population average acceptance level  $\bar{m}$ , the effect on the fitness of female and male brood in the home nest due to a small increase in drifting by workers on the home nest is:

$$\left. \frac{\partial W_{\Diamond}}{\partial Y_{\mathbf{f}}} \right|_{Y_{\mathbf{f}} = \bar{y}} = \left. \frac{\partial W_{\circlearrowleft}}{\partial Y_{\mathbf{f}}} \right|_{Y_{\mathbf{f}} = \bar{y}} = -a \frac{\bar{m}x}{\bar{K}}$$
(2.9)

As long as a recipient colony does not reject incoming drifters ( $\bar{m} > 0$ ), a small increase in drifting by workers leads to a loss in relative fitness for their sibling brood. Again substituting the relevant consanguinities ( $p_{sis}$  and  $p_{bro}$ ) and Equation 2.9 into the condition for selection on drifting (Equation 2.4) gives:

$$a\frac{\overline{m}x}{3\overline{K}} < 0 \tag{2.10}$$

Since  $\frac{\overline{m}x}{3\overline{K}}$  has a positive real value, drifting is favoured if:

$$a < 0 \tag{2.11}$$

That is, if home workers are actively deleterious to the home colony's productivity. Thus if workers provide any benefit to their home colony (a > 0), they should not be emitted as drifters to unrelated colonies.

Selection favours an increase in acceptance (m) of any incoming drifters from other colonies when:

$$c_{\mathcal{Q}} \frac{\partial W_{\mathcal{Q}}}{\partial M_{\mathbf{f}}} p_{\mathbf{sis}} + c_{\mathcal{O}} \frac{\partial W_{\mathcal{O}}}{\partial M_{\mathbf{f}}} p_{\mathbf{bro}} > 0$$
 (2.12)

A small increase by workers in acceptance to the home nest of incoming drifters increases the relative fitness of their brother and sister brood on the home nest:

$$\left. \frac{\partial W_{\Diamond}}{\partial M_{\rm f}} \right|_{M_{\epsilon} = \overline{m}} = \left. \frac{\partial W_{\varnothing}}{\partial M_{\rm f}} \right|_{M_{\epsilon} = \overline{m}} = u \frac{x \overline{y}}{\overline{K}} \tag{2.13}$$

Substituting Equation 2.13 into Equation 2.12 gives the condition for an increase in acceptance (m):

$$\bar{y}\frac{ux}{3\bar{K}} > 0 \tag{2.14}$$

Assuming incoming drifters would contribute positively to productivity (u > 0),  $\frac{ux}{3\overline{K}}$  has a positive real value, and acceptance is favoured if:

$$\bar{y} > 0 \tag{2.15}$$

Selection favours acceptance (m > 0) whenever drifters are being emitted by other colonies  $(\bar{y} > 0)$ . Whilst selection favours the minimisation of drifting to unrelated colonies in all circumstances in which workers are helpful (Equation 2.15), it favours maximum acceptance m of foreign cooperative drifters (i.e., free-riding) (Equation 2.11). Accordingly, willingness to accept drifters (higher m) is not an honest signal of willingness to emit drifters (higher y). In the absence of complex and implausible social insect colony versions of the stabilising mechanisms known to sustain indirect reciprocity (including between-colony monitoring of whole-colony reputation, uncheatable physical greenbeards, or punishment of free-riding colonies by cooperative colonies), the proposed effect of indirect reciprocity<sup>5</sup> cannot drive drifting. We plot the direction of selection in **Fig. 1e** of the main text.

Finally, we note that in principle 'social heterosis' (u > a) between *related* nests could drive drifting by kin selection, rather than by indirect reciprocity: a large benefit that could be provided to distant kin would compensate for their more distant relatedness, and so satisfy Hamilton's rule. However, this would require acutely strong social heterosis at levels unknown in any social insect: high synergies between genotypes would need to ensure that halving relatedness to recipients would more than double the benefit a worker can provide.

#### 3. Diminishing returns

We consider a situation in which colonies differ in their worker-to-brood ratio  $\psi$ , in the range  $0 < \psi \le 1$ . For simplicity, we consider neighbour colonies with equal absolute brood numbers. Variation in worker-to-brood ratio can arise for any reason. For instance, *Polistes canadensis* colonies show considerable variation in worker number for a given brood size (as shown in the horizontal range of the scatter-points of **Fig. 3** in the main text), often through chance worker mortality (estimated at 7% per day<sup>12</sup>). We ask when workers

on a home colony with  $\psi$  should invest in a partner colony that has a proportion g < 1 of worker-to-brood ratio of the home colony. The diminishing returns hypothesis predicts that workers should not invest in other colonies, despite  $\psi$  variation, when the returns from cooperation are linear: the increment in the payoffs of altruism is the same regardless of  $\psi$ . When there are diminishing returns to cooperation, however, a worker on a home colony with high  $\psi$  may experience weaker indirect fitness returns on its home colony than on a neighbour colony.

We assume that  $\psi$  variation between colonies occurs unpredictably from the perspective of an individual. Accordingly, wasps are blind to their colony type when deciding whether to pursue altruistic sterility (worker phenotype) or reproduction. Let the sex ratio (z) and proportion of females that are sterile altruists (x) be common to all nests. We assume that the partner colony will emit no drifters and ask what the equilibrium level of drifting will be from the home colony to the partner colony. The average level of drifting y from the home colony to the partner colony is Y. Both y and Y are evaluated at population average value  $\bar{y}$ .

We consider the extent of drifting that should evolve between two types of nest ('1' and '2'), which differ in their worker-to-brood ratio ( $\psi$ ). Using the framework of Davies *et al.*<sup>22</sup> (described in the Supplementary Information), we let the absolute fitnesses of focal female brood on the home type 1 colony (subscript '1') and the partner type 2 colony (subscript '2') respectively be:

$$w_{0,1} = (1 - x)K_1 \tag{3.1}$$

$$w_{9,2} = (1 - x)K_2 \tag{3.2}$$

The rate of producing reproductives on a home colony of type s is  $K_s = 1 - (1 - h_s)^T$  (see Supplementary Information for details). The total help received by broad on each colony type is:

$$h_1 = x(1-z)\psi(1-Y) \tag{3.3}$$

$$h_2 = x(1-z)\psi(g+Y)$$
 (3.4)

The population average levels of help on colonies of each type in the population:

$$\bar{h}_1 = x(1-z)\psi(1-\bar{y}) \tag{3.5}$$

$$\bar{h}_2 = x(1-z)\psi(g+\bar{y})$$
 (3.6)

The population average relative fitnesses on each nest type, for both male and female brood, assuming colony types 1 and 2 are equally frequent in the population, are (**Supplementary Information**):

$$W_{\vec{O},1} = W_{\mathcal{Q},1} = \frac{2K_1}{\overline{K}_1 + \overline{K}_2} \tag{3.7}$$

$$W_{\vec{O},2} = W_{\mathcal{Q},2} = \frac{2K_2}{\overline{K}_1 + \overline{K}_2} \tag{3.8}$$

where  $\overline{K}_s$  is the population average value of  $K_s$  for colonies of type s.

As above (Equation 2.4), let  $c_{\circlearrowleft}$  be the class reproductive value for females  $(\frac{2}{3})$ . Let  $c_{\circlearrowleft}$  be the class reproductive value for males  $(\frac{1}{3})$ . Selection favours an increase in cooperative drifting from nest type 1 to nest type 2 (y) when the net effect on all potentially-affected parties (sisters, brothers, partner-colony female brood, and partner-colony male brood) leads to an expected increase in the success of a mutant allele for drifting (we assume no effect on self fitness, since drifters are already committed to being behaviourally-sterile helpers):

$$c_{\mathcal{Q}}\left(\frac{\partial W_{\mathcal{Q},1}}{\partial Y}p_{\text{sis}} + \frac{\partial W_{\mathcal{Q},2}}{\partial Y}d_{\mathcal{Q}}p_{\text{sis}}\right) + c_{\mathcal{Q}}\left(\frac{\partial W_{\mathcal{Q},1}}{\partial Y}p_{\text{bro}} + \frac{\partial W_{\mathcal{Q},2}}{\partial Y}d_{\mathcal{Q}}p_{\text{bro}}\right) > 0$$
(3.9)

with traits evaluated at their population average values  $(y = Y = \overline{y})$ , and where  $d_{\updownarrow}$  is the devaluation in consanguinity to female brood on the partner colony (relative to sisters on the home colony) and  $d_{\circlearrowleft}$  is the devaluation in consanguinity to male brood on the partner colony (relative to brothers on the home colony). Although we are focusing on scenarios in which a worker's consanguinity to brood is lower on the partner colony than on the home colony, and hence 'devalued'  $(d_{\updownarrow}, d_{\circlearrowleft} < 1)$ , there can also be scenarios in which a worker is more consanguineous with brood on the partner colony, which may only apply to one sex. For instance, nephews are more consanguineous to a female than brothers are  $(p_{\text{nephew}} = \frac{3}{8} \text{ but } p_{\text{bro}} = \frac{1}{4})$ . If the partner-colony male brood are nephews,  $d_{\circlearrowleft}$  would be greater than 1. If the partner-colony female brood are nieces,  $d_{\Lsh}$  remains below 1 (since  $p_{\text{niece}} = \frac{3}{16} \text{ but } p_{\text{sis}} = \frac{3}{8}$ ).

A small increase in drifting (Y) by workers from the home type 1 nest leads to a loss of relative fitness for the workers' sibling broad on their home nest:

$$\left. \frac{\partial W_{\circlearrowleft,1}}{\partial Y} \right|_{Y=\bar{y}} = \left. \frac{\partial W_{\circlearrowleft,1}}{\partial Y} \right|_{Y=\bar{y}} = -\frac{2\psi T x (1-z) \left(1-\psi x (1-z) (1-\bar{y})\right)^{T-1}}{\bar{K}_1 + \bar{K}_2}$$
(3.10)

A small increase in drifting (Y) by workers from the home type 1 nest leads to an increase in relative fitness for the related broad on the partner type 2 nest to which they drift:

$$\left. \frac{\partial W_{\circlearrowleft,2}}{\partial Y} \right|_{Y=\bar{y}} = \left. \frac{\partial W_{\circlearrowleft,2}}{\partial Y} \right|_{Y=\bar{y}} = \frac{2\psi T x (1-z) \left(1 - \psi x (1-z) (g+\bar{y})\right)^{T-1}}{\bar{K}_1 + \bar{K}_2}$$
(3.11)

Substituting Equations 3.10 and 3.11 into Inequality 3.9, and dividing both sides by  $\frac{\psi T x(1-z)}{6(\overline{K}_1+\overline{K}_2)}$ , gives the condition for selection to favour a small increase in drifting:

$$\frac{\left(1 - \psi x (1 - z)(g + y)\right)^{T - 1}}{\left(1 - \psi x (1 - z)(1 - y)\right)^{T - 1}} > \frac{4}{\left(3d_{\varphi} + d_{\Diamond}\right)}$$
(3.12)

which is Inequality 2 in the main text.

To find candidate evolutionarily stable strategies (ESS) for drifting  $(y^*)$ , we set the left and right sides of Inequality 3.12 equal to one another, and rearrange for y. The ESS level of drifting is given by:

$$y^* = \frac{1}{1 + \sqrt[T-1]{\left(\frac{3}{4}d_{\varphi} + \frac{1}{4}d_{\mathring{o}}\right)}} \left[1 - \sqrt[T-1]{\left(\frac{3}{4}d_{\varphi} + \frac{1}{4}d_{\mathring{o}}\right)}g - \frac{1 - \sqrt[T-1]{\left(\frac{3}{4}d_{\varphi} + \frac{1}{4}d_{\mathring{o}}\right)}}{\psi x(1-z)}\right]$$
(3.13)

which is Equation 3 in the main text (where, for simplicity, we assume that consanguinity to male and female brood is devalued by the same proportion;  $d_{\varphi} = d_{\varnothing} = d$ ). Higher levels of relatedness (higher  $d_{\varphi}$  and  $d_{\varnothing}$ ) and stronger diminishing returns (higher T) select for higher levels of drifting  $y^*$  at equilibrium (Extended Data 1). We plot the equilibria at illustrative values in Fig. 2f of the main text. In Extended Data 2, we plot the candidate equilibria at different values of sex ratio z and female helping x.

# 4. Polistes canadensis payoffs

To obtain empirical measures of productivity in *Polistes canadensis*, we tracked a cohort of developing brood on 91 free-living post-emergence colonies over a 56-day period (from 14th June to 8th August 2016). Colonies were clustered in six aggregations on the north coast of Panama (15.2 ± 7.7 colonies per aggregation, mean ± S.D.). Five aggregations were in clearings between lowland tropical forest and the Panama Canal (former US Army Base Fort Sherman, San Lorenzo National Park, Colón Province) and one aggregation was in a clearing in a mangrove swamp (Galeta Point, Colón Province). We use an observational approach to quantify productivity. Key parameters of brood-rearing can be quantified effectively in unmanipulated colonies, including natural rates of stochastic failure, predation, parasitism, queen turnover, workforce fluctuations, and male production. Associations between brood transition rates and workforce size are correlational, and so we view our results as plausible ranges within which causal effects can exist.

To measure the impact of workers at different developmental states, we split brood into stereotypical categories in a sequence. Each category, and its notation, is listed in **Extended Data 3**. We examined each brood cell at 5-day intervals, using a ladder to access colonies and a flashlight to illuminate each cell. A small number of observations were made on the following morning due to issues with field site access. The current classification of the brood cell was dictated to a second observer, who recorded it on a hexagonal grid of the nest. Accordingly, brood classification was done by a single observer blind to the previous state of the cell.

We censused adult group size by recording total adult numbers at night (8 pm–11 pm, 6–7 censuses per colony across the monitoring period). We used a red light (which wasps are unable to see) to avoid disturbance. Nests that were difficult to observe were counted multiply on a given night and averages taken.

A small number of males emerged and stayed on nests. Adult males observed in daylight surveys during the brood counts on 5-day intervals were used to interpolate male number through time; otherwise, males were assumed to be absent. Female number (henceforth, 'group size') during each 5-day interval was defined as interpolated total number of adults minus interpolated male number. We assumed that any changes in group size between night-censuses occur gradually without sudden jumps, to avoid imposing artificial step changes in the model. We therefore estimated group size during each interval as the mean (across 5 days) of the fitted group sizes generate by a cubic spline interpolation through the night-censuses. For 5-day brood-observation intervals in which the first night census occurred within 1 day of the beginning of the interval, we allowed a limited extrapolation of 1 day in order to approximate mean group size over the 5-day interval. All intervals that would require any other extrapolation of group size were excluded from the analysis below in which group size is used as a predictor. To obtain estimates of the payoff rates, we fit a Markov model (shown in Fig. 3) to the brood development data, asking how colony productivity changes in association with different worker and brood numbers.

#### 4.1 Statistical methods

The (relatedness-weighted) marginal effect of a worker on the development time of a larva is an incomplete measure of indirect fitness payoff. This is because *Polistes* nests experience considerable individual brood death<sup>12</sup>, which returns the cell to the start of the developmental process (once the queen has redeposited an egg). A worker's major contribution might be to prevent inefficiency by minimising the rate of brood death. To accommodate both effects, we therefore analysed brood development as a Markov model, where the target parameter for inference is the expected time for a brood cell to produce a new adult ('time to absorption' of the Markov model<sup>57</sup>). The cell may cycle through repeat deaths before finally producing an adult. We treated the Markov transition matrix as a function of the predictor variables worker number, worker-to-brood ratio, brood-cell emptiness, and interactions. To estimate the transition matrix, we used Bayesian mixed models with colony as a random effect. Using the resulting estimate of the 'time to absorption' for *P. canadensis* colonies of different worker and brood sizes, we then obtained the predicted colony productivity rate, defined as the expected number of new adults produced per unit time. Because we are interested in the marginal effect of each additional worker on brood production, we took as an estimate of Hamiltonian benefit *B* the partial derivative of this rate with respect to worker number, evaluated at the point in demographic parameter space characterising any focal colony.

In principle, workers absconding due to a colony entering the declining phase and productivity falling might generate the appearance of a decline in worker number causing a decline in productivity, whilst in reality the decline in productivity caused the decline in worker number. To mitigate this as far as is possible, we include brood-cell emptiness as a measure of colony state and allow for its interactions.

Because (1) brood states are categorical and (2) longitudinal data were collected in discrete time steps, we give a multinomial likelihood for the number of transitions observed from each development state i into each state j, which provides information on the probability  $\pi_{ij}$  of a brood transition from developmental state i to state j. We model the number of transitions  $\phi_{i \to j}$  (defined over a 5-day interval) where the self-transition  $i \to i$  provides the reference category ( $\phi_{i \to i} = 1$ ) and all other transitions are described by log-linear functions of the predictors  $x_p$  in the vector  $\mathbf{x}$  (i.e.,  $\ln(\phi_{i \to (j \neq i)}) = \alpha_{ij} + \beta_{ij,p} \cdot \mathbf{x}_p$ ):

$$\pi_{ij} = \frac{\phi_{i \to j}}{1 + \sum_{j \neq i}^{J} \phi_{i \to j}} = \frac{e^{(\alpha_{ij} + \beta_{ij,p} \cdot \mathbf{x}_p)}}{1 + \sum_{j \neq i}^{J} e^{(\alpha_{ij} + \beta_{ij,p} \cdot \mathbf{x}_p)}}$$
(4.1)

The predictors include random effects for colony ID (defined below, Equation 4.3). The  $\phi_{i\to j}$  transitions are the entries in the following matrix:

State 1 is egg, states 2-6 are larval states, state 7 is pupa, state 8 is adult (successful pupation), and state 9 is death during development. The final two rows, adulthood (8) and death (9), are absorbing.

The model is solved in discrete time because brood were observed at intervals. Accordingly, brood are free to transition from one state to a state further downstream than the next step in the sequence; they have passed through the transitional states during the 5-day window. Some transitions are not biologically possible during a 5-day window (such as 1 to 6 or 3 to 7), so are not permitted in the Markov transition matrix (represented by a dash). We present two models with increasing complexity. 'Model 1' focuses on the baseline transition rates (i.e., intercepts and random effects only) for the complete dataset of observed transitions between live-brood-containing cells, which allows us to estimate the baseline productivity rate of *P. canadensis* colonies (**Fig. 2c** of the main text).

Next, we present an indicator-variable-selection model ('Model 2') to identify the marginal change in productivity associated with each additional worker at different points in the parameter space typifying P. *canadensis* colonies. We exclude all colony-observations lying outside of the night-census range, and a small number of colony observations with fewer than 10 (2-to-6 state) larvae. These steps focus the payoff model onto a core dataset of 471 colony observations on 85 colonies, comprising 123,116 observed brood transitions (from the total dataset of 168,811 observed brood transitions on 1,027 colony observations on 91 colonies). The model parameters were estimated using a Bayesian approach computed using Markov chain Monte Carlo (MCMC) simulation in JAGS (using rJAGS<sup>58</sup> in R). We confirmed MCMC convergence using potential scale reduction factors (PSRF;  $\hat{R}$ ) for five chains ( $\hat{R}$  < 1.1 for the large majority of parameters). After an adaptation phase of 5,000 iterations and a burn-in of 1,000 iterations, we sampled parameter slope values for 15,000 iterations with a thinning of 4. For the posterior predictive model (described below) to generate the inference about payoff rates using simulated input values for colony worker and brood numbers, we continued running MCMC sampling for a further 10,000 iterations with a thinning of 10.

#### 4.2 Within and between effects

We use a 'within-between' formulation<sup>59</sup>, which follows the 'de-meaning' procedure suggested by Mundlak<sup>60</sup>. We split each time-varying predictor (i.e., predictors whose values can differ between different colony-observations within the same colony) into a 'between-colony' component and a 'within-colony' component. The between-colony component is the mean value  $\bar{x}_C$  of the relevant predictor for the focal colony C (across its colony-observations). The within-colony component is the deviation from this mean in any one colony-observation.

We denote the slope dealing with a between-colony component with the subscript 'b' (e.g.,  $\beta_{ij,\text{workers}|b}$ ) and the slope dealing with a within-colony component with the subscript 'w' (e.g.,  $\beta_{ij,\text{workers}|w}$ ). We also permit  $x_{\text{workers}}$  to interact with the two other predictors:  $\beta_{ij,\text{empty,workers}}$  denotes an interaction between  $x_{\text{empty}}$  and  $x_{\text{workers}}$ , and  $\beta_{ij,\text{ratio,workers}}$  denotes an interaction between  $x_{\text{ratio}}$  and  $x_{\text{workers}}$ . We allow these interactions at both the between-colony and within-colony levels. The random intercept for the transition  $i \to j$  for colony C is  $\beta_{ij,C}$ . Thus, the fully saturated model for the transition from state i to state j (subject to pruning of variables during the indicator-variable-selection process detailed in the next section) is:

$$\ln(\phi_{i\to(j\neq i)}) = \alpha_{ij} + \beta_{ij,C} + \beta_{ij,\text{empty}|w}(x_{\text{empty}} - \bar{x}_{\text{empty},C}) + \beta_{ij,\text{empty}|b}\bar{x}_{\text{empty},C}$$

$$+ \beta_{ij,\text{workers}|w}(x_{\text{workers}} - \bar{x}_{\text{workers},C}) + \beta_{ij,\text{workers}|b}\bar{x}_{\text{workers},C}$$

$$+ \beta_{ij,\text{ratio}|w}(x_{\text{ratio}} - \bar{x}_{\text{ratio},C}) + \beta_{ij,\text{ratio}|b}\bar{x}_{\text{ratio},C}$$

$$+ \beta_{ij,\text{empty},\text{workers}|w}\left((x_{\text{empty}} - \bar{x}_{\text{empty},C})(x_{\text{workers}} - \bar{x}_{\text{workers},C})\right)$$

$$+ \beta_{ij,\text{empty},\text{workers}|b}(\bar{x}_{\text{empty},C}\bar{x}_{\text{workers},C})$$

$$+ \beta_{ij,\text{ratio},\text{workers}|w}\left((x_{\text{ratio}} - \bar{x}_{\text{ratio},C})(x_{\text{workers}} - \bar{x}_{\text{workers},C})\right)$$

$$+ \beta_{ij,\text{ratio},\text{workers}|b}(\bar{x}_{\text{ratio},C}\bar{x}_{\text{workers},C})$$

$$(4.3)$$

The random effect of colony ID is:

$$\beta_{ij,C} \sim \mathcal{N}\left(0, \frac{1}{\tau_{ij}}\right)$$
 (4.4)

where the precision  $\tau$  is drawn from a gamma distribution:

$$\tau_{ij} \sim \text{Gamma}(0.001, 0.001)$$
 (4.5)

#### 4.3 Model averaging

In an indicator-variable-selection procedure, we construct the saturated model (Equation 4.3), and introduce a series of binary indicator variables that 'switch on' or 'switch off' each predictor throughout MCMC<sup>61</sup>. The duration of MCMC time with each coefficient switched on is proportional to the coefficient's marginal likelihood of contribution. For each predictor p's regression slope  $\beta_{ij,p}$  (within the vector  $\beta_{ij,p}$  in Equation 4.1), we can therefore annex a binary coefficient  $g_{ij,p}$  that switches between 0 and 1, and then track the mean of the posterior distribution for  $g_{ij,p}$ . The closer this mean is to 1, the more frequently the corresponding regression slope  $\beta_{ij,p}$  is retained in the model.

For the random effects, we used independent Bernoulli priors for  $g_{ij,p}$  to allow the MCMC sampler to turn the specific random effect on or off directly. For other predictors, we used a product of the independent priors for  $g_{ij,p}$  and  $\beta_{ij,p}$ <sup>62</sup>:

$$f(\beta_{ij,p}|g_{ij,p}) = g_{ij,p}f(\beta_{ij,p}|g_{ij,p} = 1) + (1 - g_{ij,p})f(\beta_{ij,p}|g_{ij,p} = 0)$$
(4.6)

Following Ref.62, we used normal priors for the conditional distribution of  $\beta_{ij,p}$  given  $g_{ij,p}$ .

 $\beta_{ij,p}$  is indistinguishable from zero when  $g_{ij,p}$  switches the predictor off. We model the indicators as Bernoulli random variables with a 50:50 prior ( $g_{ij,p} \sim \text{Bernoulli}(0.5)$ ), representing our starting point of indifference between either including or dropping the slope  $\beta_{ij,p}$ . For interactions, we set a prior of Bernouilli(0.2). To accommodate the within-between formulation, we apply the indicator variables to the whole fixed effect predictor: the predictor cannot be turned on for one component (e.g., the between colony component) and switched off for the other component (e.g., the within-colony component).

The effects on the transition probabilities whose 95% credible intervals do not overlap zero are shown in **Extended Data 3**.

## 4.4 Model checking: residual deviance

The residual deviances  $D_i$  for each multinomial model corresponding to the seven initial states i are given by summing over the residual deviance contributions for each colony-observation k, where  $E_{kij}$  is the observed number of transitions from state i to state j for colony-observation k and  $\hat{E}_{kij}$  is the corresponding expected number of transitions from the fitted model<sup>63</sup>:

$$D_{i} = \sum_{k} 2 \left( \sum_{j} E_{kij} \log \left( \frac{E_{kij}}{\hat{E}_{kij}} \right) \right)$$
 (4.7)

The deviance explained by the fixed effects is presented in Extended Data Fig. S6.

In Extended Data 6–8, we plot residual deviance contributions for each colony-observation against predictors. There is no evidence of structure in the plots: most colony-observations fit relatively well, but some colony-observations exhibit very high deviance. In Supplementary Information Table S3, we provide the mean residual deviances for each of the seven starting state models (residual deviances divided by  $471n_i$ , where  $n_i$  is the number of states into which a cell can transition from starting state i, and 471 is the number of colony-observations. The models for starting states 1 and 6 have high mean residual deviances. Inspection of potential outlier colony-observations (colony-observations with unusually high residual deviance contributions) suggests that these high residual deviances may be partly reflective of episodes of unusually high mortality on colonies, with substantial death of large larvae (starting state 6) and substantial cannibalism of eggs (starting state 1). Future models focusing on predictors of severe mortality episodes and other on-colony dynamics may be required to reduce residual deviance in these scenarios.

# 4.5 Expected time to adulthood

To calculate the worker effect on the transition matrix, controlling for oviposition rate, we submit all estimated transition-to-death probabilities ( $\phi_{i\rightarrow 9} \forall i$ ) to the prediction matrices as transitions to new eggs. This isolates the potential effect of workers (as opposed to the egg-layer) to obtain per-cell efficiency without the confounding effect of variation between queens in the rate at which replacement eggs are laid following the death of larvae.

The expected time to absorption ('expected mean first passage time' in **Fig. 3a** of the main text), in which a brood cell transitions from egg to adulthood ( $\widehat{m}_{1\rightarrow 8}$ ) via intermediate states, can be obtained using the linear algebra for a discrete Markov process via the 'fundamental matrix' method of Kemeny and Snell<sup>57</sup>. Following Grinstead & Snell<sup>64</sup>, we obtain the fundamental matrix **N** by inverting the matrix  $\mathbf{I}_t - \mathbf{Q}$ , where  $\mathbf{I}_t$  is the identity matrix for the transient states and  $\mathbf{Q}$  is a square matrix of transition probabilities between each transient brood state with length equal to the number of transient states (i.e., all states apart from adulthood and death). Accordingly:

$$\mathbf{N} = (\mathbf{I}_{t} - \mathbf{Q})^{-1} \tag{4.8}$$

We solve for **N** for each of the recorded iterations in MCMC separately, in order to sample the posterior predictive distribution using simulated colonies across the parameter space of worker and brood number, with the number of empty cells set to the population-average value. The element  $n_{i,j}$  in **N** is the frequency with which the brood cell is expected to visit state j given a current state i. The vector **t** of times to absorption (using the values at any one recorded iteration of MCMC) is then:

$$\mathbf{t} = \mathbf{N}\mathbf{1} \tag{4.9}$$

where  $\mathbf{1}$  is a column vector of 1s. The ith element of  $\mathbf{t}$  is the duration (in step numbers) from state i to successful production of a new adult.

To obtain the  $I_t - Q$  matrix within MCMC, we employ the between-colony effects, which provide a measure of the quantities of interest (workforce size and worker-to-brood ratio) that abstracts away extraneous between-colony variation (in the random intercepts terms  $\beta_{ij,C}$ ) and within-colony variation that is likely to be confounded by colony ageing (in the within-colony effects). We then derive the predicted whole-colony rate of brood production by converting the per-cell time to absorption to a productivity rate per unit time and scaling this rate by the brood cohort size on each colony (using the total number of eggs to large larvae, states 1–6, a measure of cohort size).

In the main text, we plot the expected mean first passage time in **Fig. 3a** and the estimated whole-colony productivity in **Fig. 3c**.

## 4.6 Indirect fitness payoff calculation

Extracting predictions at the population mean level of cell emptiness (**Extended Data 9**), we plot the posterior predictive distribution for whole-colony productivity using the 1,000 simulated points (in *Mathematica*). That is, to extract the shape of the posterior predictive distribution, we fit the smooth interpolation to 1,000 closely-packed samples monitored in MCMC.

The diminishing returns hypothesis predicts that a worker can maximise inclusive fitness by shifting from a home colony 1 with worker-to-brood ratio  $\psi_1$  and brood number  $t_1$  (where she is related to the brood by  $r_1$ ) to a partner colony 2 with worker-to-brood ratio  $\psi_2$  and brood number  $t_2$  (where she is related to the brood by  $r_2$ ):

$$r_2 \frac{\partial w(\psi, t)}{\partial \psi} \bigg|_{\substack{\psi = \psi_1 \\ t = t_1}} > r_1 \frac{\partial w(\psi, t)}{\partial \psi} \bigg|_{\substack{\psi = \psi_2 \\ t = t_2}}$$

$$(4.10)$$

where she has a larger effect on the production rate w of offspring-equivalents by changing  $\psi_2$  on 2 than she would have by changing  $\psi_1$  on 1. The end result of the Markov model is a model of these partial derivatives (estimates of the payoffs driving the benefit term in Hamilton's rule) for colonies of different worker numbers and brood numbers (**Fig. 3e** in the main text).

#### Individual-based simulation

To simulate the spatial invasion of cooperative drifting under nonlinear returns to cooperation, we consider the evolution of a decision rule in a haplodiploid population on a square lattice in a spatially explicit individual-based simulation. Nodes represent colony sites. Each colony has a single monogamous

reproductive pair of a diploid queen and haploid drone. We simulate a death-birth updating process: each time-step, 10% of nodes are selected, and their resident colonies die. A new colony at each updating node is then founded by a new queen drawn from within the dispersal range for females, with probabilities proportional to the payoffs of each of the surrounding eight colonies. We consider two discrete conditions for the female dispersal range: (1) female philopatry (local dispersal within the Moore neighbourhood [the surrounding eight nodes]); and (2) global female dispersal (dispersal from any distance on the lattice with equal probability).

Drones are selected from anywhere in the lattice<sup>37</sup>, with probabilities proportional to the queen's fitness at each colony on the lattice. Colonies produce female workers as they age: every three time-steps, they gain a unit of workers, until reaching a maximum number of units (set to 10 in the main text results; **Fig. 4**). Queen fitness  $\boldsymbol{w}$  (production of reproductively destined brood) is determined by:

$$w = 1 - (1 - h)^T (5.1)$$

where h is the total worker effort received by the colony rescaled between 0 (no workers) and 1 (maximum number of workers possible). We consider two discrete conditions for diminishing returns: no diminishing returns (T = 1) and moderate diminishing returns (T = 3).

Each worker has a finite amount of help to distribute. We consider selection on a single locus y, which is the probability of drifting by workers on colonies above a threshold worker number (5 in **Fig. 4**). Drifter help is distributed at random within each worker's local helping window. We consider two discrete conditions for the size of the helping window: (1) workers help indiscriminately within the local Moore neighbourhood; (2) workers restrict the helping window only to colonies with immediate genealogical ties to the home colony (defined as the parent colony from which the mother queen came or a daughter colony on which a sister is queen).

Main Text Fig. 4 summarises the role of local cooperation.

## Additional information

Additional detail on methods is provided in the **Supplementary Information**, and statistical and simulation code is provided in the electronic material. Additional figures are provided in **Extended Data**.