¹¹⁰⁵ The Role of Schools in Early Adolescents' Mental Health: ¹¹⁰² Findings From the MYRIAD Study

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Objective: Recent studies suggest mental health in youths is deteriorating. The current policy in the United Kingdom emphasizes the role of schools for mental health promotion and prevention, but little data exist on what aspects of schools influence mental health in pupils. This study explored school-level influences on the mental health of young people in a large school-based sample from the United Kingdom.

Method: Baseline data from a large cluster randomized controlled trial collected between 2016 and 2018 from mainstream secondary schools selected to be representative in relation to their quality rating, size, deprivation, mixed or single-sex pupil population, and country were analyzed. Participants were pupils in their first or second year of secondary school. The study assessed whether school-level factors were associated with pupil mental health.

Results: The study included 26,885 pupils (response rate = 90%; age range, 11–14 years; 55% female) attending 85 schools in the United Kingdom. Schools accounted for 2.4% (95% CI: 2.0%–2.8%; p < .0001) of the variation in psychopathology, 1.6% (95% CI: 1.2%–2.1%; p < .0001) of depression, and 1.4% (95% CI: 1.0%–1.7%; p < .0001) of well-being. Schools in urban locations, with a higher percentage of free school meals and of White British, were associated with poorer pupil mental health. A more positive school climate was associated with better mental health.

Conclusion: School-level variables, primarily related to contextual factors, characteristics of pupil population, and school climate, explain a small but significant amount of variability in mental health of young people. This information might be used to identify schools that are in need of more resources to support mental health of young people.

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Key words: adolescents, mental health, school climate, schools, well-being

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significant proportion of children and adolescents are affected by mental health conditions, with some studies suggesting increased anxiety, depression, and self-injury in young people.^{1,2} Approximately 75% of adults who experience poor mental health in adulthood first experience difficulties before age 18.³ People affected by mental health problems during this developmental window pay a heavy price in terms of poorer educational and occupational outcomes, relationship difficulties, and recurring depression.^{4,5} So it is particularly worrying that evidence suggests worse outcomes in recent cohorts, even before the 2019 novel coronavirus disease (COVID-19) pandemic.⁶

Different aspects of school experience may influence mental health and well-being in young people through various mechanisms (see Figure S1, available online). Some factors, such as the experience of pervasive bullying in the school environment, may directly impact a young person's mental health, while others may act indirectly-for instance, the quality and character of the school as an institution, often referred to as school climate.7 Further-more, some potential influences will be outside the school's control, for example, the socioeconomic profile of the school catchment area, yet may still be important influences on pupils' mental health and therefore could be an indicator of need for additional resources.⁸ Given the long-term and Q6

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near-universal access that education provides, schools are a potentially powerful setting for delivering effective in-117 terventions to support well-being, to prevent mental health 118 119 problems, and to triage identified difficulties.9 Mental 120 health provision in schools is highly variable within as well 121 as between countries and is a current policy focus in the 122 United Kingdom, which traditionally has not had a strong 123 school-based mental health service.¹⁰ 124 125

The limited literature suggests that school has a small but significant influence on pupils' mental health, explain-ing 1%–6% of the variation.^{2,11} For example, the sense of school-connectedness is associated with mental health and educational outcomes.⁷ A relationship between school-level sense of community and the well-being of the pupils has been observed¹¹: young adolescents attending schools with higher levels of bullying are more likely to have poor mental health,¹² while school-level collective efficacy is more strongly related to adolescent alcohol use than neighbourhood-level collective efficacy.¹³

Nevertheless, schools operate in a wider structural or socioeconomic context, with factors such as deprivation directly and consistently affecting mental health.¹⁴ Even though schools may not be able to alter the broader context of the catchment area from which their pupils come, there is some evidence that they can still affect mental health of pupils over and above these powerful structural influences. For example, the UK National Longitudinal Study of Adolescent Health suggested that school-level variables influence symptoms of depression in adolescents over and above structural neighborhood factors.¹⁵ Similarly, a Scot-150 **Q7** tish cohort study that followed subjects from childhood into middle age reported school-level effects on adult self-rated health, after accounting for structural socioeconomic factors.¹⁶ Together, this limited literature suggests that while schools operate in a wider context, they may nonetheless have a specific role to play in the mental health of their students. At minimum, understanding these factors and mechanisms could help target prevention and intervention, using the school as a vehicle for evidence-based programs.⁸

In this study, we aimed to determine the extent to which variability in mental health of pupils is attributable to schools and describe which school-related factors are associated with pupils' mental health, including wider structural socioeconomic factors (urbanity, area-level deprivation), characteristics of the school community (free school meals, special educational needs or disabilities support, ethnicity), and operational features of the school (school size, pupil-toteacher ratio, mixed/single sex, school quality, social and emotional learning [SEL] provision, and school climate). We used a large (N = 26,885) sample of pupils attending 85 secondary schools from the United Kingdom, collecting

data on psychopathology, depression, and well-being using well-established continuous measures.

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METHOD

179 This study is a cross-sectional secondary analysis of baseline 180 data collected as part of the MYRIAD Project, a cluster 181 randomized controlled trial evaluating whether school-based 182 mindfulness training improves mental health of young 183 184 people (ISRCTN Registry reference 86619085).¹⁷ Data 185 used in this study were collected before randomization of 186 the schools and at least 1 year before the delivery of any 187 intervention, and thus the current analysis is not part of the 188 intervention study. The rationale for the trial is explained in 189 the study protocol.¹⁷ Administrative data were linked and 190 191 collected from the 85 UK schools participating in the trial 192 (75 in England, 4 in Northern Ireland, 3 in Scotland, and 3 193 in Wales), 739 teachers, and 26,885 pupils 11-14 years of 194 age who were in their first or second year of secondary 195 school during the 2016-2017 and 2017-2018 academic 196 197 years. The study was approved by the University of Oxford 198 Medical Sciences Division Ethics Committee. 199

We recruited schools (N = 85) in 2 cohorts: pupils provided baseline data in the academic year 2016-2017 (cohort 1; n = 13) or 2017–2018 (cohort 2; n = 72). Participant flow is described in Figure S2, available online, and additional details about study design, recruitment, and procedure are provided in Supplement 1, available online. All mainstream UK secondary schools, including private schools, were eligible if they had a substantive appointed headteacher, had not been judged inadequate in their most recent official inspection (to mitigate any risk for trial implementation), and had a strategy and structure in place for delivery of SEL (which is usually taught in Personal, Social, Health, and Economic Education in England; see Supplement 2, available online).

Three groups of school-level factors were identified: factors that related to the broader school context; characteristics of the school community, and operational features of the school (Figure S1, available online). Measures that were directly comparable across England, Northern Ireland, Scotland, and Wales were selected, where possible; otherwise, measures were mapped to their English equivalent. Pupil-level measures included mental health and demographics.

226 The broader school context represented wider structural 227 socioeconomic factors in the area which the school was 228 located, including whether a school was in a rural or urban 229 230 area, and area-level deprivation (Index of Multiple Depri-231 vation decile rating; see Supplements 1 and 2, available 232 online) obtained by linking to the school's post code. In 233 terms of characteristics of school community, we obtained 234

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the number of pupils in each school who were eligible for free school meals (as an indicator of socioeconomic status), received support for special educational needs or disabilities, and were White British (see Supplement 2, available online). The operational features of the school were the total number of pupils and the pupil-to-teacher ratio for all schools, which were also classified as mixed or single sex. An ordinal variable described overall school quality based on inspection ratings (Office for Standards in Education for England; see Supplement 2, available online), which was analyzed as an ordinal categorical variable (0 = requires improvement; 1 = good; 2 = outstanding). SEL provision was assessed against 16 quality indicators via a semistructured interview with the staff member with overall responsibility for the subject (see Supplement 2, available online). Participating teachers within each school completed 3 subscales from the Alaska School Climate and Connectedness Survey (School Leadership and Involvement, Staff Attitudes, and Respectful Climate) to provide a rating of school climate (data sources and further details are provided in Supplements 1 and 2, available online).

Mental health of pupils (eg, psychopathology, depression, and well-being) was measured with 3 validated selfreport questionnaires: the Strengths and Difficulties Questionnaire (SDQ),18 the Center for Epidemiologic Studies-Depression (CES-D) Scale,¹⁹ and the Warwick-Edinburgh Mental Well-Being Scale (WEMWBS).²⁰ The SDQ is a 25-item questionnaire that assesses psychopathology over the previous 6 months and is validated for use in school-age children. The 5 subscales assess emotional symptoms, conduct problems, hyperactivity/inattention, peer problems, and prosocial behavior. We report a total score (range, 0-40) derived by summing the first 4 subscales, where higher scores indicate higher levels of psychopathology. The CES-D Scale is a 20item questionnaire that assesses depressive symptoms and has been validated for use in adolescents. Each item is rated on a scale from 0 to 3, yielding a total score between 0 and 60, where higher scores indicate more symptoms of depression. The WEMWBS is a 14-item measure assessing mental wellbeing that has been validated for use in adolescents. Each item is scored on a scale from 1 to 5, yielding a total score between 14 and 70 (higher scores indicate greater well-being). Pupils also provided data on their gender (male, female, other/ prefer not to say) and ethnicity (White, Asian, Black, and Mixed and other ethnic minorities [eg, Arab]). Pupils' ages were obtained from school.

Analytic Approach

Multilevel linear regression models were fitted using lme4 in R 3.5.2²¹ to estimate school-level variance in pupils' mental health—psychopathology, depression, and well-being,

which were analyzed separately throughout. We reported the intracluster (intraschool) correlation coefficient (ICC), 294 which is the proportion of the total variance in the outcome 295 296 attributed at the school level. We fitted variance compo-297 nents (empty) multilevel models with no fixed predictors 298 to estimate the ICCs for pupils' mental health. We 299 then fitted multilevel models to estimate the ICCs for pu-300 pils' mental health, while using pupils' gender, age, and 301 ethnicity as predictors to control differences across clusters 302 303 on these individual level variables. The 95% CIs and p 304 values for the ICCs were obtained using nonparametric 305 bootstrapping. 306

We explored whether school factors accounted for 307 any school-level variation in pupils' mental health. First, 308 309 we examined the unique associations between each school 310 factor and pupils' mental health, while accounting for 311 pupils' nesting within schools using multilevel regression 312 models, with random intercepts only. Next, we fitted our 313 3 main multilevel models corresponding to the 3 types of 314 315 school-level factors, as described above and in Figure S1, 316 available online. School-related factors that belonged to 317 the same type were entered as covariates in the same 318 multivariable model. We further adjusted for gender, age, 319 and ethnicity at the pupil level to verify that the asso-320 ciations between school factors and pupils' mental health 321 322 remained stable. We report sensitivity analyses to test for 323 possible differences between pupils who were in their first 324 year of secondary school compared with pupils who were 325 in their second year as well as between pupils scoring 326 above and below cutoff for probable caseness of psy-327 chopathology. Thus, we stratified by year group and 328 329 separately by SDQ caseness (SDQ ≥ 18),²² and we reran 330 the analyses on the different subsamples and descriptively 331 compared them to spot any potential substantial differ-332 ence. We also used a similar approach to run restricted 333 subanalyses for schools in England only (n = 75 schools;334 335 n = 24,842 pupils).

336 To assist the interpretation of results, we grand mean 337 centered all continuous pupil (age) and school factors. 338 Multilevel models were fitted using restricted maximum 339 likelihood estimation, and model assumptions and fit were 340 341 checked via absolute model fit indices (root mean square 342 error of approximation <0.10 and standardized root mean 343 square residual <0.08).²³ We conducted complete case 344 analyses, as there were minimal missing data (range, 0.0%-345 2.8%) (Table S1 and Table S2, available online), and used 346 2-sided contrasts with a significance level of .05. Although 347 348 the study was exploratory, we checked for inflation of type I 349 errors from multiple testing by controlling for the false 350 discovery rate and calculating Benjamini-Hochberg adjusted 351 p values.²⁴ 352

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252	TABLE 1 Characteristics of Sch 24 885	nools (N $=$ 85) a	and Pupils (N $=$
373 354	20,885)		
355	Characteristic	١	/alue
356	School context		
357	Urbanicity, n (%)		
358	Rural	13	(15.29%)
359	Urban	72	(84 71%)
360	Area-level deprivation IMD	5.82	(2 73)
361	mean (SD)	5.62	(2.70)
362	Characteristics of school		
363			
364	Descente a francile	10.01	(0, 2, 2)
365		12.21	(9.33)
366	eligible for free school		
367	meals, mean (SD)		<i>i</i> = = <i>i</i> ,
368	Percentage of pupils	9.99	(5.56)
369	receiving SEND support,		
370	mean (SD)		
371	Percentage of pupils who	76.15	(24.58)
372	are White British,		
373	mean (SD)		
3/4	Operational features of the		
3/5	school		
3/6	Mixed or single sex school		
3//	(n, %)		
270	Mixed	74	(87.06%)
200	Female only	11	(12 94%)
300	Number of pupils	1016 15	(337 02)
382	moon (SD)	1010.15	(337.02)
383	Pupil to too hor ratio	15.02	(1 05)
384	Fupil-to-teacher ratio,	13.72	(1.03)
385			
386	School quality, OFSTED		
387	rating", n (%)		(47.000())
388	Requires improvement	14	(17.28%)
389	Good	47	(58.02%)
390	Outstanding	20	(24.69%)
391	SEL provision quality rating,	11.99	(2.58)
392	mean (SD)		
393	Teacher-rated school	3.94	(0.28)
394	climate, SCCS, mean (SD)		
395	Pupil sociodemographics		
396	Gender, n (%)		
397	Female	14,499	(55.25%)
398	Male	11,201	(42.68%)
399	Other/prefer not to say	543	(2.07%)
400	Age v mean (range)	12 20	(10.90 - 14.73)
401	Ethnicity n (%)	12.20	(10.70 11.70)
402	White British	10 452	(75 18%)
403	Acian	2 7 2 1	(10.45%)
404	Black	2,731 1 /22	(10.43/0)
403		1,40Z	(J.40 /0) /0.000/)
400	ivilxed and other ethnic	2,323	(0.87%)
408	minorities (eg, Arab)		
409	rupil mental health		·· =
410	Psychopathology, SDQ ² ,	11.85	(6.50)
411	mean (SD)		

Normal, n (%)	17,781	(67.60%)	
Borderline, n (%)	3,309	(12.58%)	
High, n (%)	1,657	(6.30%)	
Very high, n (%)	3,554	(13.51%)	
Depression, CES-D ^c ,	13.62	(10.06)	
mean (SD)			
Normal, n (%)	17,844	(67.21%)	
At risk, n (%)	5,910	(22.26%)	
Caseness, n (%)	2,796	(10.53%)	
Well-being, WEMWBS,	49.57	(9.87%)	
mean (SD)			

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Note: Sample size (n) and percentage (%) are given for categorical variables, and mean and SD are given for continuous variables; complete sample (N = 85 schools; N = 26,885 pupils), but number varies owing to missing data. CES-D = Center for Epidemiological Studies-Depression; IMD = index of multiple deprivation; OFSTED = Office for Standards in Education; SCCS = School Climate and Connectedness Survey; SDQ = Strengths and Difficulties Questionnaire; SEL = social and emotional learning; SEND = special educational needs and disability; WEMWBS = Warwick-Edinburgh Mental Well-Being Scale. ^aOFSTED operates in England only. ^bSDQ cutoff points: normal (0–14); borderline (15–17); high (18–19); very high (20–40).²⁵ ^cCES-D cutoff points: low (0–15); at risk of depression (16–27); caseness (28–60).²³

RESULTS

Table 1 presents the characteristics of the sample of schools and pupils. Most schools were in an urban area (85%). Inspection quality ratings suggested that 17% required improvement, 58% were good, and 25% were outstanding. There was, however, considerable variation between schools in terms of pupil ethnicity, levels of pupil eligibility for free school meals, and receipt of support for special educational needs or disabilities. School area-level deprivation also differed markedly between schools, and there was variation between schools in size, pupil-to-teacher ratio, and SEL provision. Eleven (13%) schools were single gender, all of which were girls' schools. Mental health of pupils was in line with national estimates for this age group (range, 10–14 years old).^{19,20,22}

456 A small but statistically significant proportion of the 457 total variance in pupils' mental health was explained at the 458 school level (Table 2). The amount of variance attributable 459 to schools was highest for psychopathology at 2.4% (95% 460 461 CI: 2.0%–2.8%), followed by depression at 1.6% (95% CI: 462 1.2%-2.1%) and well-being at 1.4% (95% CI: 1.0%-463 1.7%). All 3 ICCs were similar after including pupils' in-464 dividual characteristics (gender, age, and ethnicity) 465 (Table 2) as predictors in the model. A sensitivity analysis 466 showed no difference between pupils who were in their first 467 468 year of secondary school compared with pupils who were in 469 their second year or between pupils' scoring above and 470

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Journal of the American Academy of Child & Adolescent Psychiatry Volume ■ / Number ■ / ■ 2021 below cutoff for caseness of psychopathology (Table S3 and Table S4, available online). Restricted analyses for England showed a similar pattern of results (Table S5, available online).

Associations for the 3 types of school-related factors and psychopathology, depression, and psychological well-being in pupils are presented in Table 3 (the unique associations are presented in Table 4). Among school context variables, urban location was positively associated with depression in pupils (regression coefficient [B] = 0.90; 95% CI: 0.05 to 1.74; p = .04), even when adjusting for school area-level deprivation and individual confounders. School area-level deprivation, in contrast, was not associated with psychopathology, depression, or psychological well-being in pupils, suggesting better mental health and well-being among pupils attending schools located in rural areas, regardless of whether the area surrounding the school is affluent or deprived.

In the school community, a higher percentage of free school meal eligibility was associated with higher levels of psychopathology in pupils (B = 0.06; 95% CI: 0.03 to 0.09; p < .001), even while accounting for the percentage of pupils receiving special educational needs or disabilities support and school ethnic composition. A higher proportion of White British pupils in schools was correlated with higher levels of psychopathology (B = 0.02; 95%) CI: 0.01 to 0.03; p < .001) and lower levels of well-being (B = -0.02; 95% CI: -0.03 to -0.01; p = .001), when accounting for the percentage of pupils receiving special educational needs or disabilities support and free school meal eligibility. The association with well-being remained after adjusting for individual-level confounders but was attenuated for psychopathology (B = 0.01; 95%) CI: 0.00 to 0.032; p = .054). There was no association between the percentage of pupils receiving support for special educational needs or disabilities and pupils' mental health.

Among operational features of the school, teacher-rated school climate was the only school-level factor to show as-sociations with mental health of pupils. In schools with a more positive school climate, pupils reported less psycho-pathology, less depression, and greater mental well-being (Table 4). Teacher-rated positive school climate remained associated with lower levels of psychopathology (B = -1.11; 95% CI: -2.19 to -0.03; p = .046) after adjusting for other operational variables (mixed/single sex school, school quality, school size, pupil-to-teacher ratio, and SEL provision) and after adjusting for individual con-founders (Tables 2 and 3). However, the associations be-tween school climate and depression or well-being were attenuated when adjusted for other operational variables and confounders (Tables 2 and 3). Some associations were attenuated when using p values adjusted for multiple testing (eg, school urbanity and higher depression), but differences were minimal (Table 5). Results also did not significantly change when restricting the analyses to En-gland only (see Table S6, available online). The only potentially meaningful difference was that school size was negatively associated with higher levels of depression in English schools, after controlling for individual characteristics.

To assess whether these relationships were influenced by how long pupils had been in the school, we compared pupil year groups (eg, pupils in their first year who had recently joined the school and pupils in their second year who had typically been immersed in the school culture for 12 months longer). We found no evidence to suggest that there were systematic differences in school-level variance across these 2-year groups.

DISCUSSION

Given the increasing recent focus of policy makers and researchers on the role of schools in mental health of young

TABLE 2 Intraclass Correlation Coefficients for School-Level Variance of Pupils' Mental Health

		N		Unadjusted mod	els		N	Adju age	sted models for , gender, and et	pupil's hnicity
Pupil's mental health	Pupils	Schools	ICC	(95% CI)	р	Pupils	Schools	ICC	(95% CI)	р
Psychopathology, SDQ	26,303	85	0.024	(0.020 to 0.028)	< .0001	26,127	85	0.022	(0.017 to 0.026)	< .0001
Depression, CES-D Scale	26,549	85	0.016	(0.012 to 0.021)	< .0001	26,078	85	0.015	(0.011 to 0.018)	< .0001
Well-being, WEMWBS	26,463	85	0.014	(0.010 to 0.017)	< .0001	26,073	85	0.014	(0.010 to 0.017)	< .0001

Note: Multilevel models are based on complete case analysis; total sample (N = 85 schools; N = 26,885 pupils), but number varies owing to missing data. CES-D = Center for Epidemiologic Studies-Depression; ICC = intraclass correlation coefficient; SDQ = Strengths and Difficulties Questionnaire; WEMWBS = Warwick-Edinburgh Mental Well-Being Scale.

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TABLE 3 Results From Multilevel Models With Random Intercepts Showing Grouped Associations Between Different Types of School Factors and Pupils' Mental

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		Psy	chopatl	hology (SDC	2)			Depress	ion	CES-D Scal	e)		Well-being (WEMWBS)						
	U	nadjusted models		Adju: fc age and	sted model or pupil's e, gender, d ethnicity	s	- Una m	djusted odels		Adjust for age, and	ed models pupil's gender, ethnicity		Una	djusted odels		Adjust for pi gen et	ed models upil's age, der, and hnicity		
School		(95%			(95 %			(95 %			(95 %			(9 5%			(95 %		
factors Broader school	Coefficie	nt Cl)	Р	Coefficien	t CI)	P	Coefficient	CI)	р	Coefficient	t CI)	р	Coefficient	CI)	р	Coefficient	: CI)	p	
context																			
Urban vs rural	0.36	(-0.29 to 1.01)	.29	0.49	(-0.12 to 1.10)	.12	0.90	(0.05 to 1.74)	.040	0.89	(0.09 to 1.69)	.032	-0.65	(-1.44 to 0.14)	.11	-0.73	(-1.51 to 0.05)	.07	
Area-level deprivation	-0.07	(-0.15 to 0.02)	.13	-0.08	(-0.16 to 0.00)	.055	-0.06	(—0.17 to 0.05)	.30	-0.06	(-0.17 to 0.04)	.26	-0.01	(-0.11 to 0.10)	.87	0.00	(-0.10 to 0.10)	.99	
of school community																			
Pupils eligible for free school meals	0.06	(0.03 to 0.09)	< .001	0.06	(0.03 to 0.09)	< .001	0.04	(0.00 to 0.09)	.05	0.05	(0.01 to 0.09)	.011	-0.03	(-0.06 to 0.01)	.17	-0.04	(-0.07 to 0.00)	.041	
(%) SEND support (%)	-0.01	(-0.06)	.70	0.00	(-0.05 to	.89	-0.04	(-0.11 to	.26	-0.03	(-0.09 to	.36	0.01	(-0.05 to	.63	0.01	(-0.05 to	.86	
Ethnicity of pupils (%): White	0.02	(0.01 to 0.03)	< .001	0.01	(0.00 to 0.02)	.054	0.01	(-0.01 to 0.02)	.33	0.01	(0.00 to 0.02)	.10	-0.02	(-0.03 to -0.01)	.001	-0.02	(-0.03 to -0.01)	.00	
Operational features of																			
Mixed or single-sex	-0.01	(-0.77 to 0.75)	.98	0.00	-0.73 to 0.73)	.99	0.80	(-0.22 to 1.82)	.13	-0.16	(—1.16 to 0.84)	.76	0.01	(-0.95 to 0.97)	.99	0.69	(—0.25 to 1.63)	.15	
school School	-0.13	(-0.66	.62	-0.04	(—0.53 to	.87	0.02	(-0.69 to	.97	0.09	(-0.60 to	.80	0.40	(-0.27 to	.24	0.27	(-0.35 to	.40	
quality		to 0.40)			0.45)			0.72)			0.77)			1.06)			0.90)		
School size (per 100	-0.06	(-0.14 to 0.02)	.15	-0.06	(-0.13 to 0.02)	.16	-0.11	(-0.22 to 0.01)	.07	-0.10	(-0.22 to 0.02)	.08	0.03	(-0.06 to 0.13)	.53	0.03	(-0.07 to 0.13)	.60	
Pupil-to-teacher ratio	-0.06	(-0.19 to 0.08)	.44	-0.06	(-0.19 to 0.08)	.40	-0.05	(-0.25 to 0.14)	.58	-0.08	(-0.28 to 0.12)	.40	0.00	(-0.18 to 0.17)	.98	0.04	(-0.14 to 0.21)	.69	
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tors Coefficient CI) p Coe	oefficient	CI)	Coefficient	c) t	p Coefficié	ant CI)	p Coe	fficient	cI)	Coefficie	C)	٩
EL provision 0.00 (-0.10 .92 -	-0.01 (-().09 to .8	3 – 0.02	(-0.14 to	71 -0.02	? (-0.13 to	.81	·0.05 (-	-0.17 to .4	1 -0.04	(-0.16 to	49
to 0.09)		0.07)		0.09)		0.10)			0.07)		0.08)	
eacher1.11 (-2.19 .046 -	-1.22 (-2	2.22 to .02	20 -1.19	(-2.64 to .	11 -1.20) (-2.61	.10 ().58 (-	-0.77 to .4	0.69 0	(-0.60 to	8.
rated to -0.03)	·	-0.22)		0.26)		to 0.21)			1.94)		1.99)	
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SCHOOL INFLUENCES ON YOUTH MENTAL HEALTH

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ople,^{9,10} we examined the extent to which variation in mental health of young people could be explained by variables operating at the school level in current secondary schools in the United Kingdom. We considered wider structural socioeconomic factors, characteristics of the school community, and operational features of the school. We used data obtained from a sample of 26,885 pupils attending 85 schools from across the United Kingdom.

Consistent with the limited previous research,^{2,6,25} we found that schools accounted for only 1.4%–2.4% of the variability in mental health of early adolescents. Several factors explained this between-school variability; most related to the broader school context and characteristics of the pupil population, rather than operational features of the school. Specifically, schools in urban locations, with a greater proportion of adolescents eligible for free school meals and with more White British pupils, were attended by pupils with poorer mental health.

Urban living is associated with greater income 786 787 inequality, familial isolation, and exposure to substance 788 abuse, violence, and crime as well as lower community 789 cohesion, all of which are related to the higher prevalence of 790 mental health problems often detected in urban pop-791 ulations.²⁶ There is similarly a long-established relationship 792 between socioeconomic adversity and poor childhood 793 mental health.^{4,5,27} The mechanisms by which deprivation 794 795 influences mental health in childhood are multifaceted and 796 incompletely understood, but likely involve parental mental 797 health, family function, nutrition, and sleep, among 798 others.²⁷ The increase in mental health inequalities seen in 799 the 21st century in higher-income countries, particularly in 800 801 relation to emotional problems, is likely to be exacerbated 802 by the disproportional impact of COVID-19 on youths and 803 families and facing debt and financial strain.^{28,29} Further-804 more, socioeconomic and health inequalities may be even 805 wider in urban areas²⁶ and are anticipated to increase as a 806 result of the COVID-19 pandemic.^{27–29} A public mental 807 808 health approach that encompasses community as well as 809 school mental health is essential to prevent further deteri-810 oration in the mental health of children and adolescents. 811

The finding that children attending schools with a 812 higher proportion of White pupils had poorer mental 813 814 health than children in schools with more ethnically diverse 815 pupil populations is surprising. Earlier studies from the 816 United Kingdom suggest that young people from ethnic 817 minorities had a higher prevalence of mental health con-818 ditions,⁴ but the results of the present study echo recent 819 large mental health surveys of children and adolescents in the United Kingdom.^{5,29,30} Recent austerity policies in the 820 821 822 United Kingdom have resulted in drastic reductions in 823 support for children, families, and schools, which were 824

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TABLE 4 Unique Associations From Multilevel Models With Random Intercepts Between School Factors and Pupil's Mental Health

		Psych	opath	ology (SDQ	2)			Depress	ion	(CES-D Scale	e)		Well-being (WEMWBS)					
School	Unadj	usted mode	els	Adjuste pupil's and	ed models f age, gende l ethnicity	for er,	Unadjus	ted model	s	Adjusted pupil's a and	d models fo age, gender ethnicity	or *,	Unadjus	ted model	5	Adjusted pup gender, a	l models fo il's age, and ethnic	or
factors	Coefficien	t (95% CI)	р	Coefficient	: (95% CI)	p	Coefficient	(95% CI)	р	Coefficient	: (95% CI)	р	Coefficient	(95% CI)	р	Coefficient	(95% CI)	p
Urban vs rural	0.49	(-0.14 to 1.12)	.13	0.64	(0.05 to 1.24)	.037	1.02	(0.20 to 1.83)	.017	1.01	(0.24 to 1.79)	.012	-0.63	(-1.39 to 0.13)	.11	-0.73	(-1.48 to 0.02)	.06
Area-level deprivation	-0.08	(-0.16 to 0.00)	.06	-0.10	(-0.18 to -0.02)	.018	-0.09	(-0.20 to 0.02)	.11	-0.09	(-0.20 to 0.01)	.09	0.01	(-0.09 to 0.12)	.79	0.03	(-0.08 to 0.13)	.62
Pupils eligible for free school meals (%)	0.03	(0.01 to 0.06)	.016	0.03	(0.01 to 0.06)	.010	0.02	(-0.02 to 0.05)	.29	0.02	(—0.01 to 0.05)	.26	0.00	(-0.03 to 0.03)	.89	-0.01	(-0.04 to 0.02)	.65
SEND support (%)	0.02	(-0.02 to 0.07)	.32	0.02	(-0.02 to 0.07)	.28	-0.01	(-0.07 to 0.05)	.75	0.00	(-0.06 to 0.06)	.94	-0.01	(-0.06 to 0.05)	.83	-0.01	(-0.07 to 0.04)	.65
Ethnicity of pupils (%): White	0.01	(0.00 to 0.02)	.048	0.00	(-0.01 to 0.01)	.63	0.00	(-0.01 to 0.01)	.77	0.01	(-0.01 to 0.02)	.42	-0.02	(-0.03 to -0.01)	.004	-0.01	(-0.02 to 0.00)	.032
Mixed or single sex school	-0.31	(-0.99 to 0.37)	.37	-0.19	(-0.85 to 0.47)	.57	0.61	(-0.28 to 1.50)	.18	-0.24	(— 1.11 to 0.63)	.59	0.38	(-0.44 to 1.20)	.37	0.95	(0.14 to 1.75)	.024
School quality	-0.48	(-0.83 to -0.13)	.009	-0.41	(-0.75 to -0.07)	.019	-0.32	(-0.80 to 0.16)	.20	-0.45	(-0.90 to -0.01)	.06	0.55	(0.12 to 0.97)	.014	0.61	(0.19 to 1.02)	.005
School size (per 100 pupils)	-0.06	(-0.13 to 0.00)	.071	-0.06	(-0.13 to 0.00)	.054	-0.10	(-0.19 to -0.01)	.035	-0.09	(-0.18 to -0.01)	.032	0.03	(-0.05 to 0.12)	.43	0.03	(-0.05 to 0.11)	.51
Pupil-to- teacher ratio	-0.1	(-0.23 to 0.03)	.14	-0.09	(-0.22 to 0.03)	.16	-0.09	(-0.27 to 0.08)	.31	-0.09	(-0.26 to 0.08)	.29	0.04	(-0.12 to 0.20)	.62	0.06	(-0.10 to 0.22)	.46
SEL provision	-0.02	(-0.11 to 0.07)	.67	-0.02	(-0.10 to 0.07)	.72	-0.01	(-0.13 to 0.11)	.84	-0.02	(-0.13 to 0.10)	.75	-0.04	(-0.15 to 0.07)	.49	-0.03	(-0.13 to 0.08)	.65
Teacher- rated SCCS	-1.48	(-2.27 to -0.70)	< .001	-1.35	(-2.10 to -0.59)	< .001	-1.22	(-2.30 to -0.13)	.030	- 1.45	(-2.47 to -0.44)	.006	1.31	(0.32 to 2.29)	.011	1.50	(0.54 to 2.47)	.003

Note: Estimates are based on complete case analyses; total sample (N = 85 schools; N = 26,885 pupils), but N varies owing to missing data. CES-D = Center for Epidemiologic Studies-Depression; SCCS = School Climate and Connectedness Survey; SEL = social and emotional learning; SEND = special educational needs and disability; SDQ = Strengths and Difficulties Questionnaire; WEMWBS = Warwick-Edinburgh Mental Well-Being Scale.

TABLE 5 Results From Multilevel Models With Random Intercepts Showing Grouped Associations Between Different Types of School Factors and Pupils' Mental Health Using Adjusted p Values for Multiple Comparisons

	Psychopathology (SDQ)								De (CE	pression S-D Scale)				Well-being (WEMWBS)						
	l	Jnadjusted models		Ad	ljusted mode for pupil's age, gender, and ethnicity	els ,	Ui	nadjusted models		Adjusted ag an	models for e, gender, d ethnicity	pupil's		Jnadjusted models		Ad	djusted mode for pupil's age, gender, and ethnicity	els		
School factors	Coeffi	(95 % CI)	B-H	Coeffi	(95% CI)	B-H	Coeffi	(95% CI)	B-H	Coeffi	(95% CI)	B-H	Coeffi	(95% CI)	B-H	Coeffi	(95% CI)	B-H p		
Broader school context	cione				.,	P	cione	0.,	٢		0.)	٢	cione	0.,	٢	cient	0.,	2.1.9		
Urban vs	0.36	(-0.29	.40	0.49	(-0.12	.20	0.90	(0.05 to	.080	0.89	(0.09	.065	-0.65	(-1.44	.19	-0.73	(-1.51	.13		
rural		to 1.01)			to 1.10)			1.74)			to 1.69)			to 0.14)			to 0.05)			
Area-level	-0.07	(-0.15	.21	-0.08	(-0.16	.10	-0.06	(-0.17	.41	-0.06	(-0.17	.37	-0.01	(-0.11	.94	0.00	(-0.10)	.99		
deprivation Characteristics of school community		to 0.02)			to 0.00)			to 0.05)			to 0.04)			to 0.10)			to 0.10)			
Pupils eligible	0.06	(0.03	< .001	0.06	(0.03	< .001	0.04	(0.00	.10	0.05	(0.01	.023	-0.03	(-0.06	.27	-0.04	(-0.07	.081		
for free school meals (%)		to 0.09)			to 0.09)			to 0.09)			to 0.09)			to 0.01)			to 0.00)			
SEND support (%)	-0.01	(-0.06 to 0.04)	.81	0.00	(-0.05 to 0.04)	.96	-0.04	(—0.11 to 0.03)	.37	-0.03	(-0.09 to 0.03)	.48	0.01	(-0.05 to 0.07)	.75	0.01	(-0.05 to 0.06)	.94		
Ethnicity of pupils (%): White	0.02	(0.01 to 0.03)	.002	0.01	(0.00 to 0.02)	.10	0.01	(-0.01 to 0.02)	.45	0.01	(0.00 to 0.02)	.18	-0.02	(-0.03 to -0.01)	.002	-0.02	(-0.03 to -0.01)	.010		
Operational features of the school																				
Mixed or single-sex school	-0.01	(—0.77 to 0.75)	.99	0.00	(-0.72 to 0.71)	.99	0.80	(-0.23 to 1.82)	.22	-0.16	(— 1.17 to 0.85)	.86	0.01	(-0.95 to 0.97)	.99	0.69	(-0.24 to 1.62)	.25		
School	-0.13	(-0.65	.74	-0.04	(-0.53	.94	0.02	(-0.69	.99	0.09	(-0.60	.90	0.40	(-0.26	.36	0.27	(-0.36	.51		
quality		to 0.39)			to 0.45)			to 0.72)			to 0.77)			to 1.06)			to 0.91))		
School size (per 100 pupils)	-0.06	(-0.14 to 0.02)	.25	-0.06	(-0.13 to 0.02)	.26	-0.11	(-0.22 to 0.01)	.12	-0.10	(-0.21 to 0.01)	.14	0.03	(-0.07 to 0.14)	.64	0.03	(-0.07 to 0.13)	.72		

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SCHOOL INFLUENCES ON YOUTH MENTAL HEALTH

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TABLE 5 Col	ntinued																	
		Ľ	'sycho (S	patholog) (DQ)	>				Del (CES	pression -D Scale)					Wel (WE	l-being MWBS)		
	5 	nadjusted models		ara ∀	justed models for pupil's ige, gender, nd ethnicity	<u>ه</u>	5 [*]	adjusted nodels		Adjusted ag	models for e, gender, d ethnicity	pupil's		nadjusted models		Adjus fo age and	ted models r pupil's · gender, ethnicity	
School factors	Coeffi cient	(95 % Cl)	Ha a	Coeffi	(95% CI)	불	Coeffi cient	(95% CI)	Ha a	Coeffi cient	(95% CI)	Ha a	Coeffi	(95% CI)	Ha a	Coeffi cient	(95% CI)	H-8
Pupil-to-	-0.06	(-0.20	.55	-0.06	(-0.19	r 10	-0.05	(-0.25	- 1-	-0.08	(-0.27	r 12	0.00	(-0.18	6 6.	0.04	(-0.14	1 60.
teacher		to 0.09)			to 0.08)			to 0.14)			to 0.11)			to 0.18)			to 0.21)	
ratio SEI	1001	(-010	07	1001	(-0.10	6		15	c a	-002	(-014	00	-0.05	(-017	۲ ۲		15	09
provision	5	to 0.09)		0.0	to 0.08)	1	20.0	to 0.10)	2	20.0	to 0.11)	2	0.00	to 0.07)	<u>;</u>	5.00	to 0.07)	3
Teacher-rated	-1.11	(-2.18	60.	-1.22	(-2.22	.041	-1.19	(-2.64	.19	- 1.20	(-2.6	.18	0.58	(-0.77	.51	0.69	(-0.61	.41
SCCS		to -0.04)			to -0.22)			to 0.26)			to 0.21)			to 1.94)			to 1.99)	
Note : Estimates false discovery r social and emot.	are based ate from m ional learnii	on complet iultiple test ng; SEND =	te case ing. B-t = speci	analyses; 1 H = Benja al educati	total sample (N imini-Hochberç onal needs anc	V = 85 ; g; CES d disab	schools; N -D = Cen ility; SDC	I = 26,885 ter for Epi = Streng	pupils, idemio ths and), but N vari logic Studi 1 Difficultie:	ies owing to I es-Depressic s Questionné	missing >n; SCC ∍ire; WE	t data. B-ł SS = Schc EMWBS =	H adjusted _F ool Climate : Warwick-E	o value and C Edinbui	is are preser Connectedn rgh Mental	nted to cont ess Survey; ' Well-Being	rol for SEL = Scale.

previously less accessed or accessible to ethnic minorities.³¹ Young people from ethnic minorities may therefore have been less adversely affected by these policies. In addition, there is some evidence that psychological distress may be related to ethnic density. Specifically, there could be a possible beneficial effect of more culturally diverse environments for minority students, but majority students seem to be insensitive to this effect.^{32,33} Finally, the meaning of ethnicity varies greatly with culture, time, and geography, and our findings raise interesting questions about the roles of ethnic diversity and ethnic minority status as influences on mental health of pupils, which require further empirical study.

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The only operational, and thus obviously tractable, feature of schools associated with mental health of young people was teacher-rated school climate. Researchers are increasingly encouraged to define school climate as a construct that encompasses school engagement, safety, and environment, both physical and social.³⁴ School climate predicts key educational outcomes⁷ as well as mental health⁷ and well-being¹² of both staff and pupils.³⁵ A recent sys-tematic review of school climate interventions concluded that interventions aiming to promote social-emotional learning and school-wide positive behavior programs seemed more effective than those focusing on bullying, community development, or teachers' working condi-tions.³⁵ However, few of the 18 experimental studies detected were sufficiently methodologically rigorous, and the outcome of primary interest was perception of teachers and pupils of school climate. Another systematic review concluded that there was a clear association between school climate and pupils' mental health, but as most of the 48 studies were observational and cross-sectional, we cannot claim a causal relationship.³⁶ The authors also suggest that future research should pay greater attention to the compo-nents that comprise both constructs, well-being and poor mental health, and school connectedness, safety, academic environment, and peer relationships and examine how these interact.

As suggested, theory-driven studies are needed that follow children over several years to examine how broader school context (eg, deprivation), school characteristics (eg, ethnic composition), school operational features (eg, school climate), and pupils' individual factors (eg, psychopathology) interact to shape the trajectory of mental health of young people over time (Figure S1, available online).³⁶ Such frameworks could also be used to examine how SEL and targeted interventions may be more or less effective in certain contexts, schools, and subpopulations of pupils. In this sense, studies should ideally be designed to enable inferences about causality that can shape both policy and intervention development.

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While the direct influence of schools on mental health seems to be small, this does not negate schools as a setting in 1179 1180 which mental health can be improved via universal and 1181 targeted interventions. Furthermore, these small school-level 1182 effects may translate into more significant impacts if the 1183 substantial future health, economic, and societal costs of 1184 poor mental health in adolescence were modeled.^{4,6,37} 1185 Indeed, there is a growing evidence base that school-level 1186 1187 interventions can enhance resilience and functioning of 1188 young people, and for young people living in deprived areas, 1189 such interventions may be particularly important.^{1,35} Pro-1190 spective interventional research is needed to explore how 1191 broader contextual and school variables interact with in-1192 terventions to effect changes in mental health of young 1193 people during key developmental windows.^{5-9,12–17,36,38} 1194 1195 This is something we are doing in our larger MYRIAD 1196 study,¹⁷ which is collecting data from these schools over 2 1197 years so that we will be able to examine the associations over 1198 time between the broader school context, school charac-1199 1200 teristics and operational features, and mental health and 1201 well-being of young people. 1202

Regarding study limitations, we recognize that our 1203 sample excluded schools that inspections had classified as 1204 inadequate or that had no SEL strategy. The inclusion of 1205 1206 these poorly functioning schools might have increased the 1207 proportion of variation in pupils' mental health attributable 1208 to the school level. Schools were representative of schools 1209 across the United Kingdom, but these were schools that had 1210 demonstrably good Personal, Social, Health, and Economic 1211 1212 Education and participated in a trial. We included private 1213 schools, but in the United Kingdom, these institutions serve 1214 only 5%-7% of the population, an insufficient number to 1215 support a subgroup analysis. Future studies should over-1216 sample from uncommon types of schools to study if 1217 different types of provision may differ in their influence on 1218 1219 mental health.

1220 The usual caveats of how populations vary across a 1221 country apply to generalizing outside the United Kingdom. 1222 However, our findings are consistent with the reported pro-1223 portion of variation at the school level in other similar studies, 1224 including some in other countries.^{13,14,32,38,39} School-level 1225 1226 influences on pupil mental health may be observable only in 1227 pupils with significant problems, although this was not sup-1228 ported by our sensitivity analysis. Our sample cannot represent 1229 pupils who were excluded before commencement of the study 1230 1231**Q8** by their parents or by their school. Furthermore, we lacked 1232 data on some potentially important variables, such as family 1233 socioeconomic status, academic attainment, school-level 1234 violence, and pubertal status, all of which might influence 1235 mental health and well-being. Finally, our measure of school 1236 climate was based on teacher ratings alone, while a measure 1237

that included pupil, parent, and teacher ratings might have added different and valuable perspectives. 35

Our findings converge with others to suggest that in early adolescents 11-14 years of age, school influences explain 1.4%-2.4% of the variance in mental health and well-being. These small school-level effects may reflect a relative uniformity across schools in the United Kingdom in current approaches to pupil mental health. In schools located in urban areas, with pupils from predominantly White, disadvantaged backgrounds, poorer mental health in early adolescence is observed. At a population level, such findings are potentially important. Policy and system interventions focused on deprivation are likely to yield improvements in mental health of young people. In terms of schools, our findings converge with others to suggest the importance of school climate to support mental health and well-being in young people. In summary, this study has examined school structural and social features, both of which have important implications for guiding policy and the targeting of interventions.

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The corresponding study protocol can be found at https://trialsjournal. biomedcentral.com/articles/10.1186/s13063-017-1917-4. R code is available from the Open Science Framework at https://osf.io/s63fm/?view_only=5ae58f6 c053c4a16b5ddfccd0e6e1ece. The baseline data and codebook from the MYRIAD Project are available from Prof. Kuyken upon request (release of data is subject to an approved proposal and a signed data access agreement).

Data collection for the MYRIAD Project is ongoing, and the data used for this article were from an interim cut taken on April 18, 2019. Data may be subject to change for future publications owing to retrospective data deletion requests.

Dr. Ukoumunne served as the statistical expert for this research.

Author Contributions

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