

1 **Educational level as a cause of type 2 diabetes mellitus: Caution from triangulation of**
2 **observational and genetic evidence**

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1 **Abstract**

2 **Background:** Education might be causal to type 2 diabetes mellitus (T2DM). We triangulated
3 cohort and genetic evidence to consolidate the causality between education and T2DM.

4 **Methods:** We obtained observational evidence from the English Longitudinal Study of Ageing
5 (ELSA). Self-reporting educational attainment was categorised as high (post-secondary and
6 higher), middle (secondary), and low (below secondary or no academic qualifications) in 6,786
7 community-dwelling individuals aged ≥ 50 years without diabetes at ELSA wave 2, who were
8 followed until wave 8 for the first diabetes diagnosis. Additionally, we performed two-sample
9 Mendelian randomisation (MR) using an inverse-variance weighted (IVW), MR-Egger,
10 weighted median (WM), and weighted mode-based estimate (WMBE) method. Steiger
11 filtering was further applied to exclude single-nucleotide polymorphisms (SNPs) that were
12 correlated with an outcome (T2DM) stronger than exposure (education attainment).

13 **Results:** We observed 598 new diabetes cases after 10.4 years of follow-up. The adjusted
14 hazard ratios (95%CI) of T2DM were 1.20 (0.97-1.49) and 1.58 (1.28-1.96) in the middle- and
15 low-education groups, respectively, compared to the high-education group. Low education
16 was also associated with increased glycated haemoglobin levels. Psychosocial resources,
17 occupation, and health behaviours fully explained these inverse associations. In the MR
18 analysis of 210 SNPs ($R^2=0.0161$), the odds ratio of having T2DM per standard deviation-
19 decreasing years (4.2 years) of schooling was 1.33 (1.01-1.75; IVW), 1.23 (0.37-4.17; MR-
20 Egger), 1.56 (1.09-2.27; WM), and 2.94 (0.98-9.09; WMBE). However, applying Steiger
21 filtering attenuated most MR results toward the null.

22 **Conclusions:** Our inconsistent findings between cohort and genetic evidence did not support
23 the causality between education and T2DM.

24 **Keywords**

25 Educational level, Type 2 diabetes mellitus, Glycated haemoglobin, Prospective cohort,
26 Mendelian randomisation

27 **Key messages**

28 **What is already known on this subject?**

- 29 - Several pieces of evidence suggested that education attainment might play a causal
30 role in the occurrence of T2DM.

31 **What does this study add?**

- 32 - Our observational evidence suggested no direct impact of education on the risk of
33 T2DM. The observed inverse associations were mediated through insufficient
34 psychosocial resources, low occupation class, and unhealthy behaviours due to low
35 education.
36 - In contrast, the genetic evidence suggested no causal association between education
37 and the risk of T2DM. Notably, the significant associations from our genetic evidence
38 resulted from the invalid genetic instrument used in the analysis.

1 - The observational and genetic evidence was inconsistent; therefore, our triangulated
2 evidence did not support a causal role of education in the occurrence of T2DM.

3 **Introduction**

4 Type 2 diabetes mellitus (T2DM) is a significant global burden affecting more than 422 million
5 people worldwide, and its prevalence will reach 7,079 per 100,000 by 2030.[1] In some
6 countries, the increased prevalence reflects a better survival rate, while the incidence of
7 T2DM is still rising in others.[2]

8 Currently, T2DM is incurable. Thus, prevention is crucial. An effort has been put into clinical
9 risk modification, such as weight reduction and smoking cessation.[3] However, it has been
10 suggested that these clinical factors attributed to only one-third of total diabetes risks.[4]
11 Therefore, the residual factors are worth further investigation.

12 Previous observational studies have shown an inverse association between educational level
13 and the risk of T2DM.[5–8] However, residual confounders and reverse causality limited the
14 establishment of causality. Moreover, scarce evidence had examined the relationship
15 between education and glycated haemoglobin (HbA1c) levels,[9] which are the biomarker of
16 prediabetes and well-established cardiovascular risk.[10] Furthermore, findings from genetic
17 (Mendelian randomisation) studies are equivocal.[11–14]

18 This study aims to investigate the causal effect of education on the risk of T2DM and HbA1c
19 levels by comparing results from two different study designs – an approach called
20 ‘triangulation of evidence’.[15] Triangulated findings may complement the limitations of each
21 other and provide a more solid conclusion. The two methods being used here are cohort study
22 and Mendelian randomisation (MR). In brief, MR uses single-nucleotide polymorphisms
23 (SNPs) as a proxy of exposure. This genetic proxy is less likely to be associated with
24 confounders due to its random allocation according to Mendel’s law of independent
25 assortment.[16–18] Additionally, we also examine the causal pathway linking educational
26 level with the risk of T2DM and HbA1c levels.

27 **Methods**

28 This report followed the STrengthening the Reporting of OBservational Studies in
29 Epidemiology (STROBE) guidance of cohort studies and its extension to Mendelian
30 randomisation (STROBE-MR) (Table S1-S2).[19]

31 Cohort evidence

32 *Data source and study population*

33 We used the English Longitudinal Study of Ageing (ELSA) data: a prospective cohort study of
34 nationally representative community-dwelling individuals aged ≥50 years. At ELSA wave 1
35 (2002-03), samples included all consenting people who participated in the Health Survey for
36 England (HSE) in 1998, 1999 and 2001. Subsequent follow-up interviews and health
37 examinations take place regularly at two- and four-year intervals, respectively. More
38 information on ELSA can be found at <http://www.elsa-project.ac.uk/>.[20]

1 We used data from ELSA wave 2 (2004–05), which comprised follow-up interviews and health
2 examinations and constituted the baseline of our cohort study. Of 11,391 participants in ELSA
3 wave 1, 8,780 participated in ELSA wave 2, of whom 7,666 consented to the health
4 examination. Our final analysis included 6,786 individuals without a history of diabetes in
5 ELSA wave 2. To make the results comparable to genetic evidence, only white participants
6 (97.7% of core samples in ELSA wave 2) were included in the analyses (Figure S1).

7 *Educational level*

8 Educational level was the self-reported highest educational qualification obtained by ELSA
9 wave 2, further classified into three groups as implemented by a previous ELSA study.[5] A
10 high educational level was defined as a university degree, other higher or post-secondary
11 education, and A-level education (n=2,218), whereas a middle educational level included a
12 Certificate of Secondary Education (CSE) and similar foreign qualifications (n=2,106).
13 Individuals with below secondary education or without educational qualifications were
14 grouped as low educational level (n=2,462).

15 *Type 2 diabetes mellitus (T2DM) and glycated haemoglobin (HbA1c) levels*

16 The primary outcome was the self-report physician-diagnosed diabetes up to ELSA wave 8
17 (2016/17). To minimise misclassification bias, we included participants with HbA1c levels
18 $\geq 6.5\%$ at least twice in a diabetic group as suggested clinically.[10] The secondary outcome
19 was the trajectory of HbA1c levels measured at ELSA wave 2, 4, 6, and 8. Notably, HbA1c
20 measured in ELSA before October 2011 was calibrated using Diabetes Control and
21 Complications Trial (DCCT) standards, replaced by the International Federation of Clinical
22 Chemistry (IFCC) standardisation afterwards. Details of quality control of HbA1c measured in
23 ELSA has been published elsewhere.[21]

24 *Covariates*

25 We collected all covariates at baseline, mostly self-reported, except for body mass index
26 (BMI). These covariates included age (years), age², sex (i.e., male and female), marital status
27 (i.e., single, married, and separated/divorced or widowed), depressive symptom (i.e., Center
28 for Epidemiologic Studies Depression [CESD] score ≥ 4), occupational class (i.e., managerial or
29 professional, intermediate, and routine or manual occupation), BMI (i.e., normal, overweight,
30 and obese), smoking status (i.e., never, ex-, and current smoker), alcohol drinking (i.e., never
31 or almost never, 1-2 times a month, 1-2 times a week, and daily or almost daily). Moreover,
32 childhood socioeconomic position (SEP) was obtained and categorised into four groups
33 according to father's main job when participants aged 14 years: high (i.e., managerial-,
34 professional-, administrative occupations, or business owners); middle (i.e., trade- or services
35 related occupations); low (manual or casual occupations, unemployed, sick and disabled); and
36 miscellaneous (i.e., armed forces and retired). According to directed acyclic graphs (DAGs)
37 adapted from Hamad *et al.*[22] and Liang *et al.*[14] only age, sex, and childhood SEP were
38 considered confounders, whereas the rest were mediators (Figure S2).

39

1 *Statistical analysis*

2 Sample baseline characteristics were explored according to educational groups using
3 descriptive and inferential statistics, as appropriate. We created a Kaplan-Meier plot for the
4 cumulative incidence of T2DM of each group and compared it by log-rank test.

5 The association between educational levels and the risk of T2DM was examined using a Cox-
6 proportional hazards model with high education as a reference group. To investigate
7 potential causal pathways, models were adjusted for each set of covariates as follows:
8 confounding factors (i.e., age, sex, and childhood SEP); psychosocial resources (i.e.,
9 depressive symptom and marital status); occupational class (i.e., occupation); health
10 behaviours (i.e., BMI, smoking, alcohol drinking, and physical activity); and a final model that
11 was accounted for all covariates. The proportional hazards assumption was checked by
12 Schoenfeld residual statistic and log-minus-log plots. The multicollinearity of covariates was
13 examined by calculating variance inflation factor (VIF). Covariates with missing data (mostly
14 missed <5%, Table 1) were multiply imputed by chain equation (MICE) (see supplementary
15 materials).

16 Additionally, we performed the following sensitivity analyses: First, we analysed only
17 complete-case samples; second, we calculated a Bonferroni adjusted (97.5%) confidence
18 interval to account for multiplicity. Moreover, we performed subgroup analyses according to
19 sex, age groups (i.e., <75 and ≥75 years old), BMI groups, and smoking status.

20 To examine the association between education and the trajectory of HbA1c levels, we used a
21 multilevel linear (growth curve) model, allowing for random intercepts and random slopes
22 with unstructured covariance. The adjustment was similar to the T2DM outcome but based
23 on a complete-case approach. The model's validity was checked from the distribution of
24 intercepts and slopes. Sensitivity analysis was performed by excluding participants with
25 reporting diabetes during follow-up since they might receive antidiabetic agents that can
26 modify HbA1c levels and distort the actual effect of education on HbA1c.

27 Genetic evidence

28 *Data source*

29 All SNPs used in our study were derived from an MR-based platform as summary-level data
30 publicly available from <https://www.mrbase.org/>. [23] Specific ethical approval and consents
31 were already obtained in the original studies. Details of each genetic consortia can be found
32 in supplementary appendices (Table S6).

33 *Selection of instrumental variants*

34 We obtained SNPs associated with years of schooling from the Social Science Genetic
35 Association Consortium (SSGAC, n=1,131,881). [24] SNPs that reached genome-wide-
36 significance (i.e., $P\text{-value} < 5 \times 10^{-8}$) were selected and further pruned using linkage
37 disequilibrium (LD)- $r^2 < 0.001$ within a 10,000 kb window. The measuring unit of education in
38 SSGAC was per standard deviation (SD) increase in years of schooling (4.2 years).

1 *Outcomes and variants harmonisation*

2 T2DM and HbA1c variants were taken from the DIAbetes Genetics Replication And Meta-
3 analysis (DIAGRAM) consortium (34,840 cases 114,981 controls) [25] and the UK Household
4 Longitudinal Study (UKHLS, n=9,961),[26] respectively. Variants from different consortia were
5 harmonised, allowing for both palindromic SNPs (i.e., a minor allele frequency [MAF]
6 threshold of 0.3) and proxy SNPs (i.e., LD $r^2 > 0.8$ within 10,000 kb window; Figure S6).

7 *Statistical analysis*

8 The main analysis was performed using a multiplicative random-effect inverse-variance
9 weighted (IVW) method. For sensitivity analyses, we used the MR-Egger approach to examine
10 and account for unbalanced horizontal pleiotropy, if any.[27] Also, we performed a weighted
11 median and weighted mode MR. The former allowed for the invalidity of half of SNPs[28],
12 whereas the latter minimised the false-positive rates of findings.[29] These three additional
13 analyses were performed according to the guidelines for conducting MR investigations.[30]

14 Moreover, during a preliminary analysis, we observed that some of our selected SNPs had a
15 stronger association with the outcome than exposure. Therefore, we further applied MR
16 Steiger filtering to remove those SNPs and performed analyses accordingly.[31] To ensure the
17 validity of the genetic instruments and processes used in our MR analyses, we also examined
18 the association between education levels and the risk of Alzheimer's disease (AD: obtained
19 from the International Genomics of Alzheimer's Project [IGAP] consortium[32]) as a positive
20 control. This is because evidence suggested that higher education is causally related to a
21 decreased risk of AD.[33,34]

22 The power of derived effect size was estimated using the method given by Hermani *et al.*[23]
23 and Deng *et al.*[35] All analyses were performed using R version 3.6 and STATA version
24 16.1MP (StataCorp, LLC) with a two-sided alpha error of 5%. Since we considered T2DM and
25 HbA1c clinically correlated, we did not adjust for multiple testing in the MR analyses.[36]

26 *A conceptual framework of using the triangulation approach in this study*

- 27 - If findings from the cohort study show a significant association after adjusting for the
28 main confounders, then the true association is likely. Explanatory pathways will be
29 further elucidated to provide insight regarding a direct path between exposure and
30 outcome.
- 31 - The MR study is implemented to explore whether the observed association is due to
32 causation according to our conceptual framework of MR (supplementary appendices).
33 When evidence of causation is shown and the direction of the associations between
34 the cohort and MR is consistent throughout, the causality can be firmly established;
35 otherwise, the observed association might be alternatively explained by biases or
36 residual confounders.

37 **Results**

38 Cohort evidence

1 Among the 6,786 participants, most were female (56.4%), with a mean age of 66.3±9.8 years
2 old. People in the high education group were likely to be male and have higher occupation
3 classes and childhood SEPs. Also, they were likely to be non-smokers and had increased
4 physical activity levels and slightly lower BMIs than those in other groups (P-value<0.001). In
5 contrast, those in the low education group tended to be separated, divorced, or widowed and
6 have elevated depressive symptoms and a history of cardiovascular diseases (P-value<0.001,
7 Table 1).

8 After a median follow-up of 10.4 years, 598 out of the 6,786 participants reported diabetes
9 (10.1 [95%CI 9.3-10.9] per 1,000 person-years). A Kaplan-Meier plot had shown that low
10 education was associated with a significantly higher T2DM incidence (log-rank P-value<0.001,
11 Figure S3). Moreover, we observed a gradient inverse association between the education
12 levels and the risk of T2DM: The hazard ratios (95%CI) of T2DM in the middle and low
13 education groups were 1.20 (0.97-1.49) and 1.58 (1.28-1.96), respectively, compared to the
14 high education group in age, sex, and childhood SEP adjusted models (Table 2). The
15 significance remained after individually adjusting for health behaviours, psychosocial
16 resources, and occupational classes, but the association became null after simultaneous
17 adjustment. Admittedly, sex, age group, BMI, and smoking status did not significantly modify
18 the associations (Figure S4). Furthermore, the observed inverse associations were consistent
19 across sensitivity analyses (Table S4).

20 Regarding HbA1c levels (Table 2), we noticed that people in a low-education group had
21 slightly higher HbA1c levels than those in a high-education group ($\beta=0.0833$, 95%CI 0.0492-
22 0.1174) after controlling for age, sex, and childhood SEP. Additionally, the results were robust
23 after excluding diabetes participants (Table S5). The trajectory of HbA1c levels in each
24 educational group is illustrated in Figure S5.

25 Genetic evidence

26 From 1,271 schooling-associated SNPs, 210 and 184 SNPs were selected and harmonised with
27 T2DM and HbA1c levels, respectively (Figure S6). These can respectively explain 1.6% (F-
28 statistic=88.18) and 1.4% (F-statistic=74.07) of the variability in schooling years.

29 Although an inverse association between years of schooling and the risk of T2DM was initially
30 observed in the IVW model (Table 3), the results were not robust across sensitivity analyses.
31 In the IVW model, the odds of having T2DM decreased as schooling years increased: OR 0.75
32 (95%CI 0.57-0.99). The results were consistent with WM: OR 0.64 (95%CI 0.43-0.95) but not
33 with MR Egger (OR 0.81 [95%CI 0.24-2.69]) nor weighted mode MR (OR 0.34 [95%CI 0.11-
34 1.02]). We found no apparent evidence of heterogeneity on T2DM outcome ($I^2=12\%$, P-
35 value=0.09). Nevertheless, applying Steiger filtering attenuated most results towards the null.
36 Additionally, a scatter plot between SNPs-education and SNPs-T2DM did not show any
37 apparent pattern of the association (Figure S7). We also found a similar way of the
38 associations in HbA1c outcome (Table 3 and Figure S8).

1 Additionally, our positive control showed consistent findings with established evidence,
2 indicating the validity of instruments and processes used in our MR analyses (Table S8).

3 Triangulation of evidence

4 Importantly, when we triangulated pieces of evidence (Figure 1), we found inconsistent
5 results between observational study and MR. While cohort findings suggested inverse
6 associations between education level and the risk of T2DM and HbA1c levels, MR findings
7 suggested null associations.

8 **Discussion**

9 Summary of key findings

10 In the cohort study, we observed that low education was associated with an increased risk of
11 T2DM, possibly owing to inadequate psychosocial resources, unhealthy behaviours, and a
12 lower occupational class. Moreover, an observed inversed association was the same for
13 HbA1c levels, regardless of T2DM status. Nonetheless, findings from MR did not support a
14 causal association between education and the risk of T2DM and HbA1c levels. Further, they
15 indicated that significant MR results were dominated by SNPs directly associated with the
16 outcome and, therefore, not a good proxy of education.

17 Comparing with previous studies

18 Our observational findings are concordant with previous works indicating that education was
19 inversely associated with incident T2DM, and there is no direct pathway linked to T2DM.
20 However, in contrast to the previous ELSA report,[5] we did not observe different sex-specific
21 associations. This might be because we followed the participants for a more extended period
22 and used HbA1c as an additional criterion to define T2DM. So we could identify more T2DM
23 events in both sexes and gain better statistical power to detect slight differences. Moreover,
24 previous work also used antidiabetic medication data to ascertain diabetes, whereas, in this
25 study, we used only self-reporting diagnosis and HbA1c levels. Nevertheless, when we
26 restricted the analysis to wave 4, we found a trend of the association that was similar to the
27 previous ELSA study, where the association is more substantial in females than in males
28 (results not shown).[5] Also, our results were coherent with previous observational
29 studies.[6–8]

30 In terms of genetic evidence, an MR study by Hagenaars *et al.* suggested no causal link
31 between educational attainment and the risk of T2DM.[13] However, the study used only 9
32 SNPs, and the null findings might be due to statistical underpowering. In contrast to ours, two
33 recent MR studies have shown a causal association between education and the risk of
34 T2DM.[11,12] It should be noted that, before applying Steiger filtering, we produced relatively
35 similar (but slightly weaker) results as those works. However, after applying Steiger filtering,
36 almost all MR findings became null. Thus, we cannot exclude the potential direct effect of
37 genetic instruments used in their analyses. In addition, underpowering is unlikely to be a case
38 for T2DM outcome in our MR study (Table S7).

39 Interestingly, our null findings on HbA1c were consistent with the most recent work by Liang
40 *et al.* despite the contradiction in T2DM outcomes.[14] The discrepancy in results might be
41 due to different methods used for weighting SNPs in the IVW model. Rather than using an

1 additive model, we used a multiplicative one instead, as recommended.[30] The former may
2 upweight outlier SNPs and consequently erroneously strengthen the association, which can
3 be noticed by the similarity of scatter plots of SNPs exposure and SNPs outcome between our
4 work and the previous one.

5 Strengths and limitations

6 To our best knowledge, this is the first report that triangulated cohort and genetic evidence
7 on education and T2DM and HbA1c. However, there are some caveats worth noticing. First,
8 our outcome derived from self-reporting diabetes, which cannot differentiate between T1DM
9 and T2DM, and might be prone to misclassification bias. However, since incident diabetes
10 cases in our study were identified in participants aged ≥ 50 years, most events were clinically
11 assumed to be T2DM.[10] Additionally, it was shown that self-report diabetes had a very high
12 specificity (99.7%) but low sensitivity (66%).[37] Thus, we used HbA1c as an additional
13 criterion to define the outcome to improve false-negative cases. Second, we cannot exclude
14 the possibility that factors treated as mediators in our analysis can also be confounders since
15 we did not have the exact temporal sequence of each variable. For instance, some
16 participants might already be obese or active smokers before their graduation, and these risks
17 were carried over until age 50 when they participated in ELSA. Additionally, no evidence of
18 causal effect estimated from MR study could also be due to weak instrument bias from using
19 two different samples even if $F\text{-statistics} > 10$. [38] Lastly, the generalisability of our findings is
20 limited to European ancestry populations.

21 Implications of findings

22 The validity of SNPs used in the MR analysis should be a significant concern for both readers
23 and researchers when interpreting and implementing findings from MR studies. Meanwhile,
24 results from observational research alone prone to being misleading due to biases and
25 residual confounders. Hence, we encouraged using the triangulation approach to gain more
26 confidence in the causality inferences. Also, future works on different ethnicities might
27 warrant generalisability. According to our findings, targeting education might not directly
28 decrease the incidence of T2DM. However, education is a key to improve psychosocial
29 resources, healthy behaviours, and occupation, which might delay the occurrence of T2DM
30 and have a positive impact on health in the long run.[3] Therefore, improving education
31 should still be encouraged, although its causality to T2DM might not exist.

32 In summary, education did not directly affect T2DM and HbA1c levels. Inadequate
33 psychosocial resources, low occupational class, and unhealthy behaviours could explain the
34 observed inverse associations. Moreover, our triangulation of evidence did not support a
35 causal role of education in the risk of T2DM and HbA1c levels.

36

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9 publicly available.

10 **Declarations**

11 **Competing interests:** The authors have no conflicts of interest to declare that are relevant to
12 the content of this article.

13 **Availability of data and material:**

14 ELSA data were made available through the UK Data Archive ([https://www.ukdataservice.
15 ac.uk/](https://www.ukdataservice.ac.uk/)). Genetic data used in this research are publicly available from
16 <https://www.mrbase.org/>.

17 **Code availability:** In this study, all analyses were performed using STATA version 16MR
18 (StataCorp, LLC) package "stcox", "mixed", and "mrrobust". We also used R version 3.6
19 package "TwoSampleMR" for the genetic instrument extraction and harmonisation.
20 Additional R script and STATA do-file for the analyses were available upon request.

21 **Disclaimers:** The original data creators, depositors or copyright holders, the funders of the
22 Data Collections, and the UK Data Service/UK Data Archive bear no responsibility for analysing
23 or interpreting this study.

24 **Authors' contributions:** NN conceived the study aims and design and obtained access to ELSA
25 data. NN, JS, and AA contributed to the literature reviewing, data cleaning, data analyses,
26 interpretation of the findings. NN and SB developed the initial and subsequent manuscripts.
27 PC and PD critically revised the initial manuscript, and all authors participated in further
28 revisions. The final manuscript was read and approved by all authors before submission.

29 **Ethics approval:** The English Longitudinal Study of Ageing has been approved by the National
30 Research Ethics Service (London Multicentre Research Ethics Committee (MREC/01/2/91)).
31 For the MR study, specific ethical approval has been obtained individually in the original
32 genome-wide association studies (GWAS).

33 **Consent to participate:** Not applicable (specific consent was obtained in the original studies)

34 **Consent for publication:** Not applicable

1 **Table 1** Baseline characteristics of included participants

Characteristics	Educational levels			Total
	High	Middle	Low	
N	2,218	2,106	2,462	6,786
Male	1,220 (55.0)	851 (40.4)	890 (36.2)	2,961 (43.6)
Age (years)	63.6 ± 8.8	65.2 ± 9.2	69.6 ± 10.1	66.3 ± 9.8
High occupational class	1,400 (63.1)	503 (23.9)	240 (9.8)	2,143 (31.6)
Missing	11 (0.5)	16 (0.8)	65 (2.6)	92 (1.4)
High childhood SEP	1,039 (46.8)	660 (31.3)	388 (15.8)	2,087 (30.8)
Missing	82 (3.7)	88 (4.2)	96 (3.9)	266 (3.9)
Marital status				
Single	129 (5.8)	87 (4.1)	137 (5.6)	353 (5.2)
Married	1,625 (73.3)	1,436 (68.2)	1,421 (57.7)	4,482 (66.1)
Separated [§]	464 (20.9)	583 (27.7)	904 (36.7)	1,951 (28.8)
Elevated depressive symptoms	213 (9.6)	283 (13.4)	457 (18.6)	953 (14.0)
Missing	10 (0.5)	10 (0.5)	24 (1.0)	44 (0.7)
Never smoked	874 (39.4)	784 (37.2)	852 (34.6)	2,510 (37.0)
Missing	10 (0.5)	5 (0.2)	13 (0.5)	28 (0.4)
Never/ almost never drunk	352 (15.9)	507 (24.1)	948 (38.5)	1,807 (26.6)
Missing	10 (0.5)	5 (0.2)	15 (0.6)	30 (0.4)
High physical activity level	596 (26.9)	423 (20.1)	345 (14.0)	1,364 (20.1)
Missing	10 (0.5)	12 (0.6)	25 (1.0)	47 (0.7)
Body mass index (kg/m ²)	27.3 ± 4.6	27.5 ± 4.7	28.1 ± 4.9	27.7 ± 4.7
Missing	451 (20.3)	427 (20.3)	699 (28.4)	1,577 (23.2)
Having a history of CVDs	1,017 (45.9)	1,075 (51.0)	1,408 (57.2)	3,500 (51.6)

2 **Notes** All P-values (not including a missing group) were < 0.001. P-values were based on
3 different sample sizes for each variable with missing data as follows: 6,694 (occupation class),
4 6,520 (childhood SEP), 6,742 (depressive symptoms), 6,758 (smoking), 6,756 (alcohol
5 drinking), 6,739 (physical activity), and 5,209 (body mass index). Figures represent frequency
6 (%) or mean ± SD. Abbreviations: CVDs; Cardiovascular diseases, SEP; Socioeconomic position.
7 [§]Also included divorced and widowed

8

1 **Table 2** The association between education levels and the incidence of type 2 diabetes
 2 mellitus (n=6,786) and the trajectory of HbA1c levels (n=5,158)

Outcomes	High education	Middle education	Low education
Hazard ratios of incident T2DM (95%CI)			
Model 1	1.00 (ref)	1.22 (0.99 to 1.51)	1.71 (1.41 to 2.09)
<i>Confounder-adjusting models</i>			
Model 2	1.00 (ref)	1.28 (1.03 to 1.58)	1.78 (1.45 to 2.18)
Model 3 [§]	1.00 (ref)	1.20 (0.97 to 1.49)	1.58 (1.28 to 1.96)
<i>Confounder- and mediator-adjusting models</i>			
Model 4	1.00 (ref)	1.10 (0.89 to 1.37)	1.24 (0.99 to 1.54)
Model 5	1.00 (ref)	1.19 (0.96 to 1.48)	1.54 (1.24 to 1.91)
Model 6	1.00 (ref)	1.16 (0.92 to 1.46)	1.45 (1.14 to 1.84)
Model 7	1.00 (ref)	1.08 (0.86 to 1.36)	1.17 (0.92 to 1.50)
β-coefficients of HbA1c levels (95%CI)[‡]			
Model 1	0.00 (ref)	0.0263 (-0.0048 to 0.0575)	0.1088 (0.0772 to 0.1404)
<i>Confounder-adjusting models</i>			
Model 2	0.00 (ref)	0.0231 (-0.0081 to 0.0544)	0.0869 (0.0543 to 0.1194)
Model 3 [§]	0.00 (ref)	0.0217 (-0.0100 to 0.0533)	0.0833 (0.0492 to 0.1174)
<i>Confounder- and mediator-adjusting models</i>			
Model 4	0.00 (ref)	0.0028 (-0.0283 to 0.0338)	0.0396 (0.0055 to 0.0737)
Model 5	0.00 (ref)	0.0210 (-0.0107 to 0.0527)	0.0815 (0.0474 to 0.1157)
Model 6	0.00 (ref)	0.0058 (-0.0279 to 0.0394)	0.0554 (0.0173 to 0.0936)
Model 7	0.00 (ref)	-0.0074 (-0.0403 to 0.0255)	0.0230 (-0.0148 to 0.0607)

3 **Notes:** [§]Represent the main results. [‡]Random-intercept and-slope model with unstructured
 4 covariance. Embolden figures to represent statistical significance. **Abbreviations:** HbA1c;
 5 Glycated haemoglobin, T2DM; Type-2 diabetes mellitus. **Model 1:** Unadjusted models, **Model**
 6 **2:** Age and sex-adjusted models, **Model 3:** Model 2 + childhood SEP adjusted models, **Model**
 7 **4:** Model 3 + health behaviours adjusted models, **Model 5:** Model 4 + psychosocial resources
 8 adjusted models, **Model 6:** Model 3 + occupational class adjusted models, **Model 7:** Model 4
 9 + Model 5 + Model 6 adjusted models

1 **Table 3** The association between years of schooling, risk of T2DM, and HbA1c levels

MR model	Without Steiger filtering			With Steiger filtering		
T2DM	SNPs	OR (95%CI)	P-value	SNPs	OR (95%CI)	P-value
IVW [§]	210	0.75 (0.57, 0.99)^a	0.041	195	0.81 (0.62, 1.05) ^b	0.12
MR-Egger	210	0.81 (0.24, 2.69) ^c	0.73	195	0.98 (0.31, 3.10) ^d	0.97
WM	210	0.64 (0.44, 0.92)	0.017	195	0.74 (0.50, 1.11)	0.15
WMBE	210	0.34 (0.11, 1.02)	0.05	195	0.31 (0.09, 0.99)	0.049
HbA1c	SNPs	β (95%CI)	P-value	SNPs	β (95%CI)	P-value
IVW [§]	184	-0.1639^e (-0.2863, -0.0414)	0.009	46	-0.0782 ^f (-0.2584, 0.1019)	0.40
MR-Egger	184	-0.2319 ^g (-0.7198, 0.2560)	0.35	46	-0.0188 ^h (-0.9386, 0.9009)	0.97
WM	184	-0.5048 (-0.6755, -0.3342)	<0.001	46	-0.4422 (-0.7060, -0.1784)	0.001
WMBE	184	-0.7063 (-1.3318, -0.0807)	0.027	46	-0.5886 (-1.1761, -0.0012)	0.05

2 **Notes** Effect sizes were per standard deviation (SD) increase in years of schooling (4.2 years).
3 Embolden figures represent significant results. ^aI²=12% (P-value=0.09), ^bI²=0% (P-value=0.95),
4 ^cEgger-intercept = -0.001 (P-value = 0.90), ^dEgger intercept -0.003 (P-value=0.73), ^eI²=31% (P-
5 value<0.001), ^fI²=0% (P-value 0.53), ^gEgger-intercept = 0.001 (P-value = 0.78), ^hEgger intercept
6 -0.001 (P-value=0.90), [§]Represent the main results.

7 **Abbreviations** HbA1c: Glycated haemoglobin, IVW: Inverse-variance weighted, MR:
8 Mendelian randomisation, OR: Odds ratio, SNPs: Single Nucleotide Polymorphisms, T2DM:
9 Type 2 diabetes mellitus, WM: Weighted median, WMBE: Weighted mode-based estimates,

10

Type 2 DM

Study design
Cohort

Education Events/ Total Effect size [95%CI]*

High	166/ 2,218	1.00 [Ref]
Middle	182/ 2,106	1.20 [0.97, 1.49]
Low	250/ 2,462	1.58 [1.28, 1.96]

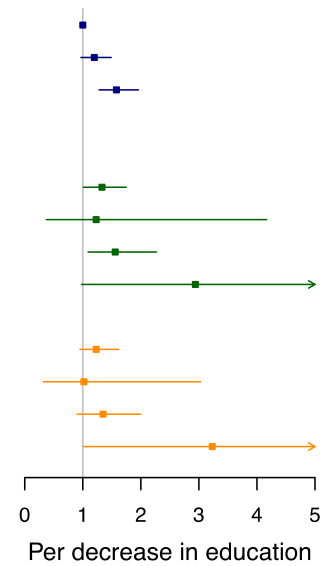
MR

Without Steiger filtering

IVW	210 SNPs	1.33 [1.01, 1.75]
MR-Egger	210 SNPs	1.23 [0.37, 4.17]
WM	210 SNPs	1.56 [1.09, 2.27]
WMBE	210 SNPs	2.94 [0.98, 9.09]

With Steiger filtering

IVW	195 SNPs	1.23 [0.95, 1.62]
MR-Egger	195 SNPs	1.02 [0.32, 3.03]
WM	195 SNPs	1.35 [0.90, 2.00]
WMBE	195 SNPs	3.23 [1.01, 11.11]



HbA1c

Study design
Cohort

Education Total Effect size [95%CI]*

High	1,735	0.00 [Ref]
Middle	1,653	0.02 [-0.01, 0.05]
Low	1,770	0.08 [0.05, 0.12]

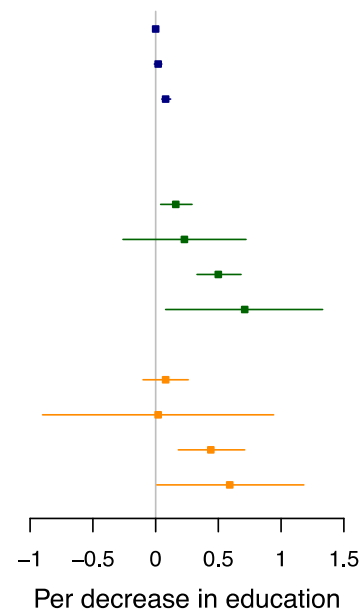
MR

Without Steiger filtering

IVW	184 SNPs	0.16 [0.04, 0.29]
MR-Egger	184 SNPs	0.23 [-0.26, 0.72]
WM	184 SNPs	0.50 [0.33, 0.68]
WMBE	184 SNPs	0.71 [0.08, 1.33]

With Steiger filtering

IVW	46 SNPs	0.08 [-0.10, 0.26]
MR-Egger	46 SNPs	0.02 [-0.90, 0.94]
WM	46 SNPs	0.44 [0.18, 0.71]
WMBE	46 SNPs	0.59 [0.01, 1.18]



1 **Figure 1** Triangulation of observational and genetic evidence on the association between
2 educational levels and the risk of type 2 diabetes mellitus and HbA1c levels

3 **Notes:** *Effect sizes are hazard ratio (adjusted for age, sex, and childhood SEP) for prospective
4 cohort design and odds ratio for Mendelian randomisation. Effect sizes from MR findings were
5 transformed from the original values to reflect per SD decrease in years of schooling.

6 **Abbreviations:** IVW; Inverse variance weighted, WM; Weighted median, WMBE; Weighted
7 mode-based estimate

8

1 **References**

- 2 1. Khan MAB, Hashim MJ, King JK, Govender RD, Mustafa H, Al Kaabi J. Epidemiology of
3 type 2 diabetes – Global burden of disease and forecasted trends. *J Epidemiol Glob*
4 *Health*. 2019;10(1):107.
- 5 2. Magliano DJ, Chen L, Islam RM, Carstensen B, Gregg EW, Pavkov ME, et al. Trends in
6 the incidence of diagnosed diabetes: A multicountry analysis of aggregate data from
7 22 million diagnoses in high-income and middle-income settings. *Lancet Diabetes*
8 *Endocrinol*. 2021;9(4):203–11.
- 9 3. American Diabetes Association. 3. Prevention or delay of type 2 diabetes: Standards
10 of medical care in diabetes-2021. *Diabetes Care*. 2021;44(Suppl 1):S34–9.
- 11 4. Steele CJ, Schottker B, Marshall AH, Kouvonen A, O’Doherty MG, Mons U, et al.
12 Education achievement and type 2 diabetes-what mediates the relationship in older
13 adults? Data from the ESTHER study: A population-based cohort study. *BMJ Open*.
14 2017;7(4):e013569.
- 15 5. Demakakos P, Marmot M, Steptoe A. Socioeconomic position and the incidence of
16 type 2 diabetes: The ELSA study. *Eur J Epidemiol*. 2012;27(5):367–78.
- 17 6. Lee TC, Glynn RJ, Peña JM, Paynter NP, Conen D, Ridker PM, et al. Socioeconomic
18 status and incident type 2 diabetes mellitus: Data from the women’s health study.
19 *PLoS One*. 2011;6(12):e27670.
- 20 7. Maty SC, Everson-Rose SA, Haan MN, Raghunathan TE, Kaplan GA. Education, income,
21 occupation, and the 34-year incidence (1965-99) of type 2 diabetes in the Alameda
22 County Study. *Int J Epidemiol*. 2005;34(6):1274–81.
- 23 8. Qi Y, Koster A, van Boxtel M, Köhler S, Schram M, Schaper N, et al. Adulthood
24 socioeconomic position and type 2 diabetes mellitus—A comparison of education,
25 occupation, income, and material deprivation: The Maastricht Study. *Int J Environ Res*
26 *Public Health*. 2019;16(8):1435.
- 27 9. Moody A, Cowley G, Ng Fat L, Mindell JS. Social inequalities in prevalence of
28 diagnosed and undiagnosed diabetes and impaired glucose regulation in participants
29 in the Health Surveys for England series. *BMJ Open*. 2016;6(2):e010155.
- 30 10. American Diabetes Association. 2. Classification and diagnosis of diabetes: Standards
31 of medical care in diabetes—2021. *Diabetes Care*. 2021;44(Supplement 1):S15–33.
- 32 11. Cao M, Cui B. Association of educational attainment with adiposity, type 2 diabetes,
33 and coronary artery diseases: a Mendelian randomization study. *Front Public Health*
34 [Internet]. 2020 Apr 22;8. Available from:
35 <https://www.frontiersin.org/article/10.3389/fpubh.2020.00112/full>
- 36 12. Adams CD, Boutwell BB. Can increasing years of schooling reduce type 2 diabetes
37 (T2D)? Evidence from a Mendelian randomization of T2D and 10 of its risk factors. *Sci*
38 *Rep*. 2020;10(1):12908.
- 39 13. Hagenaars SP, Gale CR, Deary IJ, Harris SE. Cognitive ability and physical health: A
40 Mendelian randomization study. *Sci Rep*. 2017;7(1):2651.
- 41 14. Liang J, Cai H, Liang G, Liu Z, Fang L, Zhu B, et al. Educational attainment protects

- 1 against type 2 diabetes independently of cognitive performance: a Mendelian
2 randomization study. *Acta Diabetol.* 2021;58(5):567–74.
- 3 15. Lawlor DA, Tilling K, Smith GD. Triangulation in aetiological epidemiology. *Int J*
4 *Epidemiol.* 2016;45(6):1866–86.
- 5 16. Na-Ek N. Mendelian Randomisation (MR): From a naturally randomised process to a
6 trendy research design. *Chiangrai Med J [Internet].* 2021;12(3):155–79. Available
7 from: <https://he01.tci-thaijo.org/index.php/crmjournal/article/view/242585>
- 8 17. Walker VM, Davey Smith G, Davies NM, Martin RM. Mendelian randomization: A
9 novel approach for the prediction of adverse drug events and drug repurposing
10 opportunities. *Int J Epidemiol.* 2017;46(6):2078–89.
- 11 18. Lawlor DA, Harbord RM, Sterne JAC, Timpson N, Davey Smith G. Mendelian
12 randomization: Using genes as instruments for making causal inferences in
13 epidemiology. *Stat Med.* 2008;27(8):1133–63.
- 14 19. Smith GD, Davies NM, Dimou N, Egger M, Gallo V, Golub R, et al. STROBE-MR:
15 Guidelines for strengthening the reporting of Mendelian randomization studies. *PeerJ*
16 *Prepr [Internet].* 2019;7:e27857v1. Available from:
17 <https://doi.org/10.7287/peerj.preprints.27857v1>
- 18 20. Clemens S, Phelps A, Oldfield Z, Blake M, Oskala A, Marmot M, et al. English
19 Longitudinal Study of Ageing: Waves 0-8, 1998-2017. [data collection]. 30th ed. UK
20 Data Service. 2019. SN: 5050, <http://doi.org/10.5255/UKDA-SN-5050-17>.
- 21 21. NatCen Social Research and UCL. Health Survey for England 2016: Methods [Internet].
22 2017 [cited 2021 Apr 17]. Available from:
23 <http://healthsurvey.hscic.gov.uk/media/63778/HSE2016-Methods-text.pdf>
- 24 22. Hamad R, Nguyen TT, Bhattacharya J, Glymour MM, Rehkopf DH. Educational
25 attainment and cardiovascular disease in the United States: A quasi-experimental
26 instrumental variables analysis. Rahimi K, editor. *PLOS Med.* 2019;16(6):e1002834.
- 27 23. Hemani G, Zheng J, Elsworth B, Wade KH, Haberland V, Baird D, et al. The MR-Base
28 platform supports systematic causal inference across the human phenome. *Elife.*
29 2018;7:e34408.
- 30 24. Lee JJ, Wedow R, Okbay A, Kong E, Maghzian O, Zacher M, et al. Gene discovery and
31 polygenic prediction from a genome-wide association study of educational
32 attainment in 1.1 million individuals. *Nat Genet.* 2018;50(8):1112–21.
- 33 25. Morris AP, Voight BF, Teslovich TM, Ferreira T, Segrè A V, Steinthorsdottir V, et al.
34 Large-scale association analysis provides insights into the genetic architecture and
35 pathophysiology of type 2 diabetes. *Nat Genet.* 2012;44(9):981–90.
- 36 26. Prins BP, Kuchenbaecker KB, Bao Y, Smart M, Zabaneh D, Fatemifar G, et al. Genome-
37 wide analysis of health-related biomarkers in the UK Household Longitudinal Study
38 reveals novel associations. *Sci Rep.* 2017;7(1):1–9.
- 39 27. Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid
40 instruments: effect estimation and bias detection through Egger regression. *Int J*
41 *Epidemiol.* 2015;44(2):512–25.

- 1 28. Bowden J, Smith GD, Haycock PC, Burgess S. Consistent estimation in Mendelian
2 Randomization with some invalid instruments using a weighted median estimator.
3 *Genet Epidemiol.* 2016;40(4):304–14.
- 4 29. Hartwig FP, Smith GD, Bowden J. Robust inference in summary data Mendelian
5 randomization via the zero modal pleiotropy assumption. *Int J Epidemiol.*
6 2017;46(6):1985–98.
- 7 30. Burgess S, Davey Smith G, Davies NM, Dudbridge F, Gill D, Glymour MM, et al.
8 Guidelines for performing Mendelian randomization investigations. *Wellcome Open*
9 *Res [Internet].* 2020 Apr 28;4:186. Available from:
10 <https://wellcomeopenresearch.org/articles/4-186/v2>
- 11 31. Hemani G, Tilling K, Smith GD. Orienting the causal relationship between imprecisely
12 measured traits using GWAS summary data. *PLOS Genet.* 2017;13(11):e1007081.
- 13 32. Lambert JC, Ibrahim-Verbaas CA, Harold D, Naj AC, Sims R, Bellenguez C, et al. Meta-
14 analysis of 74,046 individuals identifies 11 new susceptibility loci for Alzheimer’s
15 disease. *Nat Genet.* 2013;45(12):1452–8.
- 16 33. Anderson EL, Howe LD, Wade KH, Ben-Shlomo Y, Hill WD, Deary IJ, et al. Education,
17 intelligence and Alzheimer’s disease: evidence from a multivariable two-sample
18 Mendelian randomization study. *Int J Epidemiol.* 2020;49(4):1163–72.
- 19 34. Larsson SC, Traylor M, Malik R, Dichgans M, Burgess S, Markus HS. Modifiable
20 pathways in Alzheimer’s disease: Mendelian randomisation analysis. *BMJ.* 2017 Dec
21 6;359:j5375.
- 22 35. Deng L, Zhang H, Yu K. Power calculation for the general two-sample Mendelian
23 randomization analysis. *Genet Epidemiol.* 2020;44(3):290–9.
- 24 36. Proschan MA, Waclawiw MA. Practical guidelines for multiplicity adjustment in
25 clinical trials. *Control Clin Trials.* 2000;21(6):527–39.
- 26 37. Okura Y, Urban LH, Mahoney DW, Jacobsen SJ, Rodeheffer RJ. Agreement between
27 self-report questionnaires and medical record data was substantial for diabetes,
28 hypertension, myocardial infarction and stroke but not for heart failure. *J Clin*
29 *Epidemiol.* 2004;57(10):1096–103.
- 30 38. Burgess S, Thompson SG. Avoiding bias from weak instruments in Mendelian
31 randomization studies. *Int J Epidemiol.* 2011;40(3):755–64.
- 32