

Title: Impact of low maternal education on early childhood overweight and obesity in Europe

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Abbreviations: BMI – body mass index; CI – confidence interval; CZ-ELSPAC – The European Longitudinal Study of Pregnancy and Childhood, Czech Republic; DRIVERS – Determinants to Reduce Health Inequity Via Early Childhood, Realising Fair Employment, and Social Protection; ES-INMA – The Environment and Childhood Project, Spain; ESP – European Standard Population; FI-NFBC8586 – The Northern Finland Birth Cohort 1985/1986 Study, Finland; FR-EDEN – The Study of Pre- and Post-natal Determinants of Child Growth, Development, and Health; GR-GBC – The Greek Birth Cohort, Greece; IOTF – International Obesity Task Force; ISCED – International Standard Classification of Education; IT-GASPII – The Gene and Environment Prospective Study on Infancy in Italy, Italy; NL-ABCD – The Amsterdam Born Children and their Development Study, Netherlands; PT-G21 – The Generation XXI Study, Portugal; RII – Relative Index of Inequality; SE-ABIS – The All Babies in Southeast Sweden Study, Sweden; SII – Slope Index of Inequality; UK-MCS – The Millennium Cohort Study, United Kingdom; UA-FCOU – The Family and Children of Ukraine Study, Ukraine

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Abstract (250 words)

Background: Comparable evidence on adiposity inequalities in early life is lacking across a range of European countries. This study investigates whether low maternal education is associated with overweight and obesity risk in children from distinct European settings during early childhood.

Methods: Prospective data of 45,413 children from 11 European cohorts were used. Children's height and weight obtained at ages 4 to 7 years were used to assess prevalent overweight and obesity according to the IOTF definition. The Relative/Slope Indices of Inequality (RII/SII) were estimated within each cohort and by gender to investigate adiposity risk among children born to mothers with low education as compared to counterparts born to mothers with high education. Individual-data meta-analyses were conducted to obtain aggregate estimates and to assess heterogeneity between cohorts.

Results: Low maternal education yielded a substantial risk of early childhood adiposity across 11 European countries. Low maternal education yielded a mean risk ratio of 1.58 [95% confidence interval (CI) 1.34, 1.85] and a mean risk difference of 7.78% [5.34, 10.22] in early childhood overweight, respectively measured by the RII and SII. Early childhood obesity risk by low maternal education was as substantial for all cohorts combined (RII=2.61 [2.10, 3.23]) and (SII=4.01% [3.14, 4.88]). Inequalities in early childhood adiposity were consistent among boys, but varied among girls in a few cohorts.

Conclusions: Considerable inequalities in overweight and obesity are evident among European children in early life. Tackling early childhood adiposity is necessary to promote children's immediate health and well-being and throughout the life course.

Introduction

Early childhood adiposity generally tracks through the life course and is linked with poor metabolic risk profiles and chronic diseases in later life.¹⁻³ While social inequalities in overweight and obesity risk have been well established among school-age and adolescent European children, there is limited comparable evidence available for younger (3-7 years) children during early childhood across Europe.^{4,5}

The complex determinants of early childhood overweight and obesity are depicted by the Total Environment Assessment Model of Early Childhood.⁶ Social and environmental factors at the individual-, family-, residential- and broader national and regional levels influence overweight and obesity risk in both an independent and intertwined manner. Established risk factors across these levels include: poor maternal health and particularly maternal obesity; unhealthy behaviours during pregnancy and during childrearing; impaired foetal development, feeding practices at infancy and related diet; sedentary and sleep behaviour; alongside neighbourhood conditions such as safety and access to green spaces; to the availability, cost and formulation of food choices and the marketing environment.⁷⁻⁹ Extensive research has documented how these factors lead to inequalities in early childhood overweight and obesity in high-income countries, including those in Europe.^{5,7,10,11} While there is recognition that these established risk factors are socially patterned, it is unknown whether boys and girls differ in their susceptibility to certain social and environmental conditions.

Greater educational attainment among mothers may curtail the child's vulnerability to the cumulative nature of these obesogenic factors, through greater use of positive parenting practices and enhanced acquisition of the familial and residential environment, and a reduced likelihood of other disadvantages, such as belonging to a lone parent family and living in a deprived household or community.^{7,12} Low maternal education has been closely linked to

well-established life course predictors of early childhood adiposity, including smoking during pregnancy, foetal growth restriction, limited breast feeding, and food insecurity; as well as psycho-social stressors related to familial functioning and material insecurity.^{5,10,12,13}

High maternal education has been linked with a reduced likelihood of adiposity among pre-school aged children in some European populations, according to a recent systematic review of inequalities in early childhood health and development in Europe.⁵ As current evidence is primarily limited to individual-country studies and is disproportionately based on Western and Northern European populations, it is difficult to ascertain whether associations are consistent during the early childhood period across countries with unique obesogenic landscapes. Moreover, the potential role of gender upon the effect of established risk factors on early childhood adiposity has been overlooked in the literature.¹⁴ The objective of this study is to undertake a comparative examination of the association between maternal education and subsequent risk of both overweight and obesity among children aged 4-7 in eleven European birth cohorts, and to assess potential effect modification by gender.

Methods

Data sources

The DRIVERS for Health Equity research programme included a European cross-cohort study of inequalities in early child health and development. The present analysis draws on this component of DRIVERS, and is based on eleven European birth cohorts (in ascending order of child's age at measurement): the Gene and Environment Prospective Study on Infancy in Italy, Italy (IT-GASPII, N=543); the Environment and Childhood Project, Spain (ES-INMA, N=1726); the Generation XXI Study, Portugal (PT-G21, N=5686); the Millennium Cohort Study, United Kingdom (UK-MCS, N=14186); the All Babies in

Southeast Sweden Study (SE-ABIS, N=6494); the mother-child Study of Pre- and Post-natal Determinants of Child Growth, Development, and Health, France (FR-EDEN, N=1139); the Amsterdam Born Children and their Development Study, Netherlands (NL-ABCD, N=3245); the European Longitudinal Study of Pregnancy and Childhood, Czech Republic (CZ-ELSPAC, N=3665); the Family and Children of Ukraine Study, Ukraine (UA-FCOU, N=925); the Greek Birth Cohort, Greece (GR-GBC, N=1843); and the Northern Finland Birth Cohort 1985/1986 Study, Finland (FI-NFBC8586, N=5961). Details of participating cohorts including eligibility criteria have been previously reported.¹⁵ Participants consisted of 45,413 children born between April 1983 and October 2006, and for whom comparable data of maternal education, height and weight during early childhood were available.

Children's and maternal variables

Height and weight were obtained when children were aged between 4 to 7 years and measured by trained personnel in all cohorts, with the exception of GR-GBC and SE-ABIS. Data on these two items in both cohorts were based on parental reports, which were found to be comparable with measured height and weight in sub-sample studies.^{16,17} Body mass index (BMI) was calculated by dividing the child's weight in kilograms by the square of their height (kg/m^2). Overweight cases were defined using the age- and gender-specific thresholds for BMI recommended by The International Obesity Task Force (IOTF) in each cohort, which correspond to a BMI value of $25 \text{ kg}/\text{m}^2$ at 18 years of age. Supporting analyses on obese cases in each cohort were also defined by the IOTF as the age- and gender-specific BMI values that are comparable to an adult BMI value of $30 \text{ kg}/\text{m}^2$.¹⁸

Mother's education was ascertained at entry to each cohort study, either during pregnancy or near the time of birth. Years of schooling were available in FR-EDEN, GR-GBC, NL-ABCD and PT-G21. Levels of completed schooling were collected in IT-GASPII,

UA-FCOU, ES-INMA, CZ-ELSPAC, FI-NFBC8586 and SE-ABIS. Highest obtained qualifications were obtained in UK-MCS. The International Standard Classification of Education (ISCED) is an international categorization managed by the United Nations Educational, Scientific and Cultural Organization (UNESCO) to facilitate comparison of educational data across countries. The country-specific coding scheme provided by ISCED-1997 was used to classify mothers into the following categories: 1) Post-secondary non-tertiary to second stage of tertiary education (ISCED 4-6), 2) Upper secondary education (ISCED 3), and 3) Pre-primary to lower secondary or second stage of basic education (ISCED 0-2).¹⁹

Age, gender, and maternal age at study baseline were available in all participating cohorts and were included in the present analysis. Due to differences in the inclusion criteria between cohorts, foreign-born mothers or those belonging to an ethnic minority group were present in 7 of the 11 analytic samples. Data of mother's country of birth were available in IT-GASPII, FR-EDEN, ES-INMA and NL-ABCD. PT-G21 assessed mother's country of birth and migration history of mothers and her parents. In UK-MCS, perceived ethnicity was ascertained. These data were dichotomised and used as a proxy of mother's foreign-born or ethnic minority status.

Statistical analyses

Analyses were carried out among singleton participants, and study characteristics were assessed in each cohort (Table 1). The cohort-specific distribution of mother's education was directly age-standardised using the WHO European Standard Population (ESP) to account for variation in maternal age across samples.²⁰ Early childhood overweight and obesity prevalence by mother's education was calculated for each cohort. The X^2 test for

trend assessed linearity across educational groups, and separately for boys and girls (Table 2, Table 1S).

Associations between maternal education and early childhood overweight and obesity were estimated to infer relative and absolute social inequalities in each cohort sample, by using the Relative (and Slope) Indices of Inequality (RII/SII), respectively. These regression-based indices take into account both the size and distribution of social groups across the population.²¹ The RII is a summary measure of relative inequality, defined as the risk ratio of the outcome between children at the lowest and those at the highest end of the maternal education hierarchy. The SII is the corresponding measure of absolute inequality, defined as the risk difference in prevalence of the child outcome between the two ends of this hierarchy. A RII score greater than 1 (and a SII score greater than 0) indicates the presence of inequality between low and high positions. For both indices, higher scores denote a larger magnitude of inequality. Given the extent of cohort differences in maternal education, these indices account for such differences as a source of variation in the magnitude of health inequalities and facilitate comparison of estimates between cohorts.

Generalised linear models were performed to obtain the RII and SII by respectively specifying a logarithmic or an identity link function. Educational categories were ranked from high to low and were each assigned a value between 0 and 1, based on the cumulative percentage of the midpoint of the ranges observed for each cohort. Firstly, models were adjusted for age, gender, maternal age and ethnicity, as appropriate, for each outcome across cohorts. Sensitivity testing of these results were compared to associations with body mass index (kg/m^2) and adjusted for the same co-variables. Secondly, interaction terms between maternal education and gender were added to the models to assess effect modification by gender. Thirdly, as gender moderated the effect estimates among a few cohorts, models were then stratified by gender. Cohort- and gender-specific scores were pooled together to obtain

mean RII and SII scores for boys and girls at the aggregate level using random effects meta-analysis procedures (Figures 1 and 1S). The degree of heterogeneity was tested using the I^2 statistic and confirmed the a priori expectation of heterogeneity between cohorts that is not attributable to sampling variation. Meta-analysis findings were subsequently studied by random-effects meta-regression, in order to further assess effects moderated by gender.

Descriptive and regression analyses of the UK-MCS employed survey weights to account for the cohort's sampling design.²² Regression analyses of ES-INMA included geographic area as an additional co-variate to incorporate the multi-centre design of the cohort. All analyses were carried out using Stata version 13 (Stata Corp, College Station, TX).

Ethics

Study protocols were approved by ethics committees for each cohort. All participating mothers provided informed consent for themselves and their children. Maternal characteristics obtained at study baseline and child height and weight acquired during the preschool age period were provided from each cohort in accordance with the ethical procedures approved for each site.

Results

Mean mother's age ranged between 23.8 to 33.2 years at the time of childbirth (Table 1). Due to differences in study design, the proportion of foreign-born mothers was nil in four analytic samples, and elsewhere ranged from 1.1% in CZ-ELSPAC to 19% in NL-ABCD. Considerable proportions of mothers with low education were present in GR-GBC (68.3%), PT-G21 (60.6%) and ES-INMA (42.6%). Mothers with high education comprised a third or

more of samples in SE-ABIS, FR-EDEN, NL-ABCD and UA-FCOU. The prevalence of overweight children was lowest in UA-FCOU (7.6%) and highest in GR-GBC and IT-GASPII (24.1%).

Table 2 reports the prevalence distribution of overweight by mother's educational level by cohort and gender. Early childhood gradients in prevalent overweight among boys were statistically significant in IT-GASPII, PT-G21, UK-MCS, SE-ABIS, NL-ABCD, CZ-ELSPAC and FI-NFBC8586. Among girls, gradients in early childhood overweight were significant in PT-G21, UK-MCS, FR-EDEN, NL-ABCD, CZ-ELSPAC, UA-FCOU and FI-NFBC8586. In the case of UA-FCOU, the gradient indicated a higher risk of overweight among girls with mothers with higher levels of education. Elsewhere, overweight gradients were not detected by the X^2 test for trend, which may be due to limited study power among some cohorts.

Associations between mother's education and the prevalence of early childhood overweight were assessed within each cohort and adjusted for age and mother's age and ethnicity (as appropriate). Significant effect modification by gender was observed in relative (RII) and absolute (SII) inequalities among a few cohorts. As compared to boys, SII estimates were notably higher among girls in PT-G21 for overweight ($p=0.033$, test for interaction). SII estimates for overweight status also indicated gender differences in IT-GASPII ($p=0.016$, test for interaction) and UA-FCOU ($p=0.045$, test for interaction). However, in these two cases, the SIIs indicated a higher risk of overweight among girls with mothers with higher level of education. Gender differences were also found for relative (RII) inequalities in early childhood overweight in IT-GASPII ($p=0.016$, test for interaction) and SE-ABIS ($p=0.048$, test for interaction).

Sensitivity analyses using BMI (kg/m²) as a continuous outcome generated similar results to those of overweight prevalence in each cohort (available upon request) and addressed the concern of Type II error among smaller samples or those with fewer cases.

Figure 1 displays the combined RII and SII analyses across cohorts and by gender. Overall, children of low-educated mothers were 1.58 times (RII) more likely to be overweight than counterparts of mothers with high education, and corresponds to a mean risk difference of 7.8% (SII). Among boys, RIIs ranged from 1.22 in GR-GBC to 5.01 in NL-ABCD, combining to a mean risk ratio of 1.72. SIIs varied between 3.0% in UA-FCOU to 19.8% in IT-GASPII, and resulted in an average risk difference of 8.0%. Between cohort heterogeneity (I^2) was modest among boys, as shown by the proportion and 95% confidence interval estimates of 49.4% [95% CI: 0, 74.7] and 15.5% [95% CI: 0, 56.0], respectively for the RII/SII. Among girls, positive associations between low maternal education and lower risk of overweight were shown in IT-GASPII, UA-FCOU and GR-GBC. RIIs of overweight among girls extended from 0.23 in UA-FCOU to 3.8 in FR-EDEN, which aggregated to an average risk ratio of 1.44. Absolute inequalities (SIIs) in prevalent overweight spanned from a risk difference of -15.4% among girls in IT-GASPII to 19.4% among girls in PT-G21, which pooled to a mean risk difference of 6.86% across cohorts. Between cohort differences (I^2) were appreciable among girls, as denoted by the per cent and interval estimates for the mean RII/SII: 65.4% [95% CI: 45.7, 78.0] and 75.6% [95% CI: 56.1, 86.5].

Despite these differences in the combined overweight analysis, meta-regression found no substantial source of between cohort heterogeneity that was attributable to gender (p=0.21 for RII; p=0.90 for SII).

Supporting obesity analyses showed that the prevalence ranged from 1.1% in FR-EDEN to 6.5% in PT-G21. Gradients in obesity risk among boys and girls in each cohort were generally similar to those for overweight status (Table 1S). On average, children of

mothers with low education were 2.61 times (RII) more likely to be obese than those born to mothers with high education, which relates to an average risk difference in obesity of 4.0% in absolute terms (SII) (Figure 1S). Among girls, RIIs culminated to a 2.42 greater mean likelihood of obesity, and SIIs combined to mean risk difference of 3.98% for all cohorts combined. Statistical evidence of between cohort heterogeneity (I^2) in relative (RII) and absolute (SII) obesity inequalities was not found among boys or girls.

Discussion

To our knowledge, this is the first comparative study to examine the association between mother's education and early childhood adiposity among a multitude of birth cohorts, and to prominently demonstrate these inequalities in the European region. Substantial inequalities in both overweight and obesity risk were observed at the aggregate level. Non-significant findings observed among some cohorts may be related to factors of statistical power, such as analytic sample sizes, prevalence rates of early childhood overweight/obesity and the true magnitude of associations.

Previous regional work by the IDEFICS Consortium used cross-sectional surveys to assess adiposity risk by maternal education among children aged 4-5 in Belgium, Cyprus and Italy, and found that inequalities varied by gender in a few cases.²³ In parallel with these and other individual-level findings, the present study has demonstrated some inconsistency by gender in overweight inequalities among several cohorts.²³⁻²⁵ The social patterning of early childhood adiposity risk factors has been shown to differ throughout Europe.²⁶ For example, children of mothers with higher levels of education in Ukraine have shown to be more likely to consume greater quantities of expensive foods, such as red and processed meats, and thus exceed recommended levels of protein and cholesterol.²⁵ Increased protein intake has been

shown to predict increased adiposity among children in early childhood, and the endocrine response to a high-protein diet appears to be stronger among girls.^{27,28} Physical inactivity has also been linked to overweight and obesity during early childhood, and increased sedentary behaviour among young girls is well-established. Cultural norms and gendered expectations of physical activity during the early years may explain the extent to which inequalities differ across participating cohorts.^{29,30}

Obesogenic factors vary by broader environmental conditions that promote high energy intake and sedentary behaviour. These health behaviours are widely influenced by the built landscape, economic and political determinants and socio-cultural traditions.³¹ Taking the Ukrainian findings as example, obesity has been linked to socio-economic privilege preceding the fall of communism and to susceptibility to the free-market economy in Eastern Europe.³² Elsewhere, the general consistency of findings across cohorts from unique obesogenic environments and at different stages of the nutrition transition is noteworthy. As early childhood adiposity inequalities were demonstrated among children born over a twenty-year period, children and their mothers were subject to different environmental conditions that may have influenced the magnitude of inequalities by maternal education. Despite the observed differences in the RII and SII estimates between cohorts, the meta-analyses showed that estimated inequalities were not significantly heterogeneous for obese children, which underscores the impermeable risk attributed to low maternal education in early life. Estimated inequalities in early childhood overweight were shown to significantly differ between cohorts only among girls, which suggest some modification due to changes in obesogenic factors and broader conditions.

Preeminent inequalities in the Netherlands may relate to the large proportion of foreign-born mothers, which was linked with a higher risk of early childhood adiposity in this cohort.³³ Bias may occur as the cohorts required participation from birth to pre-school age,

where follow-up rates were lower in the Netherlands and Ukraine. Missing height and weight data were greater for children of mother's with low education in the Czech Republic, Greece and Portugal; which may underestimate inequalities in these cohorts. Parsimonious adjustment by child age, maternal age and foreign-born/ethnic minority status, does not eliminate residual confounding by factors unavailable for this analysis, such as markers of maternal health and other characteristics. For example, inequalities in early childhood adiposity by mother's education may be explained by maternal obesity, as rates tend to be higher among mothers with less education. Although data on mother's BMI prior to pregnancy were not available in all analytic cohort samples, this was accessible in the NL-ABCD and UK-MCS for the present study. RII and SII estimates were marginally attenuated and remained statistically significant after inclusion of maternal BMI in these cohorts. Analysis of children born over a twenty-year period may pose a methodological limitation, but these cohorts are among the most viable in the region, including some of the first to be established in their respective countries. Comparable findings of early childhood overweight and obesity using the same case definitions and methods of analyses on eleven European birth cohorts overcame the acknowledged gap in this area of evidence for the region, due to insufficient data sources and contrasting definitions employed between countries that make surveillance at the regional level problematic. Assessment by birth year of the children did not indicate a particular trend of inequalities in early childhood overweight and obesity.

This comparative study has established a greater risk of early childhood overweight and obesity by mother's education at the European level. Children's risk of adiposity firmly illustrates the strength of the social environment, and the limited reach of traditional prevention strategies.²⁶ The extent to which these inequalities manifest in children's early life has important policy implications. Firstly, the educational attainment of women and girls is a crucial social determinant for the health and well-being of offspring, given greater socio-

economic resources and enhanced facilities among mothers with higher levels of education. Our findings suggest that progress toward the Europe 2020 education targets to curtail the rates of early school leaving and to expand the share of young and older adults with tertiary education may have a beneficial impact on children's health of future mothers.³⁴ Actions to reduce overweight and obesity inequalities in women of childbearing age may work to reduce the inter-generational transmission of obesity risk from mothers to children. Secondly, in addition to school-based interventions, coordinated efforts need to be tailored to the family and the home environment to minimise the social gradient in overweight and obesity present well before children enter compulsory education.³⁵ Interventions targeted to mothers with lower levels of education may be implemented during pregnancy and following childbirth to offset risk factors of early childhood adiposity. For example, efforts aimed to lower prenatal growth restriction and rapid growth acceleration during infancy may improve the trajectory of adiposity throughout the life course, as well as those that promote prolonged and exclusive breast feeding, and timely and complementary feeding may also have a substantial effect on weight gain and body fatness during early childhood.

Overweight and obesity risk among European children during early childhood was greater among those born to mothers with low education. Social care and public health efforts must be accelerated to reduce these inequalities and to curtail further gaps through the life course.

Author contributions

MR designed the study, carried out all analyses, wrote the initial manuscript, and approved the final manuscript as submitted. PG critically reviewed the study, provided important theoretical contributions to the initial manuscript, and approved the final manuscript as submitted. JM and MM critically reviewed and revised the manuscript, and

approved the final manuscript as submitted. DP, FF, DH, YA, M-JS-C, SL, MV, MT, CI, IL, CB, AV, ME, TGMV, LA, LD, HB, SC, M-RJ, AT, JL and TF procured and coordinated data for this study, critically reviewed the manuscript, made important contributions to its intellectual content, and approved the final manuscript as submitted. HP conceptualised and designed the study, oversaw all stages of the analyses, guided the initial manuscript, and approved the final manuscript as submitted.

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